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# Placebo Effects on the Neurologic Pain Signature A Meta-analysis of Individual Participant Functional Magnetic Resonance Imaging Data

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**IMPORTANCE** Placebo effects reduce pain and contribute to clinical analgesia, but after decades of research, it remains unclear whether placebo treatments mainly affect nociceptive processes or other processes associated with pain evaluation.

**OBJECTIVE** We conducted a systematic, participant-level meta-analysis to test the effect of placebo treatments on pain-associated functional neuroimaging responses in the neurologic pain signature (NPS), a multivariate brain pattern tracking nociceptive pain.

**DATA SOURCES** Medline (PubMed) was searched from inception to May 2015; the search was augmented with results from previous meta-analyses and expert recommendations.

**STUDY SELECTION** Eligible studies were original investigations that were published in English in peer-reviewed journals and that involved functional neuroimaging of the human brain with evoked pain delivered under stimulus intensity-matched placebo and control conditions. The authors of all eligible studies were contacted and asked to provide single-participant data.

DATA EXTRACTION AND SYNTHESIS Data were collected between December 2015 and November 2017 following the Preferred Reporting Items for Systematic Review and Meta-Analyses of individual participant data guidelines. Results were summarized across participants and studies in a random-effects model.

MAIN OUTCOMES AND MEASURES The main, a priori outcome was NPS response; pain reports were assessed as a secondary outcome.

**RESULTS** We obtained data from 20 of 28 identified eligible studies, resulting in a total sample size of 603 healthy individuals. The NPS responses to painful stimulation compared with baseline conditions were positive in 575 participants (95.4%), with a very large effect size (g = 2.30 [95% Cl, 1.92 to 2.69]), confirming its sensitivity to nociceptive pain in this sample. Placebo treatments showed significant behavioral outcomes on pain ratings in 17 of 20 studies (85%) and in the combined sample (g = -0.66 [95% Cl, -0.80 to -0.53]). However, placebo effects on the NPS response were significant in only 3 of 20 studies (15%) and were very small in the combined sample (g = -0.08 [95% Cl, -0.15 to -0.01]). Similarly, analyses restricted to studies with low risk of bias (g = -0.07 [95% Cl, -0.15 to 0.00]) indicated very small effects, and analyses of just placebo responders (g = -0.22 [95% Cl, -0.34 to -0.11]) indicated small effects, as well.

**CONCLUSIONS AND RELEVANCE** Placebo treatments have moderate analgesic effects on pain reports. The very small effects on NPS, a validated measure that tracks levels of nociceptive pain, indicate that placebo treatments affect pain via brain mechanisms largely independent of effects on bottom-up nociceptive processing.

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Supplemental content

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Corresponding Author: Matthias Zunhammer, PhD, Klinik für Neurologie, Universitätsklinikum Essen, Hufelandstr. 55, 45147 Essen, Germany (matthias.zunhammer @uk-essen.de). **P** lacebo treatments are treatments with no intrinsic physical or pharmacological benefit. Nevertheless, they affect symptoms and physiology through patients' conceptions of the therapeutic context.<sup>1,2</sup> Placebo effects can be elicited by sham treatments, but they are also a substantial and beneficial part of the overall response to verum treatments, including those involving drugs and/or surgery.<sup>3,4</sup> In addition to conferring clinical benefits (and harms, in the case of nocebo effects), placebo effects reduce effect sizes for drug vs placebo differences in clinical trials, which may cause an increasing number of trials to fail<sup>5</sup> and thereby impede drug development.<sup>6,7</sup> Thus there is an urgent need to better understand placebo effects and to develop biomarkers for active drug responses, as well as placebo responses, to improve decision making in early clinical trials.<sup>8</sup>

Though placebo treatments can affect a variety of clinical outcomes,<sup>9</sup> placebo analgesia is the most robust and well studied.<sup>7,10,11</sup> Placebo analgesia has been linked with multiple psychological processes, including expectations and beliefs,<sup>12</sup> associative learning,<sup>13</sup> and social cognition.<sup>14</sup> Neurophysiological studies have suggested the involvement of descending inhibition of nociceptive afferents, with some studies supporting influences on spinal mechanisms<sup>15-19</sup> and others supporting higher-level cortical effects that cannot be explained by nociceptive input modulation alone, indicating affective or evaluative mechanisms.<sup>20-23</sup>

Nevertheless, the mechanisms by which placebo treatments change the perception of pain remain poorly understood, in part because of 2 limitations. First, previous studies were based on small sample sizes. Second, many brain areas associated with placebo analgesia, such as the anterior midcingulate cortex, the insula, and limbic regions, are involved in a range of functions, including cognitive decision making,<sup>24</sup> motor processes,<sup>25</sup> and emotion.<sup>26,27</sup> Thus, previous studies have not been able to establish whether placebo analgesia affects nociception-associated and pain-associated processing specifically or other cognitive and affective processes associated with the multidimensional experience of pain.

Recently, studies have begun to identify patterns of functional magnetic resonance imaging activity that yield objective and reliable<sup>28,29</sup> brain measures associated with evoked pain.<sup>30-33</sup> While they do not measure pain, which is by definition a subjective experience, they capture neurophysiological patterns associated with specific aspects of pain with high sensitivity.<sup>34,35</sup> Among ongoing efforts, the neurologic pain signature (NPS)<sup>33</sup> is a measure that has been shown to reliably track the intensity of evoked experimental pain across multiple studies with high sensitivity while responding only minimally to nonpainful somatic stimuli and other salient, aversive events, thus exhibiting high specificity.<sup>35-39</sup> Although fully understanding the neurophysiological processes captured by the NPS is a matter of ongoing investigation, previous results suggest that the NPS predominantly reflects changes in nociceptive input and the pain that arises from it, while being insensitive to higher cognitive pain modulation.<sup>36-39</sup>

In this study, we harness these methodological advances in a systematic meta-analysis of single-participant data testing placebo effects on NPS responses. If placebo treatments pre**Question** How do placebo treatments affect pain processing in the brain?

**Findings** This systematic meta-analysis of single-participant functional magnetic resonance imaging data of 603 healthy participants from 20 studies found that placebo treatments against experimental pain have moderate effects on pain reports, but very small effects on the neurologic pain signature, a cerebral measure of nociceptive pain.

Meaning Placebo analgesia seems to be predominantly mediated by networks different from those underlying the primary processing of noxious stimuli.

dominantly affected early nociceptive processes, they would be expected to reduce activity in the NPS. If so, placebo effects and the endogenous pain-regulatory processes they engage may have pervasive effects on pain generation, making it hard to dissociate pain-associated and placebo-associated processes and outcomes. Conversely, if placebos mainly affect later-stage affective and evaluative processes, they may have little influence on NPS responses. In this case, it may be possible to develop meaningful measures of nociception that are placebo insensitive.

## Methods

We followed the Preferred Reporting Items for Systematic Reviews and Meta-Analyses of Individual Participant Data.<sup>40</sup> The study protocol and hypotheses were registered (https://osf.io/n9mb3/) on December 10, 2015, which was after eligibility criteria and study search were defined, but before data collection and analysis began.

### **Eligibility Criteria and Study Search**

Criteria for study eligibility were peer-reviewed publication in the English language of an original investigation involving human participants who underwent functional neuroimaging of the brain during evoked pain and pain delivered under stimulus intensity-matched placebo and control conditions. Placebo treatment was defined as any condition where the experimental context suggested that an effective analgesic treatment was applied, including verbal suggestions and conditioning procedures that reinforced participants' expectations of reduced pain,<sup>41</sup> following the categorization of placebo paradigms introduced in Wager and Atlas in 2015.<sup>2</sup> Accordingly, nonplacebo control conditions that involved no treatment, ineffective treatment, hidden treatment (in contrast to open treatment), and unconditioned treatment (in contrast to conditioned treatment) were considered eligible. We identified potentially eligible studies on Medline on May 21, 2015 and December 6, 2015. This search was augmented by results from previous meta-analyses<sup>2,42</sup> and by requests for authors of individual placebo studies to identify additional studies missed in the online search. All 3 authors (M.Z., U.B., and

T.D.W.) screened the titles and abstracts of all records retrieved; studies that provisionally met eligibility criteria were assessed for eligibility by examining the full text. (Details appear in the eMethods and eTable 1 in the Supplement.)

### **Data Acquisition**

Authors of all eligible studies were contacted and asked to contribute participant-level neuroimaging, behavioral data (pain report), and demographic data. We requested data from the final, published analyses. Substudies with distinct participant samples and methods were treated as independent (studies 1 and 2 in Wager et al<sup>43</sup>). Reanalyses were excluded, but in the case of reanalyses with extended samples, <sup>44,45</sup> we included the study with the largest sample. Studies with partially missing data were included and analyzed based on the available data.<sup>46,47</sup>

## **Outcome Definition**

We obtained NPS responses<sup>33</sup> for each individual participant in each experimental condition and contrasted pain and placebo with pain and nonplacebo control conditions. The NPS responses were calculated as the dot product of each image with the NPS pattern, yielding a weighted average of activity across the image, where the NPS specifies the weights.<sup>33</sup> High NPS responses reflect both higher similarity of the individual images with the NPS pattern and higher-magnitude functional magnetic resonance imaging activity in the specified pattern.

To keep baseline conditions of studies with within-participant crossover designs and between-group placebo manipulations comparable, we limited our analysis to posttreatment conditions; in other words, additional baseline measurements from within-participant or mixed-design studies were excluded.<sup>46,48-50</sup> Further, nonpainful<sup>51</sup> or low-intensity<sup>43,46,48,52-54</sup> stimulus conditions were excluded from analysis to maximize detection sensitivity for placebo effects.

In several studies, images were provided for separate subconditions within placebo and control categories (eg, for leftlateralized and right-lateralized stimulation, <sup>55</sup> for strong and weak placebo conditions, <sup>56,57</sup> or for early-heat-pain and lateheat-pain periods).<sup>16,57,58</sup> In these cases, we summarized the NPS responses by calculating an average response under placebo and under control treatment for each participant (details appear in eTable 2 in the Supplement).

The scale of NPS responses depends on the scale of underlying imaging data and therefore on the image acquisition and analysis parameters used in the original studies. To avoid scaling issues, we based our analysis on the standardized effect size measure Hedges *g*, as is common in meta-analyses.<sup>59</sup> Similarly to the Cohen *d*, Hedges *g* is based on the mean difference between conditions divided by standard deviation, but with an additional correction for small sample bias. For withinparticipant studies we used Hedges *g*<sub>rm</sub>, which is based on the SD of within-participant differences corrected for withinparticipant correlations.<sup>60,61</sup>

#### **Risk of Bias Assessment**

We used the Cochrane risk of bias tool<sup>62</sup> to evaluate the risk of bias for studies included in the present meta-analysis (de-

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tails are in the eMethods and eTable 3 in the Supplement). We assessed biases from selection (which arises via insufficient randomization), performance (via insufficient blinding of participants or treatment providers), detection (via insufficient blinding of analysts), attrition (by missing data), reporting (via underreporting of nonsignificant studies), and sequence (which is potentially introduced by within-participant designs).

#### Analysis

Our main analysis followed a 3-part strategy. First, we tested the research question using all data available. Second, we conducted a conservative analysis excluding studies with high risk of bias. Third, we performed a responder analysis, in which we (1) only included participants who showed a behavioral placebo response (a pain report under placebo condition compared with the report under the control condition) above the study median, (2) excluded any experimental subconditions that may have diminished placebo effects (eg, placebo conditions deemed to be low efficacy or placebo conditions tested under pharmacological modulation), and (3) excluded any participants who were suspected outliers (the eMethods and eTable 2 in the Supplement).

Effects were summarized across studies using the generic inverse-variance weighting method with DerSimonian and Laird random effects,<sup>60</sup> meaning studies were weighted by 1/SE<sup>2</sup> (where SE is the standard error). We estimated heterogeneity in results using the  $\tau$  statistic, which represents the standard deviation of effect sizes between studies.<sup>59</sup> We tested against the null hypothesis of no effect at an error level of  $\alpha$  < .05 (2-tailed), with additional inferences based on Bayes factors.<sup>63</sup> In brief, Bayes factors represent the relative likelihood of the null and the alternative hypotheses and have the advantage (compared with P values) that support for the null hypothesis can be concluded. Further analysis details and procedures used to check image quality are provided in the eMethods in the Supplement. Analysis was completed with MATLAB 2016b (MathWorks). The analysis code is available at https://github.com/mzunhammer /PlaceboImagingMetaAnalysis.

## Results

## Data Acquisition

We identified 96 published articles, of which 28 were selected as eligible. These included a total of 759 participants. Data from 20 studies and 603 participants were obtained, thus including the majority of eligible studies (71%) and participants (79.4%) published until 2015. Details on data acquisition are provided in eFigure 1 and eTable 1 in the Supplement. Eligible studies that could not be obtained were generally similar to the studies obtained in terms of pain stimulation, placebo induction, and study design; notably, 8 eligible studies with patient samples could not be obtained (eTable 4 in the Supplement). An update of the study search in March 2018 indicated that at least 6 eligible studies (with a combined patient population of 196 individuals) were published after data collection (in eMethods and eTable 1 in the Supplement). Even when considering these additional stud-

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Table. Included Studies

| 2012     influsion       2016     19     Within     24     15 (79)     Laser     Inert topical<br>cream     Yes     Yes     Effective us<br>ineffective       Bingel et al, <sup>49</sup> 22     Within     28     15 (68)     Contact heat     Remifentanil<br>infusion     Yes     Yes     Open vs hidder       Choi et al, <sup>496</sup> 15     Within     25     15 (100)     Electrical     Saline infusion     Yes     Yes     Effective vs<br>ineffective       2011     15     Within     26     40 (100)     Contact heat     Inert topical     Yes     Yes     Effective vs<br>ineffective       2009     28     Within     26     19 (68)     Contact heat     Inert topical     Yes     No     Treatment vs n<br>ineffective vs<br>ineffective vs       213     24     Within     26     15 (42)     Distension     Saline infusion     Yes     No     Effective vs<br>ineffective vs       214, <sup>52</sup> 24     Within     27     12 (50)     Contact heat     Inert topical     Yes     Yes     Effective vs<br>ineffective vs   | Source                              | Participants,<br>No. | Design        | Mean<br>Age, y | Male,<br>No. (%) | Pain Stimulus   | Treatment Type    | Placebo<br>Suggestions | Placebo<br>Conditioning | Placebo Compariso                  |
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| 2014heatand inert topical<br>cream(lidocaine),<br>effective vs<br>ineffective (ine<br>cream)Theysohn et<br>al, 45 201430Within3515 (50)DistensionSaline infusionYesNoEffective vs<br>ineffectiveWager et al, 43<br>2004b524WithinnanaElectricalInert topical<br>creamYesNoEffective vs<br>ineffectiveWager et al, 43<br>2004b523WithinnanaContact heatInert topical<br>creamYesYesEffective vs<br>ineffectiveWrobel et al, 58<br>201438Within2622 (58)Contact heatInert topical<br>creamYesYesEffective vs<br>ineffectiveZeidan et al, 47<br>17a17aWithin288 (47)Contact heatInert topical<br>creamYesYesTreatment vs no   |                                     | 102                  | Between       | 25             | 32 (31)          | Electrical      | Inert pill        | Yes                    | Yes                     | Treatment vs no<br>treatment       |
| al, <sup>45</sup> 2014   ineffective     Wager et al, <sup>43</sup> 24   Within   na   na   Electrical   Inert topical cream   Yes   No   Effective vs ineffective     Wager et al, <sup>43</sup> 23   Within   na   na   Contact heat   Inert topical cream   Yes   Yes   Effective vs ineffective     2004b <sup>c</sup> 23   Within   na   na   Contact heat   Inert topical cream   Yes   Yes   Effective vs ineffective     Wrobel et al, <sup>58</sup> 38   Within   26   22 (58)   Contact heat   Inert topical cream   Yes   Yes   Effective vs ineffective     2014   26   22 (58)   Contact heat   Inert topical cream   Yes   Yes   Treatment vs mineffective     Zeidan et al, <sup>47</sup> 17 <sup>a</sup> Within   28   8 (47)   Contact heat   Inert topical   Yes   Yes   Treatment vs mineffective   |                                     | 32                   | Within        | 26             | 17 (53)          |                 | and inert topical | Yes                    | No                      | effective vs<br>ineffective (inert |
| 2004ab cream ineffective   Wager et al, <sup>43</sup> 23 Within na na Contact heat Inert topical cream Yes Yes Effective vs ineffective   2004b <sup>c</sup> Nobel et al, <sup>58</sup> 38 Within 26 22 (58) Contact heat Inert topical cream Yes Yes Effective vs ineffective vs ineffective   2014 28 8 (47) Contact heat Inert topical Yes Yes Treatment vs m   |                                     | 30                   | Within        | 35             | 15 (50)          | Distension      | Saline infusion   | Yes                    | No                      |                                    |
| 2004b <sup>c</sup> cream ineffective   Wrobel et al, <sup>58</sup> 38 Within 26 22 (58) Contact heat Inert topical Yes Yes Effective vs   2014 cream cream ineffective ineffective   Zeidan et al, <sup>47</sup> 17 <sup>a</sup> Within 28 8 (47) Contact heat Inert topical Yes Yes Treatment vs n  |                                     | 24                   | Within        | na             | na               | Electrical      |                   | Yes                    | No                      |                                    |
| 2014 cream ineffective   Zeidan et al, <sup>47</sup> 17 <sup>a</sup> Within 28 8 (47) Contact heat Inert topical Yes Yes Treatment vs m  |                                     | 23                   | Within        | na             | na               | Contact heat    |                   | Yes                    | Yes                     |                                    |
|  |                                     | 38                   | Within        | 26             | 22 (58)          | Contact heat    |                   | Yes                    | Yes                     |                                    |
|  |                                     | 17 <sup>a</sup>      | Within        | 28             | 8 (47)           | Contact heat    |                   | Yes                    | Yes                     | Treatment vs no treatment          |
| bbreviations: NA, not available; TENS, transcutaneous electrical nerve <sup>b</sup> Substudy 1.  | bbreviations: NA                    | , not available;     | TENS, transcu | taneous ele    | ctrical nerve    | <sup>b</sup> Su | ostudy 1.         |                        |                         |                                    |

ies, the present meta-analysis covers most eligible studies (59%) and participants (63.1%).

## **Risk of Bias**

For pain ratings, the assessment of risk of bias (eResults, eTable 3, and eFigure 2 in the Supplement) indicated a high risk of performance (self-report) and detection bias, as well as unknown levels of reporting (publication) bias. For NPS responses, we found low risk of bias, because this measure does not depend on self-report and was unknown when the original studies were performed.

# Sample Description

Included studies are listed in the **Table**, and key sample characteristics are shown in eFigure 3 in the Supplement. Details on pain stimulation, placebo treatment, image acquisition, and imaging analyses are provided in the eMethods and eTables 5, 6, 7, and 8 in the Supplement. Image alignment to Montreal Neurological Institute space was satisfactory, and the coverage of the voxels making up the NPS was near optimum levels (98.4% across all participants; eMethods and eFigure 4 in the Supplement). Four participants showed mean pain ratings less than 5% of the pain scale, indicating insufficient pain stimulation; evidence for imaging artifacts was found in 12 participants (2.0%) (eMethods in the Supplement). These participants were defined as outliers and excluded from the responder analysis, but retained in the primary and the conservative analysis.

## NPS Responses to Painful Stimulation

The NPS responses to painful stimulation, compared with low-level baseline of placebo and control conditions pooled,

| Study                                    | Participants,<br>No. | Effect<br>(95% CI)  | Reduction | Increase  | Weight<br>% |
|--|----------------------|---------------------|-----------|---|-------------|
| Atlas et al, <sup>52</sup> 2012          | 21                   | 2.52 (1.29 to 3.74) |           | <b></b>   | 4.5         |
| Bingel et al, <sup>55</sup> 2006         | 19                   | 3.20 (1.64 to 4.77) |           | <b>-</b>  | 3.5         |
| Bingel et al, <sup>49</sup> 2011         | 22                   | 2.50 (1.31 to 3.69) |           |   | 4.6         |
| Choi et al, <sup>56</sup> 2011           | 15                   | 1.88 (0.71 to 3.04) |           | _ <b></b>                                       | 4.7         |
| Eippert et al, <sup>16</sup> 2009        | 40                   | 1.27 (0.69 to 1.86) |           |   | 7.0         |
| Ellingsen et al, <sup>51</sup> 2013      | 28                   | 4.01 (2.44 to 5.57) |           | <b>_</b>  | 3.5         |
| Elsenbruch et al, <sup>64</sup> 2012     | 36                   | 2.27 (1.40 to 3.13) |           |   | 5.8         |
| Freeman et al, <sup>53</sup> 2015        | 24                   | 2.09 (1.09 to 3.09) |           |   | 5.3         |
| Geuter et al, <sup>57</sup> 2013         | 40                   | 1.94 (1.20 to 2.67) |           |   | 6.3         |
| Kessner et al, <sup>50</sup> 2014        | 39                   | 1.45 (0.82 to 2.07) |           |   | 6.8         |
| Kong et al, <sup>46</sup> 2006           | 10                   | 3.10 (1.02 to 5.19) |           |   | 2.4         |
| Kong et al, <sup>48</sup> 2009           | 12                   | 4.37 (1.79 to 6.96) |           |   | 1.7         |
| Lui et al, <sup>65</sup> 2010            | 31                   | 3.48 (2.16 to 4.80) |           | <b></b>   | 4.2         |
| Rütgen et al, <sup>54</sup> 2015         | 102                  | 3.61 (2.86 to 4.36) |           |   | 6.3         |
| Schenk et al, <sup>79</sup> 2015         | 32                   | 2.34 (1.40 to 3.28) |           |   | 5.5         |
| Theysohn et al, <sup>45</sup> 2009       | 30                   | 2.64 (1.57 to 3.70) |           |   | 5.1         |
| Wager et al, <sup>43</sup> 2004, study 1 | 24                   | 2.81 (1.56 to 4.07) |           | — <b>—</b>                                      | 4.4         |
| Wager et al, <sup>43</sup> 2004, study 2 | 23                   | 1.82 (0.89 to 2.75) |           |   | 5.6         |
| Wrobel et al, <sup>58</sup> 2014         | 38                   | 1.63 (0.95 to 2.31) |           |   | 6.6         |
| Zeidan et al, <sup>47</sup> 2015         | 17                   | 0.98 (0.18 to 1.77) |           |   | 6.1         |
| Total                                    | 603                  | 2.30 (1.92 to 2.69) |           | $\diamond$                                      | 100.0       |
|  |                      | -5                  | Neurologi | c Pain Signature Respons<br>dges g) With 95% Cl | 10<br>e     |

Figure 1. Neurologic Pain Signature Response to Noxious Stimulation vs Baseline in All Studies

Placebo and control conditions were pooled and studies weighted according to inverse variance. Total z = 11.74 (P < .001); heterogeneity  $\chi^2_{19} = 57.15$  (P < .001);  $\tau^2 = 0.47$ ;  $l^2 = 6.76\%$ .

were very large (g = 2.30 [95% CI, 1.92 to 2.69]) and significant in all included studies (Figure 1). Positive NPS responses were observed in 576 individuals (95.4%) (95.6% when outliers were excluded). The estimated between-study standard deviation of effect sizes  $\tau$  was 0.68, indicating that the outcomes of painful stimulation on NPS responses varied considerably among studies, which was as expected, given the broad range of different imaging and pain stimulation protocols used (Table; eFigure 3 and eTables 5, 6, 7, and 8 in the Supplement). Moreover, the NPS responded more strongly to high vs low levels of stimulus intensity in 6 of 7 studies (86%) in which this comparison was possible (g = 1.32 [95% CI, 0.86 to 1.79]; eFigure 5 in the Supplement). Together, these findings validate that the NPS was sensitive to noxious stimulation in this sample.

#### Placebo Effects on Pain Ratings and NPS Responses

Placebo treatments, compared with matched control conditions, showed moderate<sup>66</sup> analgesic effects on pain ratings (g = -0.66 [95% CI, -0.80 to -0.53]; Figure 2), which corresponds to a reduction of -11.3 (95% CI, -14.0 to -8.56) units on a 101-point visual analogue scale. The Bayes factor ( $B_{N[O, 0.5]}$ ) obtained for a normal null-prior (g = 0 [SD, 0.5], 2-tailed) was 9.4 × 10<sup>18</sup>, indicating overwhelming support for the hypothesis of nonnull placebo effects on pain ratings.<sup>63</sup> Effect sizes varied considerably among studies ( $\tau = 0.24$ ), which may be explained by the variation in placebo paradigms used (Table; eTable 6 in the Supplement).

In contrast, effects of placebo treatments on NPS responses were small (g = -0.08 [95% CI, -0.15 to -0.01]; Figure 3), with little between-study heterogeneity ( $\tau = 0.015$ ).

Thus, placebo effects on the NPS were 12.1% as large as effects on pain ratings, and 3.7% as large as the effects of painful stimulation on the NPS. The Bayes factor was less than 1 ( $B_{N[0, 0.5]} = 0.805$ ), indicating that these data provide very weak support in favor of the null hypothesis of no effect.<sup>63</sup>

A conservative analysis excluding all studies with high risk of bias and therefore including 15 studies with 429 participants yielded similar results (placebo - control on NPS,  $g = -0.07 [95\% \text{ CI}, -0.15 \text{ to } 0.00]; B_{N(0, 0.5)} = 0.787; eFigure 6$ in the Supplement). Somewhat larger effects were found in the responder analysis, which included only participants showing a behavioral placebo response greater than the study median and excluded potentially ineffective placebo treatments and outliers (196 participants from 18 studies were included). In this sample of so-called placebo responders, the Bayes factor (B<sub>NI0, 0.51</sub> = 113.8) indicated robust support for the hypothesis of a nonnull placebo effect on NPS responses. However, the effects remained in the small range (g = -0.22 [95% CI, -0.34 to -0.11]; eFigure 7 in the Supplement). Thus, even in placebo responders, effects of placebo on the NPS were only 4% to 14% as large as the overall NPS response to painful stimulation (Figure 1).

## Associations Between Placebo Effects on Pain Reports and NPS Responses

For studies with crossover designs, which included withinparticipant testing of both placebo and control treatments, we performed a meta-analysis of within-study correlations between placebo effects on pain report and NPS responses across individuals. A Bayes factor of 894.5 ( $B_{N[0, 0.5]}$ ) indicated robust support for the hypothesis of a nonnull correlation, with greater placebo analgesia associated with greater

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| Participants,<br>No. | Effect<br>(95% CI)   | Reduction  | Increase Weight,<br>%  |
|----------------------|--|--|--|
| 19                   | -0.48 (-0.90 to 0.05)  | <b>-</b> _   | 4.5  |
| 19                   | -1.05 (-1.37 to -0.72)   |  | 5.4  |
| 22                   | -1.10 (-1.52 to -0.67)   |  | 4.4  |
| 15                   | -0.81 (-1.08 to -0.53)   |  | 6.0  |
| 40                   | -0.67 (-0.95 to -0.39)   |  | 5.9  |
| 28                   | -0.46 (-0.77 to -0.14)   |  | 5.6  |
| 36                   | -0.37 (-0.61 to -0.14)   |  | 6.5  |
| 24                   | -0.96 (-1.36 to -0.57)   |  | 4.8  |
| 40                   | -0.70 (-1.01 to -0.38)   |  | 5.6  |
| 39                   | -0.94 (-1.59 to -0.29)   | <b>-</b>   | 2.8  |
| 10                   | -0.40 (-0.84 to 0.05)  | <b>_</b>   | 4.3  |
| 12                   | -0.49 (-1.07 to 0.09)  |  | 3.2  |
| 31                   | -0.48 (-0.70 to -0.26)   |  | 6.6  |
| 102                  | -1.00 (-1.41 to -0.59)   |  | 4.6  |
| 32                   | -0.17 (-0.53 to 0.19)  |  | - 5.1  |
| 30                   | -0.35 (-0.70 to -0.01)   |  | 5.3  |
| 24                   | -0.36 (-0.70 to -0.03)   |  | 5.4  |
| 23                   | -1.27 (-1.72 to -0.82)   |  | 4.2  |
| 38                   | -1.12 (-1.48 to -0.75)   | <b></b>  | 5.1  |
| 17                   | -0.43 (-0.83 to -0.03)   |  | 4.7  |
| 601                  | -0.66 (-0.80 to -0.53)   | $\diamond$   | 100.0  |
|                      | -2   | -1 0   | 1 2  |
|                      | 19     19     22     15     40     28     36     24     40     39     10     12     31     102     32     30     24     23     38     17 | $\begin{array}{c c c c c c c c c c c c c c c c c c c $ | 19 $-0.48 (-0.90 \text{ to } 0.05)$ 19 $-1.05 (-1.37 \text{ to } -0.72)$ 22 $-1.10 (-1.52 \text{ to } -0.67)$ 15 $-0.81 (-1.08 \text{ to } -0.53)$ 40 $-0.67 (-0.95 \text{ to } -0.39)$ 28 $-0.46 (-0.77 \text{ to } -0.14)$ 36 $-0.37 (-0.61 \text{ to } -0.14)$ 24 $-0.96 (-1.36 \text{ to } -0.57)$ 40 $-0.70 (-1.01 \text{ to } -0.38)$ 39 $-0.94 (-1.59 \text{ to } -0.29)$ 10 $-0.48 (-0.70 \text{ to } -0.26)$ 12 $-0.49 (-1.07 \text{ to } 0.9)$ 31 $-0.48 (-0.70 \text{ to } -0.26)$ 32 $-0.17 (-0.53 \text{ to } 0.19)$ 30 $-0.35 (-0.70 \text{ to } -0.01)$ 24 $-0.36 (-0.70 \text{ to } -0.03)$ 32 $-1.17 (-1.72 \text{ to } -0.82)$ 38 $-1.12 (-1.48 \text{ to } -0.75)$ 17 $-0.43 (-0.83 \text{ to } -0.33)$ 601 $-0.66 (-0.80 \text{ to } -0.53)$ |

Studies were weighted according to inverse variance. Total z = 9.65 (P < .001); heterogeneity  $\chi^2_{19}$  = 55.35 (P < .001);  $\tau^2$  = 0.06;  $l^2$  = 65.68%.

Figure 3. Changes in Neurologic Pain Signature Responses After Experimental Placebo Treatments

| Study                                    | Participants,<br>No. | Effect<br>(95% CI)     | Reduction Increase   | Weight<br>% |
|--|----------------------|------------------------|--|-------------|
| Atlas et al, <sup>52</sup> 2012          | 21                   | -0.01 (-0.55 to 0.52)  |  | 1.9         |
| Bingel et al, <sup>55</sup> 2006         | 19                   | -0.44 (-0.90 to 0.02)  |  | 2.5         |
| Bingel et al, <sup>49</sup> 2011         | 22                   | -0.14 (-0.45 to 0.18)  | <b>_</b>   | 5.4         |
| Choi et al, <sup>56</sup> 2011           | 15                   | -0.72 (-1.37 to -0.08) |  | 1.3         |
| Eippert et al, <sup>16</sup> 2009        | 40                   | -0.08 (-0.27 to 0.11)  |  | 14.6        |
| Ellingsen et al, <sup>51</sup> 2013      | 28                   | -0.08 (-0.51 to 0.35)  |  | 2.9         |
| Elsenbruch et al, <sup>64</sup> 2012     | 36                   | -0.02 (-0.34 to 0.30)  |  | 5.2         |
| Freeman et al, <sup>53</sup> 2015        | 24                   | -0.07 (-0.33 to 0.19)  | _ <b>_</b>   | 8.0         |
| Geuter et al, <sup>57</sup> 2013         | 40                   | 0.07 (-0.15 to 0.29)   |  | 10.6        |
| Kessner et al, <sup>50</sup> 2014        | 39                   | -0.25 (-0.87 to 0.36)  |  | 1.4         |
| Kong et al, <sup>46</sup> 2006           | 10                   | 0.26 (-0.14 to 0.65)   |  | 3.4         |
| Kong et al, <sup>48</sup> 2009           | 12                   | -0.14 (-0.69 to 0.40)  |  | 1.8         |
| Lui et al, <sup>65</sup> 2010            | 31                   | 0.04 (-0.24 to 0.31)   | _ <b></b>  | 7.0         |
| Rütgen et al, <sup>54</sup> 2015         | 102                  | -0.06 (-0.45 to 0.33)  |  | 3.6         |
| Schenk et al, <sup>79</sup> 2015         | 32                   | -0.27 (-0.54 to 0.00)  |  | 7.2         |
| Theysohn et al, <sup>45</sup> 2009       | 30                   | 0.03 (-0.25 to 0.31)   | _ <b></b>  | 6.7         |
| Wager et al, <sup>43</sup> 2004, study 1 | 24                   | 0.03 (-0.32 to 0.39)   |  | 4.1         |
| Wager et al, <sup>43</sup> 2004, study 2 | 23                   | 0.00 (-0.36 to 0.36)   |  | 4.1         |
| Wrobel et al, <sup>58</sup> 2014         | 38                   | -0.36 (-0.66 to -0.05) |  | 5.6         |
| Zeidan et al, <sup>47</sup> 2015         | 17                   | -0.28 (-0.71 to 0.15)  | <b>_</b>   | 2.9         |
| Total                                    | 603                  | -0.08 (-0.15 to -0.01) | $\diamond$   | 100.0       |
|  |                      | -2                     | -1 0 1<br>Neurologic Pain Signature Response<br>(Hedges g) | 2           |

Studies were weighted according to inverse variance. Total z = -2.19(P = .03); heterogeneity  $\chi_{19}^2 = 19.15$ (P = .45);  $\tau^2 = 0.00$ ;  $l^2 = 0.77\%$ .

NPS downregulation (Pearson r = 0.23 [95% CI, 0.13 to 0.33]; P < .001; eFigure 8 in the Supplement). While statistically significant, the effect was in the small to moderate range, suggesting that placebo analgesia is weakly associated with NPS downregulation.

# Comparing Placebo Effects on the NPS With Effect of Reduced Stimulus Intensity

These findings suggest that the effects of placebo are small in terms of effective changes in nociceptive input. To further quantify placebo effects on the NPS in terms of equivalent changes in noxious stimulus intensity, we compared placebo effects with effects of fixed increases in physical stimulus intensity. For this purpose, we used a convenience sample of 3 available out-of-sample studies<sup>67-69</sup> that tested the effects of fixed increments in noxious heat (steps of 0.4°C in a range from 45.9°C to 47.5°C<sup>67</sup> and steps of 1°C in a range from 46°C to 48°C<sup>69</sup>) or noxious pressure (reduction of 1.5 kg/cm<sup>2</sup> from a baseline of 6.0 kg/cm<sup>2</sup>).<sup>68</sup> These independent studies were used because stimulus intensities in our meta-analysis sample varied across individuals and produced behavioral effects several times larger than those of placebo, making estimation of equivalent stimulus intensity differences unreliable (for example, compare Figure 2 and eFigure 5 in the Supplement). Effect sizes (g) for stimulus intensity on pain ratings were -0.63 (95% CI, -0.68 to -0.58) at -0.4°C, -1.10 (95% CI, -1.30 to -0.90] at -1.0°C, and -1.24 (95% CI, -1.66 to -0.81) at -1.5 kg/ cm<sup>2</sup> (eFigure 9 in the Supplement), compared with -0.66 (95% CI, -0.80 to -0.53) for placebo.

Placebo effects on NPS responses in the full metaanalysis sample were 5.0 and 8.9 times smaller than the effect of reducing heat by  $-0.4^{\circ}C^{67}$  and  $-1.0^{\circ}C$ , <sup>69</sup> respectively, and 10.9 times smaller than the effect of reducing painful nail bed pressure by  $-1.5 \text{ kg/cm}^{2.68}$  (eFigure 9 in the Supplement), corresponding to an effective reduction in noxious heat of approximately 0.1°C and in noxious pressure of approximately 0.14 kg/cm<sup>2</sup>. Thus, placebo effects on the NPS are small compared with modest reductions in stimulus intensity, despite similar effects on pain ratings.

#### **Comparing Placebo and Remifentanil Effects on the NPS**

Two of the studies in the sample used analgesic, nonsedative doses of the  $\mu$ -opioid agonist remifentanil<sup>49,52</sup> at similar doses (with brain remifentanil concentrations of 0.76 ng/mL<sup>52</sup> and 0.80 ng/mL<sup>49</sup>). This allowed comparison of the effects of opioids and placebo treatment on NPS responses within these studies. Remifentanil (g = -0.77 [95% CI, -1.39 to -0.18]) and placebo treatments (g = -0.79 [95% CI, -1.39 to -0.18]) induced comparable analgesic effects at the behavioral level. However, the effects of remifentanil on the NPS (-1.10 [95% CI, -1.44 to -0.76] g) were about 10 times larger than the mean effect of placebo treatments in these studies (g = -0.11 [95% CI, -0.38 to 0.16]; eFigure 10 in the Supplement). These estimates indicate that the effects of placebo treatment on NPS responses are small compared with the analgesic effects of opioids, despite comparable behavioral effects.

## Discussion

This large-scale meta-analysis of participant-level data revealed that placebo treatments have moderate effects on subjective reports of pain, but minimal effects on responses in the NPS, a central nervous system marker that tracks the intensity of nociceptive pain. These findings are based on most of the neuroimaging data on placebo analgesia published in the field until 2015. Results were consistent across a variety of paininduction and placebo-induction methods and across 3 parallel analyses varying in risk of bias, including an analysis limited to placebo responders.

These results extend our understanding of placebo analgesia by suggesting that the effects of placebos on cerebral pain stimulus intensity processing are limited. We did observe small reductions in NPS response, which scaled with individual analgesia. Such effects may reflect descending inhibition of nociceptive systems, consistent with that findings of placebo, nocebo, and cognitive effects on spinal functional magnetic resonance imaging signals and brainstem nuclei involved in descending modulation that have been reported in previous studies.<sup>15,17-19,70</sup> However, the very small size of the effects on the NPS argues against a strong and pervasive early influence and point to stronger influences of other systems independent of the NPS. Our results emphasize that placebo analgesia is a phenomenon not based on a single mechanism but rather multiple mechanisms that have yet to be fully understood.

Importantly, the NPS is not a complete model of pain and pain-associated functionality and was not intended as such.33 Lack of effects on the NPS does not imply that placebo effects do not influence pain perception or pain-associated behavior. Indeed, the insensitivity of the NPS to manipulations that affect reported pain<sup>35-39</sup> implies that there must be other processes that contribute to pain report. However, the high sensitivity of the NPS to variation in nociceptive input (details may be found in Figure 1; eFigure 5 and eFigure 9 in the Supplement; and previous publications<sup>33,35-39</sup>), and its sensitivity to known analgesics (eg, the µ-opioid agonist remifentanil; eFigure 10 in the Supplement), suggest that if placebo treatments had pervasive early effects on pain processing, they should have been reflected in the NPS. They did not, leading us to infer that the placebo treatments studied here affect processes that are largely consequent to activation of nociceptive systems. Such processes include cognitive evaluation,<sup>71</sup> pain affect, pain-associated decisionmaking, and mesolimbic reward processing.72 These processes are likely important for behavior and subjective well-being in their own right, and indeed, placebo treatments can impact long-term symptom perception and functionality in clinically meaningful ways,73,74 whether they impact nociceptive pain signaling or not.

These findings also have implications for the objective assessment of treatment effects on pain-associated neurophysiology. Patients evaluate their pain within a complex set of personal and cultural factors,<sup>75</sup> which poses challenges for clinical trials that use self-reported pain as a primary outcome. Whether patients feel better is paramount for overall well-being, but it does not guarantee that a treatment impacts the intended physiological mechanisms in the brain and elsewhere in the body. Objective neurophysiological measures do not replace reported pain and well-being, but they can provide measures of pharmacodynamic efficacy on specific brain targets.<sup>8,76</sup> The present study further establishes the NPS as a brain measure that is sensitive to multiple types of evoked pain and insensitive to cognitive factors<sup>33,35-39</sup> in a large and geographically diverse sample. Because the NPS was found sensitive to opioid drugs but not placebo treatment, this may make it an appeal-

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ing target for evaluating pharmacodynamics and efficacy in early-stage clinical trials.<sup>8,76</sup> Our findings suggest that brain patterns such as the NPS can be used to assess efficacy in modulating nociceptive systems for drugs or devices that are intended to modulate nociceptive input at the peripheral or spinal level or influence descending facilitation or inhibition. As suggested by Duff et al,<sup>8</sup> the NPS and associated markers could be used to make early stop or go decisions in phase II clinical trials, before much more costly large-scale tests in phase III studies.

#### Limitations

Several caveats deserve mention. First, this meta-analysis included only studies testing experimental placebo treatments for evoked pain in healthy participants. They may not generalize to clinical pain, which likely involves a complex mix of nociceptive and extranociceptive processes.<sup>77</sup> Second, we analyzed summary images from published analyses. While this helps to ensure careful quality control and is advantageous in ensuring broad generalizability of results, it likely increases interstudy heterogeneity and reduces overall effect sizes. However, these issues are unlikely to compromise our conclusions regarding placebo effects, because we compare them with strong positive controls. Third, the present metaanalysis only covers the relevant literature until mid-2015; more recent studies were not sought because of the time demands of collecting and analyzing participant-level imaging data. Advances in data sharing and standardization

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will hopefully make it possible to perform participant-level meta-analyses more quickly in the future.

Importantly, not all placebo manipulations are likely to be equally effective, as indicated by the heterogeneity in placebo effects on pain ratings (as shown in Figure 2 and described in previous publications<sup>10,11</sup>). Although the studies tested here were fairly homogenous in terms of placebo effect on the NPS (Figure 3), there are multiple pathways to cognitive pain modulation,<sup>37,78</sup> and some may affect the NPS more strongly than others. Treatment contexts not studied here may still influence NPS responses.

## Conclusions

In sum, we have shown that placebo treatments have only small effects on a cerebral pattern tracking nociceptive pain in what is to our knowledge the largest meta-analysis of singleparticipant neuroimaging data on this topic to date. This suggests that placebo analgesia is largely mediated by networks different from those underlying the primary processing of noxious stimuli. Further studies are necessary to better understand which aspects of pain processing are affected by placebo treatments, and the significance of those processes for long-term clinical outcomes and wellbeing. This work serves as a starting point for the development of brain models that track pain-associated outcomes and other clinical and behavioral endpoints.

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