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# Pre-trial predictors of conflict response efficacy in human dorsolateral prefrontal cortex

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37	ABSTRACT
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39	The ability to perform motor actions depends, in part, on the brain's initial state, that is,
40	the ensemble firing rate pattern prior to the initiation of action. We hypothesized that the same
41	principle would apply to cognitive functions as well. To test this idea, we examined a unique set
42	of single unit data collected in human dorsolateral prefrontal (dlPFC) cortex. Data were collected
43	in a conflict task that interleaves Simon (motor-type) and Eriksen (flanker-type) conflict trials. In
44	dlPFC, variability in pre-trial firing rate predicted the ability to resolve conflict, as inferred from
45	reaction times. Ensemble patterns that predicted faster Simon reaction times overlapped slightly
46	with those predicting Erikson performance, indicating that the two conflict types are associated
47	with near-orthogonal initial states, and suggesting that there is a weak abstract or amodal conflict
48	preparatory state in this region. These codes became fully orthogonalized in the response state.
49	We interpret these results in light of the initial state hypothesis, arguing that the firing patterns in
50	dlPFC immediately preceding the start of a task predispose it for the efficient implementation of
51	cognitive action.
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#### INTRODUCTION

59 The ability to respond effectively to a stimulus can depend on the state of the brain even before the stimulus appears <sup>1–3</sup>. In other words, our responses are determined not only by the 60 61 neural activity driven by the response-driving stimulus, but by the way that activity interacts with ongoing neural activity<sup>4</sup>. In the motor system, one expression of this idea is the *initial state* 62 hypothesis, which holds that motor control involves a series of dynamical states and that 63 initiation of motor control requires a particular state <sup>5-10</sup>. Variability in performance, typically 64 65 assessed with reaction times, corresponds in part to variability in pre-trial firing rates because those reflect the response of the system relative to the optimal response-driving initial state. We 66 and others have proposed that dynamical principles relevant to the motor system may apply to 67 non-motor processes, including higher level cognitive processes <sup>11–15</sup>. We hypothesized, 68 therefore, that the ability to implement a cognitive process may likewise depend on the initial 69 70 state of the system.

71 We are particularly interested in conflict detection and resolution, a pair of complementary and relatively well-studied cognitive behaviors whose neuronal basis is 72 73 beginning to be understood <sup>16–21</sup>. Conflict typically refers to a competition between possible 74 stimuli for attention and/or action and generally evokes slower reaction times, increased error rates, and disengagement from alternative tasks <sup>22,23</sup>. This refocusing of mental resources is the 75 basis of conflict resolution and presumably occurs in response to an internal detection of conflict 76 77 and generation and propagation of a conflict signal. We hypothesized that the ability to deal 78 effectively with conflict depends in part on variability in neural processes in conflict-relevant 79 brain regions before the appearance of the conflicting stimuli.

The dorsolateral prefrontal cortex (dlPFC) is among the most studied brain regions for cognitive control, along with anterior cingulate cortex <sup>16,20,24–26</sup>. dlPFC shows systematic changes in hemodynamic response, local field potential (LFP), and firing rate in the face of conflict (ibid.). We have recently proposed that these two regions play somewhat distinct, albeit complementary roles in conflict detection and resolution <sup>16</sup>. We proposed that dlPFC is more associated with implementation, and thus potentially a closer cognitive analogue to motor areas (see also <sup>25–27</sup>).

87 Here, we examined firing rates of single neurons in dlPFC while humans performed the multi-source interference task (MSIT), a task that manipulates two different forms of conflict, a 88 motor (Simon) type and a perceptual (Eriksen flanker) type  $^{16,19}$ . We found that activity patterns 89 90 in the period preceding the start of the trial predicted reaction time in dlPFC, even after 91 regressing out prior trial reaction time and conflict type. At the individual cell level, we found evidence for a neural code for response times specific to each conflict type. These codes 92 93 overlapped slightly but significantly at the population level, indicating the presence of a weak 94 shared conflict-amodal code. These results endorse the idea that conflict resolution reflects the interaction between stimulus-driven activity and ongoing fluctuations in pre-trial activity, 95 96 support the initial state hypothesis for cognitive actions, and suggests a mechanism by which the 97 brain can respond both flexibly and efficiently to different conflict conditions. 98

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103	RESULTS
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105	Behavior
106	We examined responses of single neurons recorded in dlPFC in 9 human patients (Figure
107	1A). Task-related responses in this dataset were described in an earlier study, but pre-trial
108	responses, the focus of the present study, were not analyzed <sup>16,28</sup> . Participants performed the
109	multi-source interference task (MSIT), a task that involves two independently manipulated types
110	of conflict (Figure 1B).
111	The validity of this task as a manipulation of conflict has been demonstrated <sup>16,19,29,30</sup> . We
112	therefore only briefly summarize the evidence that the task manipulates conflict. Most
113	importantly, median reaction time in the Simon trials (1.5 sec) was significantly slower than no
114	conflict trials (1.26 sec, p=0.017, z=2.402, ranksum=8438). Likewise, reaction times in the
115	Eriksen trials (1.63 sec) was slower than in no conflict trials (p<0.001, z=4.51, ranksum=259).
116	Finally, reaction times on both-conflict trials (1.71 sec) were longer than on Simon trials
117	(p=0.006, z=2.71, ranksum=11428) although not compared to Eriksen: p=0.353, z=0.927,
118	ranksum=11134). (Note that the difference between both and Simon survives Bonferroni
119	correction). Despite the non-significant difference between both conflict trials and Eriksen-only
120	trials, when the effects of either single type of conflict (Eriksen or Simon) are averaged, the
121	effect of both types of conflict occurring together is still larger than the effect of either one
122	(p=0.012, z=2.514, ranksum=15093).







- Figure 1. Multi-source interference task (MSIT) design, recording locations, and behavioral
- results. (A) Basic task design. Participants fixate on a central cross and then see a visual cue
- 127 consisting of three numbers and has to identify the unique number with a button push.
- 128 "correct response" is the left button if the target is 1, middle if 2, right if 3. Four example cues
- are shown here, and in each case, the target is "2" and the middle button is the correct
- 130 response. This is most obvious for the first cue ("none"), where there is no conflicting
- information. In the other three examples, conflicting information makes the task more difficult.
- 132 First, incongruence between the location of the target number in the 3-digit sequence and

- 133 location of the correct button in the 3-button pad produces spatial (Simon) conflict (orange).
- 134 Second, the distracting presence of numbers that are valid button choices ("1", "2", "3")
- 135 produces flanker (Eriksen) conflict (green). Trials can also simultaneously have both types
- 136 (blue). B. The visual cues are associated with one or more sensorimotor responses. Every cue
- has a correct response, meaning the button press that corresponds to the unique target. Cues
- can also have one or more distractor responses, meaning the button press that corresponds to
   task-irrelevant spatial information (Simon) or flanking distractors (Eriksen). If and only if the
- 140 correct response and distractor response do not match, then the cue causes conflict because
- 141 only one button response can ultimately be chosen. C. Diagram of the intracranial implant
- showing the UMA and tungsten microelectrode recoding locations schematized as a purple
- 143 square on the surface of dIPFC. sulcus. D. The average (mean) response times across subjects
- in each of the four task conditions and (right) the mean response times within each subject.
  Bars = standard error across subjects.
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## Pre-trial single neuron correlates of conflict

We recorded from 378 neurons from 9 patients in dlPFC. Our goal was to determine
whether responses of neurons before the start of the trial predict subsequent reaction time.
Consider, first, responses of an example neurons shown in Figure 2.

In our example neuron (**Figure 2**), taken from dlPFC, both Simon responses and Eriksen responses were significantly greater before the start of faster reaction time trials than before slower reaction time trials (fast Simon RT trials: 3.2 spikes/sec, slow Simon RT trials: 2.5 spikes/sec, p=0.005, t-test on z-scored data; fast Eriksen RT trials: 3.2 spikes/sec, slow Eriksen RT trials: 2.5 spikes/sec, p=0.008, t-test on z-scored data). Specifically, we ran a median split on reaction times post-hoc and separated firing rates on those two categories. As described below, firing rate in this neuron also predicted the reaction time in a continuous model.

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## Pre-trial population correlates of conflict

160 To explore these effects at the population level, we fit generalized linear models (GLMs) 161 to the pretrial firing rates and reaction times for all trials. Our analyses controlled for prior trial 162 conflict type, because conflict level on the previous trial can modulate preparatory neural 163 responses and lead to trial-to-trial adjustments, such as post-error slowing and conflict-164 adaptation/trial congruency effects <sup>19,31,32</sup>. Our analyses also controlled for previous trial reaction 165 time (RT), to remove possible effects of slow drifts in arousal. We analyzed the 500ms epoch 166 between the fixation and cue onset in the current trial.

167 We first asked whether pre-trial activity predicted reaction times on all trials without 168 regard to conflict or history. Pre-trial firing rates in 13.2% of neurons in dlPFC were predictive 169 of the RT on the upcoming trial (n = 50/378 neurons). This proportion is greater than that 170 predicted by chance (p < 0.001, one-sided binomial test) and remained the same after controlling 171 for prior trial reaction time and conflict type. These results indicate that a small but statistically 172 significant fraction of neurons have firing rates that predict upcoming reaction times.

After establishing that pretrial activity predicts reaction time overall, we next examined
whether it modulated reaction times differently depending on the upcoming conflict condition.
We compared a model with a single parameter for any conflict type ("conflict type-amodal",
valued at 1 for any conflict type and 0 for no conflict) to a model with separate parameters for

177 Eriksen and Simon conflict types ("conflict type-specific"), including trials in which the

distractor stimuli appeared separately as well as together, and controlling for previous trial
conflict and RT. We also compared these models to the "no-conflict" model described above. In
both areas, we found significant proportions of cells with a significant main effect of firing rate
on response time.

At the single cell level, we found evidence for both conflict type-amodal tuning and 182 183 conflict type-specific tuning. In the conflict type-amodal model, 7.7% of cells exhibited 184 significant single conflict coefficients, a proportion that was significantly greater than chance (n 185 = 29/378, p = 0.009, binomial test with chance rate of 5%) and 11.1% of cells showed a main 186 effect (Wald test) of firing rate (p < 0.001, one sided binomial test with chance rate of 5%). With the conflict type-specific model, 14.6 % of all cells (n=55/378) exhibited significant coefficients 187 188 (Wald tests) for predicting RT on either Eriksen trials or on Simon Trials (p < 0.001, binomial 189 test with chance rate of 9.75%). Examining the main effect of overall firing rate in this model, 13.2% of cells showed a significant effect (n = 50/378, p < 0.001, binomial test). 190

When we examined the evidence for conflict tuning in dIPFC with model comparison, we 191 found that 73% of cells preferred the conflict-specific model (n = 277/378, median model 3:1 192 193 BIC weight ratio = 23, median model 3:2 BIC weight ratio = 40), 22% preferred the no-conflict 194 model (n = 83/378, median model 1:3 BIC weight ratio = 2.9, median model 1:2 BIC weight 195 ratio = 44) and only 3% preferred the conflict type-amodal model (n = 10/378, median model 2:3 196 BIC weight ratio = 2.8, median model 2:1 BIC weight ratio =  $2.4 \times 10^6$ ). Of the 29 cells with 197 significant conflict type-amodal coefficients, 93% (n = 27/29) preferred the conflict type-specific model. These BIC results indicate that cells tuned for conflict type-specific responding dominate 198 199 in dlPFC.

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Figure 2. Individual dlPFC neurons signal the speed of upcoming responses in a conflict-specific
 manner. (A-B) PSTHs of example neuron 369 showing significantly higher pre-trial activity
 before fast (red) responses than slow (blue) responses in both (A) Eriksen and (B) Simon conflict
 trials. Gray shading indicates the 500ms analysis window. The dotted line indicates the response
 on the last trial and the solid vertical line indicates the stimulus onset.

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#### Semi-orthogonal coding for two forms of conflict preparation in dlPFC

211 Our observation that a significant portion of individual cells code for conflict type-212 amodal reaction times but yet were mostly better described with the conflict type-specific model 213 led us to wonder if *population-level* activity might contain a conflict type-amodal signal. To test 214 this possibility, we asked how ensemble codes for predicting resolution of Simon and Eriksen 215 conflict, derived from the conflict type-specific model, were related to each other. To do this, we calculated a vector of GLM regression weights for each distractor type across neurons. We call 216 217 these vectors the *pre-trial tuning weight vectors*. We then compared these vectors by computing 218 the Spearman correlation of the vectors corresponding to Eriksen and Simon coefficients (cf. <sup>33</sup>). 219 (Note that the Spearman test does not assume linearity and is thus more general than the more 220 common Pearson and is also less sensitive to potential outliers). We tested these analyses on all 221 cells that preferred either conflict type-amodal or conflict type-specific models by BIC (n =222 287/738) as well as on all cells; doing so allows us to include contributions from all relevant 223 neurons, even those with real effects that do not pass the strict significance threshold; this 224 approach thus has better signal-to-noise (and moderately less susceptible to Type II errors) than 225 analysis approaches that focus on cells that cross a significance threshold (Blanchard et al., 226 2018).

227 We found that in the conflict-preferring population, the codes for Simon and Eriksen 228 were weakly positively correlated ( $\rho = 0.12$ , p = 0.048; permutation test). Relatedly, we found 229 the angle  $\theta$  between these vectors to be 83° and significantly less than 90° (p = 0.017, 230 permutation test). This result is consistent with the population-level superposition of collinear 231 and orthogonal coding for optimal initial conditions for conflict responding. In the entire 232 population, the angle  $\theta$  between these vectors was 86° and less than 90° at a trend level (p =233 0.079, permutation test), and the Spearman's  $\rho$  was 0.093 (p = 0.038, permutation test). In other 234 words, this result indicates that there is a conflict type-amodal code recoverable from dlPFC 235 ensemble activity.

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## Firing rates in the response period explain response time variability

238 If neural firing patterns in the preparatory period bias response speed in a conflict type-239 specific manner, we would expect the neural state in the response period to reflect this. To test 240 this hypothesis, we ran the conflict type-specific model described above on the 500ms period 241 after stimulus onset. We found that 16.9 % of all cells (n = 64/378) exhibited significant 242 coefficients (Wald Test) for predicting RT on either Eriksen trials or Simon Trials (p < 0.001, 243 binomial test with chance rate of 9.75%). To examine the degree to which the preparatory state 244 resembled the response state, we computed similarity measures for their respective coding 245 vectors, which we call the *response tuning weight vectors*. For both areas, we found that the pre-246 trial and response tuning weight vectors were partially co-linear. The Spearman correlation 247 between Eriksen codes for conflict-preferring cells was r = 0.22 (p < 0.001, permutation test) and between Simon codes was  $\rho = 0.23$  (p < 0.001, permutation test); the angles were 77° for the 248 249 Eriksen (p < 0.001, permutation test) and 78° for the Simon condition (p < 0.001, permutation

test). The results for all cells were similar (Eriksen  $\rho = 0.22$ , p < 0.001 and  $\theta = 78^{\circ}$ , p < 0.001; 250 251 Simon  $\rho = 0.23$ , p < 0.001 and  $\theta = 75^{\circ}$ , p < 0.001). These results support the notion that the 252 neural states in dlPFC that predispose for more efficient conflict resolution overlap with the states of efficient responding themselves. Next, we queried the association between the Eriksen 253 254 and Simon response tuning weight vectors. We found that the codes for both the conflict-255 preferring and all-cell populations were fully orthogonalized in the response period, in contrast to the pre-trial period (conflict-preferring cells  $\rho = -0.01$ , p = 0.882 and  $\theta = 90^{\circ}$ , p > 0.05; all 256 cells  $\rho = -0.04$ , p > 0.05 and  $\theta = 92^{\circ}$ , p > 0.05, permutation tests). To determine whether the 257 pre-trial and response correlation coefficients and angles differed, we computed bootstrap 258 distributions for each and compared the medians. The correlation coefficient for Eriksen versus 259 Simon pre-trial tuning weight vectors differed significantly from those for the response tuning 260 weight vectors for both the conflict-preferring population and all cells (Wilcoxon rank sums, all 261 p < 0.001). This is important to confirm because the difference between a significant and non-262 263 significant effect is not necessarily itself significant <sup>34</sup>. 264



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**Figure 3.** The angle between conflict-response tuning weight vectors in dlPFC orthogonalizes

between the pre-trial and response period. A zero-degree angle represents complete collinearity.

269 (A-D) the angle between Eriksen and Simon coding vectors (blue circle, blue numbering and red

270 line). The significance of the difference from a  $90^{\circ}$  (blue italics) was computed from permutation

- 271 testing. The null distribution is shown with grey circles. (A) Vector angle between conflict
- 272 preferring cells (amodal or conflict-type specific) showing superimposed collinear and
- 273 orthogonal coding. (B) The angle between all cells shows a similar trend. (C) Same as (A) for
- 274 response codes, showing purely orthogonal coding. (D) Same as for (B) for response codes,
- 275 showing orthogonal coding.

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Figure 4. The angle between pre-trial and response conflict-response tuning weight vectors in
dlPFC shows partially co-linear coding in dlPFC. A zero-degree angle represents complete
collinearity. (A-B) Conflict-preferring coding vector angles (blue circle, blue numbering and red
line); the significance of the difference from a 90° (blue italics) was computed from permutation
testing (the null distribution is shown with grey circles). (A) the angle between pre-trial and
response Eriksen coding vectors shows significant collinearity as does (B) the angle between
pre-trial and response Simon coding vectors.

#### DISCUSSION

286 We examined the responses of single neurons in human dlPFC during a task that 287 interleaved two kinds of conflict, Simon and Eriksen. We found that firing rates of a modest but 288 significant proportion of neurons in both areas before the trial predicts the efficiency of cognitive 289 control, as inferred from reaction times. The fact that ensemble responses predict reaction time 290 before the trial, and presumably task-driven cognition, supports the hypothesis that successful cognitive control reflects, in part, the ability to transition through specific brain states. By 291 292 controlling for prior trial condition and reaction time, we showed that these brains states do not 293 simply reflect adaptation or drifts in arousal, but rather history-independent patterns that 294 predispose to efficient responding.

295 We also compared the patterns that predicted the efficiency of upcoming responses on 296 trials with either Simon, Eriksen or both types of conflict. At the individual cell level, we found 297 that these responses were consistent with three types of codes: a conflict-independent (no-298 conflict), a conflict type-amodal and conflict type-specific code. At the population level, we 299 found a weak general conflict type-amodal code. These results demonstrate, first, that neurons in 300 dlPFC process both type-specific and domain-general neural computations. While the domain-301 general element was observed both at the individual cell level and the population level, model 302 comparison suggests it is more likely a population level phenomenon: in most cells the full model with separate conflict terms significantly outperformed a model with one single term for 303 any type of conflict. A sizable proportion of individual cells (22%) did prefer a model with no 304 305 conflict term, predicting response speed across all conditions. However, the domain-306 general/conflict type-amodal element emerged at the population level from the correlation 307 between the conflict-specific regression coefficients across neurons.

308 Studies of motor control in rodents and non-human primates have shown that variability 309 in the firing rates of neurons in motor cortex during movement preparation predicts variability in subsequent movements <sup>6,9,10</sup>. This body of work has given rise to the "initial condition 310 311 hypothesis", which posits that neural firing patterns behave like dynamical systems, and the trajectory of neural dynamics therefore depends on the initial state of the system <sup>3,8,35</sup>. One 312 313 implication of this theory is that preparatory neural states can therefore be optimized to 314 efficiently produce a desired behavioral outcome <sup>36</sup>. If similar principles apply to neural 315 dynamics during cognitive tasks, this suggests that preparatory activity may also be optimized to support the efficient application of cognitive control, consistent with the notion of pro-active 316 317 control <sup>37</sup>.

Indeed, a growing literature suggests neural patterns subserving cognition are also 318 319 consistent with dynamical systems models <sup>38</sup>. Using the same task studied, here, we recently 320 showed that neural population activity in prefrontal regions during conflict resolution follows 321 low-dimensional trajectories that differ depending on the type of conflict <sup>28</sup>. The divergence of 322 these trajectories raises the question of whether they have different optimal initial states, or if 323 similar neural starting conditions give rise to similar task performance. The results we present 324 provide a somewhat nuanced answer: the optimal initial states overlap slightly, suggesting a 325 weak mechanism for shared conflict-responding, but orthogonalize over the course of response, 326 supporting the notion of trajectory divergence.

327 What are the implications of preparatory brain states for conflict resolution? One 328 possibility is that initial states that predict response times reflect changes in arousal or attention 329 that are either spontaneous or the result of adaptations to the previous trial. While attractive for 330 its simplicity, this explanation would not explain the persistence of the effect after controlling for 331 trial history nor would it explain the conflict type-specific nature of the initial states. Rather, our 332 results suggest that the computations that perform conflict resolution are not only distinct, as 333 implied by the divergent neural state space trajectories, but are facilitated with both shared and 334 independent factors. In this task, participants do not have knowledge of whether the upcoming 335 trial will have one form of conflict or the other or both. Maintaining both shared and independent 336 preparatory factors for conflict types, independent of the past, may allow the brain to respond 337 flexibly to unpredictable challenges while minimizing interference between the processes needed 338 to respond to those challenges. Such a factorized model of representation has been proposed as a 339 mechanism to minimize interference and facilitate generalization in learning <sup>39</sup>. Why would 340 preparatory states vary from trial to trial? The simplest explanation is spontaneous fluctuations 341 due to noise. Another possibility is that participants are subtly making predictions about the 342 condition of the upcoming trial, perhaps based on longer trial history than the recent past 343 controlled for here.

344 Recent years have seen the emergence of the dynamical systems perspective in motor 345 neurophysiology. This view sees the aggregate activity of neurons in a region as constituting a 346 state and that the execution of motor actions is driven by the lawful progression across states in 347 motor regions. It has been further proposed that cognitive performance may reflect a similar 348 dynamical system view, however, this view has been difficult to test. We recently demonstrated 349 that, in asynchronous choice, neurons in two core reward areas show subspace orthogonalization, a neural process previously associated only with motor cortex <sup>40</sup>. Here we sought to test a critical 350 prediction of the initial state hypothesis. In the motor cortex, the *initial state hypothesis* holds 351 352 that successful implementation of a motor actions cannot begin until motor cortex enters into a 353 specific ensemble state, defined in practice as an ensemble firing rate pattern. We hypothesized

354 that an analogous idea in cognition would be that implementation of a cognitive act would 355 require implementation of an initial state. Our results suggest that not only does the initial state of neural activity prior to the cognitive act of conflict resolution support efficient responding in a 356 357 pro-active manner <sup>37</sup>, it does so in a largely conflict type-specific manner. The small degree of collinearity we observe in dlPFC may contribute to untangling of stimulus-action processes 358 359 before full orthogonalization during implementation. Previous work in premotor cortex has 360 shown that the neural state space during responding predicts reaction times in motor tasks <sup>41</sup>. We 361 found that these optimal initial states share structure with the optimal responding states, suggesting they support optimal response trajectories. Future studies may examine whether 362 363 neural state-space trajectories separate by response time differently depending on both the initial 364 states and the specific cognitive computations the brain performs. 365

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# **METHODS**

371372 Subjects and ethics statement

373 374 We studied one cohort of 9 patients: 8 (2 female) with movement disorders (Parkinson's disease or essential tremor) who were undergoing deep brain stimulation (DBS) surgery, and one male 375 376 patient with epilepsy undergoing intracranial seizure monitoring. The entry point for the 377 trajectory of the DBS electrode is typically in the inferior portion of the superior frontal gyrus or superior portion of the middle frontal gyrus, within 2 cm of the coronal suture. This area 378 379 corresponds to dIPFC (Brodmann's areas 9 and 46). The single epilepsy patient in this cohort 380 underwent a craniotomy for placement of subdural grid/strip electrodes in a prefrontal area 381 including dlPFC.

- 382 All decisions regarding sEEG and DBS trajectories and craniotomy location were made solely
- 383based on clinical criteria. The Columbia University Medical Center Institutional Review Board
- 384 approved these experiments, and all subjects provided informed consent prior to participating in
- the study.

- 386 Behavioral Task
- 387 All subjects performed the multi-source interference task (MSIT; Figure 1A). In this task, each
- trial began with a 500-millisecond fixation period. This was followed by a cue indicating
- the *correct response* as well as the *distractor response*. The cue consisted of three integers
- drawn from {0, 1, 2, 3}. One of these three numbers (the "*correct response cue*") was different
  from the other two numbers (the "*distractor response cues*"). Subjects were instructed to
- indicate the identity of the correct response number on a 3-button pad. The three buttons on this
- 393 pad corresponded to the numbers 1 (left button), 2 (middle) and 3 (right), respectively.
- 394 The MSIT task therefore presented two types of conflict. Simon (motor spatial) conflict occurred
- if the correct response cue was located in a different position in the cue than the corresponding
- position on the 3-button pad (e.g. '0 0 1'; target in right position, but left button is correct
- 397 choice). Eriksen (flanker) conflict occurred if the distractor numbers were possible button
- 398 choices (e.g. '3 2 3', in which "3" corresponds to a possible button choice; vs. '0 2 0', in which
- 399 "0" does not correspond to a possible button choice).
- 400 After each subject registered his or her response, the cue disappeared, and feedback appeared.
- 401 The feedback consisted of the target number, but it appeared in a different color. The duration of
- 402 the feedback was variable (300 to 800 milliseconds, drawn from a uniform distribution therein).
- 403 The inter-trial interval varied uniformly randomly between 1 and 1.5 seconds.
- 404 The task was presented on a computer monitor controlled by the Psychophysics Matlab Toolbox
- 405 (<u>www.psychtoolbox.org</u>; The MathWorks, Inc). This software interfaced with data acquisition
- 406 cards (National Instruments,) that allowed for synchronization of behavioral events and neural
- 407 data with sub-millisecond precision.

# 408 Data Acquisition and preprocessing

409 Single unit activity (SUA) was recorded from a combination of two techniques. The DBS

410 surgeries were performed according to standard clinical procedure, using clinical microelectrode

411 recording (Frederick Haer Corp.). Prior to inserting the guide tubes for the clinical recordings,

- 412 we placed the microelectrodes in the cortex under direct vision to record from dlPFC, (IRB-
- 413 AAAK2104). The epilepsy implant in Cohort 2 included a Utah-style microelectrode array
- (UMA) implanted in dlPFC (IRB-AAAB6324). In all cases, data were amplified, high-pass
   filtered, and digitized at 30 kilosamples per second on a neural signal processor (Blackrock
- 416 Microsystems, LLC).
- 417 SUA data were re-thresholded offline at negative four times the root mean square of the 250 Hz
- 418 high-pass filtered signal. Well-isolated action potential waveforms were then segregated in a
- 419 semi-supervised manner using the T-distribution expectation-maximization method on a feature
- 420 space comprised of the first three principal components using Offline Sorter (OLS) software
- 421 (Plexon Inc, Dallas, TX; USA). The times of threshold crossing for identified single units were
- 422 retained for further analysis.
- 423 Data Analysis
- 424

425 We determined the effect variations in pre-stimulus firing rates had on reaction times by

- 426 comparing four generalized linear models. We first fit gamma distributions to the reaction times
- 427 and excluded reaction times with a less than 0.005 probability following  $^{42}$ . For each model, we
- 428 centered and scaled the continuous predictor variables (firing rate and reaction time) by z-
- scoring. We analyzed correct and incorrect trials to prevent false-positives from data-censoring
- 430 effects. We pre-selected the pre-trial analysis interval as the 500ms period between fixation and
- 431 stimulus onset and the response analysis interval as the 500ms following stimulus onset. To
- determine the overall effect of firing rate marginalized over condition, we first fit the following
- 433 generalized linear model:

$$RT \sim \beta_0 + \beta_1 \frac{FR}{436}$$

We then compared several alternative models, while controlling for reaction time and priorconflict: a single firing rate coefficient for all trials model,

$$RT \sim \beta_0 + \beta_1 FR + pC + pRT$$

a model with an additional term for the firing rate on trials with any conflict (Simon, Eriksen orboth),

$$RT \sim \beta_0 + \beta_1 FR + \beta_2 FR * C + C + pC + \dot{p}\hat{R}T$$
441
441
444

and a model with separate, additive conflict terms for Simon and Eriksen conditions,

446 447  $RT \sim \beta_0 + \beta_1 FR + \beta_2 FR * C_E + \beta_3 FR * C_S + C_E + C_S + pC + pRT$  bioRxiv preprint doi: https://doi.org/10.1101/2021.07.07.451322; this version posted July 9, 2021. The copyright holder for this preprint (which was not certified by peer review) is the author/funder. All rights reserved. No reuse allowed without permission.

#### 448

where FR is firing rate on all trials, C is an indicator variable for trials with any conflict, CE and
CS are indicator variables for Eriksen and Simon trials, respectively, pC is a categorical variable
for prior conflict type and pRT is the previous trial reaction time. We used a normal distribution
with a log link function because the reaction time data was well described by a log-normal
distribution.

454

We compared these models by their BIC weight ratios <sup>43</sup>. We identified the best fitting model for each cell as the model with a greater than one BIC weight ratio for all pairwise comparisons. The BIC penalty for model complexity is greater than that for Akaike Information Criteria (AIC) as the number of model parameters exceeds  $e^2 \sim 7$  and thus more appropriate here (and more conservative) <sup>43,44</sup>. To assess the significance of overall model fits we performed deviance tests relative to a constant model. To assess the significance of coefficients we performed Wald tests (REF).

462

463 To compare the neural codes for conflict, we entered the regression coefficients into individual 464 vectors for each conflict condition. We then computed the Spearman correlation between those 465 vectors as well as the angle between the vectors. We excluded points more than 3 median 466 absolute deviations from the median because angle measurements; this approach excluded 6 467 cells. We then randomized the vector entries and computed correlations between the randomized 468 vectors to form null distributions (2000 permutations). We computed p-values for the real measurement relative to the corresponding null distribution (permutation test). To compare 469 470 angles and correlations, we constructed bootstrap distributions with resampling (5000 samples) 471 and compared the medians of the resultant distributions with the non-parametric Wilcoxon rank 472 sum test.

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