Volume 20, Number 1, 2010 © Mary Ann Liebert, Inc.

Pp. 39-47 DOI: 10.1089/cap.2009.0047

Predictors of Placebo Response in Randomized Controlled Trials of Psychotropic Drugs for Children and Adolescents with Internalizing Disorders

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Abstract

Objective: The aim of this study was to assess predictors of placebo response in all available short-term, placebo-controlled trials of psychotropic drugs for children and adolescents with internalizing disorders, major depressive disorder (MDD), obsessive compulsive disorder (OCD,) and anxiety disorders (ANX) exclusive of OCD and posttraumatic stress disorder (PTSD).

Method: We reviewed the literature relevant to the use of psychotropic medication in children and adolescents with internalizing disorders, restricting our review to double-blind studies including a placebo arm. Placebo response, defined according to each trial's primary response outcome variable and Clinical Global Impressions–Improvement, when available, and potential predictive variables were extracted from 40 studies.

Results: From 1972 to 2007, we found 23 trials that evaluated the efficacy of psychotropic medication involving youth with MDD, 7 pertaining to youths with OCD, and 10 pertaining to youths with ANX (N = 2,533 patients in placebo arms). For all internalizing disorders combined, predictors of nonresponse to placebo were the percentage of Caucasian patients included in the study and the duration of the disorder: Both variables were negatively correlated with the percent of placebo responders. The type of disorder was found to predict the robustness of placebo response: (OCD < ANX < MDD). For a subset of MDD studies, we found that baseline illness severity tended to be negatively correlated with placebo response. Finally, trial "success" was significantly associated with lower placebo response rate.

Conclusion: Predictors of placebo response in internalizing disorders of youths parallel those in adult studies, with the exception of race. These predictors should be considered when designing placebo-controlled trials in youths to enhance findings of true drug-placebo differences.

Introduction

Internalizing disorders are a public health concern because of their frequency and morbidity; child and adolescent depression, in particular, is a concern because of its implication for suicidal acts and youth mortality (Flament et al. 2001; Birmaher et al. 2007). During the 1980s, the arrival of the selective serotonin reuptake inhibitors (SSRIs), which are safer in overdose and associated with fewer side effects than tricyclic antidepressants (TCAs) or monoamine oxidase inhibitors, was viewed as an important step in the treatment of affective and anxiety disorders, first in adults, and then in children and adolescents. In the early 1990s, SSRI use in

children and adolescents increased rapidly in developed countries, sometimes to a higher degree than the prevalence rate of depression or anxiety disorder in this age range (Brophy 1995). Prior to the introduction of SSRIs, placebo-controlled trials in youths rarely demonstrated the superiority of TCAs over placebo, owing in part to high placebo response rates, which made the establishment of drug efficacy difficult. The problem of high placebo response rate appears in many SSRI trials in youth major depressive disorder (MDD) trials. In fact, the evidence of SSRI efficacy has been more robustly demonstrated in child and adolescent obsessive compulsive disorder (OCD) (Flament et al. 2000; Geller et al. 2003) and anxiety disorders (ANX) than in MDD (Bridge et al. 2007).

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Dr. Donnelly has received research funding from the National Institute for Mental Health (NIMH) (STAART [Citalopram for Children With Autism and Repetitive Behavior] study #U54 MH066398-04).

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Recently, there has been a substantial increase in interest in placebo response rates in children participating in clinical trials (Rohde et al. 2008; Bridge et al. 2009; Emslie 2009). The focus on placebo response in youth medication trials for MDD has important implications regarding future trial design and selection criteria for study subjects and also has the potential for informing clinical practice decisions regarding when to use medication and in what types of clinical scenarios and patient presentations. We have sought to identify characteristics that may contribute to the placebo response in youths treated with medication for several internalizing disorders in addition to MDD. Previously, we hypothesized that the response to placebo was high in child and adolescent depression because of specific psychopathological factors associated with MDD in youths that promote psychotherapeutic effects in pharmacological trials (Cohen 2007). To test whether response to placebo was different according to type of disorder, we performed a pooled analysis of all randomized double-blind, placebo-controlled studies that have been conducted in the last 25 years on internalizing disorders and reported that the placebo response rates in pharmacological trials for MDD were higher than in those for OCD and other ANX (Cohen et al. 2008).

Identification of other potential predictors of the placebo response in children and adolescents with internalizing disorders could be helpful both in terms of psychopharmacological research, because the placebo response rate is key in detecting drug versus placebo differences within a trial, and in terms of clinical practice, to better detect those children and adolescents that would "truly" benefit from medication. Indeed, in adults, an analysis of the Food and Drug Administration (FDA) Summary Basis of Approval reports for 11 approved antidepressants for MDD between 1985 and 2005 found that the magnitude of placebo response was the single most powerful variable associated with the outcome of an antidepressant trial (Khan et al. 2003). Lower placebo response rates were more likely to result in "positive" treatment trials. For studies with placebo response rate over 30%, the occurrence of significant separation from an antidepressant was 1 in 5; in contrast, for studies with lower placebo response rates, the occurrence of a statistically significance effect favoring drug over placebo was close to 3 in 4.

Placebo response in youths with internalizing disorders has only recently received attention in the literature. Mayes et al. (2007) reported that response to placebo in trials of fluoxetine was more pronounced in children than in adolescents. Bridge et al. (2009) reported in a review of 12 published and unpublished placebo-controlled trials of second-generation antidepressants for child and adolescent depression that the number of study sites was the single best predictor of placebo response. The more study sites, the higher was the placebo response. Baseline severity of illness also emerged as a significant inverse predictor of placebo response. However, to keep the same criterion of response across trials, i.e., Clinical Global Impressions–Improvement (CGI-I), this review did not include all available studies.

In adults, placebo response has received greater attention (Enck et al. 2008), particularly in the context of MDD studies finding a high placebo response rate both in pharmacological and non-pharmacological trials (Brunoni et al. 2009). Adult population predictors of placebo response include younger age, male sex, longer trial duration, fixed-dose protocol (versus flexible-dose), shorter duration of illness, lower baseline severity, fewer prior episodes, study location, and time of research publication (Khan et al. 1991; Wilcox et al. 1992; Charney et al. 2002; Walsh et al. 2002; Stein et al. 2006; Khan et al. 2007; Posternak et al. 2007; Kirsch et al. 2008). Surprisingly, placebo lead-in strategies did not

lower subsequent placebo response rates (Trivedi and Trush 1994). This may be a consequence of short intervention lead-in periods. It remains to be seen if alternate lead-in strategies, such as longer periods of placebo therapy or using nonresponse to an initial course of psychotherapy, diminish subsequent placebo response after pharmacological trial inclusion (Gelenberg et al. 2008).

In this article, we assess characteristics and predictors (classified as sociodemographic, methodological, psychopathological, and other) of placebo response in all available short-term, randomized, double-blind, placebo-controlled trials of psychotropic drugs for child and adolescent internalizing disorders, MDD, OCD, and other ANX, exclusive of OCD and posttraumatic stress disorder (PTSD).

Method

Search and study selection

Search strategy and study selection methods have been detailed elsewhere (Cohen et al. 2008). We searched the Medline database for articles describing randomized, placebo-controlled trials of medication for children and adolescents diagnosed with MDD, OCD, and ANX. Searches included combinations of the following keywords: "Major depression," "obsessive compulsive disorder," and "anxiety disorders" and/or "children/adolescents" and/or "placebo-controlled." In addition, references from identified articles and reviews on the same conditions were examined. In particular, recent meta-analyses and reviews that explored the relationship between non-TCAs and suicide were of particular interest, given that some included unpublished trials and detailed response rates according to age (Kapczinski et al. 2003; Ryan 2005; Jureidini et al. 2004; Whittington et al. 2004; Hammad et al. 2006; Hetrick et al. 2007; Bridge et al. 2009). Using these methods, we found 70 publications between January, 1972, and October, 2007. Exclusion criteria were: (1) Crossover design; (2) no indication of the number of responders in either the original report or the available reviews; (3) fewer than 10 individuals in the placebo arms; (4) lack of placebo arms; (5) no randomization; (6) no doubleblind evaluation of response; (7) studies on PTSD; and (8) secondary analysis of prior reported trials. Table 1 lists all of the published controlled randomized trials of psychotropic medications that were included in the current analysis. Details on excluded studies and trial flow are given in Cohen et al. (2008).

Response criteria

Given that analyses were performed across studies, we kept each trial's definition of responders as indicated in Table 1. Owing to the diversity of definition of "response" across studies, we performed a subgroup analysis that used only the CGI-I, when used in the trials as either a primary or secondary variable (e.g., responder = CGI-I \leq 2), as our response criterion.

Predictors of placebo response

We examined the following predictors of placebo response. (1) Sociodemographic variables: Age (mean age of the placebo arm); sex (% of males); race (% of caucasian patients); socioeconomic status (SES) (% of low SES). (2) Methodological characteristics of the trials: *N* of patients included in the placebo arm; type of active treatment (TCA vs. non-TCA), duration of the trial; number of study sites; number of meetings during the study protocol; Detsky et al.'s short version quality score of the trial that includes items related to randomization, blindness, inclusion/exclusion criteria, outcome measures, treatment description, and statistical analysis

(Detsky et al. 1992); use of a placebo run-in; use of a screening and/or washout period; location of the study (U.S. only vs. not U.S. only). (3) Psychopathological characteristics: Disorder type (MDD, OCD, or ANX); duration of the disorder prior to trial entrance; prior history (% of first episodes); CGI–Severity at baseline. We added two other variables: The year of publication, as it has been shown that the time of research publication, with more recent studies showing higher placebo response rates, may be important in determining the placebo response (Keck et al. 2000; Walsh et al. 2002; Sysko and Walsh 2007), and the "positive" (or negative) outcome of the difference between drug and placebo at the end of the trial (Khan et al. 2003). Of note, regarding the year of publication, we chose: (1) The first publication when trials were reported in several papers; (2) the year of FDA or drug company synopsis for unpublished trials.

Statistical analysis

Descriptive statistics were used to characterize the features of each of the included trials. The assumption of normality was assessed using the Shapiro-Wilk test. The assumption of homogeneity of variance was assessed using the Barlett test. Then, depending on the nature of each variable, we used the following method to test statistical links with response to placebo. First, an analysis of variance (ANOVA) was applied for comparison of response rates according to diagnosis (MDD, OCD, and ANX). Second, Pearson's correlations (r) were used for continuous variables that showed a normal distribution. For those that did not show a normal distribution, we examined their link using the Spearman rank correlation (ρ) . As a sensitivity analysis, we also iteratively deleted each trial and recomputed the correlation estimates to assess the stability of the results. Finally, two-group comparisons were made using Student t-tests. Secondary analyses following the same methodology were conducted on the subsample of studies that used CGI-I as response criterion (N=29), on the subsample of studies that included patients with MDD (N=23).

Results

Among the 70 studies (N = 5,894), we found 23 trials that met our inclusion criteria and evaluated the efficacy of psychotropic medication (mainly non-TCAs) involving youths with MDD, 7 on youths with OCD, and 10 on youths with other ANX. Altogether, the studies included 2,533 patients in placebo arms (n = 1,528 in MDD, n = 371 in OCD, n = 634 in ANX). Table 1 summarizes study characteristics, including potential predictive variables. Our goal to include SES in the potential predictors was unsuccessful owing to both the absence of inclusion in many trials and the diversity of reporting methods in others.

Regarding placebo response rates according to diagnosis, the assumption of homogeneity of variance was assessed with the Barlett test, yielding a p value of 0.53. The assumption of normality was assessed with the Shapiro–Wilk test, which yielded p values of 0.65 (MDD), 0.77 (OCD), and 0.3 (ANX), respectively. Given that it was not possible to reject the hypothesis of the normality of residuals, we tested each potential predictor as indicated in Table 2, depending on the type of variable (continuous vs. dichotomous) and its distribution (normal or not) (see Methods). The same analysis was conducted on the subgroup of studies (n = 29) that used the CGI to indicate responder status. All results are shown in Table 2.

Among sociodemographic characteristics, the percentage of Caucasian patients included in the studies was significantly negatively correlated with the percent of placebo responders whatever the definition; i.e., the more Caucasian patients, the lower the

placebo response. To assess whether one single trial strongly influenced this finding, we performed a sensitivity analysis by iteratively deleting each trial and recomputing the correlation estimates. The correlation remained significant with a Spearman correlation coefficient ρ ranging from -0.43 to -0.55 $(0.003 \leq p \leq 0.014)$, depending on the excluded trial. The percentage of males tended to be negatively correlated with the percentage of placebo responders defined with CGI, i.e., the more male patients, the lower the placebo response.

Among the methodological characteristics, the use of a washout or screening period is the only variable that was significantly associated with the CGI placebo response. When a washout/screening period occurred prior to entry into the trial, the placebo response was lower (33.44% \pm 13.5% versus 45.8% \pm 10.1% on CGI-I placebo response of 1 or 2, respectively). However, the difference did not reach statistical significance when the definition of responder using other criteria was the primary variable (38.8% \pm 20.1% versus 44.3% \pm 15.7%).

Among the psychopathological characteristics, two variables were strong predictors of response: The type of disorder and the duration of the index episode of illness prior to trial entrance. Placebo response was lower in OCD than ANX and MDD trials (31% [range, 4–41%] vs. 39.6% [range, 9–53] vs. 49.6% [range, 17–90%], respectively). Disorder duration was significantly negatively correlated with the percent of placebo responders whatever the definition, i.e., the longer the duration of illness, the lower the placebo response. The sensitivity analysis showed that the correlation remained significant with a Pearson correlation coefficient r ranging from -0.42 to -0.56 (0.007 $\leq p \leq$ 0.05) depending on the excluded trial (Fig. 1).

Superiority of the drug versus placebo on the primary outcome variable was associated with the degree of placebo response whatever the definition of response. The lowest placebo responses yielded trials showing greater superiority of the tested drugs. When the trial conclusion was that the drug effect was superior to the placebo effect, the percentage of placebo response (mean \pm SD) was 32.7 \pm 16.4 and 33.1 \pm 13.4 according to either the primary variable or the CGI-I definitions, respectively. When the trial conclusion was that drug effect failed to separate from the placebo effect, the percentage of placebo response (mean \pm SD) was 47 \pm 17.1 and 43.6 \pm 11.6, respectively. Finally, time of publication was not significantly associated with placebo response.

When the same analysis was conducted on the subgroup of studies testing pharmacological treatments in child and adolescent MDD (n=23), only baseline CGI–Severity showed a tendency to a significant negative correlation with the percentage of placebo responders. The more severe the disorder at baseline, the lower the placebo response (r=-0.54, p=0.09, n=11; and r=-0.61, p=0.08, n=9, for definition of response with the primary variable and the CGI-I, respectively).

Discussion

The main limitation of this study is the fact that we pooled all placebo arms from studies that varied in methodology (e.g., inclusion criteria, initial placebo washout period, definition of responders, co-morbidity inclusions, and duration of treatment). In particular, the main analysis was conducted with a variety of outcome variables, and, for a common outcome variable, a diversity of responder definitions were used (e.g., in OCD studies, responders were defined as patients showing a 25% or 40% decrease of the Children's Yale-Brown Obsessive Compulsive Scale [CYBOCS];

TABLE 1. PLACEBO RESPONSE RATES AND MAIN CHARACTERISTICS IN DOUBLE-BLIND PLACEBO-CONTROLLED TRIALS FOR INTERNALIZED DISORDERS IN CHILDREN AND ADOLESCENTS

| Kramer and Feiguine 1981 | | | × | White Sit | в | only S | Screening 1 | run-in | Drug | duration | episode | (CGI) | of responders | Responders | $placebo^{a}$ |
|-----------------------------|--------|----------------------|-------|-----------|------|----------------|----------------------------|------------------|------------------------|----------|---------|-------|--|---------------|----------------|
| Gramer and Feiguine 1981 | | | | | | | Major depressive disorder | ressive | disorder | | | | | | |
| | 0.67 | 20 (13–18) 6 weeks | 15 | 09 | _ | Yes | $ m N_0$ | $^{ m No}$ | Amitryptiline | 36 | NA | NA | Authors-scale | | No |
| Preskorn et al. 1987 | 0.64 | 22 (6–12) 6 weeks | NA | NA | - | Yes | Yes | Yes | Imipramine | NA V | NA | NA | CGI | | Š |
| Puig-Antich et al. 1987 | 0.77 | 53 (NA) 5 weeks | 9.2 | 59 | 1 | Yes | Yes | $^{ m N}_{ m o}$ | Imipramine | NA | NA | NA | K-SADS-dep/ | 15/22 (68%) | Š |
| | | | | | | | | | | | | | anhedonia ≤ 2 | | |
| Hughes et al. 1990 | 0.5 | 31 (6–12) 6 weeks | NA | NA V | NA > | Yes | NA | NA | Imipramine | NA | NA | NA | CDRS 50% | | Š |
| Geller et al. 1990 | 9.0 | 31 (12–17) 8 weeks | | | | Yes | Yes | Yes | Nortryptiline | 41 | NA | NA | CDRS < 25% | 4/19 (21%) | $^{ m N}$ |
| Geller et al. 1992 | 0.89 | 50 (6–12) 8 weeks | | 2.99 | 1 | Yes | Yes | Yes | Nortryptiline | 40 | NA | NA | CDRS < 20 | 5/24 (17%) | Š |
| Kutcher et al. 1994* | 0.7 | 42 (15–19) 6 weeks | 17.8 | 30 | 2 | Yes | Yes | Yes | Desipramine | Ϋ́ | NA | NA | $\frac{1}{10000000000000000000000000000000000$ | | No |
| Kve et al. 1996 | 0.77 | 31 (12–18) 6 weeks | 15.1 | 69 | _ | Yes | Yes | Yes | Amitroptiline | 1 | N | 3.5 | HDRS 50% | | Ž |
| Emslie et al. 1997 | 0.97 | 96 (8–18) 8 weeks | 12.5 | 2 2 | - | Yes | Yes | Yes | Fluoxetine | 3.2 | 48 | 4.9 | CGI < 2 | 16/48 (33%) | Yes |
| Birmaher et al. 1998 | 0.73 | 27 (13–17) 10 weeks | 16.1 | 29 | _ | Yes | Yes | S | Amitryptiline | 13.6 | 50 | 4.6 | HDRS 50% | 11/14 (79%) | Š |
| Klein et al. 1998 | 0.83 | 45 (13–18) 6 weeks | 15.7 | 33 | | Yes | Yes | Yes | Desipramine | 6 | = | NA | CGI < 2 | | Š |
| Milin et al. 2000** | 0.97 | 286 (13–18) 12 weeks | 15.6 | 4 | 33 | No | Yes | Yes | Paroxetine | 19.1 | 70 | 4. | MADRS 50% | | N _o |
| Keller et al. 2001 | | 275 (13–17) 8 weeks | 15.1 | 46 | | No | Yes | No | Paroxetine | 13 | 77 | 4.4 | HDRS < 8 or 50% | | No |
| Emslie et al. 2002 | 0.89 | 219 (8–18) 8 weeks | 12.7 | 51 | | Yes | Yes | Yes | Fluoxetine | 14.1 | 79 | NA | CDRS 30% | | No |
| Wagner et al. 2003 | | 376 (6–17) 10 weeks | 12 | 49 | 53 | Yes | Yes | N _o | Sertraline | 22 | 88 | 4.5 | CDRS 40%° | 105/179 (59%) | Xes^b |
| March et al. (TADS) 2004 | · — | 439 (12-17) 12 weeks | 14.6 | 46 | | Yes | No | No | Fluoxetine | 17.3 | 98 | 4.8 | $CGI \le 2$ | 39/112 (35%) | Yes |
| Wagner et al. 2004 | 0.67 | 174 (7–17) 8 weeks | 12 | 47 | | S _N | Yes | Yes | Citalopram | 18.6 | 81 | 4.3 | $CDRS \le 28^{c,d}$ | 20/85 (24%) | Yes |
| Emslie et al. 2006 | | 206 (7–17) 8 weeks | 12 | 53 | | °Z | N _o | N _o | Paroxetine | 24.9 | 53 | 4.3 | $CGI \leq \overline{2}$ | | Š |
| Von Knorring et al. 2006 | | 244 (13-18) 12 weeks | 16 | NA | | No | No | No | Citalopram | NA | 36 | NA | K-SADS-dep/ | 47/77 (61%) | No |
|) | | | | | | | | | • | | | | anhedonia ≤ 2 | | |
| Wagner et al. 2006 | 0.67 | 268 (6–17) 8 weeks | 12.6 | 48 | 25 \ | Yes | N _o | Yes | Escitalopram | 15.6 | NA | 4.2 | CGI≤2 | 69/132 (52%) | Š |
| Emslie et al. 2007 | | 367 (7–17) 8 weeks | 12.2 | 54 | | Yes | Yes | Yes | Venlafaxine | 21.4 | 98 | 4.5 | CDRS 35% ^{c,d} | 99/165 (60%) | $^{\rm q}$ |
| FDA: CN104-141 2007 | | 206 (12–17) 8 weeks | 14.7 | 41 | | Yes | NA | $^{ m N}$ | Nefazodone | 18.5 | 73 | NA | CGI≤2 | | No |
| FDA: 003-045 2007 | N Q | 259 (7-17) 8 weeks | 12.2 | 48 | | Yes | NA | No | Mirtazapine | NA | NA | NA | CGI ≤ 2 | 42/85 (49%) | No |
| | | | | | | Õ | 9 9 | npulsi | mpulsive disorder | | | | | | |
| DeVeaugh et al. 1992 | 0.79 | 60 (10-17) 8 weeks | 14 | 55 | , | Yes | Yes | Yes | Clomipramine | 49 | NA | NA | $CGI \le 2$ | 5/29 (17%) | Yes |
| March et al. 1998 | 0.86 | 189 (6–17) 12 weeks | 12.6 | NA | | Yes | Yes | Yes | Sertraline | 99 | NA | 3.8 | CYBOCS 25% | | Yes |
| Geller et al. 2001 | 0.92 | 103 (6–17) 13 weeks | 11.4 | 47 | 21 | Yes | Yes | Š | Fluoxetine | NA | NA | 4.8 | CYBOCS 40% | | Yes |
| Riddle et al. 2001 | 0.86 | 120 (8-17) 10 weeks | 12.7 | 55 | | Yes | Yes | Yes | Fluvoxamine | 41 | NA | 3.9 | CYBOCS 25% | 17/63 (27%) | Š |
| Liebowitz et al. 2002 | | 43 (8-17) 8 weeks | 12.3 | 64 | 2 | Xes | Yes | $^{ m N}_{ m o}$ | Fluoxetine | NA | NA | 4.6 | $CGI \le 2$ | 7/22 (32%) | Š |
| Geller et al. 2004 | 0.96 | 207 (7–17) 10 weeks | 11.6 | 61 | | Yes | $\overset{	ext{No}}{\sim}$ | S _o | Paroxetine | 53 | 09 | 4.5 | CYBOCS 25% | 42/102 (41%) | Yes |
| POTS 2004 | - | 112 (7-17) 12 weeks | 12.3 | 50 | | Yes | No No | $^{ m N}$ | Sertraline | NA | NA | 4.7 | CYBOCS < 10 | | Yes |
| | | | | | Non | -opse | com | oulsive | pulsive anxiety disord | ler | | | | | |
| Gittelman-Klein et al. 1973 | 3 0.86 | 35 (6–15) 6 weeks | 10.8 | 46 | | Yes | | N | Imipramine | NA | NA | NA | School attendance | 9/19 (47%) | No |
| Berney et al. 1981 | | 46 (9–15) 12 weeks1 | NA | 41 | 1 | Yes | No | N _o | Clomipramine | NA | NA | 3.8 | CGI < 2 | | No |
| Klein et al. 1992 | 0.7 | 20 (6–15) 6 weeks | 9.5 | 99 | 1 | Yes | Yes | N _o | Imipramine | NA | NA | NA | Global improvement | 4/9 (44%) | Š |
| Simeon et al. 1992 | ~ | 30 (NA) 4 weeks | 11.8 | 9/ | 2 | °Z | Yes | Yes | Alprazolam | NA | NA | NA | CGI≤2 | 5/13 (38%) | Š |
| Rynn et al. 2001 | | 22 (5–17) 9 weeks | 11.7 | 77 | | Yes | Yes | N _o | Sertraline | NA | NA | 4 | CGI < 2 | 1/11 (9%) | Yes |
| R ÚPP 2001 | | 128 (6–17) 8 weeks | 10.3 | 51 | 5 | Yes | Yes | No | Fluvoxamine | NA | NA | NA | $CGI \leq 2$ | 19/65 (29%) | Yes |
| Birmaher et al. 2003 | | 74 (7–17) 12 weeks | 11.9 | 46 | - | Yes | Yes | S _o | Fluoxetine | NA | NA | 3.9 | CGI < 2 | 13/37 (35%) | Yes |
| Wagner et al. 2004 | | 322 (8–17) 16 weeks | 13.3 | 57 | 38 | Š | Z | N | Paroxetine | Ϋ́Z | Ν | 4 | $CGI \leq 2$ | 59/154 (38%) | Yes |
| Rynn et al 2007 | | 323 (6–17) 8 weeks | 12.5 | 52 | | , S | Yes | Yes | Venlafaxine | 39 | ΥN | 4 ک | | 77/159 (48%) | Yes |
| Kymii et ai: 2007 | | CALCO (11 0) CAC | 1.1.1 | 1 | | 3 | 2 | 3 | | | | | 7/ 150 | | 3 |

Site = number of sites; Disease duration = in months; First episode = % of first episode included in the PBO arm; CDRS = Children's Depression Rating Scale; CGI = Clinical Global Impression—Severity; HDRS = Hamilton Depression Rating Scale; K-SADS-dep = Schedule for Affective Disorder and Schizophrenia—depression/anhedonia subscore; MADRS—Montgomery—Asberg Depression Rating Scale; CYBOCS = Child Yale-Brown Obsessive Compulsive Scale; Placebo responders = N of responders in the placebo arm/N of subjects randomised in the placebo arm (%). Abbreviations: N = Number of subjects randomized in the study; Q = quality score of the report; NA = not appropriate; M = % of male included in the PBO arm; White = % of white included in the PBO arm;

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^aPossible superiority of drug vs. placebo concerns the primary variable.

These reports are pooled analysis in which individual trials did not reveal significant treatment effect.

CGI responders are also indicated as secondary variables in these studies.

^dThis study used a continuous variable as the primary outcome variable; the same variable as a dichotomous secondary variable was chosen in the current analysis for definition of responders.

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| TABLE 2. PREDICTOR OF PLACEBO RESPONSE RATES IN DOUBLE-BLIND PLACEBO-CONTROLLED TRIALS |
|--|
| FOR INTERNALIZED DISORDERS IN CHILDREN AND ADOLESCENTS |

| | Statistics | Prima | ry variable response | CGI-I response | |
|-----------------------------|------------|------------|-------------------------------|----------------|-------------------------------|
| Variables | | N of trial | Results | N of trial | Results |
| Sociodemographic charac | teristics | | | | |
| Mean age | Pearson | 37 | $r = 0.3 \ (p = 0.076)$ | 27 | r = 0.27 (N.S.) |
| Male | Pearson | 36 | r = -0.16 (N.S.) | 27 | $r = -0.4 \ (p = 0.0538)$ |
| White | Spearman | 32 | $\rho = -0.457 \ (p = 0.006)$ | 24 | $\rho = -0.52 \ (p = 0.009)$ |
| Methodological character | istics | | | | |
| N placebo arm | Spearman | 40 | $\rho = 0.1 \text{ (N.S.)}$ | 29 | $\rho = 0.25 \text{ (N.S.)}$ |
| Study duration | Spearman | 40 | $\rho = -0.21$ (N.S.) | 29 | $\rho = -0.17$ (N.S.) |
| TCA vs. other drug | Student | 40 | $t = -0.43 \ (p = 0.67)$ | 29 | $t = -0.34 \ (p = 0.74)$ |
| Washout | Student | 36 | $t = 0.84 \ (p = 0.4)$ | 29 | $t = 2.4 \ (p = 0.024)$ |
| Placebo run-in | Student | 39 | t = -0.44 (N.S.) | 29 | t = 0.21 (N.S.) |
| U.S. only | Student | 40 | t = 0.69 (N.S.) | 29 | t = 1.43 (N.S.) |
| N of meeting | Pearson | 30 | r = 0.0 (N.S.) | 22 | r = -0.323 (N.S.) |
| Quality score | Spearman | 38 | $\rho = -0.02$ (N.S.) | 27 | $\rho = -0.08 \text{ (N.S.)}$ |
| Psychopathological chara- | cteristics | | • | | • |
| Disorder duration | Pearson | 23 | $r = -0.46 \ (p = 0.027)$ | 19 | $r = -0.46 \ (p = 0.04)$ |
| First episode | Pearson | 14 | r = -0.14 (N.S.) | 12 | r = 0.04 (N.S.) |
| Baseline CGI-S | Pearson | 23 | r = -0.16 (N.S.) | 19 | r = 0.33 (N.S.) |
| Disease ^a | ANOVA | 40 | $F = 6.07 \ (p = 0.005)$ | 29 | F = 7.37 (p = 0.003) |
| Other variables | | | • | | • |
| Year of publication | Spearman | 40 | $\rho = 0.05 \text{ (N.S.)}$ | 29 | $\rho = 0.23 \text{ (N.S.)}$ |
| Drug > placebo ^b | Student | 40 | $t = 2.61 \ (p = 0.013)$ | 29 | $t = 2.25 \ (p = 0.033)$ |

^aMajor depressive disorder versus obsessive compulsive disorder versus other anxiety disorder.

see Table 1). However, the fact we found overall similar results with the CGI criterion of response is somewhat reassuring. Most of the studies were sponsored by the pharmaceutical industry, and, in many cases, either broad inclusion criteria (e.g., including youth with co-morbid anxiety in depression trials and *vice versa*), or

narrow inclusion criteria were used (excluding suicidal youth from depression trials), both of which may limit generalizability to real-world clinical populations.

In our study, pooled analysis does not address the fact that MDD is likely to be highly heterogeneous and that some patients may

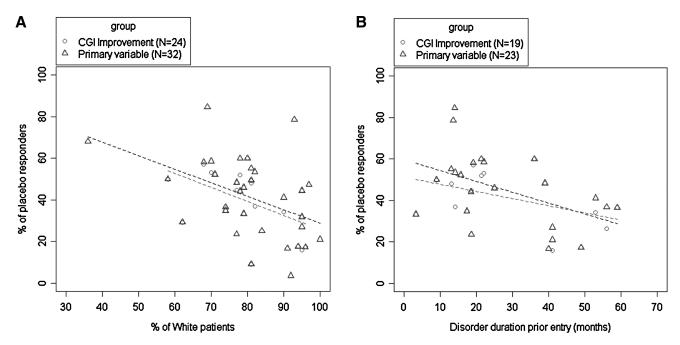


FIG. 1. Proportion of placebo response in child and adolescent internalizing disorders according to the percent of **(A)** white/Caucasian patients and **(B)** the disorder duration prior entry. CGI = Clinical Global Impressions.

^bMeans that the drug was significantly superior to the placebo on the primary response variable.

Abbreviations: CGI-I = Clinical Global Impressions-Improvement; N = Number; N.S. = Not significant; TCA = tricyclic antidepressant; CGI-S = Clinical Global Impressions-Severity; ANOVA = analysis of variance.

differ in their response to nonspecific professional attention (as evidenced in the wide range of placebo response rates across studies) (Cohen et al. 2008). Three studies conducted before 1980 did not use Diagnostic and Statistical Manual of Mental Disorders, 3rd edition (DSM-III) (or later) criteria (American Psychiatric Association 1980) and may have had substantial differences in the disorders' definitions. However, they only accounted for 48 patients in the total analysis. Compared to the Bridge analysis (2009), we did not consider predictors of drug response as a comparator to the analyses of predictors of placebo response. This choice was based on our primary hypothesis related to differential placebo response according to psychopathology. We decided to focus on placebo arms only because we hypothesized that response to drug may be dependent on psychopathology and other factors beyond our set of hypotheses as well. Finally, we were limited to trial-level summary data, which may fail to identify important individual patient factors influencing the response to placebo. Therefore, important factors such as SES, psychiatric family history, family adherence, history of sexual abuse or of other maltreatment, and time course of response over study weeks could not be investigated.

Despite these limitations, and in contrast to what has been shown in adult mood disorders (Keck et al. 2000; Walsh et al. 2002; Sysko and Walsh 2007), we did not find a positive correlation between placebo response rate and year of publication despite apparent variability in both placebo rates and publication dates. Our study supports the findings of Bridge et al. (2009) in a wider range of disorders, that increasing placebo response rates over time most likely is a publication artifact, more likely associated with increasing number of study sites over time. Second, age did not appear to modify placebo response rates in children versus adolescents across internalizing disorders (Cohen et al. 2008; Bridge et al. 2009). Most of the data used in these analyses came from studies on SSRIs that were conducted in the late 1990s and early 2000s, after several consensus conferences and guidelines had been done on youth psychopharmacological trials, and the methodologies were very similar for these studies (Bridge et al. 2007). When the CGI-I was used to define placebo response rates (either as the primary or secondary variable), we obtained the same results. There is some evidence that CGI severity scores at baseline may be capturing different aspects of illness severity versus other measures (e.g., Children's Depression Rating Scale, Revised [CDRS-R]) and that the CGI-I may serve to inflate placebo response rates (Bridge et al. 2009).

MDD in children and adolescents appears to be more responsive to placebo than do other internalizing disorders, which may suggest differential features underlying the psychopathology and pathophysiology of the disorders, discussed in depth elsewhere (Cohen et al. 2008). Our findings reveal that, as in adults (Khan et al. 2003), the magnitude of placebo response was the single most powerful predictor of the outcome of an antidepressant trial for all internalizing disorders studied. This is not surprising, given that the percentage of responders in the active compound arm of most studies in youth with internalizing disorders varies fairly widely: 36%-71% (MDD), 21%-65% (OCD), and 56%-91% (ANX). The ranges are narrower when extreme values are excluded: 49%-67% (MDD), 42%-57% (OCD), and 56%-78% (ANX) (Bridge et al. 2007). Therefore, the "success" of a trial is more dependent on the placebo response amplitude than the response to the active treatment itself. This result confirms the Bridge et al. (2009) analysis on trials for youth with depression.

We identified another inverse predictor of placebo response related to psychopathological characteristics: The duration of the disorder prior to trial inclusion. When the analysis was limited to the subgroup of studies on MDD, we only found a tendency for baseline severity of illness to correlate with lower placebo response rate, as in the Bridge et al. (2009) analysis. This is consistent with a recent meta-analysis conducted by Kirsch et al. (2008) showing that drug-placebo differences in antidepressant efficacy in adults increase as a function of baseline severity, and that the relationship between initial severity and antidepressant efficacy was attributable to decreased responsiveness to placebo among very severely depressed patients rather than to increased responsiveness to medication. These last points raise questions about the benefit-to-risk profile of using medications in treating youths with internalizing disorders who have only mild-to-moderate functional impairment. The use of psychotherapy and/or psychosocial approaches as firstline treatment for mild-to-moderate severity pediatric internalizing disorders may be a better first choice (Muratori et al. 2003; Brent et al. 1997; Goodyear et al. 2007). Pharmacotherapy should be restricted to patients showing insufficient improvement of their mood or anxiety symptoms with psychosocial treatment or psychotherapy and to patients showing severe impairments (Cohen 2007; Emslie 2009). Indeed, a recent finding from the Treatment of Adolescent Depression Study (TADS) group (Kennard et al. 2009) is reassuring in that patients who do not initially respond to placebo subsequently respond to active treatment at the same rate as those initially assigned to active treatment and that there were no differences between these groups in terms of rates of suicidal events. For children and adolescents with low-to-moderate levels of internalizing symptoms, it seems that being in regular clinical care, and not medication treatment per se, is the most important initial treatment intervention.

Methodological characteristics were not significant predictors of placebo response. Use of a placebo run-in, study duration, and Detsky et al. quality scores all failed to emerge as predictors. The only moderate contribution to reducing placebo response was the use of a washout-screening period, which was significantly associated with a lower placebo CGI-I response in the subgroup analysis. However, this result should be regarded with caution because it was not confirmed in the larger analysis. This may be a consequence of the brief washout-screening and placebo run-in periods, typically lasting less than 2 weeks. In contrast to the Bridge et al. (2009) report on MDD, we did not find a correlation of placebo response with the number of study sites, neither in our large analysis, nor in the MDD subanalysis. This may be due to our larger sample of studies (23 in MDD in our sample vs. 12 in Bridge et al.). Although as expressed by these authors, the tendency to increase the number of study sites in trials for youth MDD to enhance the number of patients, and therefore the statistical power, has been remarkable in recent years, has resulted in the tendency to include less severe patients and may hinder the ability to detect true drugplacebo differences in future studies in youth.

Finally, an intriguing result of the present analysis is the strong negative correlation found between the percentage of Caucasian patients included in the trials and the magnitude of the placebo response. This correlation appeared with both definitions of response, as defined by the primary study variable and CGI-I responder status, respectively ($\rho = -0.457$, p = 0.006; $\rho = -0.52$, p = 0.009). We cannot account for this finding based on genetic differences between Caucasians and other racial groups because we are not aware of such findings having been reported in either the child or adult depression treatment literature. We hypothesize that Caucasian race may serve as a proxy for other socioeconomic variables rather than the construct of ethnicity *per se*. Indeed, low SES and early life adversities are risk factors for internalizing

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disorders (Goodyear et al. 2007) and may serve as response predictors to different types of treatment interventions. Alternatively, the meaning of medication treatment may be perceived differently across cultural/ethnic/racial groups, leading to different expectations and placebo response rates. For example, differences in placebo response rates by the origin of patients has been reported in adult migraine prophylaxis (Macedo et al. 2008) and by patient's personal expectations in chronic pain (Linde et al. 2007). In a recent study, Stevens et al. (2009) showed that African-American parents may hold more negative perceptions than other parents regarding antidepressant prescription. Or, some other factor, such as unidentified co-morbidities (e.g., conduct disorder), was differentially represented across ethnic groups in the trials, the latter of which we were unable to capture in our analysis.

Our study has several implications for future research and for clinical decision making. Because differences in drug effectiveness appear to be due more to placebo rather than active treatment factors and to patient characteristics such as illness severity and duration, greater attention should be paid to recruiting study subjects with severe levels of illness. The use of longer placebo run-ins or assessment periods should be considered prior to actual drug treatment in clinical trials. Studies with patients resistant to psychotherapy should be considered as well. Consideration should be given to limiting future clinical trials to fewer sites using highly trained researchers to more tightly manage recruitment, assessment, and conduct of treatment methodology. Future studies should also examine the time course of response to identify potential differences in the time-action effect of placebo versus drug. The true differences in placebo responsivity between adolescents and preadolescents remains an open question, and future stratified study designs may elucidate developmental characteristics important to ultimate treatment response. So too, there have been almost no direct drug-drug-placebo comparator trials in child and adolescent populations. Undertaking comparator trials may begin to elucidate true differences in efficacy between the various drugs on the market for particular types of patients and disorders. Greater attention should be paid to systematically assessing socioeconomic and ethnic variables because this may identify modifiers of treatment response to drug versus placebo in select groups.

Finally, clinicians need to understand the strengths and limitations of research on placebo response. Children and adolescents with mild-to-moderate levels of internalizing symptoms, or of short duration illness, may be better served, and not appreciably harmed, by more conservative approaches to initial therapy using psychosocial treatments before consideration of medication. This was recently confirmed in the analysis of long-term outcomes of adolescents initially treated with placebo in the TADS short-term trial (Kennard et al. 2009).

Disclosures

Previously, David Cohen was an investigator in two industry-sponsored trials, DEROXADO and ADOKOTE, conducted by SmithKline Beecham and Sanofi-Synthélabo, respectively. He also received honoraria for speaking fees from Janssen. Craig Donnelly serves as a consultant for Eli Lilly, and has received grant and research funds from Eli Lilly. He has also served as a consultant to Johnson & Johnson, is on a speakers' bureau for Johnson & Johnson, and serves on a speakers' bureau for Shire. David Cohen, Angèle Consoli, Nicolas Bodeau, Diane Purper-Ouakil, Emmanuelle Deniau, Jean-Marc Guilé report no competing interests during the last 36 months.

Acknowledgments

The authors would like to thank the Centre d'Activités et de Recherches en Psychiatrie Infanto-Juvénile (CARPIJ), the Programme Hospitalier de Recherche Clinique (PHRC) (2006AOM098), and the Fondation Whyeth for their support in promoting research on child and adolescent mood disorders.

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