



Prepregnancy Habitual Intakes of Total, Supplemental, and Food Folate and Risk of Gestational Diabetes Mellitus: A Prospective Cohort Study

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OBJECTIVE

To identify novel modifiable risk factors of gestational diabetes mellitus (GDM) by examining the association between prepregnancy habitual folate intake and GDM risk.

RESEARCH DESIGN AND METHODS

The study included 14,553 women in the Nurses' Health Study II who reported at least one singleton pregnancy between the 1991 and 2001 questionnaires. Prepregnancy intakes of total folate, supplemental folate, and food folate were assessed using a food frequency questionnaire administered every 4 years. Incident GDM was ascertained from a self-reported physician diagnosis. Relative risks (RRs) of GDM were estimated using log-binomial models, with adjustment for demographic, lifestyle, and dietary factors.

RESULTS

Over the study follow-up, 824 incident GDM cases were reported among 20,199 pregnancies. Women with adequate total folate intake (≥ 400 $\mu\text{g}/\text{day}$) had an RR of GDM of 0.83 (95% CI 0.72, 0.95, $P = 0.007$) compared with women with inadequate intake (< 400 $\mu\text{g}/\text{day}$). This association was entirely driven by supplemental folate intake. The RRs of GDM for 1–399, 400–599, and ≥ 600 $\mu\text{g}/\text{day}$ of supplemental folate intake were 0.83, 0.77, and 0.70, respectively, compared with no supplemental folate intake ($P_{\text{trend}} = 0.002$). The association between supplemental folate intake and GDM risk largely persisted after additional adjustment for intake of multivitamins and other micronutrients, as well as among women who likely planned for the pregnancy.

CONCLUSIONS

Higher habitual intakes of supplemental folate before pregnancy were significantly associated with lower GDM risk. If confirmed, these findings indicate that prepregnancy folic acid supplementation could offer a novel and low-cost avenue to reduce GDM risk.

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Gestational diabetes mellitus (GDM) is a common complication in pregnancy, estimated to affect 5–9% of pregnancies in the U.S. in 2009–2010 (1). It is associated not only with adverse perinatal outcomes but also with long-term cardiometabolic risk in mothers and their offspring (2). Identifying potentially modifiable factors that contribute to the prevention of GDM may improve the health and well-being of both women and their children.

Folate is a B vitamin that occurs naturally in foods such as vegetables, fruits, and beans. Its synthetic form, folic acid, is commonly used in dietary supplements and fortified foods. Folate serves as a cofactor in one-carbon metabolism essential for nucleotide synthesis and methylation. Low folate intake leads to increased homocysteine levels, which is a risk factor for cardiovascular disease and stroke (3). Emerging evidence suggests that low folate intake and high homocysteine levels may play a role in metabolic disturbances, including insulin resistance (4), dyslipidemia (4–6), and liver damage (4,7), through compromised methylation capacity and oxidative stress, whereas folic acid supplementation may improve these metabolic parameters (6,8–10). These metabolic disturbances have been implicated in the pathogenesis of GDM. However, it is unclear whether higher folate intake is associated with lower GDM risk. Epidemiologic studies addressing this question are sparse, and findings are inconsistent (11–20). Importantly, most of the studies had a cross-sectional design and a small sample size (11–16,18), and none considered potential confounding from other aspects of diet. Our study examined prepregnancy habitual intake of folate in relation to the risk of GDM in a large prospective cohort followed over 10 years accounting for potential confounding from other aspects of diet. In a subgroup of the study sample, we also explored whether the association was modified by a common single nucleotide polymorphism, methylenetetrahydrofolate reductase (*MTHFR*) C677T, which leads to lower folate and higher homocysteine levels (21).

RESEARCH DESIGN AND METHODS

Study Population

The Nurses' Health Study II (NHS II) enrolled 116,678 female nurses aged 25–44 years in 1989 and followed them using biennial questionnaires, where women reported their lifestyle and disease

outcomes. Response rates for each questionnaire cycle were >90%. The current study included women who reported at least one singleton pregnancy lasting >6 months. Only pregnancies that occurred after the return of the first dietary assessment in 1991 were included in the current analyses. GDM was last ascertained on the 2001 questionnaire, when most participants had passed their reproductive age. Pregnancies were excluded if the woman reported a prior GDM diagnosis, as they might consequently modify their diet and lifestyle in subsequent pregnancies. Pregnancies were also excluded if the woman had prior diagnosis of type 2 diabetes, cardiovascular disease, or cancer other than skin cancer or if she had not returned any food frequency questionnaire (FFQ) before the pregnancy, had >70 FFQ items missing, or reported improbable total energy intake (<600 or >3,500 kcal/day). A total of 20,199 pregnancies among 14,553 women were included in this study. This study was approved by the institutional review board of Partners Health Care, Boston, Massachusetts.

Dietary Assessment

Women responded to a semiquantitative FFQ (22) in 1991 and every 4 years thereafter, where they reported the frequencies in which they consumed a specific portion of each of the 131 food or food group items during the past year. They also reported their use of dietary supplements, including the brand, dose, and frequency of use. We obtained the nutrient contents of each item from a nutrient database derived from the U.S. Department of Agriculture with additional information from manufacturers (23). For dietary information collected after 1998, we used an updated database, which reflects the mandatory folic acid fortification of cereal and grain products. Intake of each nutrient, including folate, was estimated as the sum of the intake from foods and supplements. Nutrient intakes from food were adjusted for total energy intake using the nutrient residual method (24). This study evaluated folate intake from supplements and food, together (i.e., total folate) and separately, as the exposures of interest. Folate intake estimated by the FFQ has shown moderate to high correlation in previous studies with an estimate from prospectively collected diet records ($r = 0.71$) (25), red blood cell folate levels ($r = 0.51$) (26), and plasma folate levels ($r = 0.63$) (27). The

Alternate Healthy Eating Index 2010 (AHEI-2010) was derived and used as a measure of overall dietary quality (28).

Outcome Assessment

Women reported incident GDM diagnosis on the biennial questionnaire up to 2001. GDM diagnosis was attributed to the first pregnancy when more than one eligible pregnancy was reported during the questionnaire period. An earlier study among a subsample of NHS II participants reported that 94% of self-reported GDM diagnosis was confirmed by medical record review, while among women who had a pregnancy uncomplicated by GDM, 83% reported a glucose loading test, and 100% reported frequent urine screening in pregnancy, consistent with a high degree of GDM surveillance (29). The National Diabetes Data Group criteria (30) for diagnosing GDM were widely adopted during the study follow-up period, between 1991 and 2001.

Nondietary Covariates

Women reported their frequencies engaging in common recreational activities in 1991, 1997, and 2001, from which their total physical activity was estimated. They reported their race and family history of diabetes in 1989. They also reported their height, weight, parity, smoking status, oral contraceptive use, use of ovulation induction medication, and concurrent infertility (i.e., having tried to become pregnant for >1 year without success since the last questionnaire cycle) in each biennial questionnaire, and the values from the most recent questionnaire cycle were used in the analysis. A validation study among a subsample of NHS I participants found a high correlation between self-reported weight and weight measured by a technician ($r = 0.97$) (31). Women's BMI was calculated as weight in kilograms divided by the square of height in meters. Missing values in covariates were coded as a separate category.

Statistical Analysis

Participant characteristics at baseline in 1991 were presented by categories of folate intake in 1991. Differences across the categories were compared using ANOVA test for continuous variables and χ^2 test for categorical variables.

Cumulative average amounts of nutrient and total energy intakes and physical activities before pregnancy were derived

to reduce measurement error and to represent long-term diet (32). Relative risks (RRs) and 95% CIs of incident GDM in relation to categories of folate intake (i.e., quartiles and adequacy [<400 $\mu\text{g}/\text{day}$ and ≥ 400 $\mu\text{g}/\text{day}$] of total folate intake, categories of supplemental folate intake [0, 1–399, 400–599, and 600–2000 $\mu\text{g}/\text{day}$], and quartiles of food folate intake) were estimated using log-binomial models with generalized estimating equations and robust variance estimates. Log-Poisson models were used in instances when log-binomial models did not converge, which provides comparable estimates with wider CIs. Linear trends of GDM risk across categories of folate intake were examined by fitting the models using the median intake of each category of folate intake as a continuous variable. In a series of models, RRs and linear trends were estimated with adjustment for potential confounders, which were selected a priori, including age (months), race (white, African American, Hispanic, Asian, others), parity (0, 1, 2, ≥ 3), family history of diabetes (yes, no), prepregnancy BMI (<21.0 , 21.0–22.9, 23.0–24.9, 25.0–26.9, 27.0–28.9, 29.0–30.9, 31.0–32.9, 33.0–34.9, ≥ 35.0 , and missing), cigarette smoking (never, past, current), alcohol use (0, 0.1–5.0, 5.1–9.9, ≥ 10 g/day), and quartiles of physical activity, total energy intake, glycemic load, and intakes of saturated fat, total fiber, and heme iron. In addition, restricted cubic spline models were used to flexibly model potential nonlinear relations between continuous folate intake and GDM risk.

To evaluate whether intakes of specific micronutrients or multivitamin use confounded the associations between folate intake and GDM risk, we performed the analyses with additional adjustments for other micronutrients (i.e., quartiles of vitamins B1, B2, B3, B5, B6, B12, A, C, and D, calcium, and magnesium) and multivitamin intake (times/week). To evaluate the possibility that women who used folate supplements may have a lower GDM risk due to a higher likelihood of pregnancy planning, we restricted the analyses to women who were likely planning for pregnancy: those who were currently married and not using oral contraceptives ($n = 14,594$), who concurrently used ovulation induction medications ($n = 2,159$), and who reported concurrent infertility ($n = 1,984$). To evaluate the extent to which our findings were robust to unmeasured

confounders, we calculated the E-values (33) for the estimated associations between folate intake and GDM risk. E-value reflects the magnitude of associations an unmeasured confounder must have with both the exposure and the outcome to fully explain the observed association between the exposure and the outcome. For a protective exposure, the E-value can be estimated as follows (33):

$$E\text{-value} = \frac{1}{RR} + \sqrt{\frac{1}{RR} \times \left(\frac{1}{RR} - 1\right)}$$

To examine potential effect modification by major risk factors of GDM, we stratified the analyses by age (<35 vs. ≥ 35 years), parity (nulliparous vs. parous), family history of diabetes (yes vs. no), and prepregnancy BMI (<25 , 25–29, ≥ 30 kg/m^2). To test the extent to which the findings were sensitive to model specifications, we used the updated prepregnancy dietary intakes from the most recent dietary assessment instead of the cumulative average intakes. We also adjusted for AHEI dietary quality score instead of individual nutrients.

Last, we conducted an exploratory analysis to examine the association between folate intake and GDM risk by *MTHFR* C677T genotypes (CC/CT vs. TT) among women genotyped for studies of GDM and case-control studies of other conditions (34). This analysis is restricted to white women, as only 3.7% ($n = 38$) of this subsample was nonwhite. A total of 1,313 pregnancies reported by 999 women during follow-up were included in this analysis. Compared with women without genotyping information, those with genotyping information on average were slightly older and more likely to have a family history of diabetes. Other demographic and lifestyle characteristics measured in the study were similar between the two groups. Odds ratios (ORs) and 95% CIs were estimated using logistic regression models with generalized estimating equations and robust variance estimates. All analyses in subsets of the study sample were adjusted for the same set of covariates as was used in the analyses in the full sample but used fewer categories to ensure model convergence.

RESULTS

During the study follow-up, 824 incident GDM cases were reported among the

20,199 pregnancies. At baseline in 1991, women who had higher folate intake from any source (i.e., food folate and supplemental folate) were more likely to be white, parous, and physically active and to have overall higher dietary quality and to use multivitamins, vitamin B-complex, and folic acid; they were less likely to be a current smoker and consumed less alcohol (Table 1).

Total folate intake was inversely associated with GDM risk (Table 2). After adjustment for age, race, parity, family history of diabetes, prepregnancy BMI, cigarette smoking, alcohol use, physical activity, and intakes of total energy, saturated fat, glycemic load, total fiber, and heme iron, the RRs of GDM across increasing quartiles of total folate intake were 1.00 (reference), 1.01 (95% CI 0.84, 1.21), 0.81 (0.67, 0.99), and 0.81 (0.66, 0.98), respectively ($P_{\text{trend}} = 0.009$). Each 100 $\mu\text{g}/\text{day}$ increase in total folate intake was associated with an RR of 0.96 (0.94, 0.99) for GDM. Adequate total folate intake (≥ 400 $\mu\text{g}/\text{day}$) was associated with an RR of 0.83 (0.72, 0.95) for GDM ($P = 0.007$) compared with inadequate intake (<400 $\mu\text{g}/\text{day}$). The association between total folate intake and GDM risk was entirely driven by supplemental folate (Table 2). The RRs of GDM by supplemental folate intake were 0.83 (0.71, 0.98) for 1–399 $\mu\text{g}/\text{day}$, 0.77 (0.64, 0.93) for 400–599 $\mu\text{g}/\text{day}$, and 0.70 (0.52, 0.94) for ≥ 600 $\mu\text{g}/\text{day}$ compared with no supplemental folate intake ($P_{\text{trend}} = 0.002$). Each 100 $\mu\text{g}/\text{day}$ increase in supplemental folate intake was associated with an RR of 0.95 (0.92, 0.98) for GDM. The restricted cubic spline model demonstrated linear ($P = 0.21$ for curvature) associations between supplemental folate intake and GDM risk (Fig. 1). Food folate intake was not associated with GDM risk ($P_{\text{trend}} = 0.66$).

In the multivariate models with additional adjustment for other micronutrients, the association between supplemental folate intake and GDM risk remained significant and generally unchanged (Supplementary Table 1). These findings suggest that the association between supplemental folate intake and GDM risk was not due to other micronutrients tested. In the multivariate model with additional adjustment for multivitamin intake, the association between supplemental folate intake and GDM risk largely persisted; the RRs of

Table 1—Age-standardized characteristics of study population (n = 14,553) according to categories of folate intake in 1991 in NHS II

	Total folate*		Supplemental folate		Food folate	
	Q1	Q4	None	≥600 μg/day	Q1	Q4
No. of participants	3,600	3,413	6,791	1,161	3,418	3,601
Folate intake, μg/day, median (range)	246 (81–294)	873 (698–2,770)	0	900 (600–2,000)	189 (46–212)	342 (302–1,030)
Age, years	31.8 ± 3.2	31.9 ± 3.3	31.9 ± 3.3	31.8 ± 3.2	31.3 ± 3.1	32.4 ± 3.3
White, %	90.8	94.1	92.4	93.7	91.0	93.6
Family history of diabetes, %	11.7	10.8	11.3	9.9	11.2	11.0
Nulliparous, %	35.2	32.5	36.6	15.1	32.0	43.9
Current smoking, %	12.5	6.9	10.6	5.4	11.3	9.2
Alcohol, g/day	3.2 ± 5.6	2.5 ± 4.4	3.2 ± 5.2	2.1 ± 4.2	2.6 ± 5.1	3.4 ± 5.1
BMI, kg/m ²	23.6 ± 4.8	23.1 ± 4.0	23.4 ± 4.4	23.5 ± 4.3	23.6 ± 4.8	22.9 ± 3.9
Physical activity, MET h/week	19.0 ± 26.2	26.7 ± 33.5	22.6 ± 29.2	24.4 ± 34.1	17.8 ± 24.3	30.9 ± 35.5
AHEI score	39.2 ± 9.3	45.9 ± 10.4	43.0 ± 10.4	43.5 ± 10.7	37.3 ± 8.9	49.8 ± 10.0
Total calories, kcal/day	1,753 ± 554	1,762 ± 536	1,780 ± 545	2,029 ± 556	1,801 ± 570	1,804 ± 552
Carbohydrate, %E	48.1 ± 7.5	52.0 ± 7.1	49.9 ± 7.4	52.0 ± 6.6	49.1 ± 7.9	53.0 ± 7.2
Glycemic load†	120 ± 22	126 ± 21	122 ± 21	125 ± 19	123 ± 24	127 ± 21
Glycemic index†	54.8 ± 3.4	53.6 ± 3.2	54.2 ± 3.3	53.7 ± 3.0	55.4 ± 3.3	52.8 ± 3.1
Protein, %E	18.8 ± 3.4	19.6 ± 3.4	19.2 ± 3.4	19.5 ± 3.2	18.3 ± 3.5	19.6 ± 3.5
Total fat, %E	33.4 ± 5.5	29.5 ± 5.3	31.5 ± 5.6	29.9 ± 5.2	33.2 ± 5.8	28.3 ± 5.1
Saturated fat, %E	12.1 ± 2.4	10.7 ± 2.3	11.3 ± 2.4	11.0 ± 2.2	12.2 ± 2.6	10.0 ± 2.2
Monounsaturated fat, %E	12.8 ± 2.3	11.0 ± 2.3	11.9 ± 2.4	11.2 ± 2.2	12.7 ± 2.5	10.5 ± 2.2
Polyunsaturated fat, %E	5.6 ± 1.3	5.3 ± 1.3	5.5 ± 1.3	5.2 ± 1.4	5.5 ± 1.4	5.4 ± 1.3
Trans fat, %E	1.9 ± 0.7	1.4 ± 0.5	1.6 ± 0.6	1.5 ± 0.5	1.9 ± 0.7	1.3 ± 0.5
Cholesterol, mg/day†	245 ± 62	233 ± 67	239 ± 65	236 ± 64	236 ± 61	232 ± 74
Total fiber, g/day†	15.1 ± 3.6	19.4 ± 6.2	17.8 ± 5.3	18.4 ± 5.3	14.5 ± 4.4	22.0 ± 6.0
Heme iron, mg/day†	1.2 ± 0.4	1.0 ± 0.4	1.1 ± 0.4	1.0 ± 0.4	1.1 ± 0.4	1.0 ± 0.4
Vitamin A, μg/day†	798 ± 570	2,795 ± 1,423	1,043 ± 739	2,799 ± 1,531	1,310 ± 1,046	2,008 ± 1,381
Vitamin C, mg/day†	136 ± 173	382 ± 382	166 ± 189	349 ± 320	176 ± 235	334 ± 326
Vitamin E, mg/day†	11.3 ± 33.1	33.4 ± 61.7	12.7 ± 36.3	30.7 ± 60.1	15.6 ± 33.9	25.9 ± 57.5
Total folate (B9), μg/day†	240 ± 40	966 ± 315	313 ± 106	1215 ± 301	415 ± 308	624 ± 315
Food folate, μg/day†	216 ± 40	288 ± 79	260 ± 71	271 ± 69	186 ± 25	358 ± 59
Supplemental folate, μg/day	3.4 ± 16	570 ± 316	0 ± 0	914 ± 198	168 ± 274	209 ± 289
Thiamin (B1), mg/day†	1.7 ± 2.0	5.5 ± 10.0	1.9 ± 2.2	5.4 ± 10.0	2.8 ± 5.8	3.8 ± 7.1
Riboflavin (B2), mg/day†	2.0 ± 1.9	6.3 ± 10.0	2.2 ± 2.1	6.3 ± 9.9	3.2 ± 4.7	4.3 ± 7.1
Niacin (B3), mg/day†	23.6 ± 12.8	50.9 ± 26.6	25.6 ± 13.8	49.5 ± 26.9	31.6 ± 19.5	38.4 ± 25.2
Pantothenic acid (B5), mg/day†	5.0 ± 4.0	15.7 ± 15.0	5.6 ± 4.4	13.3 ± 17.6	8.3 ± 13.3	11.4 ± 12.7
Vitamin B6, mg/day†	3.4 ± 13.4	12.1 ± 25.4	3.7 ± 13.8	14.9 ± 26.4	5.2 ± 14.9	8.6 ± 23.7
Cobalamin (B12), μg/day†	5.9 ± 3.7	16.1 ± 16.8	6.7 ± 4.9	18.0 ± 23.4	8.5 ± 7.0	11.2 ± 14.2
Multivitamin, %	7.9	90.8	2.7	95.4	42.9	52.3
Frequency of multivitamin intake, days/week	0.2 ± 0.9	6.4 ± 2.5	0.1 ± 0.9	7.3 ± 1.9	2.3 ± 3.2	2.9 ± 3.4
Vitamin B-complex, %	1.3	5.7	1.5	4.7	2.3	5.4
Folic acid, %	0.0	2.1	0.0	4.6	0.3	0.9

Values are means ± SD for continuous variables and percentages for categorical variables. All values except age are standardized to the age distribution of the study population. Comparisons of all characteristics across quartiles of folate intake are significant, except for family history of diabetes. Q, quartile; %E, % of energy. *Includes food folate and folic acid from supplements and fortified food. †Energy adjusted using residual method.

GDM by supplemental folate intake were 0.81 (95% CI 0.66, 1.00) for 1–399 μg/day, 0.73 (0.52, 1.02) for 400–599 μg/day, and 0.65 (0.42, 1.01) for ≥600 μg/day compared with no supplemental folate intake, although the linear trend became nonsignificant ($P_{\text{trend}} = 0.10$)

(Supplementary Table 1), whereas weekly frequency of multivitamin use was not associated with GDM risk (OR 1.01 [95% CI 0.96, 1.06], $P = 0.69$). This finding suggests that the major part of the association between supplemental folate intake and GDM risk was not explained by multivitamin

intake. When the analysis was restricted to women who were likely planning for pregnancy—those who were married and not using oral contraceptives, who concurrently used ovulation induction medications, and who reported concurrent infertility—the association between supplemental folate

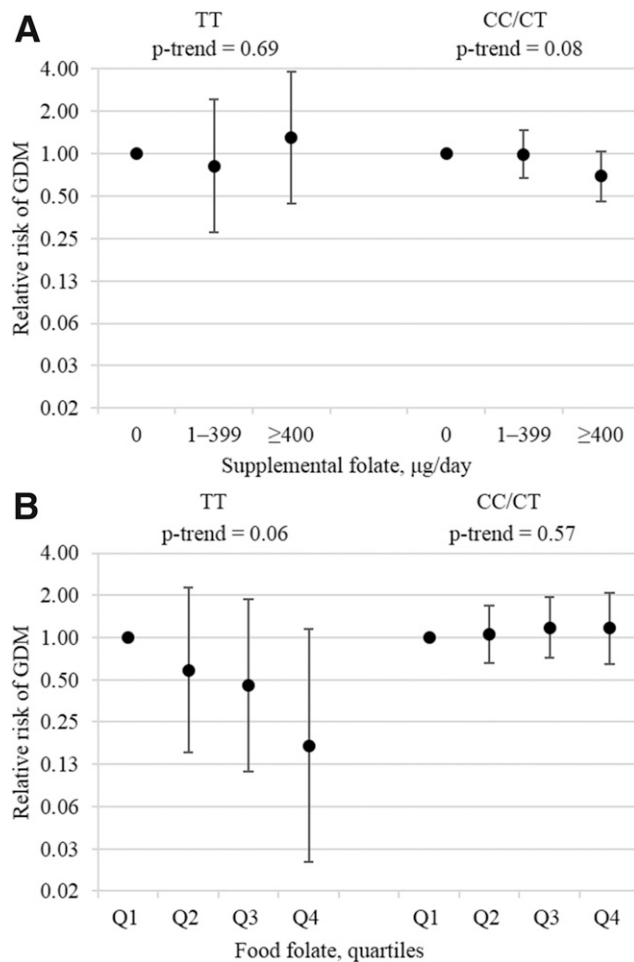


Figure 2—RRs and 95% CI of GDM according to prepregnancy supplemental (A) and food folate (B) intake by *MTHFR* C677T genotypes among white women. The CC/CT and TT groups included 1,136 and 177 pregnancies, respectively. Models adjusted for age (months), nulliparity (yes, no), family history of diabetes (yes, no), physical activity (MET h/week), prepregnancy BMI (<25.0, 25.0–29.9, and >30.0 kg/m²), current smoking (yes, no), alcohol use (yes, no), and dietary intakes of total energy (kcal/day), glycemic load (unit), saturated fat (% energy), total fiber (g/day), and heme iron (mg/day). Three categories of supplemental folate intake were used instead of four to avoid small numbers in some categories. Q, quartile.

leads to compromised methylation capacity and impaired phosphatidylcholine synthesis, which is essential for very low-density lipoprotein assembly and homeostasis (5,6). Consistent with this mechanism, inadequate folate intake has been shown to lead to liver steatosis in mice (37). Second, folic acid supplementation may upregulate AMPK among mice fed with a high-fat diet, improving insulin resistance and hepatic inflammation (8,9). Lastly, folic acid supplementation—directly (10) or via lowering blood homocysteine levels (7)—may protect against oxidative stress in mice, which is known to contribute to endothelial dysfunction and insulin resistance (38); among human populations, homocysteine levels have also been positively associated with GDM risk (39).

This study has several strengths. It was based on a large, prospectively followed cohort of women, which conferred adequate statistical power to detect modest associations and enabled us to establish a temporal relationship between the exposure and the outcome. We also had comprehensive assessments of dietary intake, which allowed the examination of food folate and supplemental folate separately and enabled us to adjust for confounding by other dietary factors; the repeated dietary assessment also likely reduced measurement error and potentially reflected long-term diet (32).

Several potential limitations merit discussion. First, measurement error in self-reported dietary intake was likely. However, folate intake measured from the

FFQ previously showed good correlation with prospectively collected diet records ($r = 0.71$) (25), red blood cell folate ($r = 0.51$) (26), and plasma folate levels ($r = 0.63$) (27). In addition, due to the prospective design, such measurement error is not likely to vary with respect to GDM status. Second, as in all observational studies, residual confounding cannot be ruled out. However, we adjusted for a comprehensive list of risk factors for GDM, including physical activity and other dietary factors. We also conducted sensitivity analyses to examine potential confounding from other micronutrients, multivitamin use, and pregnancy planning. Significant associations of folate intake from supplements with GDM persisted in these sensitivity analyses. Lastly, our exploratory analysis stratified by *MTHFR* C677T genotype was conducted in a small subsample, conferring limited statistical power. Future studies with larger sample sizes are needed to further examine whether *MTHFR* polymorphisms may modify the association between folate intake and GDM risk.

In the present large prospective cohort study, we found higher prepregnancy habitual folate intake from supplements to be significantly associated with lower risk of GDM. In addition, doses of folate supplement greater than current recommendation were associated with an even lower GDM risk. Given that pregnancy concerns both women and their fetus and that GDM is only one such relevant outcome, future data both from animal models and human epidemiologic studies are warranted to determine the appropriate supplementation dose. GDM is a substantial burden on the health of mothers and children, and folic acid supplementation has been widely recommended among all women of reproductive age to prevent birth defects. Moreover, data from one recent intervention study demonstrated that promoting folic acid supplementation can be achieved through a community-based prepregnancy care program (40). If confirmed, our findings indicate that prepregnancy folic acid supplementation could offer a feasible, novel, and low-cost avenue to reduce the risk of GDM.

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Author Contributions. M.L., S.L., and C.Z. contributed to the study design. J.E.C. and F.B.H. were involved in data collection. M.L. and S.L. analyzed data. M.L. drafted the manuscript. A.J.G. conducted technical review. C.Z. obtained funding. M.L., S.L., J.E.C., A.J.G., S.H.L., S.N.H., X.W., M.D., G.B., A.A.B., S.F.O., J.L.M., F.B.H., and C.Z. contributed to the interpretation of the results and revision of the manuscript for important intellectual content and approved the final version of the manuscript. M.L. and C.Z. are the guarantors of this work and, as such, had full access to all the data in the study and take responsibility for the integrity of the data and the accuracy of the data analysis.

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