Prognostic Indices for Older Adults

A Systematic Review

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AILURE TO CONSIDER PROGNOsis in the context of clinical decision making can lead to poor care. Hospice is underutilized for patients with nonmalignant yet lifethreatening diseases. 1 Healthy older patients with good prognosis have low rates of cancer screening.2 Older adults with advanced dementia or metastatic cancer are screened for slow-growing cancers that are unlikely to ever cause them symptoms but may lead to distress from false-positive results, invasive workups, and treatments.3,4 In recognition of these phenomena, guidelines increasingly incorporate life expectancy as a central factor in weighing the benefits and the burdens of tests and treatments (TABLE 1). Prognostic indices offer a potential role for moving beyond arbitrary age-based cutoffs in clinical decision making for older adults.2 However, little is known about the quality of prognostic indices for older adults, limiting their clinical use.

We performed a systematic review to describe the quality and limitations of validated non–disease-specific prognostic indices that predict absolute risk

For editorial comment see p 199.



CME available online at www.jamaarchivescme.com and questions on p 205.

Context To better target services to those who may benefit, many guidelines recommend incorporating life expectancy into clinical decisions.

Objective To assess the quality and limitations of prognostic indices for mortality in older adults through systematic review.

Data Sources We searched MEDLINE, EMBASE, Cochrane, and Google Scholar from their inception through November 2011.

Study Selection We included indices if they were validated and predicted absolute risk of mortality in patients whose average age was 60 years or older. We excluded indices that estimated intensive care unit, disease-specific, or in-hospital mortality.

Data Extraction For each prognostic index, we extracted data on clinical setting, potential for bias, generalizability, and accuracy.

Results We reviewed 21 593 titles to identify 16 indices that predict risk of mortality from 6 months to 5 years for older adults in a variety of clinical settings: the community (6 indices), nursing home (2 indices), and hospital (8 indices). At least 1 measure of transportability (the index is accurate in more than 1 population) was tested for all but 3 indices. By our measures, no study was free from potential bias. Although 13 indices had C statistics of 0.70 or greater, none of the indices had C statistics of 0.90 or greater. Only 2 indices were independently validated by investigators who were not involved in the index's development.

Conclusion We identified several indices for predicting overall mortality in different patient groups; future studies need to independently test their accuracy in heterogeneous populations and their ability to improve clinical outcomes before their widespread use can be recommended.

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of all-cause mortality in older adults. Recognizing that older adults are more likely to have more than 1 chronic illness than younger adults, we focused on non–disease-specific indices.

METHODS

We used broad Medical Subject Heading terms (eg, mortality, prognosis, aged) to search MEDLINE, EMBASE, Cochrane, and Google Scholar from their inception through November 2011 for English-language–validated prognostic indices that predicted absolute risk of all-cause mortality in patients whose average age was 60

years or older. Authors of included studies and experts in the field were contacted and asked for additional

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published and unpublished sources. We excluded indices that estimated intensive care unit (ICU), in-hospital, or disease-specific mortality. Two investigators (L.C.Y. and A.K.S) independently applied these inclusion and exclusion criteria to select prognostic indices and independently abstracted their data. Disagreements were resolved by consensus or, if necessary, the involvement of a third investigator (S.J.L).

There are no accepted criteria to assess the quality of prognostic indices. Therefore, we adapted criteria from previous work published by experts in medicine and epidemiology.28-35 We abstracted data on the quality of prognostic indices, including information on potential bias, generalizability, and accuracy (TABLE 2). For discrimination, we considered *C* statistics in the range of 0.50 to 0.59 to indicate poor, 0.60 to 0.69 to indicate moderate, 0.70 to 0.79 to indicate good, 0.80 to 0.89 to indicate very good, and 0.90 or greater to indicate excellent discrimination.44 For calibration, we considered 10 or more percentage points' difference between predicted and observed mortality to be evidence of poor calibration and less than 10 percentage points' difference to be evidence that the model was well calibrated. To further assess the potential limitations of these indices in clinical practice, we tracked studies that predicted greater than 50% mortality, since 50% mortality represents the median residual lifespan. We report 95% confidence intervals on measures of discrimination and calibration where available.

RESULTS

One investigator title-screened 21 593 studies to identify 4120 potentially relevant abstracts (eFigure, available at http://www.jama.com). After excluding studies with participants whose average age was less than 60 years old; studies that predicted only relative risk; or indices that predicted only disease-specific, in-hospital, or ICU mortality, there were 341 studies published between January 1987 and November

Expectancy Sample Clinical Decision Guideline Short-term (<2 y) <6 mo Discontinuation of statins^{5,6} <6 mo Referral to hospice Medicare regulations < 1-2 yNonoperative management of asymptomatic abdominal None aortic aneurysm⁷ Mid-term (2-3 y) <2-3 y Blood pressure/lipid control in diabetes mellitus unlikely California Healthcare to prevent macrovascular complications Foundation and AGS11 <2-3 y Lowering blood pressure to <140/80 mm Hg unlikely to improve cardiovascular outcomes^{5,12} Long-term (>3 y) Discontinuation of colon cancer screening 13,14 <5 y or <7 yAGS¹⁵ or USPSTF¹⁶ <5 y or "limited" AGS¹⁸ or USPSTF¹⁹ Discontinuation of breast cancer screening 13,17 <5 yStented bioprosthetic heart valve may be preferable to metallic valve²⁰ <5 yLimited benefit to lowering hemoglobin A₁₀ therapeutic California Healthcare target to <8%5 Foundation and

Table 1. Sample Clinical Decisions Influenced by Life Expectancy^a

breast-conserving surgery and hormonal therapy)^{26,27}

Abbreviations: ACS, American Cancer Society; AGS, American Geriatrics Society; AUA, American Urological Association; ER+, estrogen receptor-positive; USPSTF, US Preventive Services Task Force.

^a Prognosis is only one of many important factors to consider for these clinical decisions.

mortality benefit if life expectancy <15 y (for patients with T1, T2 ER+ breast cancer status after

Tight glycemic control in diabetes mellitus unlikely to prevent microvascular complications^{5,21,22}

Irradiation therapy to ipsilateral breast may not have

Discontinuation of prostate cancer screening²³

2011. After review of the full text of these studies, 317 studies were excluded, leaving 24 studies (eFigure). ^{36-43,45-60} Three of these studies presented updated versions of an index, ^{36,40,53} and 5 provided additional validation for an index, ^{38,43,54,58,59} resulting in a total of 16 unique indices.

<8 y

< 10 y

< 15 y

All indices were developed using secondary analysis of existing data sets of participants from the United States (11 indices)* and western Europe (4 indices). 42,47,48,52 The most common final predictors of mortality included functional status and comorbidities (each only absent in <5 indices). Three indices tested only reproducibility and did not evaluate any form of transportability (split sample validation only 47,48 and bootstrapping only⁵⁷) (Table 2). Only a single form of transportability was tested for 4 indices (geographic 39,46,52 and historical⁵¹). For 4 indices, the investigators who developed the index tested

*References 36, 39-41, 45, 46, 49, 51, 55-57.

the transportability of their index in a separate validation study. ^{37,38,42,43,49,55,58,59} Two indices were additionally validated by an investigator not involved in the index's development. ^{36,38,41,54}

California Healthcare

AGS1

None

ACS²⁴ and AUA²⁵

Foundation and

None of the examined indices had a C statistic \geq 0.90; 3 indices had C statistics between 0.80 and 0.89, suggesting very good discrimination^{39,40,49}; 10 indices had C statistics between 0.70 and 0.79, suggesting good discrimination†; and 3 indices had C statistics between 0.60 and 0.69, suggesting moderate discrimination.^{45,47,51} Indices were generally well calibrated across risk groups (TABLE 3). Two indices reported a greater than 10% difference between predicted and observed mortality.^{36,40}

We present a descriptive summary of each index by setting. Results of data abstraction regarding potential bias, generalizability, and accuracy are shown in Table 3 and TABLE 4.

†References 36, 37, 41, 42, 46, 48, 52, 55-57.

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Community-Dwelling Older Adults

Our review identified 6 indices for community-dwelling older adults. Indices estimated mortality risk from 1 year⁵⁶ to 5 years.⁵⁵ The highest-risk group from Schonberg et al at 9-year follow-up predicted 92% mortality (95% CI, 86%-96%).⁵⁸

Gagne et al⁵⁶ developed a mortality risk score to predict 1-year mortality by combining conditions in the Romano et al⁶² implementation of the Charlson et al index⁶³ and the van Walraven et al⁶⁴ implementation of the Elixhauser et al system.⁶⁵ The sample was a secondary analysis of Medicare enrollees 65 years and older who in 2004 participated in a pharmacy assistance contract for low-income seniors who did not qualify for Medicaid prescription drug coverage in Pennsylvania (development cohort, n=120 679) and New Jersey (validation cohort,

n=123855). The model had good discrimination and was well calibrated (Table 3). Reclassification measures compared the model favorably against the Romano/Charlson and van Walraven/Elixhauser indices.

The 15-month index by Mazzaglia et al⁵² is a 7-item questionnaire for primary care physicians that was developed in 2470 primary care patients who were 65 years and older residing in northwestern Florence, Italy, and validated in a sample of 2926 similar patients residing in southwestern Florence. The model was well calibrated and had good discrimination, but it predicted the narrowest range of mortality of any examined index (0%-10% risk).

Carey et al⁴⁶ developed a 2-year index for community-dwelling elderly individuals from a sample of 4516 adults 70 years and older from the eastern,

western, and central United States who had been interviewed in the Asset and Health Dynamics Among the Oldest Old (AHEAD) study in 1993. Carey et al subsequently validated the index in 2877 similar interviewees from the southern United States. The index had good discrimination and was well calibrated across all 3 risk levels but predicted only a narrow range of mortality (5%-36% risk).

The index by Carey et al for 3-year mortality⁴⁵ was developed in functionally impaired, nursing home–eligible, community-dwelling adults who were 55 years and older in the years 1988 through 1996, living in the western United States (n=2232), and enrolled in the Program of All-Inclusive Care for the Elderly (PACE), a senior daycare program providing multidisciplinary services. Validation was conducted in PACE participants from the eastern and

Term	Explanation	Measurement/Example 13% of participants in the Flacker and Kiely ³⁶ development cohort were lost to follow-up (unknown mortality at 1 y) and may have systematically differed		
Bias	Systematic variation (nonrandom error) in the development or validation of a prognostic index			
Accuracy	The degree to which predicted outcomes match observed outcomes			
Calibration	How close each level of prediction is to what is observed for that risk group	Compares predicted vs observed mortality rate; Hosmer-Lemeshow ^b		
Discrimination	How well those who die are distinguished from those who do not die	C statistic ^c		
Generalizability	Ability of a prognostic index to provide accurate predictions in a new sample of patients			
Reproducibility	The index is accurate in patients who were not included in the development cohort but who are from the same underlying population; a measure of overfitting (matching the predictive model to random noise in the data)	Data resampling (also called bootstrapping) ^d		
Transportability	The index is accurate in patients drawn from a different but related population or in data collected by using methods that differ from those used in development; a measure of both overfitting and underfitting (the omission of important predictors of mortality)	Nonrandomly split sample ^e or independent validation		
Methodological	Accuracy is maintained when the index is tested in data collected using different methods; independent validation tests the accuracy of the index by investigators not involved in the development of the index	Porock et al ³⁷ developed index and Kruse et al ³⁸ independently validated it		
Historical	Accuracy is maintained when the index is tested in data from a different calendar time	Inouye et al ³⁹ development sample was from 1989-1991; validation sample was from 1995-1996 ⁴⁰		
Geographic	Accuracy is maintained when the index is tested in data from different locations	Lee et al ³⁹ developed in eastern, western, and central US and validated in southern US		
Spectrum	Accuracy is maintained in a patient sample that is, on average, more or less advanced in disease process or that has a somewhat different disease process or trajectory	Walter et al ⁴¹ developed in tertiary care hospital and validated in community hospital		
Follow-up interval	Accuracy is maintained when the index is tested over a longer or shorter period	Pilotto et al ⁴² developed for 1-y and San Carlo et al ⁴³ validated for 1-mo mortality		

^a Adapted from Justice et al, ²⁸ Hayden et al, ³⁴ McGinn et al, ³⁵ and Steyerberg et al. ²⁹

b Higher values closest to 1 indicate better fit.

^CHigher values closest to 1 indicate better discrimination.

Develop the index in the entire data set, and then validate it in multiple bootstrap samples generated from the original sample with replacement.

^e Develop the index in one part of the data and validate it in another portion that differs on some major variable. Nonrandomly split samples measure an index's transportability better than randomly split samples.

		Genera	alizability ^a		Accuracy	
					Calibration ^c	
Source	Index	Development Cohort	Validation Cohort	Discrimination (95% CI) ^b	Predicted Mortality (95% CI), % ^b	Observed Mortality (95% CI), % ^b
Gagne et al, ⁵⁶ 2011	1-y index for low-income elders	n = 120 679 Average age 80 y 83% Female 29% Hospitalized in last year 9% Nursing home residents Median 18 distinct ICD-9 diagnoses 9% 1-y Mortality	ommunity-Dwelling Patients n = 123 855 Average age 79 y 77% Female 27% Hospitalized in last year 9% Nursing home residents Median 12 distinct ICD-9 diagnoses 8% 1-y Mortality	Validation C = 0.79 (0.79-0.79)	<7 7-17 >17	3 12 29
Mazzaglia et al, ⁵² 2007	15-mo index	n = 2470 Mean age 75 y 56% Female 5% 15-mo Mortality	n = 2926 Mean age 75 y 59% Female 4% 15-mo Mortality	Derivation $C = 0.75$ (0.72-0.78) Validation $C = 0.75$ (0.73-0.78)	0 (0.04-1.1) 1 (0.4-3.6) 1 (0.4-2.3) 10 (7.9-11.5)	0 (0.03-1.1) 1 (0.1-2.1) 1 (0.2-1.1) 8 (6.7-9.8)
Carey et al, ⁴⁶ 2004	2-y index	n = 4516 Mean age 78 y 61% Female 84% White 13% Dependent in ≥1 ADL 28% Difficulty with stairs 13% Diabetes 14% Cancer 31% Heart disease 10% Mortality	n = 2877 Mean age 78 y 61% Female 73% White 17% Dependent in ≥1 ADL 41% Difficulty with stairs 14% Diabetes 13% Cancer 32% Heart disease 12% 2-y Mortality	Derivation C = 0.76 Validation C = 0.74	3 11 34	5 12 36
Carey et al, ⁴⁵ 2008	3-y index for nursing-home eligible elders	n = 2232 Mean age 79 y 68% Female 62% Difficulty bathing on own 23% Diabetes 23% Coronary artery disease 37% 3-y Mortality	n = 1667 Mean age 79 y 76% Female 72% Difficulty bathing on own 27% Diabetes 27% Coronary artery disease 36% 3-y Mortality	Derivation $C = 0.66$ Validation $C = 0.69$	21 36 54	18 35 55
Lee et al, ³⁹ 2006	Lee 4-y index	n = 11 701 Mean age 67 y 57% Female 81% White 15% Diabetes 12% Cancer 17% Coronary artery disease 12% 4-y Mortality	n = 8009 Mean age 67 y 57% Female 71% White 16% Diabetes 11% Cancer 19% Coronary artery disease 13% 4-y Mortality	Derivation $C = 0.84$ Validation $C = 0.82$	<5 4-9 12-19 22-24 43-48 54-67	<5 6-9 15-20 20-28 44-45 59-64
Schonberg et al, ⁵⁵ 2009	5-y index	n = 16077 27% Age >80 y 60% Female 85% White 18% Dependent in at least 1 ADL or IADL 15% Diabetes 15% Cancer 11% Coronary artery disease 17% 5-y Mortality	n = 8038 Validation cohort reported as "similar" to development	Validation C = 0.75	2 (1-4) 8 (6-9) 25 (23-28) 47 (32-42) 71 (65-77)	3 (1-6) 8 (6-10) 29 (25-33) 49 (43-55) 62 (54-70)
Porock et al, ³⁷ 2005	6-mo index	n = 32 599 51% Age >85 y 74% Female 92% White 26% 6-mo Mortality	Nursing Home Patients n = 10 991 50% Age >85 y 73% Female 92% White 26% 6-mo Mortality	Development C = 0.75	9 23 43 62 81	10 23 43 58 82
Flacker and Kiely, ³⁶ 2003	1-y index for long-stay (>1 y) patients	n = 22749 49% Age >84 y 74% Female 83% White 9% Cancer 26% Heart disease 56% Dementia 21% Mortality	n = 40 328 46% Age >84 y 73% Female 82% White 9% Cancer 24% Heart disease 54% Dementia 22% Mortality	Derivation $C = 0.71$	8 13 31 52 76 80	9 13 31 59 79 100

(continued)

Table 3. Summary of 16 Validated General Prognostic Indices for Older Adults (continued)

		Genera	Accuracy				
					Calibi	ration ^c	
Source	Index	Source Index	Development Cohort	Validation Cohort	Discrimination (95% CI) ^b	Predicted Mortality (95% CI), % ^b	Observed Mortality (95% CI), % b
Di Bari et al, ⁴⁷ 2010	1-y index for ED triage	n = 5457 71% Age ≥80 y 55% Female 6% Cardiovascular disease 2% Respiratory disease 50% ≥5 Medications 34% 1-y Mortality	Hospital Patients n = 5456 Characteristics reported in development cohort are for all participants; random split sample validation not reported separately	Development C = 0.66 Validation C = 0.64	20 28 41 52	21 29 40 50	
Fischer et al, ⁴⁹ 2006	1-y index on admission	n = 435 Mean age 63 y 2% Female 23% Cancer 36% ≥2 Hospitalizations in past year 26% 1-y Mortality	n = 438 Characteristics reported in development cohort are for all participants; historical split sample validation not reported separately	Development C = 0.82	<18 18-48 >49	NR	
Inouye et al, ⁴⁰ 2003	1-y index on admission	n = 525 Mean age 79 y 56% Female 91% White 7% Nursing home resident 11% Pneumonia 24% Albumin ≤3.5 g/dL 27% Creatinine >1.5 mg/dL 29% 1-y Mortality	n = 1246 Average age 81 y 52% Female 94% white 32% Nursing home resident 100% Pneumonia 49% Albumin ≤3.5 g/dL 20% Creatinine >1.5 mg/dL 39% 1-y Mortality	Development C = 0.83 Validation C = 0.77	8 24 51 74	5 17 33 61	
Pilotto et al, ⁴² 2008	1-y index on admission	n = 838 Mean age 79 y 55% Female Mean 4/6 functional ADL Mean 3 errors on SPMSQ Mean 4 medications 18% 1-y Mortality	n = 857 Mean age 78 y 53% Female Mean 4/6 functional ADL Mean 3 errors on SPMSQ Mean 4 medications 17% 1-y Mortality	Development C = 0.75 (0.71-0.80)	8 21 43	6 23 45	

(continued)

midwestern United States (n=1667). The index was well calibrated but showed only moderate discrimination. Accuracy was similar for 1-year mortality.

Lee et al³⁹ developed a 4-year mortality index in community-dwelling adults older than 50 years from the eastern, western, and central United States who were interviewed in the Health and Retirement Survey of 1998 (81% participation rate, n = 11 701). To test geographic transportability, the index was validated in interviewees from the southern United States (n=8009). The Lee et al index was well calibrated and showed very good discrimination.

The index by Schonberg et al⁵⁵ to predict 5-year mortality was developed from a nationally representative sample of adults older than 65 years (n=16 077) who responded to the 1997-2000 National Health Interview Survey (NHIS) (74% participation rate);

it was well calibrated and had good discrimination in a random sample of n=8038 adults drawn from the same data source. Schonberg et al⁵⁸ then further validated the index in respondents to the 2001-2004 NHIS (n=22057, 25% aged >80 years, 57% female, 12% dependent in at least 1 instrumental activity of daily living, 18% with diabetes, 15% with cancer) and found no change in discrimination (C statistic, 0.75). The Kaplan-Meier method demonstrated widening separation between risk groups out to 9 years.

Nursing Home Residents

Two indices were developed for the nursing home, both using the Minimum Data Set (MDS), a clinical and administrative data set that is federally required of all US nursing homes. The MDS Mortality Rating Index by Porock et al³⁷ to estimate 6-month mor-

tality in nursing home patients was developed using data from all Missouri long-term care residents in 1999. Study authors later created a simplified version of this model using the same data set.53 The revised Flacker and Kiely36,50 long-stay index for 1-year mortality was developed and validated from the MDS using a split sample of nursing home residents who were 65 years and older and residing longer than 1 year in Medicare-certified nursing homes within New York (n=63077). Both indices demonstrated very good discrimination and were well calibrated across a wide range of mortality risk levels, except the revised Flacker and Kiely for the highest risk group (20% difference).

Kruse et al³⁸ prospectively validated indices by Porock et al and Flacker and Kiely in a small, prospective, single nursing home study in 2007 (n=130, mean age 83 years, 61% female, 24% dementia, 23% congestive heart failure).

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Table 3. Summary of 16 Validated General Prognostic Indices for Older Adults (continued)

		Genera	Accuracy			
					Calib	ration ^c
Source	Index	Development Cohort	Validation Cohort	Discrimination (95% CI) ^b	Predicted Mortality (95% CI), % ^b	Observed Mortality (95% CI), % b
Teno et al, ⁵⁷ 2000	1-y index on admission	n = 1266 Median age 85 y 25% Cancer 9% Congestive heart failure 61% Female 40% 1-y Mortality	Hospital Patients Validation performed using resampling from development cohort	Derivation C = 0.73 Validation C = 0.74	22 58 66 82 86 93	26 57 72 80 89 95
Levine et al, ⁵¹ 2007	1-y index after discharge	n = 2739 Mean age 78 y 63% Female 15% Discharged to SNF 23% CHF 27% COPD 8% Metastatic cancer 26% 1-y Mortality	n = 3643 Mean age 78 y 65% Female 17% Discharged to SNF 24% CHF 24% COPD 4% Metastatic cancer 26% 1-y Mortality	Derivation $C = 0.67$ Validation $C = 0.65$	14 (11-16) 18 (15-21) 32 (28-36) 46 (42-50)	14 (12-16) 24 (22-27) 30 (26-33) 42 (38-45)
Walter et al, ⁴¹ 2001	1-y index after discharge	n = 1495 Mean age 81 y 67% Female 60% White 40% Black 27% CHF 30% Discharged to a nursing home or SNF 27% Dependent in 5 ADL 10% Albumin <3.0 g/dL 40% Creatinine ≥1.5 mg/dL 33% 1-y Mortality	n = 1427 Mean age 79 y 61% Female 88% White 12% Black 29% CHF 14% Discharged to a nursing home or SNF 15% Dependent in 5 ADL 19% Albumin <3.0 g/dL 20% Creatinine ≥1.5 mg/dL 28% 1-y Mortality	Derivation $C = 0.75$ Validation $C = 0.79$	13 (10-16) 20 (16-24) 37 (33-41) 68 (63-73)	4 (2-6) 19 (15-23) 34 (29-39) 64 (58-70)
Dramé et al, ⁴⁸ 2008	2-y index after discharge	n = 870 Mean age 85 y 64% Female 60% Dependent in 1 ADL 15% Charlson comorbidity score ≥3 44% 2-y Mortality	n = 436 Mean age 85 y 64% Female 61% Dependent in 1 ADL 19% Charlson comorbidity score ≥3 44% 2-y Mortality	Derivation C = 0.72 (0.68-0.75) Validation C = 0.71 (0.66-0.75)	21 (15-26) 50 (45-54) 62 (59-71)	22 (14-29) 49 (42-55) 65 (55-76)

Abbreviations: ADLs, activities of daily living; CHF, congestive heart failure; COPD, chronic obstructive pulmonary disease; ED, emergency department; IADL, instrumental activity of daily living; ICD-9, International Classification of Diseases, Ninth Revision); NR, not reported; SNF, skilled nursing facility; SPMSQ, Short Portable Mental Status Questionnaire.

^aDescriptive information on age, sex, race, morbidity, and mortality is reported where available.

^b95% Cls are reported where available.

For the Porock et al index, discriminatory ability was lower in the validation study by Kruse et al (*C* statistic, 0.59; 95% CI, 0.46-0.72) than in the original derivation study by Porock et al (*C* statistic, 0.75) or using the simplified score (*C* statistic, 0.76). For the revised Flacker and Kiely index, discriminatory ability was the same in both the original derivation study by Flacker and Kiely (*C* statistic, 0.71) and the external validation by Kruse et al (*C* statistic, 0.72; 95% CI, 0.62-0.81).

Hospitalized Older Adults

We identified 8 indices that estimated mortality risk for hospitalized older

adults. Seven indices estimated 1-year mortality. Five were intended for use in the emergency department or on hospital admission^{40,42,47,49,57} and 3 after hospital discharge.^{41,48,51}

The "Silver Code" by Di Bari et al,⁴⁷ a 1-year index for emergency triage of individuals aged 75 years and older, was developed and validated using administrative records of patients admitted to the hospital via the emergency department from Florence, Italy, in 2005 (n=10913). They achieved 91% linkage across 4 administrative data sets (demographics, hospitalizations, prescription medications, and mortality). Random split sample validation was

conducted on half the cohort. The index was well calibrated and discriminatory ability was moderate.

Fischer et al⁴⁹ conducted a retrospective medical record review to develop a 1-year index for hospitalized elderly individuals using 4 prespecified predictors called the CARING criteria, collected at admission. Their sample included patients admitted to the medical service of a US Department of Veterans Affairs hospital in a 4-month period in 1999 (n=873). Participants admitted in the first 2 months of the study period were included in the development cohort; the remainder were in the validation cohort. The model had very

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^cPredicted and observed mortality for low-, intermediate-, and high-risk groups over a specified time period. Risk groups reported as in the original study; the number of mortality risk groups varied between studies.

Index	Sample Described (Participation) ^a	Prognostic Variables Defined ^b	Blinded Measurement ^c	Potential Predictors Complete ^d	Mortality Outcome Complete ^e	Conceptual Model, Stability Tested ^f
Gagne et al, ⁵⁶ 2011	Partly; race/ethnicity not described (participation not optional in this administrative data set)	Partly; ICD-9 codes have limited reproducibility	inity-Dwelling Patients Yes	NR	NR	Partly; stability not tested
Mazzaglia et al, ⁵² 2007	Partly; race/ethnicity not described; Italian sample (participants not compared with nonparticipants)	Partly; "inadequacy of income" not well described	Yes	NR	99%	Yes
Carey et al, ⁴⁶ 2004	Partly; no comparison of respondents with nonrespondents	Yes	Yes	99.3%	NR	Yes
Carey et al, ⁴⁵ 2008	Yes (participation not optional in this administrative data set)	Yes	Yes	92%	NR	No; not conceptually based; stability not tested
Lee et al, ³⁹ 2006	Partly; participants not compared with nonparticipants (81% participation rate)	Yes	Yes	NR	98%	Yes
Schonberg et al, ⁵⁵ 2009	Partly; participants not compared with nonparticipants (74% participation rate)	Yes	Yes	95%	97%	Yes
Porock et al, ³⁷ 2005	Partly; comorbidities not described (participation not optional in this administrative data set)	Nurs Yes	si ng Home Patients Yes	NR	>99% linkage to Missouri death certificates	Yes
Flacker and Kiely, ³⁶ 2003	Yes (participation not optional in this administrative data set)	Yes	Yes	NR	87%	Yes
		Н	lospital Patients			
Di Bari et al, ⁴⁷ 2010	Partly; race/ethnicity not described; Italian sample; "admitted for medical reasons" not clear (participation not optional in this administrative data set)	Partly; admission to "day hospital" not clearly defined	Yes	91% linkage across 4 data sets, 0% after linkage	91% linkage across 4 data sets, including mortality	No; not conceptually based; stability not tested
Fischer et al, ⁴⁹ 2006	Yes	Yes	Partly; for validation, 10% blinded medical record review with 100% agreement	NR	98%	Partly; final predictors for model selected a priori; stability not tested
Inouye et al, ⁴⁰ 2003	Partly; participants not compared with nonparticipants (86% participation rate)	Partly; ICD-9 codes have limited reproducibility	Yes	>99% for all predictors	100%	Yes
Pilotto et al, ⁴² 2008	Partly; race/ethnicity not described; Italian sample; participants not compared with nonparticipants (80% participation rate)	Yes	Yes	90%	82%	Yes

(continued)

Table 4. Potential Sources of Bias for 16 Validated General Prognostic Indices (continued)

Index	Sample Described (Participation) ^a	Prognostic Variables Defined ^b	Blinded Measurement ^c	Potential Predictors Complete ^d	Mortality Outcome Complete ^e	Conceptual Model, Stability Tested ^f
Teno et al, ⁵⁷ 2000	Partly; race/ethnicity and participation rate not described	Yes	Hospital Patients Yes	81%	100%	Yes
Levine et al, ⁵¹ 2007	Partly (participation rate not reported)	Partly; <i>ICD-9</i> codes have limited reproducibility	Yes	NR	>99%	No; not conceptually based; stability not tested
Walter et al, ⁴¹ 2001	Partly (participation rate not described)	Yes	Yes	96%	100%	Yes
Dramé et al, ⁴⁸ 2008	Partly; race/ethnicity not described; French sample (87% participation rate)	Yes	Yes	NR	92%	No; not conceptually based; stability not tested

d Completeness of predictors: percentage of sample with complete predictors.

e Completeness of mortality: percentage of sample with complete follow-up or percentage of successful linkage to vital statistics records (eg, National Death Index).

Model building: selection of potential predictors is conceptually based, and stability of model by varying assumptions and/or modeling techniques is tested (yes/partly/no/unsure).

good discrimination and a reported error rate of 0.26 in the validation cohort. Youngwerth et al59 later prospectively tested the external validity of the CARING criteria in a younger, sexbalanced sample from a university hospital in 2005 (n=427, average age 54 years, 50% female). No C statistic was reported for the external validation.

The Burden of Illness Score for Elderly Persons by Inouye et al40 updated previous indices developed by the same group^{60,66} by adding functional and laboratory data to diagnoses from administrative data to estimate 1-year mortality. Participants were drawn from a prospective study of individuals aged 70 years and older who were hospitalized at Yale-New Haven Hospital from 1989 through 1991 (n=525). The study was validated in a sample of 1246 participants from 27 Connecticut hospitals who were 65 years and older with a principal discharge diagnosis of pneumonia from 1995 through 1996. The investigators demonstrated improvement in the C statistic with the addition of laboratory and functional and cognitive measures to administrative data (validation C statistics, administrative alone, 0.59; all measures, 0.77). The model was well calibrated at the extremes but was less accurate in middlerisk groups (Table 4).

Pilotto et al⁴² used information from the standardized Geriatrics Assessment, performed at admission, to develop a 1-year prognostic index for hospitalized individuals aged 65 years and older in a sample of 838 consecutively admitted patients to the geriatrics unit of an Italian hospital in 2004, validating in 857 participants from 2005. They subsequently tested the model's accuracy at 1 year and 1 month in participants from the same hospital from 2005 to 2007 (n=4088).⁴³ The model was well calibrated and demonstrated good discrimination in the larger validation study (C statistic, 0.71; 95% CI, 0.70-0.74), and performance was similar at 1 month (C statistic, 0.76; 95% CI, 0.73 - 0.79).

Teno et al⁵⁷ developed a nomogram to predict 1- and 2-year mortality based on medicine and ICU patients aged 80 years and older who were enrolled in the Hospitalized Elder Longitudinal Project (HELP) from 5 different hospitals across the United States from 1993 to 1994 (n=1266). Teno et al tested the reproducibility of the index in 150 random samples from the original 1266 patients. The Teno et al nomogram is convenient in that it predicts multiple end points from a single score. The index includes the APACHE III scale, which requires arterial blood gas measurement.

Levine et al51 developed a 1-year prognostic model for hospitalized elderly individuals after discharge using data from a cohort of patients admitted to hospitalist and nonhospitalist physicians at the University of Chicago Hospitals from July 1997 through June 1999 (development cohort, n=2739) and July 1999 through June 2001 (validation cohort, n=3643). The index had moderate discriminatory ability and was well calibrated.

Walter et al41 developed a 1-year index for elderly individuals after hospital discharge using secondary data from a study of patients aged 70 years and older who were hospitalized between 1993 and 1997 at the University of Hospitals Cleveland (development cohort, n = 1495) and the Akron City Hospital (validation cohort, n = 1427). The model demonstrated good discrimination and was well calibrated across risk groups. Rozzini et al54 subsequently externally validated the index's performance predicting 6-month mortality in a retrospective analysis of 840 consecu-

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Abbreviations: ICD-9, International Classification of Diseases, Ninth Revision; NR, not reported.

a Sample description: study and source populations clearly defined and study sample clearly described (age, sex, race/ethnicity, comorbidities, baseline mortality rates); enrollment procedures clear and, unless administrative data, comparison of participants and nonparticipants (yes/partly/no/unsure). Participation rates provided for studies requiring consent. b Prognostic variables defined: clear, reproducible measures (yes/partly/no/unsure). ICD-9 codes rated partly due to concerns about reproducibility.⁶¹

CBlinding: developers of the prognostic index were blinded to the measurement of potential prognostic variables and mortality outcomes (yes/partly/no/unsure). Secondary analyses of existing data categorized as ves.

tively admitted participants to a hospital in Italy and found monotonic increases in mortality for each predicted risk level (observed 4%, 10%, 25%, and 46% 6-month mortality).

The Dramé et al⁴⁸ index for 2-year mortality was developed in hospitalized adults aged 75 years and older based on secondary data obtained in the emergency department as part of the SAFES study in France (n=870). It showed good calibration and discrimination in a split sample validation of 436 older adults.

COMMENT

Our review identified 16 validated nondisease-specific prognostic indices for older adults. Studies were abstracted for information about index quality, including potential for bias, generalizability, and accuracy.

We highlighted criteria for evaluating prognostic indices and identified several high-quality prognostic indices. Unfortunately, although these indices hold the promise of improving the targeting of interventions in older adults, there is insufficient evidence at this time to recommend the widespread use of prognostic indices in clinical practice. Only 2 indices were validated by investigators not involved in the studies' development, and no index had been prospectively tested and found to be accurate in a large diverse sample. Confidence intervals were not presented for either measures of discrimination or calibration for 14 indices. By our measures, no study was completely free from potential sources of bias. Testing of transportability was limited, raising concerns about overfitting and underfitting. These factors limit a clinician's ability to assess the accuracy of these indices across patient groups that differ according to severity of illness, methodology of data collection, geographic location, and time.

Even if quality barriers are overcome, important limitations remain. Several indices require collection of information that may not be routinely assessed in elderly patients, such as activities of daily living. Many of these indices rely on clinical information from administrative data sets, and the accuracy of codes from the International Classification of Diseases, Ninth Revision, has been called into question.⁶¹ Thus, indices by Gagne et al, Inouye et al, and Levine et al may be better suited to risk adjustment than clinical use. Moreover, coding algorithms are subject to change. The MDS has been updated to a new version (3.0) since the development of indices for nursing home patients, and some variables in indices by Porock et al and Flacker and Kiely have been changed or are no longer present.⁶⁷ Finally, PubMed has no single Medical Subject Heading term for prognostic index, making it difficult for a busy clinician to locate these studies.

Ultimately, an index will be judged not only on its accuracy across diverse settings, but also on its clinical effect. Studies that demonstrate effect on prognostic estimates, clinician behavior, and patient outcomes have a higher level of evidence for use in clinical decision making (eg, Ottawa ankle rules).35 We are aware of only 2 small studies that tested the effect of these indices on clinical outcomes.51,68 The highest level of evidence, however, would come from large prospective trials that randomize clinicians to using the index or not. evaluating the effect of the index on prognostic estimates, clinical decision making, and patient outcomes. Such large randomized trials have not been performed.

None of the *C* statistics for the included indices were higher than 0.90, suggesting unexplained variation in mortality. However, discriminatory ability of these indices is consistent with other indices that commonly drive clinical decisions, such as the CHADS2 index to help determine warfarin therapy (*C* statistic, 0.68-0.72)⁶⁹; the Framingham risk score to help determine lipid therapy (*C* statistic, 0.63-0.83)⁷⁰; and the TIMI risk score to help determine invasive therapy for unstable angina (*C* statistic, 0.65).⁷¹

There may be a limited role for the highest-quality indices in the right set-

tings. If patient characteristics align closely with those of the development or validation cohorts, clinicians may find prognostic information useful to help inform, though not replace, their clinical judgment. Prediction rules have been shown to outperform clinicians in terms of prognostication, 72,73 whereas human prediction on its own is fraught with bias.74 The indices we identified were developed from heterogeneous groups of patients. Applying this information to the individual patient, therefore, requires a nuanced use of the index. Patients are likely to have conditions that are not included in the index (eg, Parkinson disease). The clinician must account for these conditions and decide whether their effect is adequately accounted for by the indices' predictors.

Indices are most likely to be clinically useful when they predict a wide range of mortality. Clinical decisions are most likely to be influenced by either very low or very high mortality risk. Although 10 indices predicted greater than 50% mortality, only 3 predicted greater than 80% risk in the highest risk group. Midrange probabilities may still be useful in clinical decisions in which life expectancy plays a role, allowing patient preferences to drive the physician's recommendation. The following case illustrates this issue.

Case

Ms A is a 75-year-old clinic patient who has been hospitalized twice in the past year for chronic obstructive pulmonary disease and has a history of diabetes and difficulty walking a quarter mile. She has not been previously screened for colon cancer. The US Preventive Services Task Force recommends that individual factors should determine the decision to screen or not screen patients aged 75 to 85 years; patients must live at least 7 years to benefit from screening, and the net benefits in this age group are small. 16 Using indices developed for communitydwelling elderly individuals, it is determined that Ms A has a 54% to 67% mortality risk at 4 years (Lee et al in-

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dex) and 75% at 9 years (Schonberg et al index). Should Ms A undergo colorectal cancer screening?

In this case, the prognostic information may be helpful as her physician discusses the possibility of colon cancer screening in relation to other health priorities, such as maintaining mobility. Because her median life expectancy is less than 4 years, Ms A will probably not live long enough to benefit from screening. And if screening is difficult for her, there is enough uncertainty in her likelihood of benefit that she probably should focus on other priorities. However, if she feels strongly about wanting to be screened, the estimates are not strong enough on their own to refute that decision.

Limitations

We have refrained from explicitly ranking or categorizing the quality of these indices, recognizing that no agreed-on scientifically developed system for rating index quality currently exists. Some will argue that minimizing risk for potential bias is of critical importance, while others might argue that an index should be judged on its ability to perform accurately across diverse settings. Our review excluded indices that estimated only relative risk or had not been validated, and future research may find that some of these indices are generalizable and accurate. Our ability to assess publication bias was limited by our small sample size.

CONCLUSION

While neither a clinician nor an index can predict with absolute certainty how long an older adult will live, validated prognostic indices might improve the accuracy of the prognostic assumptions that influence clinical decisions. However, further research is needed before general prognostic indices for elderly individuals can be recommended for routine use. Future research should focus on prospectively testing the validity of these indices across diverse clinical settings and analyzing their effect on clinical decision making and patient outcomes.

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Study concept and design: Yourman, Lee, Schonberg, Widera, Smith.

Acquisition of data: Yourman, Smith.

Analysis and interpretation of data: Yourman, Lee, Schonberg, Smith.

Drafting of the manuscript: Yourman, Lee, Smith. Critical revision of the manuscript for important intellectual content: Yourman, Lee, Schonberg, Widera,

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