# Randomization Inference with Natural Experiments: An Analysis of Ballot Effects in the 2003 California Recall Election

Kosuke Imai

Department of Politics Princeton University

Joint work with Daniel E. Ho, Yale Law School

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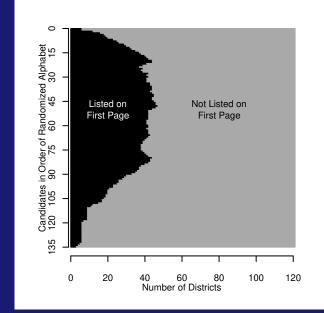
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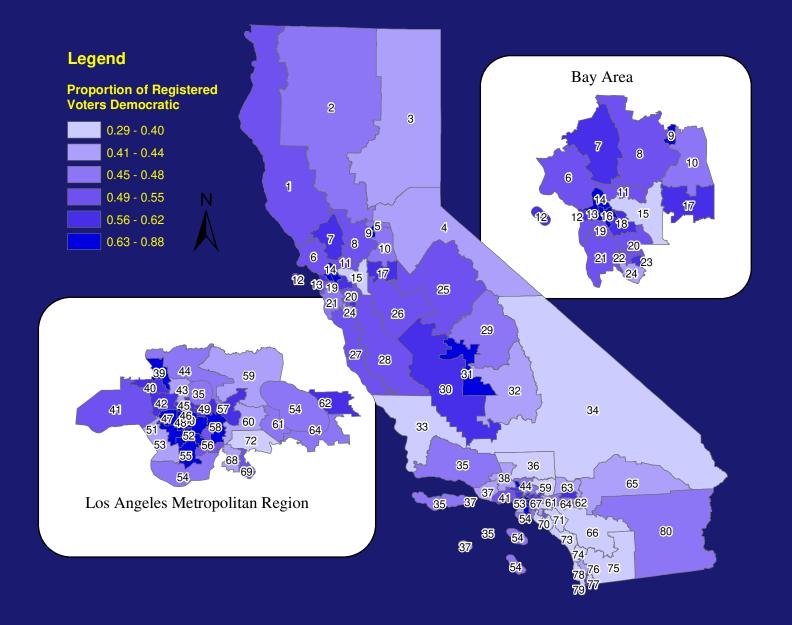
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# California Assembly Districts by Percentage of Registered Democrats



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  - $\star~T_i=1$  if the candidate is listed on the first page, and  $T_i=0$  otherwise.
  - ★  $Y_i = Y_i(1)T_i + (1 T_i)Y_i(0)$ 
    - $Y_i(1)$ : potential vote share when the candidate is placed on the first page.
    - $Y_i(0)$ : potential vote share when the candidate is not placed on the first page.
  - $\star~t_i$  and  $y_i:$  observed values of  $T_i$  and  $Y_i.$
  - \* Unit ballot page effect:  $\tau_i \equiv Y_i(1) Y_i(0)$ .
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    - ★ Sample average ballot effect:

$$W^{D}(T) = \frac{\sum_{i=1}^{121} T_{i} y_{i}}{N_{1}} - \frac{\sum_{i=1}^{121} (1 - T_{i}) y_{i}}{N_{0}},$$

corresponding to the difference-in-means estimator.

$$W^{L}(T) = (T^{\top}MT)^{-1}T^{\top}My$$

corresponding to the linear least squares estimator, where  $y = (y_1, y_2, \ldots, y_{121})$ ,  $M = I - X(X^T X)^{-1}X^T$ , and X is the matrix of the observed pretreatment covariates.

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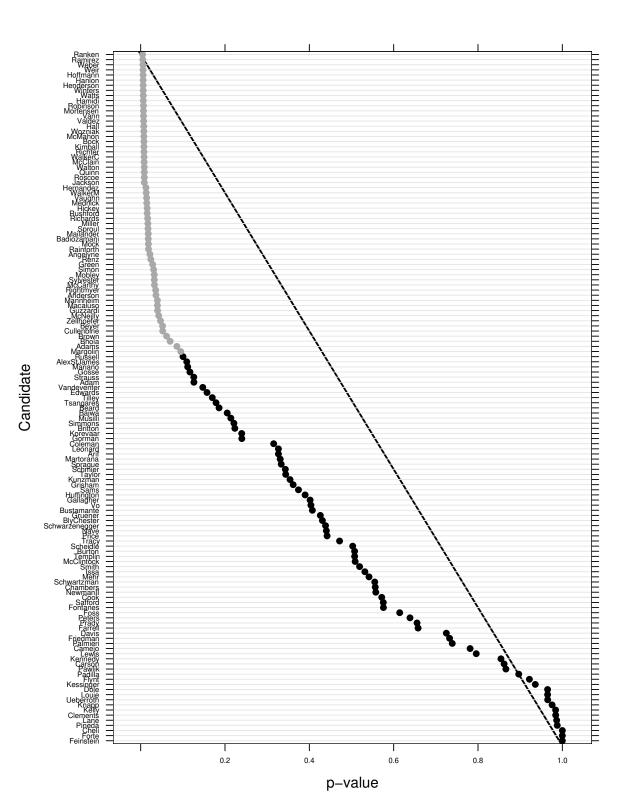
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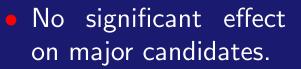
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Since the number of permutations is large, we use Monte Carlo approximation,

$$\Pr(W^{\mathsf{D}}(\mathsf{T}) \ge W^{\mathsf{D}}(\mathsf{t})) \approx \frac{1}{\mathfrak{m}} \sum_{j=1}^{\mathfrak{m}} \operatorname{I}(W^{\mathsf{D}}(\mathsf{T}^{(j)}) \ge W^{\mathsf{D}}(\mathsf{t})),$$

with m = 10,000.





 Positive ballot effect on 40% of minor candidates.

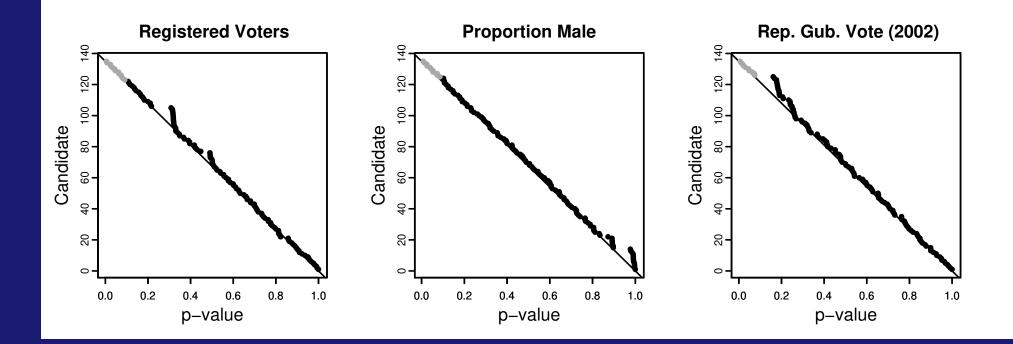
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- 2. Given the null value  $\tau_0$ , the test statistic is given by

$$W_{\tau_0}^{D}(T) = \frac{\sum_{i=1}^{121} T_i \{y_i + (1 - t_i)\tau_0\}}{\sum_{i=1}^{121} T_i} - \frac{\sum_{i=1}^{121} (1 - T_i)(y_i - t_i\tau_0)}{\sum_{i=1}^{121} (1 - T_i)},$$

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or its covariance-adjusted analogue

$$W_{\tau_0}^{L}(\mathsf{T}) = (\mathsf{T}^{\top}\mathsf{M}\mathsf{T})^{-1}\mathsf{T}^{\top}\mathsf{M}\mathfrak{y}^*,$$

where each element of  $y^*$  is  $y_i^* = T_i \{y_i + (1-t_i)\tau_0\} + (1-T_i)(y_i - t_i\tau_0).$ 

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3. Two-tailed level  $\alpha$  test; accept H<sub>0</sub> if

$$t \in A_{\alpha}(\tau_0) = \left\{ u : \frac{\alpha}{2} \leq \Pr(W^{D}_{\tau_0}(\mathsf{T}) \geq W^{D}(u)) \leq 1 - \frac{\alpha}{2} \right\},\$$

and reject  $H_0$  otherwise.

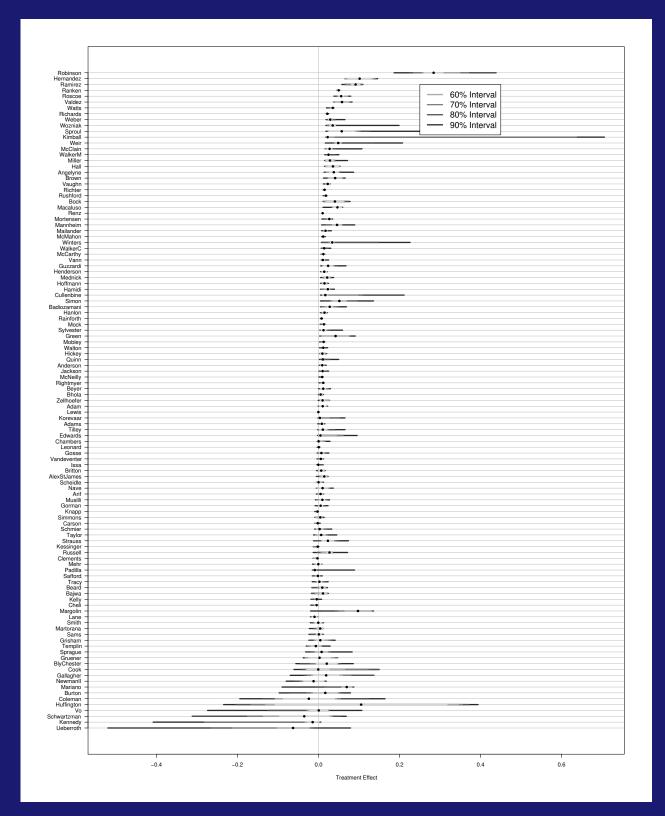
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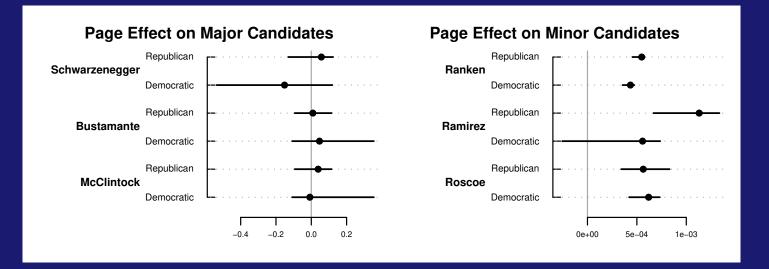
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- Nonparametric estimates of CDF (for the sampling distributions of causal effect estimators) can also be obtained by estimating  $\tau_{\rm U}$  and  $\tau_{\rm L}$  for different values of  $\alpha \in [0, 0.5]$ .



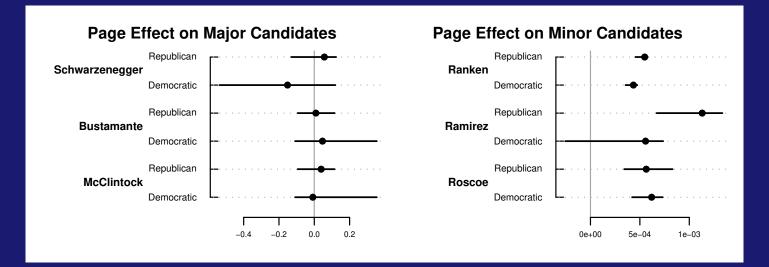
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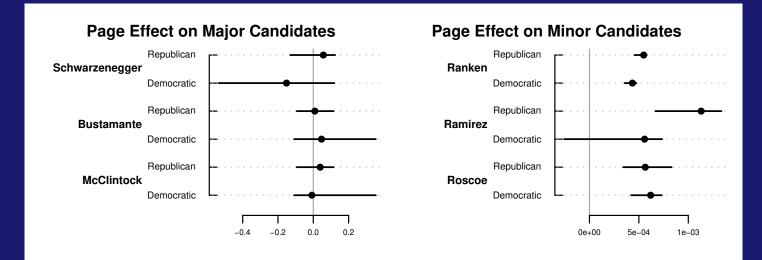
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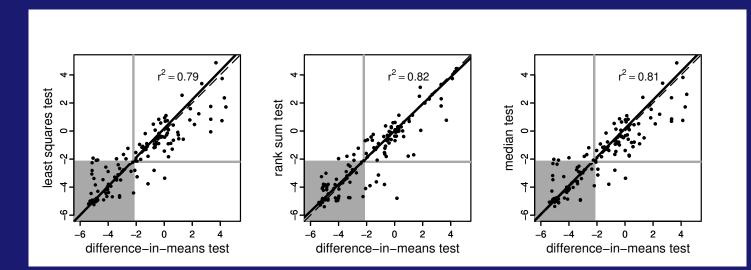
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#### **Sensitivity Analyses**

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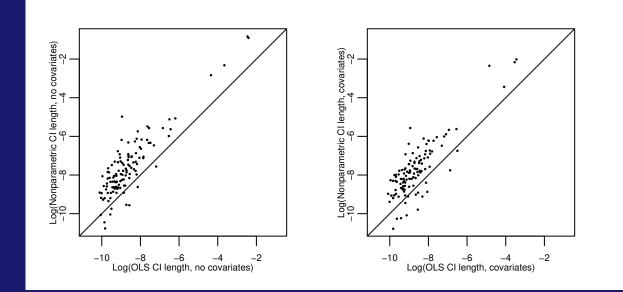
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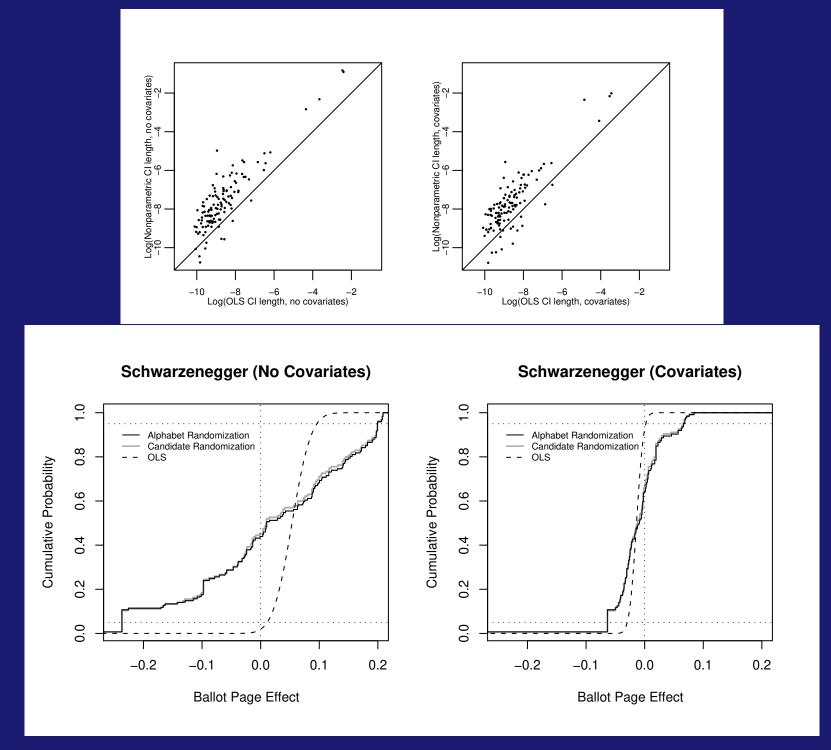
	Without Covariates			With Covariates		
	OLS	GLM logit	RI Fisher	OLS	GLM logit	RI Fisher
Major Candidate	S					
Schwarzenegger	1.09 9.63	-1.23 7.53	-23.72 19.90	-2.97 0.21	-4.98 - 2.05	-6.44 6.87
Bustamante -	-8.46 0.54 -	-5.37 4.04	-20.07 20.31	-1.12 1.78	0.96 3.01	-5.86 5.64
McClintock	0.50 3.09	-1.10 1.24	-3.47 6.36	1.56 3.25	0.29 2.05	0.36 3.57
All Candidates						
Positive effects	56(41%)	63(47%)	55(41%)	50(37%)	59(44%)	47(35%)
Negative effects	11 (8%)	8 (6%)	4 (3%)	8 (6%)	17(13%)	2 (1%)
Null effects	68(50%)	64(47%)	59(44%)	77(57%)	59(44%)	64(47%)
Unidentified	0 (0%)	0 (0%)	17(13%)	0 (0%)	0 (0%)	22(16%)
Comparison with Randomization Inference						
Agreement	89(66%)	87(64%)	108(80%)	88(65%)	74(55%)	108(80%)

# **Comparisons of Confidence Intervals and CDF**

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- Parametric inferences can be sensitive to modeling and other assumptions.
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- The randomization inference framework can directly incorporate complex randomization schemes in natural experiments.