

RATE OF APPROXIMATION OF FUNCTIONS OF BOUNDED  
VARIATION BY MODIFIED LUPAS OPERATORS

WANG YUANKWEI AND GUO SHUNSHENG

This paper discusses the rate of approximation of functions of bounded variation using the Modified Lupas operator. We obtain an approximation theorem and our estimate is essentially the best possible.

1. INTRODUCTION

Let  $f$  be a function defined on  $(0, \infty)$  with bounded variation in each finite interval and  $f(x) = O(x^r)$   $x \rightarrow \infty$ . We denote  $\{f\}$  by  $BV_{loc,r}(0, \infty)$ . The Modified Lupas operator  $M_n$  applied to  $f$  is

$$(1.1) \quad M_n(f, x) = (n-1) \sum_{k=0}^{\infty} P_{nk}(x) \int_0^{\infty} P_{nk}(t) f(t) dt$$

where

$$P_{nk}(t) = \binom{n+k-1}{k} \frac{t^k}{(1+t)^{n+k}}.$$

It is also written as

$$(1.2) \quad \int_0^{\infty} H_n(x, t) f(t) dt$$

where

$$H_n(x, t) = (n-1) \sum_{k=0}^{\infty} P_{nk}(x) P_{nk}(t).$$

This kernel is positive and

$$(1.3) \quad \int_0^{\infty} H_n(x, t) dt = 1,$$

therefore  $M_n(f, x)$  is linear positive [1].

In this paper, the main result is the following:

---

Received 5th September 1990

---

Copyright Clearance Centre, Inc. Serial-fee code: 0004-9729/91 \$A2.00+0.00.

**THEOREM.** *Let  $f$  be a function belonging to  $BV_{loc,r}(0, \infty)$ . Then for every  $x \in (0, \infty)$  and  $n$  sufficiently large, we have*

$$(1.4) \quad \left| M_n(f, x) - \frac{1}{2}(f(x+) - f(x-)) \right| \leq \frac{5(1+x)}{nx} \sum_{k=1}^n \bigvee_{x-\frac{x}{\sqrt{k}}}^{x+\frac{x}{\sqrt{k}}} (g_x) + \frac{50}{\sqrt{n}} \left( \frac{1+x}{x} \right)^{3/2} |f(x+) - f(x-)| + \frac{(1+x)^r}{x^4} O\left(\frac{1}{n^2}\right),$$

where

$$g_x(t) = \begin{cases} f(t) - f(x+) & x < t < \infty, \\ 0 & t = x, \\ f(t) - f(x-) & 0 \leq t < x. \end{cases}$$

The proof of the theorem will use the Bojanic method [2] and some results of probability theory.

2. LEMMAS

**LEMMA 1.** *If  $\{\xi_k\}$  ( $k \geq 1$ ) are independent random variables with the same distribution functions and  $0 < D\xi_k < \infty$ ,  $\beta_3 = E(\xi_r - \xi_i)^3 < \infty$ , then*

$$(2.1) \quad \max_y \left| P\left(\frac{1}{b_1\sqrt{n}} \sum_{k=1}^n (\xi_k - a_1) \leq y\right) - \frac{1}{\sqrt{2\pi}} \int_{-\infty}^y e^{-t^2/2} dt \right| < \frac{c}{\sqrt{n}} \frac{\beta_3}{b_1^3}$$

where  $a_1 = E(\xi_1)$  (expectation of  $\xi_1$ ),  $b_1^2 = D\xi_1 = E(\xi_1 - E\xi_1)^2$  and  $1/\sqrt{2\pi} \leq c < 0.82$  (see [3]).

**LEMMA 2.** [3] *If  $\{\xi_i\}$  are independent random variables with the same geometric distribution functions  $P(\xi_i = k) = x^k(1-x)$ ,  $i = 1, 2, \dots$ , then  $E\xi_i = x/(1-x)$ ,  $D\xi_i = x/(1-x)^2$ ,  $\eta_n = \sum_{i=1}^n \xi_i$  with*

$$(2.2) \quad P(\eta_n = k) = \binom{n+k-1}{k} x^k(1-x)^n.$$

**LEMMA 3.** *For every  $x \in (0, +\infty)$ ,  $k \in \mathbb{N}$ , we have*

$$(2.3) \quad P_{nk}(x) \leq \frac{33}{\sqrt{n}} \left(\frac{1+x}{x}\right)^{3/2}.$$

**PROOF:** Since

$$P_{nk}(x) = \binom{n+k-1}{k} \frac{x^k}{(1+x)^{n+k}} \quad x \in (0, \infty),$$

we take  $t = x/(1 + x)$ ; then  $t \in (0, 1)$  and

$$(2.4) \quad P_{nk}(x) = \binom{n+k-1}{k} t^k (1-t)^n.$$

Using Lemma 2 we have

$$\binom{n+k-1}{k} t^k (1-t)^n = P(k-1 < \eta_n \leq k) = P(w(k-1, n) < w(\eta_n, n) \leq w(k, n))$$

where  $w(k, n) = (x(1-t) - nt)/\sqrt{nt}$ . Using Lemma 1 we have  $a_1 = t/(1-t)$ ,  $b_1 = \sqrt{t}/(1-t)$ . Hence

$$(2.5) \quad \left| \binom{n+k-1}{k} t^k (1-t)^n - \frac{1}{\sqrt{2\pi}} \int_I e^{-t^2/2} dt \right| < 2 \frac{\beta_3}{\sqrt{n}(\sqrt{t}/(1-t))^3}$$

where  $I = [w(k-1, n), w(k, n)]$ ,

$$\begin{aligned} \beta_3 &= E\left(\xi_k - \frac{t}{1-t}\right)^3 = \sum_{k=0}^{\infty} \left(k - \frac{t}{1-t}\right)^3 t^k (1-t) \\ &\leq \sum_{k=0}^{\infty} \left[ k^3 + 3k^2 \frac{t}{1-t} + 3k \left(\frac{t}{1-t}\right)^2 + \left(\frac{t}{1-t}\right)^3 \right] t^k (1-t). \end{aligned}$$

Since

$$\begin{aligned} \sum_{k=0}^{\infty} t^k (1-t) &= 1, & \sum_{k=0}^{\infty} k t^k (1-t) &= \frac{t}{1-t}, \\ \sum_{k=0}^{\infty} k^2 t^k (1-t) &= \frac{t(1+t)}{1-t^2}, & \sum_{k=0}^{\infty} k^3 t^k (1-t) &= \frac{t^3 + 4t^2 + t}{(1-t)^3}, \end{aligned}$$

therefore  $\beta_3 \leq 16/(1-t)^3$ .

But the second term on the left side of (2.5) is not greater than  $(1-t)/\sqrt{2\pi nt}$ . Hence we have

$$\binom{n+k-1}{k} t^k (1-t)^n \leq \frac{33}{\sqrt{nt^{3/2}}};$$

therefore  $P_{nk}(x) \leq \frac{33}{\sqrt{n}} \left(\frac{1+x}{x}\right)^{3/2}$ , and (2.3) is proved. □

**LEMMA 4.** For every  $k \geq 0$  we have

$$\sum_{j=0}^k P_{n-1,j}(x) = (n-1) \int_x^{\infty} P_{nk}(t) dt.$$

PROOF: This can easily be proved by differentiating both the left-hand and right-hand sides. □

LEMMA 5. If  $n$  is sufficiently large, for every  $k \geq 0$  we have

$$\left| \sum_{j=0}^k P_{n-1,j}(x) - \sum_{j=0}^k P_{nj}(x) \right| \leq \frac{33}{\sqrt{n}} \left( \frac{1+x}{x} \right)^{3/2}.$$

PROOF: From (2.2) we have  $\sum_{j=0}^k P_{nj}(x) = P(\eta_n \leq k)$ . Again using Lemma 1 we have

$$\left| \sum_{j=0}^k P_{nj}(x) - \frac{1}{\sqrt{2\pi}} \int_{-\infty}^{w(k,n)} e^{-u^2/2} du \right| < \frac{16}{\sqrt{n}} \left( \frac{1+x}{x} \right)^{3/2};$$

hence

$$\begin{aligned} & \left| \sum_{j=0}^k P_{n-1,j}(x) - \sum_{j=0}^k P_{nj}(x) \right| \\ & < \left| \frac{1}{\sqrt{2\pi}} \int_{w(k,n)}^{w(k,n-1)} e^{-u^2/2} du \right| + \frac{32}{\sqrt{n}} \left( \frac{1+x}{x} \right)^{3/2} \leq \frac{33}{\sqrt{n}} \left( \frac{1+x}{x} \right)^{3/2}. \end{aligned}$$

□

LEMMA 6. If  $n$  is sufficiently large and  $x \in (0, \infty)$ , then

$$(2.8) \quad \frac{x(1+x)}{n} \leq M_n((t-x)^2, x) \leq \frac{3x(1+x)}{n};$$

$$(2.9) \quad M_n((t-x)^4, x) = O\left(\frac{1}{n^2}\right).$$

PROOF: Let

$$(2.10) \quad T_{nm} = (n-r+1) \sum_{k=0}^{\infty} P_{n+r,k}(x) \int_0^{\infty} P_{n-r,k+r}(t)(t-x)^m dt.$$

Then

$$\begin{aligned} T_{n2} &= (n-r+1) \sum_{k=0}^{\infty} P_{n+r,k}(x) \int_0^{\infty} P_{n+r,k}(x) \int_0^{\infty} P_{n-r,k+r}(t)(t-x)^2 dt \\ &= \frac{2(n-1)x(1+x)}{(n-r-2)(n-r-3)} + \frac{(r+1)(r+2)(1+2x)^2}{(n-r-2)(n-r-3)}. \end{aligned}$$

Putting  $r = 0$  we have

$$(2.11) \quad T_{nm} = (n + 1) \sum_{k=0}^{\infty} P_{nk}(x) \int_0^{\infty} P_{nk}(t)(t - x)^m dt = O\left(\frac{1}{n^{[(m+1)/2]}}\right),$$

$$T_{n2} = \frac{2(n - 1)x(1 + x)}{(n - 2)(n - 3)} + \frac{2(1 + 2x)^2}{(n - 2)(n - 3)}.$$

From this, (2.8) and (2.9) can be proved (see [4]). □

**LEMMA 7.** *In  $n$  is sufficiently large and  $x \in (0, \infty)$ , then for  $0 \leq y < x$ , we have*

$$(2.12) \quad \int_0^y H_n(x, t) dt \leq \frac{3x(1 + x)}{n(x - y)^2};$$

for  $x < z < \infty$ , we have

$$(2.13) \quad \int_z^{\infty} H_n(x, t) dt \leq \frac{3x(1 + x)}{n(z - x)^2}.$$

**PROOF:** Since  $0 \leq y < x$ , for  $t \in [0, y]$  we have  $(x - t)/(x - y) \geq 1$ .

From

$$M_n(f, x) = \int_0^{\infty} H_n(x, t) f(t) dt$$

we have 
$$\frac{x(1 + x)}{n} \leq \int_0^{\infty} H_n(x, t)(t - x)^2 dt \leq \frac{3x(1 + x)}{n};$$

therefore 
$$\begin{aligned} \int_0^y H_n(x, t) dt &\leq \int_0^y \left(\frac{x - t}{x - y}\right)^2 H_n(x, t) dt \\ &\leq \frac{1}{(x - y)^2} \int_0^{\infty} (x - t)^2 H_n(x, t) dt \leq \frac{3x(1 + x)}{n(x - y)^2}, \end{aligned}$$

proving (2.12). The proof of (2.13) is similar. □

### 3. PROOF OF THE THEOREM

Now

$$f(t) = \frac{f(x+) + f(x-)}{2} + g_x(t) + \frac{f(x+) - f(x-)}{2} \text{sign}(t - x);$$

hence

$$(3.1) \quad \begin{aligned} &\left| M_n(f, x) - \frac{1}{2}(f(x+) + f(x-)) \right| \\ &\leq |M_n(g_x, x)| + \frac{1}{2} |f(x+) - f(x-)| |M_n(\text{sign}(t - x), x)|. \end{aligned}$$

Thus to estimate  $|M_n(f, x) - (f(x+) - f(x-))/2|$  we need an estimate for  $M_n(g_x, x)$  and  $M_n(\text{sign}(t - x), x)$ .

To estimate  $M_n(\text{sign}(t - x), x)$ , we first decompose it into two parts as follows:

$$M_n(\text{sign}(t - x), x) = \int_0^\infty \text{sign}(t - x)H_n(x, t)dt = \int_x^\infty H_n(x, t)dt - \int_0^x H_n(x, t)dt \stackrel{\text{def}}{=} A_n(x) - B_n(x).$$

Using Lemma 4, we have

$$\begin{aligned} A_n(x) &= \int_x^\infty H_n(x, t)dt = \int_x^\infty (n - 1) \sum_{k=0}^\infty P_{nk}(x)P_{nk}(t)dt \\ &= (n - 1) \sum_{k=0}^\infty P_{nk}(x) \int_x^\infty P_{nk}(t)dt \\ &= \sum_{k=0}^\infty P_{nk}(x) \sum_{j=0}^k P_{n-1,j}(x). \end{aligned}$$

By Lemma 5, it follows that

$$\begin{aligned} (3.2) \quad & \left| A_n(x) - \sum_{k=0}^\infty P_{nk}(x) \sum_{j=0}^k P_{nj}(x) \right| \\ &= \left| \sum_{k=0}^\infty P_{nk}(x) \sum_{j=0}^k P_{nj}(x) - \sum_{k=0}^\infty P_{nk}(x) \sum_{j=0}^k P_{n-1,j}(x) \right| \\ &\leq \sum_{k=0}^\infty P_{nk}(x) \frac{33}{\sqrt{n}} \left( \frac{1+x}{x} \right)^{3/2} \leq \frac{33}{\sqrt{n}} \left( \frac{1+x}{x} \right)^{3/2}. \end{aligned}$$

Let

$$\begin{aligned} S &= \sum_{k=0}^\infty \left( P_{nk}(x) \sum_{j=0}^k P_{nj}(x) \right) \\ &= P_{n0}P_{n0} + P_{n1}(P_{n0} + P_{n1}) + \dots + P_{nn}(P_{n0} + \dots + P_{nn}) + \dots \end{aligned}$$

Since

$$\begin{aligned} 1 &= (P_{n0} + P_{n1} + \dots + P_{nm} + \dots)(P_{n0} + P_{n1} + \dots + P_{nm} + \dots) \\ &= P_{n0}(P_{n0} + P_{n1} + \dots + P_{nm} + \dots) + P_{n1}(P_{n0} + \dots) + P_{nm}(P_{n0} + \dots) + \dots, \end{aligned}$$

we have

$$\begin{aligned}
 1 - S &= P_{n0}(P_{n1} + P_{n2} + \dots) + P_{n1}(P_{n2} + P_{n3} + \dots) \\
 &\quad + P_{nm}(P_{n,m+1} + P_{n,m+2} + \dots) + \dots \\
 &= P_{n1}P_{n0} + P_{n2}(P_{n0} + P_{n1}) + \dots \\
 &\quad + P_{nm}(P_{n0} + P_{n1} + \dots + P_{n,m-1}) + \dots,
 \end{aligned}$$

and  $2S - 1 = P_{n0}^2 + P_{n1}^2 + \dots + P_{nm}^2 + \dots$ .

Using Lemma 3, we obtain

$$\begin{aligned}
 (3.3) \quad \left| S - \frac{1}{2} \right| &= \frac{1}{2}(P_{n0}^2 + P_{n1}^2 + \dots) + \frac{1}{2} \sum_{k=0}^{\infty} P_{nk}^2 \\
 &\leq \frac{1}{2} \sum_{k=0}^{\infty} P_{nk} \left[ \frac{33}{\sqrt{n}} \left( \frac{1+x}{x} \right)^{3/2} \right] \leq \frac{33}{2\sqrt{n}} \left( \frac{1+x}{x} \right)^{3/2}.
 \end{aligned}$$

By (3.2), (3.3) and  $B_n(x) = 1 - A_n(x)$ , we have

$$|A_n(x) - B_n(x)| = |2A_n(x) - 1| \leq \frac{100}{\sqrt{n}} \left( \frac{1+x}{x} \right)^{3/2}.$$

Hence

$$(3.4) \quad |M_n(\text{sign}(t-x), x)| \leq \frac{100}{\sqrt{n}} \left( \frac{1+x}{x} \right)^{3/2}.$$

The estimate of  $M_n(g_x, x)$  is similar to [4]. We first decompose  $[0, +\infty)$  into three parts, as follows:

$$I_1 = \left[ 0, x - \frac{x}{\sqrt{n}} \right], \quad I_2 = \left[ x - \frac{x}{\sqrt{n}}, x + \frac{x}{\sqrt{n}} \right], \quad I_3 = \left[ x + \frac{x}{\sqrt{n}}, \infty \right).$$

Then

$$\begin{aligned}
 M_n(g_x, x) &= \int_0^{\infty} g_x(t)H_n(x, t)dt = \left( \int_{I_1} + \int_{I_2} + \int_{I_3} \right) g_x(t)H_n(x, t)dt \\
 &\stackrel{\text{def}}{=} \Delta_{1,n}(f, x) + \Delta_{2,n}(f, x) + \Delta_{3,n}(f, x).
 \end{aligned}$$

First, we estimate  $\Delta_{2,n}(f, x)$ . For  $t \in I_2$ , we have

$$|g_x(t)| = |g_x(t) - g_x(x)| \leq \bigvee_{x-z/\sqrt{n}}^{x+z/\sqrt{n}} (g_x)$$

and so

$$(3.5) \quad |\Delta_{2,n}(f, x)| \leq \int_{x-z/\sqrt{n}}^{x+z/\sqrt{n}} (g_x) \int_{x-z/\sqrt{n}}^{x+z/\sqrt{n}} H_n(x, t) dt \leq \int_{x-z/\sqrt{n}}^{x+z/\sqrt{n}} (g_x).$$

Secondly, we estimate  $\Delta_{1,n}(f, x)$ . Let  $\lambda_n(x, t) = \int_0^t H_n(x, u) du$ . We have

$$\begin{aligned} |\Delta_{1,n}(f, x)| &= \left| \int_0^{x-z/\sqrt{n}} g_x(t) H_n(x, t) dt \right| = \left| \int_0^y g_x(t) H_n(x, t) dt \right| \\ &= \left| \int_0^y g_x(t) d_t \lambda_n(x, t) \right| \\ &= \left| g_x(y+) \lambda_n(x, y) - \int_0^y \lambda_n(x, t) d_t g_x(t) \right| \\ &\leq \int_{y+}^x (g_x) \lambda_n(x, y) + \int_0^y \lambda_n(x, t) d_t \left( - \int_t^x (g_x) \right), \end{aligned}$$

where

$$y = x - \frac{x}{\sqrt{n}}.$$

By Lemma 7, we have

$$|\Delta_{1,n}(f, x)| \leq \int_{y+}^x (g_x) \frac{3x(1+x)}{n(x-y)^2} + \frac{3x(1+x)}{n} \int_0^y \frac{1}{(x-t)^2} d_t \left( - \int_t^x (g_x) \right)$$

where  $n$  is sufficiently large.

Since

$$\int_0^y \frac{1}{(x-t)^3} d_t \left( - \int_t^x (g_x) \right) = - \frac{V_{y+}^x (g_x)}{(x-y)^2} + \frac{V_0^x (g_x)}{x^2} + 2 \int_0^y \int_t^x (g_x) \frac{dt}{(x-t)^3},$$

we have

$$\begin{aligned} |\Delta_{1,n}(f, x)| &\leq \frac{3x(1+x)}{n} \left[ \int_{y+}^x (g_x) / (x-y)^2 + \int_0^y \frac{1}{(x-t)^2} d_t \left( - \int_t^x (g_x) \right) \right] \\ &\leq \frac{3x(1+x)}{n} \left[ \int_0^x (g_x) / x^2 + 2 \int_0^{x-z/\sqrt{n}} \int_t^x (g_x) \frac{dt}{(x-t)^3} \right]. \end{aligned}$$

Furthermore, since

$$\int_0^{x-\frac{x}{\sqrt{n}}} \int_t^x (g_x) \frac{dt}{(x-t)^2} + \int_1^n \int_{x-z/\sqrt{T}}^x (g_x) \frac{1}{x^3/T^{3/2}} \frac{1}{2} x \frac{1}{T^{3/2}} dT \leq \frac{1}{2x^2} \sum_{k=1}^n \int_{x-z/\sqrt{k}}^x (g_x),$$



it follows that

$$(3.6) \quad |\Delta_{1,n}(f, x)| \leq \frac{3(1+x)}{nx} \left[ \bigvee_0^x (g_x) + \sum_{k=1}^n \bigvee_{x-x/\sqrt{k}}^x (g_x) \right].$$

Last, we estimate  $\Delta_{3,n}(f, x)$ . Since

$$\begin{aligned} |\Delta_{3,n}(f, x)| &= \left| \int_{x+x/\sqrt{n}}^{\infty} g_x(t) H_n(x, t) dt \right| \\ &= \left[ \int_{x+x/\sqrt{n}}^{2x} + \int_{2x}^{\infty} \right] g_x(t) H_n(x, t) dt \stackrel{\text{def}}{=} R_{1,n} + R_{2,n}, \end{aligned}$$

using similar methods as above, we have

$$(3.7) \quad |R_{1,n}| \leq \frac{4(1+x)}{nx} \sum_{k=1}^n \bigvee_x^{x+x/\sqrt{k}} (g_x).$$

Since  $g_x(t) = 0(t^r)$  ( $t \rightarrow \infty$ ), there exists  $M > 0$  such that

$$\begin{aligned} (3.8) \quad |R_{2,n}| &\leq M \int_{2x}^{\infty} t^r H_n(x, t) dt \leq M \int_{|t-x| \geq x} t^r H_n(x, t) dt \\ &\leq \frac{M}{x^4} \sum_{k=0}^{\infty} (n-1) \int_0^{\infty} P_{nk}(x) P_{nk}(t) t^r (t-x)^4 dt. \end{aligned}$$

From

$$\begin{aligned} &\frac{(n-1)P_{nk}(t)P_{nk}(x)t^r}{(n-r-1)P_{n+r,k}(x)P_{n-r,k+r}(t)} \\ &= (1+x)^r \frac{(k+1) \cdots (k+r)(n-1)n \cdots (n+r-1)}{(k+n) \cdots (k+n+r-1)(n-r-1) \cdots (n-1)} \end{aligned}$$

it follows that, for every  $k$  and  $n > r + 1$ , we have

$$\frac{(k+1) \cdots (k+r)}{(k+n) \cdots (k+n+r-1)} < 1, \quad \lim_n \frac{(n-1) \cdots (n+r-1)}{(n-r-1) \cdots (n-1)} = 1.$$

Hence if  $n$  is sufficiently large, then

$$(n-1)P_{nk}(x)P_{nk}(t)t^r \leq (1+x)^r (n-r-1)P_{n+r,k}(x)P_{n-r,k+r}(t).$$

Replacing above in (3.6), we have

$$\begin{aligned}
 |R_{2n}| &\leq \frac{M}{x^4} \sum_{k=0}^{\infty} \int_0^{\infty} (1+x)^r (n-r+1) P_{n+r,k}(x) P_{n-r,k+r}(t) (t-x)^4 dt \\
 &\leq \frac{M}{x^4} (1+x)^r (n-r+1) \sum_{k=0}^{\infty} P_{n+r,k}(x) \int_0^{\infty} P_{n-r,k+r}(t) (t-x)^4 dt \\
 &= \frac{M(1+x)^r}{x^4} T_{n^4}.
 \end{aligned}$$

Using (2.11), we obtain

$$(3.9) \quad |R_{2n}| \leq \frac{(1+x)^r}{x^4} O\left(\frac{1}{n^2}\right).$$

From (3.5), (3.6), (3.7) and (3.9), it follows that

$$\begin{aligned}
 (3.10) \quad |M_n(g_x, x)| &\leq \bigvee_{z-x/\sqrt{n}}^{x+z/\sqrt{n}} (g_x) + \frac{3(1+x)}{nx} \left[ \bigvee_0^x (g_x) + \sum_{k=1}^n \bigvee_{z-x/\sqrt{k}}^z (g_x) \right] \\
 &\quad + \frac{4(1+x)}{nx} \sum_{k=1}^n \bigvee_z^{x+z/\sqrt{k}} (g_x) + \frac{(1+x)^r}{x^4} O\left(\frac{1}{n^2}\right),
 \end{aligned}$$

where  $n$  is sufficiently large.

Our theorem now follows from (3.1), (3.4), (3.10); that is

$$\begin{aligned}
 &\left| M_n(f, x) - \frac{1}{2}[f(x+) + f(x-)] \right| \\
 &\leq \frac{50}{\sqrt{n}} \left(\frac{1+x}{x}\right)^{3/2} |f(x+) - f(x-)| \\
 &\quad + \frac{(1+x)^r}{x^4} O\left(\frac{1}{n^2}\right) + \frac{4(1+x)}{nx} \sum_{k=1}^n \bigvee_{z-x/\sqrt{k}}^{x+z/\sqrt{k}} (g_x) + \bigvee_{z-x/\sqrt{n}}^{x+z/\sqrt{n}} (g_x) \\
 &\leq \frac{50}{\sqrt{n}} \left(\frac{1+x}{x}\right)^{3/2} |f(x+) - f(x-)| \\
 &\quad + \frac{5(1+x)}{nx} \sum_{k=1}^n \bigvee_{z-x/\sqrt{k}}^{x+z/\sqrt{k}} (g_x) + \frac{(1+x)^r}{x^4} O\left(\frac{1}{n^2}\right).
 \end{aligned}$$

4. REMARK

We shall prove that our estimate is essentially the best possible. Consider the function  $f(t) = |t - x|$  ( $0 < x < \infty$ ) on  $[0, \infty)$ . It obviously belongs to  $BV_{loc,r}(0, \infty)$ . Since

$$M_n(f, x) = \int_0^\infty H_n(x, t) f(t) dt = \int_0^\infty (n - 1) \left[ \sum_{k=0}^\infty P_{nk}(x) P_{nk}(t) \right] f(t) dt,$$

by Lemma 6, for any small  $\delta > 0$  and  $n$  sufficiently large, we have

$$(4.1) \quad \begin{aligned} M_n(|t - x|, x) &= (n - 1) \sum_{k=0}^\infty P_{nk}(x) \left[ \int_{x-\delta}^{x+\delta} + \int_{|t-x|>\delta} \right] |t - x| P_{nk}(t) dt \\ &\leq \delta + \frac{1}{\delta} M_n((t - x)^2, x) \leq \delta + \frac{1}{n\delta} 3x(1 + x); \end{aligned}$$

and

$$\begin{aligned} M_n(|t - x|, x) &\geq \int_{x-\delta}^{x+\delta} |t - x| H_n(x, t) dt \geq \frac{1}{\delta} \int_{x-\delta}^{x+\delta} (t - x)^2 H_n(x, t) dt \\ &\geq \frac{x(1 + x)}{n\delta} - \frac{1}{\delta} \int_{|t-x|>\delta} (t - x)^2 H_n(x, t) dt. \end{aligned}$$

Furtherfore, using Lemma 6, we have

$$\int_{|t-x|>\delta} (t - x)^2 H_n(x, t) dt \leq \frac{1}{\delta^2} M_n((t - x)^4, x) \leq \frac{c_1}{\delta^2 n^2}.$$

Hence

$$(4.2) \quad M_n(|t - x|, x) \geq \frac{x(1 + x)}{n\delta} - \frac{c_1}{n^2 \delta^2}.$$

Choose  $\delta = 2\sqrt{c_1/nx(1 + x)}$ . We obtain from (4.1) and (4.2) that

$$(4.3) \quad \frac{3[x(1 + x)]^{3/2}}{8\sqrt{c_1 n}} \leq M_n(|t - x|, x) \leq \frac{2[c_1 + (x(1 + x))^2]}{\sqrt{c_1 n x}}.$$

For  $f(t) = |t - x|$ , putting  $t = x$ , we have

$$\begin{aligned} &\left| M_n(f, x) - \frac{1}{2}(f(x+) + f(x-)) \right| \\ &= |M_n(|t - x|, x)| \leq \frac{3(1 + x)}{nx} \sum_{k=1}^n \bigvee_{x-\sqrt{k}}^{x+\sqrt{k}} (f) + \frac{(1 + x)}{x^4} O\left(\frac{1}{n^2}\right). \end{aligned}$$

Since

$$\bigvee_{x-\beta}^{x+\alpha} (f) = \alpha + \beta,$$

it follows that

$$\begin{aligned} (4.4) \quad M_n(|t-x|, x) &\leq \frac{4(1+x)}{nx} \sum_{k=1}^n \bigvee_{x-z/\sqrt{k}}^{x+z/\sqrt{k}} (f) \\ &= \frac{8(1+x)}{n} \sum_{k=1}^n \frac{1}{\sqrt{k}} \leq \frac{8(1+x)}{\sqrt{n}}. \end{aligned}$$

By comparing (4.3) and (4.4), we see that (1.2) cannot be asymptotically improved for  $BV_{loc,r}(0, \infty)$ ; that is, our estimate is essentially the best possible.

REFERENCES

- [1] A. Sahai and G. Prasad, ‘On simultaneous approximation by modified Lupas operators’, *J. Approx. Theory* **45** (1985), 122–128.
- [2] R. Bojanic, ‘An estimate of the rate of convergence for Fourier series of functions of bounded variation’, *Publ. Inst. Math. Belgrade (N.S.)* **26** (40) (1979), 57–60.
- [3] Y.S. Chow and H. Teicher, *Probability theory* (Springer-Verlag, Heidelberg, Berlin, New York, 1978).
- [4] F. Cheng, ‘On the rate of convergence of Bernstein polynomials of functions of bounded variation’, *J. Approx. Theory* **39** (1983), 259–274.

Department of Mathematics  
 Hebei Teacher’s University  
 Shijiazhuang  
 Hebei province  
 The People’s Republic of China