META-ANALYSIS

Reno-protective effects of renin-angiotensin system blockade in type 2 diabetic patients: a systematic review and network meta-analysis

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Abstract

Aims/hypothesis This meta-analysis aimed to compare the renal outcomes between ACE inhibitor (ACEI)/angiotensin II receptor blocker (ARB) and other antihypertensive drugs or placebo in type 2 diabetes.

Methods Publications were identified from Medline and Embase up to July 2011. Only randomised controlled trials comparing ACEI/ARB monotherapy with other active drugs

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Centre for Clinical Epidemiology and Biostatistics, School of Medicine and Public Health, University of Newcastle and Hunter Medical Research Institute, Newcastle, NSW, Australia or placebo were eligible. The outcome of end-stage renal disease, doubling of serum creatinine, microvascular complications, microalbuminuria, macroalbuminuria and albuminuria regression were extracted. Risk ratios were pooled using a random-effects model if heterogeneity was present; a fixedeffects model was used in the absence of heterogeneity.

Results Of 673 studies identified, 28 were eligible (n=13-4,912). In direct meta-analysis, ACEI/ARB had significantly lower risk of serum creatinine doubling (pooled RR=0.66 [95% CI 0.52, 0.83]), macroalbuminuria (pooled RR=0.70 [95% CI 0.50, 1.00]) and albuminuria regression (pooled RR 1.16 [95% CI 1.00, 1.39]) than other antihypertensive drugs, mainly calcium channel blockers (CCBs). Although the risks of end-stage renal disease and microalbuminuria were lower in the ACEI/ARB group (pooled RR 0.82 [95% CI 0.64, 1.05] and 0.84 [95% CI 0.61, 1.15], respectively), the differences were not statistically significant. The ACEI/ARB benefit over placebo was significant for all outcomes except microalbuminuria. A network meta-analysis detected significant treatment effects across all outcomes for both active drugs and placebo comparisons.

Conclusions/interpretation Our review suggests a consistent reno-protective effect of ACEI/ARB over other antihypertensive drugs, mainly CCBs, and placebo in type 2 diabetes. The lack of any differences in BP decrease between ACEI/ARB and active comparators suggest this benefit is not due simply to the antihypertensive effect.

Keywords Albuminuria regression · Angiotensin receptor blockers · Angiotensin-converting enzyme inhibitors · Doubling of serum creatinine · End-stage renal disease · Macroalbuminuria · Meta-analysis · Microalbuminuria · Renin–angiotensin system blockade · Reno-protective effect · Systematic review · Type 2 diabetes

Abbreviations

| ACEI | ACE inhibitor |
|--------|---------------------------------|
| ARB | Angiotensin II receptor blocker |
| BB | Beta-blocker |
| CCB | Calcium-channel blocker |
| DBP | Diastolic BP |
| d-CCB | Dihydropyridine CCB |
| ESRD | End-stage renal disease |
| nd-CCB | Non-dihydropyridine CCB |
| RAS | Renin-angiotensin system |
| RCT | Randomised controlled trial |
| SBP | Systolic BP |
| UKPDS | UK Prospective Diabetes Study |

Introduction

Diabetic nephropathy is a significant health and economic burden across the world. The prevalence of micro- and macroalbuminuria in type 2 diabetes is as high as 37–40% in western countries and 57.4–59.8% in Asian countries [1–3]. Microalbuminuria is a well-established risk factor for cardiovascular disease and is also associated with end-stage renal disease (ESRD) [4–7]. Preventive treatments have been prescribed for type 2 diabetes with or without hypertension with the aim of lowering BP, and delaying or even preventing the progression of diabetic nephropathy.

The reno-protective effects of renin–angiotensin system (RAS) blockers in type 2 diabetes have been controversial [8]. A few systematic reviews have been conducted [9–11], but these reviews pooled studies with mixed populations of participants with type 1 and type 2 diabetes, with and without diabetic nephropathy, and the focus was mainly on surrogate rather than clinical outcomes. We therefore conducted a systematic review and meta-analysis comparing the effects of ACE inhibitor (ACEI)/angiotensin II receptor blocker (ARB) with other antihypertensive drugs and placebo on ESRD, doubling of serum creatinine, microvascular complications, micro- and macroalbuminuria and regression of albuminuria.

Methods

Study selection

English-language publications were identified from Medline (1949–July 2011) and Embase (1974–April 2011) using PubMed and Ovid search engines. Search terms and strategies for Medline were as follows: (*type 2 diabetes or type 2 diabetes mellitus or type 2 DM or non insulin dependent diabetes*) and (*ACEI or angiotensin-converting enzyme inhibitors or ACE inhibitors or ARB or angiotensin receptor blockers**) and (*nephropathy or overt nephropathy or*)

microalbuminuria or macroalbuminuria or microvascular complications or doubling serum creatinine or ESRD or end-stage renal disease or urinary albumin excretion). We also searched for any additional studies in the reference lists of identified publications. Data from the most recent or the most complete report by the same authors were used. We restricted our searches to clinical trials. Studies with the following criteria were included: type 2 diabetic adult individuals, randomised trial comparing any single ACEI/ARB with other single agents (e.g. beta-blocker [BB], calciumchannel blocker [CCB], diuretics) or placebo, had at least one of the following outcomes (microalbuminuria, macroalbuminuria, albuminuria regression, microvascular complications, serum creatinine doubling and/or ESRD), and reported number of patients and events in each treatment arm. Studies were excluded if they were crossover trials, used dual therapies or compared different dosages of ACEI/ARB.

Outcome measures The outcomes of interest were ESRD, doubling of serum creatinine, microvascular complications, macroalbuminuria, microalbuminuria and regression of albuminuria. ESRD was defined as a requirement for renal replacement therapy or dialysis. Doubling of serum creatinine was defined as an increase in serum creatinine level of at least two times compared with baseline level. Microvascular complications were defined as a composite of having nephropathy and/or retinopathy. Microalbuminuria was defined as urine albumin excretion rate of 30-300 mg/24 h for 24 h urine collection, 3.5-35 mg/mmol for urinary albumin/ creatinine ratio from a spot urine collection, or 20-199 µg/min for timed urine collection. Macroalbuminuria was defined as urine albumin excretion rate ≥300 mg/24 h, ≥25 mg/mmol creatinine or $\geq 200 \ \mu g/min$ for the same specimens, respectively. Regression of albuminuria was defined as a change from a higher to a lower stage of albuminuria.

Data extraction

Two investigators (P. Vejakama and D. Lertrattananon) independently extracted data, including study and participant characteristics (e.g. age, BP, albuminuria stage, serum glucose, and HbA_{1c}) and numbers of events across intervention groups. Discrepancies were discussed with a third party (A. Thakkinstian) and resolved by consensus.

Risk of bias assessment

Risk of bias was assessed using the Cochrane Collaboration's tool addressing six domains: sequence generation, allocation concealment, blinding of participants/outcome assessors, incomplete outcome, selective outcome reporting and other source of bias. Two investigators (P. Vejakama and D. Lertrattananon) independently completed the assessments; discrepancies were discussed with a third party (A. Thakkinstian) and resolved by consensus.

Statistical analysis

For direct meta-analysis, the intervention of interest was ACEI/ARB monotherapy, while the comparator group was any antihypertensive drug or placebo. Pooled RRs were used to compare treatment effects. The Q test and I^2 statistic were used to assess the presence and degree of heterogeneity. If heterogeneity was present or $I^2 > 25\%$, the random-effects model was applied, otherwise the fixed-effects model was used. Sources of heterogeneity were explored using meta-regression. The Harbord test and a funnel plot were applied to assess publication bias [12]. If either suggested asymmetry of the funnel, contour-enhanced funnel plots were used to distinguish the cause of asymmetry (i.e. heterogeneity and reporting bias) [13, 14].

A network meta-analysis [15–17] was performed to compare indirectly all treatment effects. Summary data for treatments and outcomes were expanded to the individual patient level. A binary regression analysis was applied and the pooled RRs and 95% CIs were estimated. All analyses were performed using STATA version 11 (Stata Corp., College Station, TX, USA). *p* values less than 0.05 were considered statistically significant, except for the test of heterogeneity where p<0.10 was used.

Results

Of the 673 articles located, full papers of 153 plus four additional studies from reference lists were reviewed, leading to data pooling for 28 studies (see Fig. 1). The characteristics of these studies are described in Table 1. The major ACEI/ARBs used were enalapril (32%), followed by lisinopril (10.7%) and ramipril (10.7%). The majority of comparators were dihydropyridine CCBs (d-CCBs; 43%), placebo (39%), or BB/diuretic/non-dihydropyridine CCBs (nd-CCBs; 18%). Numbers of individuals are summarised according to the treatments and outcomes of interest (electronic supplementary material [ESM] Tables 1–6).

The assessments of risk of bias are described in ESM Table 7. Problems with incomplete outcome reports were least frequent in 2/28 (7.1%) studies [18, 19], followed by other sources of bias in 5/28 (17.9%) [18, 20–23] and selective reporting of outcomes in 5/28 (17.9%) [18, 21, 22, 24, 25]. Although all studies were randomised controlled trials (RCTs), only 14 (50.0%) studies clearly described randomisation [21, 25–37] and allocation concealment [21, 22, 26–29, 31, 33–39].

Direct meta-analysis

ESRD Nine studies [20, 21, 24, 26, 29, 30, 33, 34, 40] reported the effect of ACEI/ARB on ESRD (n=13,295). Of these, five studies [20, 21, 26, 30, 40] compared ACEI with other antihypertensive drugs (three BB, two d-CCB), three [24, 29, 34] compared ACEI with placebo, and one [33] compared ARB with d-CCB and placebo (ESM Table 1).

Six studies [20, 21, 26, 30, 33, 40] directly compared ACEI/ARB with other active drugs (n=1,090 vs 1,055); four were trials that studied patients with macroalbuminuria whereas the other two studied mixed patients with micro-and macroalbuminuria.

The treatment effects were homogeneous ($\chi^2=1.54$, df=5, p=0.908, $I^2=0.0\%$), suggesting that ACEI/ARB reduced the risk of ESRD by 18% (pooled RR=0.82 [95% CI 0.64, 1.05]; see Fig. 2a). However, this result was not statistically significant (Table 2). Neither the Harbord test nor a funnel plot (ESM Fig. 1a) suggested publication bias (coefficient -0.001, SE=0.436, p=0.998).

For placebo controls, pooled estimates based on four studies [24, 29, 33, 34] (n=5,581 vs 5,569) demonstrated homogeneous treatment effects ($\chi^2=1.11$, df=3, p=0.774, $I^2=0.0\%$; Fig. 2b) despite a mix of patients with normo-, micro- and macroalbuminuria. ACEI/ARB significantly decreased the risk of ESRD by 20% (pooled RR=0.80 [95% CI 0.69, 0.93]; Table 2). Although the Harbord test was not significant (coefficient 1.220, SE=0.311, p=0.059), the contour-enhanced funnel plot showed asymmetry (ESM Fig. 1b). One high-precision study fell in the significant area (grey shading) whereas the other three (one high and two low precision) were in the non-significant area. Applying 'trim and fill' suggested two low-precision studies (square symbols) were missing; adding these studies yielded a pooled RR of 0.78 (95% CI 0.68, 0.91) with $I^2 = 0\%$.

Doubling of serum creatinine Six studies [20, 24, 29, 33, 34, 39] reported the effect of ACEI/ARB vs other anti-hypertensive drugs on doubling of serum creatinine (n=16,216), and five studies [24, 29, 33, 34, 39] reported ACEI/ARB vs placebo (ESM Table 2).

The treatment effects of ACEI/ARB vs antihypertensive drugs (n=597 vs 601) were homogeneous ($\chi^2=0.76$, df=1, p=0.382, $l^2=0.0\%$) with a pooled RR of 0.66 (95% CI 0.53, 0.83; Fig. 2b), suggesting a significant reduction in risk of 34%.

The ACEI/ARB effects were also present when compared with placebo (n=7,831 vs 7,784). The pooled RR was 0.76 (95% CI 0.69, 0.84), indicating a 24% lower risk of serum creatinine doubling. The pooled effect was homogeneous ($\chi^2=1.67$, df=4, p=0.796, $I^2=0.0\%$) without

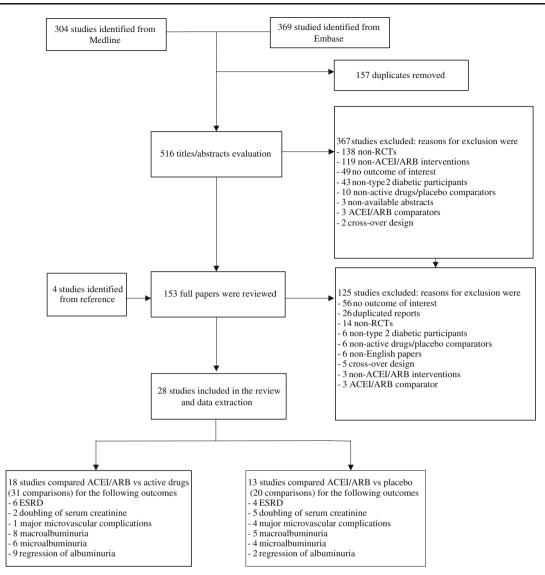


Fig. 1 Flow of study selection

publication bias (Harbord test coefficient=0.629, SE=0.908, p=0.538; ESM Fig. 2).

Microvascular complications Of five studies [24, 26, 31, 32, 37] (n=6,489), only one [26] compared ACEI/ARB with active drug; the other four [24, 31, 32, 37] compared ACEI/ARB with placebo (ESM Table 3).

Compared with placebo controls (n=2,847 vs 2,884), ACEI/ARB significantly reduced the risk of microvascular complications by 15% (pooled RR=0.85 [95% CI 0.76, 0.97]), with low heterogeneity ($\chi^2=3.34$, df=3, p=0.342, $I^2=10.3\%$; Fig. 2c) and without publication bias (Harbord test coefficient=-1.51, SE=0.53, p=0.105; ESM Fig. 3). In addition, the ACEI/ARB group (n=2,884) had a significantly lower risk of retinopathy (13% lower, pooled RR=0.87 [95% CI 0.76, 0.99]) with low heterogeneity (χ^2 =3.51, *df*=3, *p*=0.319, *I*²=14.6%).

Macroalbuminuria Twelve studies [19, 21–26, 32, 36, 38, 41, 42] reported ACEI/ARB effects on macroalbuminuria (n=5,151) with ACEI/ARB vs other therapies (five with d-CCB, one with diuretic, one with BB and one with BB/ diuretic) in eight studies [19, 21–23, 25, 26, 36, 38] and placebo in five studies [24, 32, 36, 41, 42] (ESM Table 4). Eight trials [19, 22, 25, 26, 32, 36, 38, 41] (n=1,401) studied patients with microalbuminuria, and four [21, 23, 24, 42] (n=3,761) studied mixed patients with normo- and microalbuminuria.

Direct pooled estimates of treatment effects between ACEI/ARB and other antihypertensive drugs (n=641 vs

| Author(s), year | Type of | Type of participant at baseline | ine | | | Intervention | Comparator | Outcomes | Follow-up |
|-------------------------------|----------------|--------------------------------------|--------------------|----------------------------------|-------|--------------------|-------------------------------------|--|----------------|
| | Age (years) | HbA _{1c} , % (mmol/mmol) | Renal function | Albuminuria | HT | (Ann Ann) | (IIIg) uay) | | (Jeans) |
| UKPDS, 1998 [26] | 56 | 7.0 (53) | I | Normoalb Microalb Macroalb | Yes | Captopril (25–50) | Atenolol (50–100) | Microalb, Macroalb, ESRD | 8.4 |
| HOPE, 2000 [24] | 65 | I | 93.9^{a} | Normoalb | Mix | Ramipril (10) | Placebo | Macroalb, ESRD | 4.5 |
| Ahmad et al, 1997 [41] | 50 | 8.1 (65) | 124 ^b | Microalb | No | Enalapril (10) | Placebo | Macroalb | 5 |
| Baba et al, 2001 [25] | 09 | Ι | I | Normoalb Microalb | Yes | Enalapril (5–50) | Nifedipine R (20–60) | Microalb, Macroalb, albuminuria regression | 2 |
| Bakris et al, 1996 [20] | 63 | 10.7 (93) | 63.6° | Macroalb | Yes | Lisinopril | Verapamil/diltiazem Atenolol | ESRD | 9 |
| Chan et al, 2000 [21] | 58 | I | Ι | Normoalb Microalb Macroalb | Yes | Enalapril (10–40) | Nifedipine SR (40–80) | Microalb, Macroalb, albuminuria regression, ESRD | 2 ¹ |
| Fogari et al, 2005 [27] | 60 | 6.9 (52) | I | Microalb | Yes | Lisinopril (10) | Manidipine (10) | Albuminuria regression | 2 |
| Fogari et al, 2002 [28] | 63 | 7.1 (54) | 88.4 ^a | Microalb | Yes | Fosinopril (10–30) | Amlodipine (5–15) | Albuminuria regression | 4 |
| Fogari et al, 1999 [40] | 56 | 7.2 (55) | 43.8 ^c | Macroalb | Yes | Ramipril (5) | Nitrendipine 20 | ESRD | 2 |
| Lacourciere et al, 1993 [22], | 57 | 9.3 (78) | 77.8 ^a | Macroalb | Yes | Captopril (50–100) | Metoprolol (50–100) HCTZ (25–50) | Microalb, Macroalb | ω |
| Lebovitz et al, 1994 [42] | | | | Normoalb Microalb | Yes | Enalapril (5–40) | Placebo | Macroalb | ю |
| Marre et al, 2004 [29] | 60 | 7.8 (62) | 92.5° | Microalb | Mixed | Ramipril (1.25) | Placebo | Albuminuria regression FSRD | 4 |
| Marre et al, 2004 [38] | 65 | 7.6 (60) | 89.2 ^a | Microalb | Yes | Enalapril (10) | Indapamide (1.5) | Macroalb, albuminuria | 1 |
| Mosconi et al, 1996 [44] | Ι | I | 64.6 ^b | Microalb | Yes | Enalapril (5–20) | Nitrendipine (10–40) | regression Albuminuria regression | 2 |
| Nielson et al, 1997 [30] | 61 | 8.6 (71) | 74.5 ^b | Macroalb | Yes | Lisinopril (10–20) | Atenolol (50-100) | ESRD | 3.5 |
| Ogawa et al, 2007 [45] | 61 | 6.8 (51) | 65.8 | Microalb | Yes | Temocapril (4) | Nifedipine CR (40) | Albuminuria regression | 2 |
| Ravid et al, 1998 [31] | 55 | 9.3 (78) | 107.7 ^c | Normoalb | No | Enalapril (10) | Placebo | Microalb | 9 |
| Ravid et al, 1993 [32] | 44 | 10.4 (90) | 104.6^{a} | Microalb | No | Enalapril (10) | Placebo | Macroalb | 5 |
| Ruggenenti et al, 2004 [43] | 62 | 5.8 (40) | 79.6 ^a | Normoalb | Yes | Trandolapril (2) | Verapamil SR (240), | Microalb | 3.6 |
| Shiba et al, 2000 [23] | 61 | 8.1 (65) | I | Normoalb Microalb | Yes | Delapril (60) | placebo Manidipine (10) | Macroalb | 2 |
| Brenner et al, 2001 [34] | 60 | 8.4 (68) | 168 ^a | Macroalb | | Losartan (50–100) | Placebo | ESRD, doubling Scr | 3.4 |
| Jerums et al, 2004 [36] | 53 | 8.1 (65) | 92 | Microalb | No | Perindopril | Nifedipine | Macroalb, albuminuria | 9 |
| Haller et al, 2011 [39] | 57.7 | 7.7 (61) | 84.9 | Normoalb | Mixed | Olmesartan (40) | Placebo | regression Microalb, doubling Scr | 3.2 |
| Lewis et al, 2001 [33] | 59 | 8.2 (66) | 147.6^{a} | Proteinuria | Yes | Irbesartan | Amlodipine, placebo | Doubling Scr, ESRD | 2.6 |
| Kashiwagi, 2007 [18] | 62 | Ι | I | Microalb | Yes | Valsartan | Amlodipine | Albuminuria regression | 0.5 |

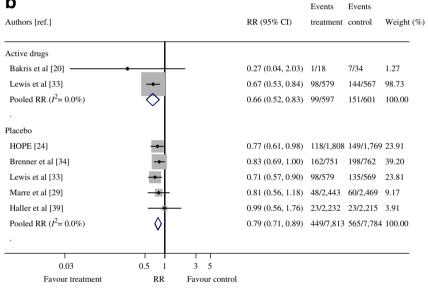
 Table 1
 Characteristics of studies in the meta-analysis

| Age HbA1c, % Renal Albuminuria HT (years) (mmol/mmol) function Hnction HT Bilous et al, 2009 [35] 56.9 - 90 Normoalb Mixed Sjolie et al, 2008 [37] 57 8.2 (66) - 90 Normoalb Mixed Estacio et al, 2000 [19] 58 11.6 (103) - Normoalb Yes ^a Serum creatinine (mmol/l) * * Normoalb Yes ^b Estimated GFR (ml min ⁻¹ 1.73 m ⁻²) * * * * | - (ung/uay) ed Candesartan ed Candesartan | (mg/may) | | Follow-up |
|--|---|--------------------------|-------------------------------|-----------|
| Bilous et al, 2009 [35] 56.9 – 90 Normoalb Mixed Sjolie et al, 2008 [37] 57 8.2 (66) – Normoalb Mixed Estacio et al, 2000 [19] 58 11.6 (103) – Normoalb Yes Microalb Macroalb Ves bestimated GFR (ml min ⁻¹ 1.73 m ⁻²) ° Creatinine clearance (ml/mi) | | | | (y cars) |
| Sjolue et al, 2008 [3/] 5/ 8.2 (66) – Normoalb Mixed Estacio et al, 2000 [19] 58 11.6 (103) – Normoalb Yes Microalb Macroalb ^a Serum creatinine (mmol/l) ^b Estimated GFR (ml min ⁻¹ 1.73 m ⁻²) ^c Creatinine clearance (ml/min) | | Placebo | Microalb | 4.7 |
| Estacio et al, 2000 [19] 58 11.6 (103) – Normoalb Yes Microalb Macroalb ^a Serum creatinine (mmol/l) ^b Estimated GFR (ml min ⁻¹ 1.73 m ⁻²) ^c Creatinine clearance (ml/min) | | Placebo | Microvascular complication | 4.7 |
| ^a Serum creatinine (mmol/l) ^b Estimated GFR (ml min ⁻¹ 1.73 m ⁻²) ^c Creatinine clearance (ml/min) | Enalapril (5–40) | Nisoldipine (10-60) | Microalb, Macroalb | 5.3 |
| ^c Creatinine clearance (ml/min) | | | | |
| HCTZ, hydrochlorothiazide; HT, hypertension; Macroalby Macroalbuminuria; Microalb, microalbuminuria; Normoalb, normoalbuminuria; Ser, serum creatinine | uminuria; Normoalb, nori | moalbuminuria; Scr, seru | un creatinine | |
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| a | | | | | Events | Events | |
|-----------------------------|------------|---|----------|--------------------|-----------|-----------|------------|
| Authors [ref.] | | | | RR (95% CI) | treatment | control | Weight (%) |
| Active drugs | | | | | | | |
| Bakris et al [20] | | | | 0.20 (0.01, 3.60) | 0/18 | 4/34 | 0.73 |
| Nielsen et al [30] | • | | | 1.12 (0.08, 16.52) | 1/17 | 1/19 | 0.83 |
| UKPDS [26] | | | | 0.89 (0.23, 3.55) | 4/400 | 4/358 | 3.16 |
| Fogari et al [40] | | | | 0.96 (0.06, 14.55) | 1/26 | 1/25 | 0.81 |
| Chan et al [21] | + | | | 1.25 (0.41, 3.83) | 6/50 | 5/52 | 4.78 |
| Lewis et al [33] | - | | | 0.80 (0.62, 1.04) | 87/579 | 106/567 | 89.69 |
| Pooled RR ($I^2 = 0.0\%$) | | | | 0.82 (0.64, 1.05) | 99/1,090 | 121/1,055 | 5 100.00 |
| | | | | | | | |
| Placebo | | | | | | | |
| HOPE [24] | | | | 1.22 (0.48, 3.09) | 10/1,808 | 8/1,769 | 2.56 |
| Brenner et al [34] | - | | | 0.77 (0.64, 0.93) | 147/751 | 194/762 | 61.48 |
| Lewis et al [33] | | | | 0.81 (0.63, 1.06) | 87/579 | 105/569 | 32.66 |
| Marre et al [29] | | - | | 0.93 (0.41, 2.10) | 1/2,443 | 12/2,469 | 3.30 |
| Pooled RR ($I^2 = 0.0\%$) | \diamond | | | 0.80 (0.69, 0.93) | 255/5,581 | 319/5,569 | 0 100.00 |
| | | | | | | | |
| | | | | | | | |
| 0.01 | 0.5 1 | 3 | 10 | | | | |
| Favour treatment | RR | | Favour o | control | | | |

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| C | | | | Events | Events | |
|----------------------------|-------------|------------|-------------------|-----------|-----------|------------|
| Authors [ref.] | | | RR (95% CI) | treatment | control | Weight (%) |
| Ravid et al [32] | | | 0.64 (0.30, 1.34) | 9/49 | 13/45 | 3.09 |
| Ravid et al [31] | | | 0.41 (0.17, 1.00) | 6/77 | 15/79 | 2.17 |
| HOPE [24] | | - | 0.86 (0.74, 0.99) | 273/1,808 | 312/1,769 | 56.49 |
| Sjolie et al [37] | | + | 0.89 (0.73, 1.08) | 161/950 | 182/954 | 38.24 |
| Pooled RR (I^2 = 10.3%) | | \Diamond | 0.85 (0.74, 0.97) | 449/2,884 | 522/2,847 | 100.00 |
| 0.15 | 0.5 | 1 1.3 | | | | |
| | r treatment | RR | Favour control | | | |

Fig. 2 Forest plots of treatment effects between ACEIs and other active drugs: (a) ESRD; (b) doubling of serum creatinine; (c) major microvascular complications; (d) macroalbuminuria; (e) microalbuminuria; and (f) albuminuria regression

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| a | | | | Events | Events | |
|------------------------------|-----------------------|--------------|--------------------|-----------|-----------|-----------|
| Authors [ref.] | | | RR (95% CI) | treatment | control | Weight (% |
| Active drugs | | | | | | |
| Lacourciere et al [22] | • | _ | 0.26 (0.01, 4.83) | 0/9 | 2/12 | 1.41 |
| UKPDS [26] | | | 0.48 (0.20, 1.15) | 7/153 | 14/146 | 15.60 |
| Chan et al [21] | | | 0.65 (0.20, 2.12) | 4/39 | 6/38 | 8.59 |
| Estacio et al [19] | | | 0.95 (0.50, 1.81) | 13/67 | 17/83 | 28.81 |
| Shiba et al [23] | <u> </u> | | 0.08 (0.00, 1.37) | 0/23 | 4/16 | 1.48 |
| Baba et al [25] | | - | 0.91 (0.21, 3.87) | 3/53 | 4/64 | 5.71 |
| Marre et al [38] | | | 0.69 (0.38, 1.22) | 18/286 | 26/283 | 36.03 |
| Jerums et al [36] | | | 2.00 (0.21, 18.98) | 2/11 | 1/11 | 2.38 |
| Pooled RR ($I^2 = 0.0\%$) | \diamond | | 0.70 (0.50, 1.00) | 47/641 | 74/653 | 100.00 |
| Placebo | | | | | | |
| Ravid et al [32] | | | 0.29 (0.13, 0.66) | 6/49 | 19/45 | 21.92 |
| Lebovetz et al [42] | • | | 0.32 (0.07, 1.38) | 2/30 | 8/38 | 10.79 |
| Ahmad et al [41] | | | 0.33 (0.11, 0.95) | 4/52 | 12/51 | 16.70 |
| HOPE [24] | + | | 0.77 (0.61, 0.97) | 117/1,808 | 149/1,769 | 38.54 |
| Jerums et al [36] | | | 0.39 (0.10, 1.53) | 2/11 | 7/15 | 12.05 |
| Pooled RR ($I^2 = 53.7\%$) | $\overline{\diamond}$ | | 0.45 (0.26, 0.79) | 131/1,950 | 195/1,918 | 100.00 |
| | Ť | | | | | |
| | | 1 1 | | | | |
| 0.002 | 0.5 1 2 | 5 10 | | | | |
| Favour treatment | RR | Favour contr | rol | | | |

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| Authors [ref.] | | RR (95% CI) | treatment | control | Weight (% |
|----------------------------|-----------------|-------------------|-----------|----------|-----------|
| Active drugs | | | | | |
| Lacourciere et al [22] | | 0.56 (0.05, 5.81) | 1/25 | 2/28 | 1.74 |
| UKPDS [26] - | • | 1.21 (0.84, 1.73) | 48/153 | 38/146 | 30.58 |
| Chan et al [21] | • | 1.11 (0.35, 3.57) | 4/18 | 5/25 | 6.35 |
| Estacio et al [19] | - | 0.87 (0.53, 1.42) | 25/123 | 25/107 | 22.79 |
| Baba et al [25] | F | 0.73 (0.40, 1.31) | 15/95 | 23/106 | 18.30 |
| Ruggenenti et al [43] | | 0.50 (0.29, 0.87) | 18/301 | 36/303 | 20.24 |
| Pooled RR (I^2 = 35.1%) | > | 0.84 (0.61, 1.15) | 111/715 | 129/715 | 100.00 |
| Placebo | | | | | |
| Ravid et al [31] | | 0.34 (0.13, 0.90) | 5/77 | 15/79 | 5.87 |
| Ruggenenti et al [43] | Ł | 0.60 (0.34, 1.05) | 18/301 | 30/300 | 14.35 |
| Bilous et al [35] | - | 0.92 (0.73, 1.17) | 114/951 | 124/954 | 37.21 |
| Haller et al [39] - | - | 0.92 (0.76, 1.12) | 178/1961 | 210/2139 | 42.58 |
| Pooled RR (I^2 =49.4%) | T | 0.82 (0.64, 1.05) | 315/3290 | 379/3472 | 100.00 |
| - | | 10.5 | | | |
| 0.054 | 1 | 18.5 | | | |
| Favour treatment RF | R Favour contro | ol | | | |

Events Events

f

| T | | | | | | | Events | Events | |
|-----------------------------|-----|--------------------|---|-------|---------|-------------------|-----------|---------|-----------|
| Authors [ref.] | | | | | | RR (95% CI) | treatment | control | Weight (% |
| Active drugs | | | | | | | | | |
| Mosconi et al [44] | | • | | | | 1.17 (0.50, 2.74) | 4/6 | 4/7 | 3.88 |
| Chan et al [21] | | | | | | 1.01 (0.35, 2.95) | 6/32 | 5/27 | 2.51 |
| Baba et al [25] | | — | | | | 0.70 (0.37, 1.34) | 11/53 | 19/64 | 6.62 |
| Fogari et al [28] | | • | - | | | 1.40 (0.99, 1.97) | 47/102 | 34/103 | 20.34 |
| Jerums et al [36 | • | | | | | 0.33 (0.04, 2.73) | 1/11 | 3/11 | 0.66 |
| Marre et al [38] | | - | | | | 1.06 (0.87, 1.29) | 120/286 | 112/283 | 45.91 |
| Fogari et al [27] | | ++ | - | | | 1.45 (0.94, 2.24) | 27/49 | 19/50 | 13.73 |
| Kashiwagi [18] | | - | ٠ | | | 1.99 (0.95, 4.18) | 17/73 | 9/77 | 5.11 |
| Ogawa et al [45] | | | | | _ | 1.32 (0.28, 6.16) | 5/34 | 2/18 | 1.23 |
| Pooled RR ($I^2 = 8.4\%$) | | \diamond | | | | 1.17 (0.99, 1.39) | 238/646 | 207/640 | 100.00 |
| Placebo | | | | | | | | | |
| Jerums et al [36] | • | _ | | _ | | 0.45 (0.05, 3.81) | 1/11 | 3/15 | 0.55 |
| Marre et al [38] | | - | | | | 1.17 (1.00, 1.37) | 253/936 | 215/933 | 99 |
| Pooled RR ($I^2 = 0.0\%$) | I | $\hat{\mathbf{P}}$ | | | | 1.17 (1.00, 1.37) | 254/947 | 218/948 | 100.00 |
| | | | _ | | _ | | | | |
| 0.03 | 0.5 | 1 | 2 | 4 | 6 | | | | |
| Favour treatment | | RR | I | Favor | ur cont | trol | | | |

Fig. 2 (continued)

 Table 2
 Summarised characteristics of studies and direct pooling treatment effects according to clinical outcomes

| Dichotomous outcomes | Number of studies | Sample size | RR (95% CI) |
|-----------------------------------|-------------------|-------------|-------------------|
| ESRD | | | |
| ACEI/ARB vs other active drugs | 6 | 2,147 | 0.82 (0.64, 1.05) |
| ACEI/ARB vs placebo | 4 | 10,581 | 0.80 (0.69, 0.93) |
| Doubling of serum creatinine | | | |
| ACEI/ARB vs other active drugs | 2 | 1,198 | 0.66 (0.53, 0.83) |
| ACEI/ARB vs placebo | 4 | 10,594 | 0.76 (0.69, 0.84) |
| Major microvascular complications | | | |
| ACEI/ARB vs other active drugs | 1 | 758 | 1.28 (0.81, 2.03) |
| ACEI/ARB vs placebo | 4 | 6,489 | 0.85 (0.74, 0.97) |
| Macroalbuminuria | | | |
| ACEI/ARB vs other active drugs | 8 | 1,211 | 0.71 (0.50, 1.00) |
| ACEI/ARB vs placebo | 5 | 3,868 | 0.67 (0.54, 0.83) |
| Microalbuminuria | | | |
| ACEI/ARB vs other active drugs | 6 | 1,430 | 0.84 (0.61, 1.15) |
| ACEI/ARB vs placebo | 4 | 6,762 | 0.82 (0.64, 1.05) |
| Albuminuria regression | | | |
| ACEI/ARB vs other active drugs | 9 | 1,286 | 1.16 (0.99, 1.39) |
| ACEI/ARB vs placebo | 2 | 1,238 | 1.17 (1.00, 1.37) |

653) of eight trials [19, 21–23, 25, 26, 36, 38] were homogeneous (χ^2 =5.24, *df*=7, *p*=0.631, *I*²=0.0%; Fig. 2d) with borderline significant risk reduction of 29% (pooled RR= 0.71 [95% CI 0.50, 1.00]). Neither the Harbord test nor the funnel plot suggested publication bias (coefficient=-0.733, SE=0.828, *p*=0.410; ESM Fig. 4a).

Among five placebo-controlled trials [24, 32, 36, 41, 42] (n=1,950 vs 1,918), the pooled RR was 0.45 (95% CI 0.26, 0.79) with moderate heterogeneity ($\chi^2 = 8.64$, df=4, p=0.071, $l^2=53.7\%$; Fig. 2d). This suggested that ACEI/ ARB significantly reduced risk of macroalbuminuria by 55%. The Harbord test suggested asymmetry of the funnel plot (coefficient=-2.043, SE=0.547, p=0.033). The contourenhanced funnel plot showed that one-half of the studies included were in the high- and non-significance areas (ESM Fig. 4b), suggesting treatment-effect heterogeneity. Metaregression analysis fitting type of patient (microalbuminuria and mixed patients) reduced the degree of heterogeneity (I^2) from 53.7% to 0%. Subgroup analysis performed in patients with microalbuminuria and mixed micro- and macroalbuminuria yielded pooled RRs of 0.31 (95% CI 0.18, 0.57) and 0.45 (95% CI 0.26, 0.79), respectively.

Microalbuminuria Nine studies [19, 21, 22, 25, 26, 31, 35, 39, 43] had microalbuminuria outcomes (n=7,891): six [19, 21, 22, 25, 26, 43] looking at ACEI vs other antihypertensive drugs, two looking at ACEI vs placebo [31, 43] and two looking at ARB vs placebo [35, 39] (ESM Table 5).

The pooled treatment effects between ACEI/ARB and other antihypertensive drugs were somewhat heterogeneous (χ^2 =7.70, df=5, p=0.174, I^2 =35.1%) with a pooled RR of 0.84 (95% CI 0.61, 1.15) (Table 2 and Fig. 2e). The Harbord test suggested no asymmetry of the funnel plot (coefficient=-0.281, SE=1.367, p=0.847; ESM Fig. 5a).

Compared with placebo in four studies [31, 35, 39, 43] (n=3,290 vs 3,472), the pooled effect of ACEI/ARBs [24, 28, 32, 37] was 0.82 (95% CI 0.64, 1.05; Fig. 2e) with moderate heterogeneity (χ^2 =5.93, df=3, p=0.115, I^2 =49.4%). Meta-regression could not identify the source of the heterogeneity. The Harbord test suggested asymmetry of the funnel (coefficient=-2.626, SE=0.360, p=0.018), and the contour-enhanced funnel plot suggested that this heterogeneity might be the cause of the asymmetry (ESM Fig. 5b).

Albuminuria regression Ten studies [18, 21, 25, 27–29, 36, 38, 44, 45] reported regression of albuminuria (n=3,710; ESM Table 6).

The treatment effect between ACEI/ARB and other antihypertensive drugs, based on nine studies [18, 21, 25, 27, 28, 36, 38, 44, 45] (n=646 vs 640), was homogeneous (χ^2 =8.74, df=8, p=0.365, f^2 =8.4%) with a borderline non-significant effect of 16% (pooled RR=1.16 [95% CI 0.99, 1.39]; see Table 2 and Fig. 2f). Neither the Harbord test nor a funnel plot suggested publication bias (coefficient=-0.086, SE=0.718, p=0.907; ESM Fig. 6). For the placebo-controlled trials (n=947 vs 948), the likelihood of albuminuria regression was 17% higher with ACEI/ARB, which was on the borderline of non-significance (RR=1.17 [95% CI 1.00, 1.37]; Table 2).

Network meta-analysis

ESRD A network meta-analysis was applied to assess all treatment comparisons, (ESM Fig. 7a and ESM Table 8). Arrows and tails refer to interventions and comparators, bold and dashed lines refer to direct and indirect comparisons, respectively. Numbers under the line refer to the number of studies for the direct comparator data, whereas numbers above the line indicate pooled RRs. The analysis suggested that the ACEI/ARB (n=6,092) significantly decreased the risk of ESRD when compared with d-CCB (n=644) and placebo (n=5,569,) with pooled RRs of 0.25 (95% CI 0.07, 0.96) and 0.77 (95% CI 0.64, 0.92), respectively. However, the risk of ESRD was 2.00 (95% CI 0.35, 11.55) times higher, but not significantly so, compared with the BB group (n=411). Comparing ACEI/ARB with all antihypertensive drugs (dCCB + BB) resulted in 62% significantly lower risk of ESRD (pooled RR=0.38 [95% CI 0.31, 0.47]).

Pooling mean BPs from four studies [21, 26, 30, 40] showed non-significant differences between groups, with mean differences of 1.27 (95% CI -1.42, 3.95) for systolic BP (SBP) and -0.71 (95% CI -3.88, 2.46) for diastolic BP (DBP).

Doubling of serum creatinine ESM Fig. 7b and ESM Table 8 show that ACEI/ARBs (n=7,831) significantly reduced the risk of serum creatinine doubling by 77% (pooled RR=0.23 [95% CI 0.09, 0.57]), 72% (pooled RR=0.28 [95% CI 0.11, 0.71]), and 21% (pooled RR=0.79 [95% CI 0.75, 0.84]) compared with d-CCB (n=567), nd-CCB/BB (n=34), and placebo (n=7,784), respectively. Combining all antihypertensive drugs (n=601) and comparing with ACEI/ARB resulted in 77% significantly lower risk (pooled RR=0.23 [95% CI 0.09, 0.58]).

Microvascular complications The number of participants in each treatment arm is described in ESM Table 3, with total number of patients in ACEI/ARB and placebo of 3,284 vs 2,847, respectively. ACEI/ARBs significantly reduced the risk of microvascular complications by 19% (pooled RR= 0.81 [95% CI 0.71, 0.92]) when compared with placebo, but increased the risk almost twofold when compared with BB (ESM Fig. 7c and ESM Table 8).

Macroalbuminuria Network meta-analysis was performed based on 12 studies [19, 21–26, 32, 36, 38, 41, 42] (n=5,151; ESM Table 4). ACEI/ARBs (n=2,580) significantly reduced macroalbuminuria by 29% (pooled RR=0.71

[95% CI 0.62, 0.83]) compared with BB/diuretic (n=441), 55% (pooled RR=0.45 [95% CI 0.29, 0.70]) vs d-CCB (n=212), and 33% (pooled RR=0.67 [95% CI 0.49, 0.92]) vs placebo (n=1,918; ESM Fig. 7d and ESM Table 8). The ACEI/ARBs significantly reduced macroalbuminuria by 40% (pooled RR=0.60 [95% CI 0.47, 0.77]) compared with all antihypertensive drugs (n=653).

Pooling SBP and DBP in three studies [21, 38, 41], the mean SBP and DBP differences were 2.82 (95% CI 0.32, 5.31) and 0.33 (95% CI -0.93, 1.60) mmHg, respectively.

Microalbuminuria Based on pooling of results from nine studies [19, 21, 22, 25, 26, 31, 35, 39, 43] (n=7,891), ACEI/ARB (n=3,704) had significantly better benefit than d-CCB (n=238) with a pooled RR of 0.50 (95% CI 0.38, 0.65), but significantly worse outcomes than BB (i.e. favouring BB) with the pooled RR of 0.48 (95% CI 0.35, 0.66; ESM Fig. 7e and ESM Table 8). When compared with nd-CCB (n=303) and placebo (n=3,472), ACEI/ARBs did not reduce microalbuminuria with pooled RRs of 0.93 (95% CI 0.71, 1.22) and 1.01 (95% CI 0.82, 1.25), respectively. Compared with all antihypertensive drugs, the ACEI/ARB significantly reduced the risk by 39% (pooled RR=0.61 [95% CI 0.45, 0.83]).

Regression of albuminuria Using data from ten studies (n=3,710) [18, 21, 25, 27–29, 36, 38, 44, 45], ACEI/ARB (n=1,582) significantly promoted regression of albuminuria compared with placebo (n=948), but not with diuretic (n=283); RRs were 1.35 (95% CI 1.07, 1.70) in favour of ACEI/ARB and 1.28 (95% CI 1.02, 1.60) in favour of diuretic, respectively (ESM Fig. 7f and ESM Table 8). The pooled RR was1.16 (95% CI 0.83, 1.63) for ACEI/ARB vs d-CCB (n=357).

In six of ten studies where data were available [21, 27–29, 44, 45], the mean SBP was significantly higher for ACEI/ARB (mean difference=2.35 [95% CI 1.07, 3.64]) than placebo but not so for DBP (mean difference=0.54 mmHg [95% CI -0.40, 1.48 mmHg]).

Discussion

The direct and indirect comparisons in our study confirm the clear benefit of ACEI/ARB over placebo for all outcomes (ESRD, serum creatinine doubling, microvascular complications, macroalbuminuria and regression of albuminuria) except microalbuminuria. The lack of a statistically significant treatment effect for microalbuminuria outcome is likely to result partly from the heterogeneity of treatment effects across the studies.

Our study extends the results of Strippoli et al. [11], who showed borderline significant benefit for ACEI/ARB vs placebo for doubling of serum creatinine (RR=0.79 [95% CI 0.75, 0.84] vs 0.60 [95% CI 0.34, 1.05]) and ESRD (RR=0.77 [95% CI 0.64, 0.92] vs 0.64 [95% CI 0.4, 1.03]). The addition of new data and the use of network meta-analysis methods increase the power and allow us to demonstrate the reno-protective effects of RAS blockade more confidently.

The major contribution of our study is in teasing out the potential reno-protective effect of ACEI/ARBs over other antihypertensive drugs. Overall, the direct meta-analysis suggests a potential benefit of ACEI/ARB in reducing macroalbuminuria and doubling of serum creatinine, and in promoting regression of albuminuria. The results of network meta-analysis suggest additional benefits of the ACEI/ARB in reducing the risk of ESRD. These benefits contradict the previous meta-analysis [9], which found renoprotective effects of ACEI/ARB in patients without diabetes but not in patients who had already developed diabetic nephropathy. This discrepancy with previous work may be due partly to the addition of new data, thereby improving power, or may be driven partly by the fact that the major active comparators were CCBs; in some comparisons, BBs and diuretics are significantly better than ACEI/ARBs, though the numbers in these comparisons are small and driven largely by the UK Prospective Diabetes Study (UKPDS).

In the studies that reported BP change, there was no significant difference in BP decrease between ACEI/ARB and other antihypertensive drugs, suggesting that there is a specific reno-protective effect of ACEI/ARBs beyond their antihypertensive effect. This runs contrary to the conclusion of Casas et al. [9], who found the effect of RAS blockade to depend on its BP-lowering effect. In their study, Casas et al. showed that the benefit of RAS blockers compared with active control was associated with the extra degree of BP decrease. These results are hard to reconcile with summarylevel data. It is of course possible for both conclusions to be true, i.e. the effect of RAS blockers is linked to the level of BP control but that for two agents that bring about the same drop in BP, RAS blockers would have an additional benefit. This possibility would be better teased out with a metaanalysis of individual patient data or a direct test of this hypothesis in a large-scale RCT.

Our study has some strengths. We have applied a network meta-analysis to increase the power of the tests and reduce type I errors [15–17]. We applied a mixed model, which is thought to be the most appropriate method for this kind of pooling. The network method 'borrows' treatment information from other studies and increases the total sample size. As a result, treatment effects that could not be detected or treatment effects of borderline significance in direct meta-analysis could be identified. In addition, all possible treatment comparisons are mapped and displayed (see ESM Table 9). The weakness of our study is that despite the large overall numbers, some comparator groups still have small

numbers, leading to poor precision of estimates. The lack of BP data for some studies also limits our ability to judge the specific reno-protective effects.

In conclusion, there appears to be a consistent benefit of ACEI/ARB on all outcomes in type 2 diabetes, including the 'hard clinical endpoints' of ESRD and doubling of serum creatinine, when compared with other antihypertensive drugs (mainly CCBs) and placebo, both in direct and indirect meta-analyses. This is seen in the context of no difference in BP decrease, suggesting a specific reno-protective effect. Other antihypertensives, particularly BBs and diuretics, may have even greater reno-protective effects but this needs to be investigated in further studies or individual patient meta-analysis.

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