

Risk-averse single machine scheduling: complexity and approximation

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Abstract

In this paper, a class of single machine scheduling problems is considered. It is assumed that job processing times and due dates can be uncertain and they are specified in the form of discrete scenario set. A probability distribution in the scenario set is known. In order to choose a schedule, some risk criteria such as the value at risk and conditional value at risk are used. Various positive and negative complexity results are provided for basic single machine scheduling problems. In this paper, new complexity results are shown and some known complexity results are strengthened.

Keywords Single machine \cdot Value at risk \cdot Conditional value at risk \cdot Computational complexity \cdot Approximation algorithms \cdot Robust scheduling

1 Introduction

Scheduling under risk and uncertainty has attracted considerable attention in the recent literature. In practical applications of scheduling models, the exact values of input parameters, such as job processing times or due dates, are often unknown in advance. Hence, a solution must be computed, before the true realization of the input data is revealed. Typically, a scenario set \mathcal{U} is a part of the input, which contains all possible realizations of the problem parameters, called *scenarios*. If the probability distribution in \mathcal{U} is unknown, then robust optimization framework can be applied and solution performance in a worst case is optimized. First robust scheduling problems have been discussed in Daniels and Kouvelis (1995), Kouvelis and Yu (1997) and Yu and Kouvelis (1993). Two uncertainty representations, namely discrete and interval ones, were considered. In the former, scenario set \mathcal{U} contains a finite number of distinct scenarios. In the latter, for each uncertain parameter an interval of

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² Department of Computer Science, Faculty of Fundamental Problems of Technology, Wrocław University of Science and Technology, Wrocław, Poland its possible values is specified and \mathcal{U} is the Cartesian product of these intervals. In order to compute a solution, the minmax and minmax regret criteria can be applied. Minmax (regret) scheduling problems have various complexity properties, depending on the cost function and the uncertainty representation [see, e.g., Averbakh (2000), Lebedev and Averbakh (2006), Kasperski (2005), Aissi et al. (2011) and Drwal and Rischke (2016)]. For a survey of minmax (regret) scheduling problems we refer the reader to Kasperski and Zieliński (2014).

The robust scheduling models have well-known drawbacks. Minimizing the maximum cost can lead to very conservative solutions. The reason is that the probability of occurrence of the worst scenario may be very small and the information connected with the remaining scenarios is ignored while computing a solution. One method of overcoming this drawback was given in Kasperski and Zieliński (2016), where the OWA criterion, proposed in Yager (1988), was applied to compute an optimal schedule. In this approach, a set of weights is specified by the decision maker, which reflect his attitude toward a risk. The OWA operator contains the maximum, average and Hurwicz criteria as special cases. However, it does not take into account a probabilistic information, which may be available for scenario set U.

In the case, when a probability distribution in \mathcal{U} is known, the stochastic scheduling models are considered. The parameters of scheduling problem are then random variables with known probability distributions. Under this assumption, the expected solution performance is typically optimized [see,

e.g., Möhring et al. (1999), Pinedo (2008), Skutella and Uetz (2005) and Skutella et al. (2016)]. However, this criterion assumes that the decision maker is risk neutral and leads to solutions that guarantee an optimal long-run performance. Such a solution may be questionable, for example, if it is implemented only once [see, e.g., Kouvelis and Yu (1997)]. In this case, the decision-maker attitude toward a risk should be taken into account.

In Krokhmal et al. (2002), a criterion called conditional value at risk (CVaR) was applied to a stochastic portfolio selection problem. Using this criterion, the decision maker provides a parameter $\alpha \in [0, 1)$, which reflects his attitude toward a risk. When $\alpha = 0$, then CVaR becomes the expectation. However, for greater value of α , more attention is paid to the worst outcomes, which fits into the robust optimization framework. The conditional value at risk is closely connected with the value at risk (VaR) criterion [see, e.g., Pflug (2000)], which is just the α -quantile of a random outcome. Both risk criteria have attracted considerable attention in stochastic optimization [see, e.g., Natarajan et al. (2014), Chang et al. (2017), Nikolova (2010) and Ogryczak (2012)]. This paper is motivated by the recent papers (Sarin et al. 2014; Atakan et al. 2017), in which the following stochastic scheduling models were discussed. We are given a scheduling problem with discrete scenario set \mathcal{U} . Each scenario $\xi_i \in \mathcal{U}$ is a realization of the problem parameters (for example, processing times and due dates), which can occur with a known positive probability $Pr[\xi_i]$. The cost of a given schedule is a discrete random variable with the probability distribution induced by the probability distribution in \mathcal{U} . The VaR and CVaR criteria, with a fixed level α , are used to compute a best solution.

In Sarin et al. (2014) and Atakan et al. (2017), solution methods, based on mixed integer programming models, were proposed to minimize VaR and CVaR in scheduling problems with the total weighted tardiness criterion. The aim of this paper is to analyze the models discussed in Sarin et al. (2014) and Atakan et al. (2017) from the complexity point of view. We will consider the class of single machine scheduling problems with basic cost functions, such as the maximum tardiness, the total flow time, the total tardiness and the number of late jobs. We will discuss also the weighted versions of these cost functions. We provide a picture of computational complexity for all these problems by proving some positive and negative complexity results. Since VaR and CVaR generalize the maximum criterion, we can use some results known from robust minmax scheduling. The complexity results for the minmax versions of single machine scheduling problems under discrete scenario set were obtained in Aissi et al. (2011), Aloulou and Croce (2008), Daniels and Kouvelis (1995) and Mastrolilli et al. (2013). In this paper, we will show that some of these results can be strengthened.

This paper is organized as follows. In Sect. 2, we recall the definitions of the VaR and CVaR criteria and show their properties, which will be used later on. In Sect. 3, the problems discussed in this paper are defined. In Sect. 4, some general relationships between the problems with various risk criteria are shown. Finally, Sects. 5 and 6 contain some new negative and positive complexity results for the considered problems. These results are summarized in the tables presented in Sect. 3.

2 The risk criteria

Let *Y* be a random variable. We will consider the following risk criteria (Pflug 2000; Rockafellar and Uryasev 2000):

• *Value at risk* (*α*-quantile of *Y*):

 $\operatorname{VaR}_{\alpha}[Y] = \inf\{t : \Pr[Y \le t] \ge \alpha\}, \alpha \in (0, 1],$

• Conditional value at risk:

$$\mathbf{CVaR}_{\alpha}[Y] = \inf \left\{ \gamma + \frac{1}{1-\alpha} \mathbf{E}[Y-\gamma]^{+} : \gamma \in \mathbb{R} \right\},\$$

$$\alpha \in [0, 1),$$

where $[x]^+ = \max\{0, x\}$. Assume that *Y* is a discrete random variable taking nonnegative values b_1, \ldots, b_K . Then, **VaR**_{α}[*Y*] and **CVaR**_{α}[*Y*] can be computed by using the following programs, respectively [see, e.g., Atakan et al. (2017), Ogryczak (2012) and Rockafellar and Uryasev (2000)]:

(a)
$$\min \theta$$

(b) $\min \gamma + \frac{1}{1-\alpha} \sum_{i \in [K]} \Pr[Y = b_k] u_i$
s.t. $b_k - \theta \le M \beta_k$, $k \in [K]$
s.t. $\gamma + u_k \ge b_k$, $k \in [K]$
 $\sum_{k \in [K]} \Pr[Y = b_k] \beta_k \le 1 - \alpha$
 $\mu_k \ge 0$, $k \in [K]$
 $\beta_k \in \{0, 1\}$, $k \in [K]$
(1)

where $M \ge \max\{b_1, \dots, b_K\}$ and $[K] = \{1, \dots, K\}$. Notice that (1)b is a linear programming problem. In the following, we will use the following dual to (1)b:

$$\max \sum_{\substack{k \in [K] \\ k \in [K]}} b_k r_k$$
s.t.
$$\sum_{\substack{k \in [K] \\ 0 \le r_k \le \frac{\Pr[Y = b_k]}{1 - \alpha}} r_k \in [K]$$
(2)

The equality constraint in (2) follows from the fact that γ is real-valued decision variable in (1)b. Substituting r_k =

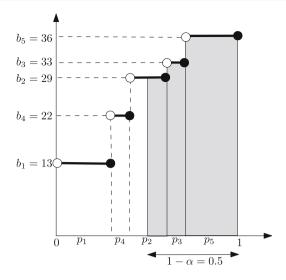


Fig. 1 A computation of **CVaR**_{0.5}[*Y*] for *Y* taking the values of 13, 29, 33, 22 and 36 with the probabilities 0.3, 0.2, 0.1, 0.1, 0.3, respectively, where $Pr[Y = b_k] = p_k$. The value of **CVaR**_{0.5}[*Y*] is the gray area divided by $1 - \alpha = 0.5$

 $q_k/(1 - \alpha)$ into (2), we get the following equivalent formulation for **CVaR**_{α}[*Y*]:

$$\max \quad \frac{1}{1-\alpha} \sum_{k \in [K]} b_k q_k$$

s.t.
$$\sum_{k \in [K]} q_k = 1-\alpha$$

$$0 \le q_k \le \Pr[Y = b_k], \quad k \in [K]$$
 (3)

Program (3) can be solved by using a greedy method, which is illustrated in Fig. 1. Namely, we fix the optimal values of q_k by greedily distributing the amount $1 - \alpha$ among the largest values of b_i . It is easy to see that $\mathbf{CVaR}_0[Y] = \mathbf{E}[Y] =$ $\sum_{k \in [K]} b_k \Pr[Y = b_k]$. On the other hand, $\mathbf{CVaR}_{1-\epsilon}[Y] =$ $\mathbf{VaR}_1[Y] = \mathbf{Max}[Y] = \max_{k \in [K]} b_k$ for sufficiently small $\epsilon > 0$ and any probability distribution.

We now show several properties of the risk measures which will be used later on in this paper.

Lemma 1 Let Y be a discrete random variable which can take K nonnegative values b_1, \ldots, b_K . The following inequalities hold for each $\alpha \in [0, 1)$:

$$\mathbf{E}[Y] \le \mathbf{CVaR}_{\alpha}[Y] \le \min\left\{\frac{1}{\Pr_{\min}}, \frac{1}{1-\alpha}\right\} \mathbf{E}[Y], \qquad (4)$$

where $\Pr_{\min} = \min_{k \in [K]} \Pr[Y = b_k]$.

Proof Fix $\alpha \in [0, 1)$. The inequality $\mathbf{E}[Y] \leq \mathbf{CVaR}_{\alpha}[Y]$ follows directly from the definition of the expected value and the conditional value at risk. We now prove the second inequality. Let $r_1^*, \ldots r_k^*$ be the optimal values in (2). Then, the inequality

$$\mathbf{CVaR}_{\alpha}[Y] = \sum_{k \in [K]} r_k^* b_k \le \sum_{k \in [K]} \frac{\Pr[Y = b_k]}{(1 - \alpha)} b_k$$
$$= \frac{1}{1 - \alpha} \mathbf{E}[Y]$$

holds. Since the value of $\mathbf{CVaR}_{\alpha}[Y]$ is a convex combination of b_1, \ldots, b_k (see (2)), we have $\mathbf{CVaR}_{\alpha}[Y] \leq \mathbf{Max}[Y] = b_{\max} \leq \sum_{k \in [K]} \frac{\Pr[Y=b_k]}{\Pr_{\min}} b_k = \frac{1}{\Pr_{\min}} \mathbf{E}[Y]$, and the lemma follows.

Lemma 2 Let X and Y be two discrete random variables taking nonnegative values a_1, \ldots, a_K , and b_1, \ldots, b_K , respectively, with $\Pr[X = a_i] = \Pr[Y = b_i]$ and $a_i \leq \gamma b_i$ for each $i \in [K]$ and some fixed $\gamma \geq 0$. Then $\operatorname{CVaR}_{\alpha}[X] \leq \gamma \operatorname{CVaR}_{\alpha}[Y]$ for each $\alpha \in [0, 1)$ and $\operatorname{VaR}_{\alpha}[X] \leq \gamma \operatorname{VaR}_{\alpha}[Y]$ for each $\alpha \in (0, 1]$.

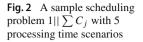
Proof Let us compute $\mathbb{CVaR}_{\alpha}[X]$ by using (2) and denote by r_k^* , $k \in [K]$, the optimal values in (2). Then, $\mathbb{CVaR}_{\alpha}[X] = \sum_{k \in [K]} r_k^* a_k \leq \gamma \sum_{k \in [K]} r_k^* b_k \leq \gamma \mathbb{CVaR}_{\alpha}[Y]$. Let us compute $\mathbb{VaR}_{\alpha}[Y]$ by solving the problem (1)a. Let θ^* , β_k^* , $k \in [K]$, be an optimal solution to (1)a. Since $\gamma \geq 0$, the constraint $\gamma b_k - \gamma \theta^* \leq \gamma M \beta_k^*$ holds for each $k \in [K]$. By $a_k \leq \gamma b_k$ for each $k \in [K]$, we get $a_k - \gamma \theta^* \leq M' \beta_k^*$, where $M' = \gamma M \geq \max\{a_1, \ldots, a_K\}, k \in [K]$. In consequence,

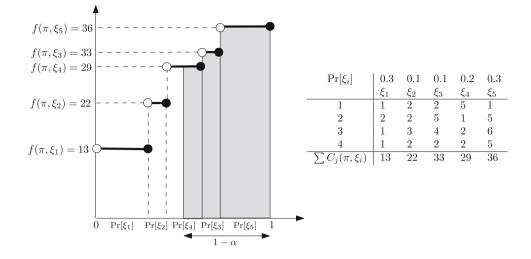
$$a_{k} - \gamma \theta^{*} \leq M' \beta_{k}^{*} \qquad k \in [K]$$
$$\sum_{k \in [K]} \Pr[X = a_{k}] \cdot \beta_{k}^{*} \leq 1 - \alpha \qquad (5)$$

and $\operatorname{VaR}_{\alpha}[X] \leq \gamma \theta^* = \gamma \operatorname{VaR}_{\alpha}[Y].$

3 Problem formulations

We are given a set J of n jobs, which can be partially ordered by some precedence constraints. Namely, $i \rightarrow j$ means that job j cannot start before job i is completed. For each job $j \in J$, a nonnegative processing time p_i , a nonnegative due date d_i and a nonnegative weight w_i can be specified. A schedule π is a feasible (i.e., preserving the precedence constraints) permutation of the jobs and Π is the set of all feasible schedules. We will use $C_i(\pi)$ to denote the completion time of job j in schedule π . Obeying the standard notation, we will use $T_i(\pi) = [C_i(\pi) - d_j]^+$ to define the tardiness of j in π , and $U_i(\pi) = 1$ if $C_i(\pi) > d_i$ (job j is late in π) and $U_i(\pi) = 0$ (job j is on-time in π), otherwise. In the deterministic case, we seek a schedule $\pi \in \Pi$ that minimizes a given *cost function* $f(\pi)$. The basic cost functions are the total flow time $\sum_{i \in J} C_j(\pi)$, the total tardiness $\sum_{i \in J} T_j(\pi)$, the maximum tardiness $\max_{j \in J} T_j(\pi)$ and the total number of late jobs $\sum_{i \in J} U_j(\pi)$. We can also consider the weighted versions of these functions. Scheduling problems \mathcal{P} will be





denoted by means of the standard Graham's notation [see, e.g., Brucker (2007)].

In this paper, we assume that job processing times and due dates can be uncertain. The uncertainty is modeled by a discrete *scenario set* $\mathcal{U} = \{\xi_1, \xi_2, \ldots, \xi_K\}$. Each realization of the parameters $\xi \in \mathcal{U}$ is called a *scenario*. For each scenario $\xi \in \mathcal{U}$, a probability $\Pr[\xi]$ of its occurrence is known. Without loss of generality, we can assume $\Pr[\xi] > 0$. We will use $p_j(\xi)$ and $d_j(\xi)$ to denote the processing time and due date of job *j* under scenario $\xi \in \mathcal{U}$, respectively. We will denote by $C_j(\pi, \xi)$, $T_j(\pi, \xi)$ and $U_j(\pi, \xi)$ the completion time, tardiness and unit penalty of job π , respectively, under scenario $\xi \in \mathcal{U}$. Also, $f(\pi, \xi)$ stands for the cost of schedule π under scenario $\xi \in \mathcal{U}$. Given a feasible schedule $\pi \in \Pi$, we denote by $F(\pi)$ a random cost of π . Notice that $F(\pi)$ is a discrete random variable with the probability distribution induced by the probability distribution in \mathcal{U} .

For a fixed value of α , we can compute a performance measure of π , namely the expected cost $\mathbf{E}[F(\pi)]$, the maximum cost $\mathbf{Max}[F(\pi)]$, the value at risk $\mathbf{VaR}_{\alpha}[F(\pi)]$ and the conditional value at risk $\mathbf{CVaR}_{\alpha}[F(\pi)]$. A sample problem 1|| $\sum C_j$ with 4 jobs and 5 processing time scenarios is shown in Fig. 2. Let $\pi = (1, 2, 3, 4)$. It is easily seen that $\mathbf{E}[F(\pi)] = 26$, $\mathbf{VaR}_{0.5}[F(\pi)] = 29$, $\mathbf{CVaR}_{0.5}[F(\pi)] = 34$ and $\mathbf{Max}[F(\pi)] = 36$.

In this paper, we will study the problems $MIN-VAR_{\alpha} \mathcal{P}$, $MIN-CVAR_{\alpha} \mathcal{P}$, $MIN-EXP \mathcal{P}$, and $MIN-MAX \mathcal{P}$, in which we minimize the corresponding performance measure for a fixed α and a specific single machine scheduling problem \mathcal{P} , under a given scenario set \mathcal{U} . Notice that the robust $MIN-MAX \mathcal{P}$ problem is a special case of both $MIN-VAR_{\alpha} \mathcal{P}$ and $MIN-CVAR_{\alpha} \mathcal{P}$. Also, $MIN-EXP \mathcal{P}$ is a special case of $MIN-CVAR_{\alpha} \mathcal{P}$.

In the next sections, we provide a number of new positive and negative complexity and approximation results for basic single machine scheduling problems \mathcal{P} . Tables 1, 2 and 3 summarize the known and new results. In Table 1, the negative results for uncertain due dates and deterministic processing times are shown. In Table 2, the negative results for uncertain processing times and deterministic due dates are presented. Finally, in Table 3, some positive results are shown.

4 Some general properties

In this section, we will show some general relationships between the problems with various performance criteria. These properties will be used later to establish some positive and negative complexity results for particular problems.

Theorem 1 The following statements hold:

- 1. If MIN-EXP \mathcal{P} is approximable within $\sigma > 1$ (for $\sigma = 1$ it is polynomially solvable), then MIN-CVAR_{α} \mathcal{P} is approximable within $\sigma\rho$, where $\rho = \min\{\frac{1}{\Pr_{\min}}, \frac{1}{1-\alpha}\}$, for each constant $\alpha \in [0, 1)$.
- 2.a If MIN–EXP \mathcal{P} with K-scenarios is (strongly) NP-hard, then MIN–CVAR_{α} \mathcal{P} with K + 1 scenarios is also (strongly) NP-hard for each constant $\alpha \in [0, 1)$.
- 2.b If MIN–EXP \mathcal{P} with K-scenarios is hard to approximate within $\rho > 1$, then MIN–CVAR $_{\alpha} \mathcal{P}$ with K + 1 scenarios is also hard to approximate within $\rho > 1$ for each constant $\alpha \in [0, 1)$.

Proof We first prove assertion 1. Let π^* minimize the expected cost and π' minimize the conditional value at risk for a fixed $\alpha \in [0, 1)$. We will denote by $\hat{\pi}$ a σ -approximation schedule for MIN-EXP \mathcal{P} . Using Lemma 1 we get

$$\begin{aligned} \mathbf{CVaR}_{\alpha}[\mathbf{F}(\hat{\pi})] &\leq \rho \mathbf{E}[\mathbf{F}(\hat{\pi})] \leq \sigma \rho \mathbf{E}[\mathbf{F}(\pi^*)] \\ &\leq \sigma \rho \mathbf{E}[\mathbf{F}(\pi')] \leq \sigma \rho \mathbf{CVaR}_{\alpha}[\mathbf{F}(\pi')], \end{aligned}$$

and the assertion follows.

| Table 1 | Complexity results | for uncertain due dates (| (processing times are o | deterministic) |
|---------|--------------------|---------------------------|-------------------------|----------------|
|---------|--------------------|---------------------------|-------------------------|----------------|

| \mathcal{P} | Min-Exp \mathcal{P} | $MIN-VAR_{\alpha} \mathcal{P}$ | $MIN-CVAR_{\alpha} \mathcal{P}$ | MIN-MAX \mathcal{P} |
|-----------------------|---|--|--|---|
| $1 p_j = 1 \max T_j$ | Str. NP-hard not appr. within $\frac{7}{6} - \epsilon, \epsilon > 0$ (Kasperski and Zieliński 2016) | Str. NP-hard not at all appr. for any $\alpha \in (0, 1)$ | Str. NP-hard not appr. within $\frac{7}{6} - \epsilon, \epsilon > 0$ for any $\alpha \in [0, 1)$ | Poly. sol. (Kasperski and Zieliński 2016) |
| $1 p_j = 1 \sum U_j$ | Poly sol. (assignment) | Str. NP-hard not at all appr. for any $\alpha \in (0, 1)$ | Str. NP-hard for any $\alpha \in (0, 1)$ | Str. NP-hard not appr. for any constant $\gamma > 1$ |
| $1 \mid\mid \sum U_j$ | NP-hard | As above | As above | As above |
| $1 p_j = 1 \sum T_j$ | Poly sol. (assignment) | Str. NP-hard not at all appr. for any $\alpha \in (0, 1)$ | Str. NP-hard for any $\alpha \in (0, 1)$ | Str. NP-hard not appr. within $\frac{5}{4} - \epsilon, \epsilon > 0$ |
| $1 \sum T_j$ | Str. NP-hard | As above | As above | As above |

 Table 2
 Complexity results for uncertain processing times (the due dates are deterministic)

| \mathcal{P} | Min–Exp \mathcal{P} | $MIN-VAR_{\alpha} \mathcal{P}$ | $MIN-CVAR_{\alpha} \mathcal{P}$ | Min-Max \mathcal{P} |
|---------------|-----------------------|---|---|---|
| $1 \sum C_j$ | Poly sol. | Str. NP-hard not appr. within $\frac{6}{5} - \epsilon, \epsilon > 0$ | Str. NP-hard for any $\alpha \in (0, 1)$ | Str. NP-hard not appr. within $\frac{6}{5} - \epsilon$, $\epsilon > 0$ (Kouvelis and Yu 1997; Mastrolilli et al. 2013) |
| $1 \sum U_j$ | Open | Str. NP-hard for any $\alpha \in (0, 1]$ | Str. NP-hard for any $\alpha \in (0, 1)$ | Str. NP-hard |
| $1 \sum T_j$ | NP-hard (Lawler 1977) | Str. NP-hard not appr. within $\frac{6}{5} - \epsilon, \epsilon > 0$ | Str. NP-hard for any $\alpha \in (0, 1)$ | Str. NP-hard not appr. within $\frac{6}{5} - \epsilon$, $\epsilon > 0$ |

Table 3 Positive complexity results

| \mathcal{P} | Min–Exp \mathcal{P} | $MIN-VAR_{\alpha} \mathcal{P}$ | $MIN-CVAR_{\alpha} \mathcal{P}$ | MIN-MAX \mathcal{P} |
|---|---|---|---|---|
| $1 prec \max w_jT_j$ with $\tilde{d}_j, \tilde{p}_j, w_j$ | $O(f_{\max}^{K} K n^2)$ FPTAS for const. <i>K</i> | $O(f_{\max}^{K} K n^2)$ FPTAS for const. <i>K</i> | $O(f_{\max}^{K}Kn^{2})$ FPTAS for const. K | O(Kn ²) (Kasperski and Zieliński 2016) |
| $1 prec \sum w_j C_j$ with \tilde{p}_j, w_j | As the determ. problem | Appr. within 2 for const. K | Appr. within 2 | Appr. within 2 (Mastrolilli et al. 2013) |
| $1 prec^* \sum w_jC_j$ with \tilde{p}_j, w_j | Poly sol. | Appr. within 2 for const. <i>K</i> | Appr. within $\min\{\frac{1}{1-\alpha}, 2\}$ | Appr. within 2 (Mastrolilli et al. 2013) |
| $1 p_j = 1 \sum w_j U_j$ with \tilde{d}_j, w_j | Poly sol. | - | Appr. within $\min\{\frac{1}{\Pr_{\min}}, \frac{1}{1-\alpha}\}$ | Appr. within <i>K</i> |
| $1 \sum w_j U_j$ with \tilde{d}_j, p_j, w_j | Appr. within $4 + \epsilon$, $\epsilon > 0$ | - | Appr. within $\min\{\frac{4+\epsilon}{\Pr_{\min}}, \frac{4+\epsilon}{1-\alpha}\}$ | Appr. within $(4 + \epsilon)K$, $\epsilon > 0$ |
| $1 p_j = 1 \sum w_j T_j$ with \tilde{d}_j, w_j | Poly sol. | - | Appr. within $\min\{\frac{1}{\Pr_{\min}}, \frac{1}{1-\alpha}\}$ | Appr. within <i>K</i> |
| $1 \mid\mid \sum w_j T_j$ with \tilde{d}_j, p_j, w_j | Appr. within $4 + \epsilon$, $\epsilon > 0$ | - | Appr. within $\min\{\frac{4+\epsilon}{\Pr_{\min}}, \frac{4+\epsilon}{1-\alpha}\}$ | Appr. within $(4 + \epsilon)K$, $\epsilon > 0$ |

 $d_j(\tilde{d}_j), p_j(\tilde{p}_j)$ and w_j stand for deterministic (uncertain) due dates, processing times and weights, respectively, in problem \mathcal{P} ; f_{max} is an upper bound on the cost of any schedule under any scenario; $prec^*$ is a polynomially solvable structure of the precedence constraints; $\Pr_{\min} = \min_{k \in [K]} \Pr[\xi_k]$

In order to prove assertion 2, consider an instance of MIN-EXP \mathcal{P} with $\mathcal{U} = \{\xi_1, \ldots, \xi_K\}$. Fix $\alpha \in (0, 1)$ (the statement trivially holds for $\alpha = 0$) and add one additional scenario ξ' under which the cost of each schedule is 0. (For example, all job processing times are 0 under ξ' .) We fix $\Pr'[\xi'] = \alpha$ and $\Pr'[\xi_i] = \Pr[\xi_i] \cdot (1 - \alpha)$ for each $i \in [K]$. Denote by $F'(\pi)$ the random cost of π under the new scenario set \mathcal{U}' . For each schedule π , we get (see Fig. 3):

$$\mathbf{CVaR}_{\alpha}[\mathbf{F}'(\pi)] = \frac{1}{1-\alpha} \sum_{i \in [K]} \Pr'[\xi_i] f(\pi, \xi_i)$$
$$= \sum_{i \in [K]} \Pr[\xi_i] f(\pi, \xi_i) = \mathbf{E}[\mathbf{F}(\pi)].$$

Hence, there is a cost preserving reduction from MIN-EXP \mathcal{P} with K scenarios to MIN-CVAR $_{\alpha} \mathcal{P}$ with K + 1 scenarios and the theorem follows.

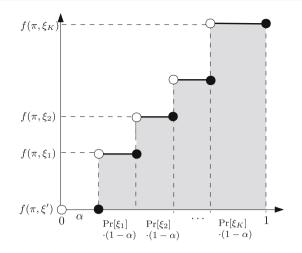


Fig. 3 Illustration of the proof of Theorem 1

Theorem 2 Assume that $w_j = 1$ for each job $j \in J$ in problem \mathcal{P} . The following statements hold:

- 1.a If MIN-MAX \mathcal{P} with $K \geq 2$ scenarios is (strongly) NP-hard, then MIN-VAR_{α} \mathcal{P} with K + 1 scenarios is (strongly) NP-hard for each constant $\alpha \in (0, 1]$.
- 1.b If MIN–MAX \mathcal{P} with $K \geq 2$ scenarios is hard to approximate within $\rho > 1$, then MIN–VAR_{α} \mathcal{P} with K + 1scenarios is hard to approximate within $\rho > 1$ for each constant $\alpha \in (0, 1]$.
- 2. If MIN–MAX \mathcal{P} with $K \ge 2$ scenarios is (strongly) NPhard, then MIN–CVAR $_{\alpha} \mathcal{P}$ with K + 1 scenarios is (strongly) NP-hard for each constant $\alpha \in (0, 1)$

Proof Choose an instance of the MIN–MAX \mathcal{P} problem with $\mathcal{U} = \{\xi_1, \ldots, \xi_K\}, K \ge 2$. Fix $\alpha \in (0, 1)$ and create \mathcal{U}' by adding to \mathcal{U} a dummy scenario ξ' such that the cost of each schedule under ξ' equals M and $M \ge f(\pi, \xi_i)$ for each $i \in [K]$ and each $\pi \in \Pi$. It is enough to fix $p_j(\xi') = p_{\max}$ and $d_j(\xi') = d_{\min}$ for each job $j \in J$, where $p_{\max} = \max_{j \in J, i \in [K]} p_j(\xi_i)$ is the maximum job processing time and $d_{\min} = \min_{j \in J, i \in [K]} d_j(\xi_i)$ is the minimum due date over all scenarios. For each of the two assertions, we define an appropriate probability distribution in \mathcal{U}' . We will use $F'(\pi)$ to denote the random cost of π under \mathcal{U}' .

In order to prove statement 1, we fix $\Pr[\xi'] = 1 - \alpha$ and $\Pr[\xi_i] = \frac{\alpha}{K}$ for each $i \in [K]$ (see Fig. 4a). The equality $\operatorname{VaR}_{\alpha}[F'(\pi)] = \operatorname{Max}[F(\pi)]$ holds. Hence, there is a cost preserving reduction from MIN–MAX \mathcal{P} with *K* scenarios to MIN–VAR $_{\alpha} \mathcal{P}$ with K + 1 scenarios and the statement follows. Note that statement 1 holds trivially for $\alpha = 1$.

Let us now prove statement 2. Assume first that $1-\alpha < \frac{1}{K}$. Define $\Pr[\xi_i] = \frac{1}{K}$ for each $i \in [K]$ and $\Pr[\xi'] = 0$. The dummy scenario is not used in this case. We get $\mathbf{CVaR}_{\alpha}[F'(\pi)] = \mathbf{Max}[F(\pi)]$ and the statement is true. Consider the case, when $1-\alpha \geq \frac{1}{K}$. Fix $\Pr[\xi'] = \gamma$ and $Pr[\xi_i] = \beta$ for each $i \in [K]$, where γ and β satisfy the following system of equations (see Fig. 4b):

$$\begin{cases} \beta + \gamma = 1 - \alpha \\ K\beta + \gamma = 1 \end{cases}$$

In consequence, $\beta = \frac{\alpha}{K-1}$ and $\gamma = 1 - \frac{K\alpha}{K-1}$. Observe that $\beta > 1$ as $\alpha \in (0, 1)$ and $K \ge 2$, and $\gamma \ge 0$, because $\frac{K\alpha}{K-1} \le 1$.

For each schedule π , we get

$$\mathbf{CVaR}_{\alpha}[\mathbf{F}'(\pi)] = \frac{1}{1-\alpha} (\beta \cdot \mathbf{Max}[\mathbf{F}(\pi)] + \gamma M).$$

Hence, MIN-MAX \mathcal{P} and the corresponding instance of MIN-CVAR $_{\alpha}$ \mathcal{P} have the same optimal solutions and the theorem follows.

5 Negative complexity results

In this section, we will prove some negative complexity results for basic single machine scheduling problems. These results are summarized in Tables 1 and 2.

5.1 Uncertain due dates

We first address the problem of minimizing the value at risk criterion. The following theorem characterizes the complexity of some basic problems:

Theorem 3 For each $\alpha \in (0, 1)$, MIN–VAR $_{\alpha} \mathcal{P}$ is strongly NP-hard and not at all approximable, when $\mathcal{P} \in \{1|p_j = 1 | \max T_j, 1|p_j = 1 | \sum T_j, 1|p_j = 1 | \sum U_j\}$ and only due dates are uncertain.

Proof Consider an instance of the following NP-hard MIN 3- SAT problem (Kohli et al. 1994; Avidor and Zwick 2002). We are given boolean variables x_1, \ldots, x_n , a collection of clauses C_1, \ldots, C_m , where each clause is a disjunction of at most 3 literals (variables or their negations), and we ask if there is an assignment to the variables which satisfies at most L < m clauses. We can ask equivalently, if there is an assignment to the variables for which at least l = m - L clauses are not satisfied.

Given an instance of MIN 3- SAT, we create two jobs J_{x_i} and $J_{\overline{x}_i}$ for each variable x_i , $i \in [n]$. A *due date scenario* ξ_i corresponds to clause $C_i = (l_1 \lor l_2 \lor l_3)$ and is formed as follows. For each q = 1, 2, 3, if $l_q = x_j$, then the due date of J_{x_j} is 2j - 1 and the due date of $J_{\overline{x}_j}$ is 2j; if $l_q = \overline{x}_j$, then the due date of J_{x_j} is 2j and the due date of $J_{\overline{x}_j}$ is 2j - 1; if neither x_j nor \overline{x}_j appears in C_i , then the due dates of J_{x_j} and $J_{\overline{x}_j}$ are set to 2j. An example is shown in Table 4.

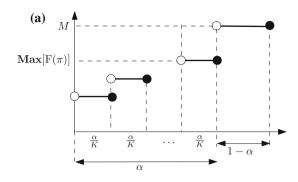


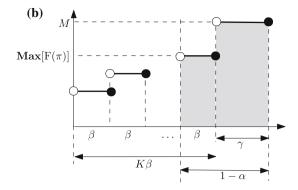
Fig. 4 Illustration of the proof of Theorem 2

Table 4 The set of jobs and the
due date scenarios for the
formula $(x_1 \vee \overline{x}_2 \vee \overline{x}_3) \wedge (\overline{x}_2 \vee \overline{x}_3 \vee x_4) \wedge (\overline{x}_1 \vee x_2 \vee \overline{x}_4) \wedge (x_1 \vee x_2 \vee x_3) \wedge (x_1 \vee x_3 \vee \overline{x}_4)$

| | ξ1 | ξ2 | ξ3 | ξ4 | ξ5 |
|----------------------|----|----|----|----|----|
| J_{x_1} | 1 | 2 | 2 | 1 | 1 |
| $J_{\overline{x}_1}$ | 2 | 2 | 1 | 2 | 2 |
| J_{x_2} | 4 | 4 | 3 | 3 | 4 |
| $J_{\overline{x}_2}$ | 3 | 3 | 4 | 4 | 4 |
| J_{x_3} | 6 | 6 | 6 | 5 | 5 |
| $J_{\overline{x}_3}$ | 5 | 5 | 6 | 6 | 6 |
| J_{x_4} | 8 | 7 | 8 | 8 | 8 |
| $J_{\overline{x}_4}$ | 8 | 8 | 7 | 8 | 7 |

Let us define a subset of the schedules $\Pi' \subseteq \Pi$ such that each schedule $\pi \in \Pi'$ is of the form $\pi =$ $(J_1, J'_1, J_2, J'_2, \dots, J_n, J'_n)$, where $J_j, J'_j \in \{J_{x_j}, J_{\overline{x}_j}\}$ for $j \in [n]$. Observe that Π' contains exactly 2^n schedules and each such a schedule corresponds to the assignment to the variables such that $x_i = 0$ if J_{x_i} is processed before $J_{\overline{x}_i}$ and $x_i = 1$ otherwise. Note that this correspondence is one-to-one. In the following, we assume that $f(\pi, \xi_i)$ is the maximum tardiness, or the total tardiness, or the sum of unit penalties in π under ξ_i . The reasoning will be the same for each of these cost functions. If $\pi \notin \Pi'$, then $f(\pi, \xi_i) > 0$ for each scenario ξ_i . Indeed, suppose that $\pi \notin \Pi'$ and let J_i (J'_i) be the last job in π which is not placed properly, i.e., J_j , $(J'_j) \notin \{J_{x_j}, J_{\overline{x}_j}\}$. Then, J_j (J'_j) is late under all scenarios. On the other hand, if $\pi \in \Pi'$, then the number of scenarios under which no job is late is equal to the number of unsatisfiable clauses for the assignment corresponding to π . Fix $\alpha \in (0, 1)$. We will add to \mathcal{U} one additional scenario ξ' and define a probability distribution in \mathcal{U} , depending on the fixed α , so that the answer to MIN 3- SAT is yes if and only if there is schedule π for which $\operatorname{VaR}_{\alpha}[F(\pi)] \leq 0$. This will prove the stated result. We consider two cases:

1. $l/m \ge \alpha$. We create dummy scenario ξ' under which the due date of all jobs is equal to 0. The probability of this scenario is equal to $\frac{l-\alpha m}{l}$. The probability of each



of the remaining scenarios is equal to $\frac{1}{m}(1 - \frac{l-\alpha m}{l}) = \frac{\alpha}{l}$. Assume that the answer to MIN 3-SAT is yes. So, there is an assignment to the variables which satisfies at most m - l clauses. By the above construction, there is a schedule $\pi \in \Pi'$ whose cost is positive under at most m - l scenarios plus the dummy one. It holds

$$\Pr[\mathbf{F}(\pi) > 0] \le \frac{l - \alpha m}{l} + (m - l)\frac{\alpha}{l} = 1 - \alpha.$$

Hence, $\Pr[F(\pi) \le 0] \ge \alpha$ and $\operatorname{VaR}_{\alpha}[F(\pi)] \le 0$. Assume that the answer to MIN 3- SAT is no. Then, for every schedule π there are more than m - l scenarios under which the cost of π is positive plus the dummy one. Hence, $\Pr[F(\pi) > 0] > (1 - \alpha)$ and $\Pr[F(\pi) \le 0] < \alpha$. In consequence, $\operatorname{VaR}_{\alpha}[F(\pi)] > 0$.

2. $l/m < \alpha$. We create dummy scenario ξ' under which the due date of each job equals 2n. The probability of the dummy scenario is $\frac{m\alpha-l}{m-l}$. The probability of each of the remaining scenarios is equal to $\frac{1}{m}(1 - \frac{m\alpha-l}{m-l}) = \frac{1-\alpha}{m-l}$. Assume that the answer to MIN 3- SAT is yes. So, there is an assignment to the variables which satisfies at most m - l clauses. By the construction, there is a schedule π whose cost is positive under at most m - l scenarios. Hence,

$$\Pr[F(\pi) \le 0] = 1 - \Pr[F(\pi) > 0]$$
$$\ge 1 - (m - l)\frac{1 - \alpha}{m - l} = \alpha$$

and $\operatorname{VaR}_{\alpha}[F(\pi)] \leq 0$. Assume that the answer to MIN 3- SAT is no. Then, for each assignment more than m-l clauses are satisfied. By the construction, for every schedule π there are more than m-l scenarios under which the cost π is positive. Therefore, $\Pr[F(\pi) > 0] >$ $(m-l)\frac{1-\alpha}{m-l} = (1-\alpha)$ and $\Pr[F(\pi) \leq 0] < \alpha$, so $\operatorname{VaR}_{\alpha}[F(\pi)] > 0$.

It follows from Theorem 3 that the problem MIN–VAR_{α} 1 $||\sum w_j T_j$, discussed in Atakan et al. (2017), is strongly

Table 5 The set of jobs and the due date scenarios for the formula $(x_1 \vee \overline{x}_2 \vee \overline{x}_3) \wedge (\overline{x}_2 \vee \overline{x}_3 \vee x_4) \wedge (\overline{x}_1 \vee x_2 \vee \overline{x}_4) \wedge (x_1 \vee x_2 \vee x_3) \wedge (x_1 \vee x_3 \vee \overline{x}_4)$. Schedule $\pi = (J_{x_1}, J_{\overline{x}_1}, J_{\overline{x}_2}, J_{x_2}, J_{x_3}, J_{\overline{x}_3}, J_{\overline{x}_4}, J_{x_4}))$ corresponds to a truth assignment

| | ξ1 | ξ2 | ξ3 | ξ4 | ξ5 | ξ_1' | ξ_2' | ξ'3 | ξ'_4 |
|----------------------|----|----|----|----|----|----------|----------|---------|----------|
| J_{x_1} | 1 | 2 | 2 | 1 | 1 | 1/2 | 8 | 8 | 8 |
| $J_{\overline{x}_1}$ | 2 | 2 | 1 | 2 | 2 | 1/2 | 8 | 8 | 8 |
| J_{x_2} | 4 | 4 | 3 | 3 | 4 | 8 | 2+1/2 | 8 | 8 |
| $J_{\overline{x}_2}$ | 3 | 3 | 4 | 4 | 4 | 8 | 2 + 1/2 | 8 | 8 |
| J_{x_3} | 6 | 6 | 6 | 5 | 5 | 8 | 8 | 4 + 1/2 | 8 |
| $J_{\overline{x}_3}$ | 5 | 5 | 6 | 6 | 6 | 8 | 8 | 4+1/2 | 8 |
| J_{x_4} | 8 | 7 | 8 | 8 | 8 | 8 | 8 | 8 | 6+1/2 |
| $J_{\overline{x}_4}$ | 8 | 8 | 7 | 8 | 7 | 8 | 8 | 8 | 6+1/2 |

NP-hard and not at all approximable even in the restrictive case, in which all job processing times and weights are equal to 1. This negative result is true for each fixed $\alpha \in (0, 1)$. It was shown in Kasperski and Zieliński (2016) that MIN-EXP $1|p_j = 1| \max T_j$ is strongly NP-hard and hard to approximate within $7/6 - \epsilon$ for any $\epsilon > 0$. Hence, we immediately get from Theorem 1 that for each constant $\alpha \in [0, 1)$, MIN-CVAR $_{\alpha} 1|p_j = 1| \max T_j$ is strongly NPhard and hard to approximate within $7/6 - \epsilon$ for any $\epsilon > 0$.

We consider now the problem with the total tardiness criterion. The deterministic $1||\sum T_j$ problem is known to be NP-hard (Lawler 1977). However, $1|p_j = 1|\sum T_j$ is polynomially solvable [see, e.g., Brucker (2007)]. The following result characterizes the complexity of the minmax version of this problem:

Theorem 4 MIN–MAX $1|p_j = 1| \sum T_j$ with uncertain due dates is strongly NP-hard and not approximable within $\frac{5}{4} - \epsilon$ for any $\epsilon > 0$.

Proof We will show a reduction from the NP-complete 3-SAT problem, in which we are given boolean variables x_1, \ldots, x_n , a collection of clauses $C_1, \ldots C_m$, where each clause is a disjunction of at most 3 literals (variables or their negations) and we ask if there is an assignment to the variables which satisfies all clauses [see, e.g., Garey and Johnson (1979)]. Given an instance of 3- SAT, we create two jobs J_{x_j} and $J_{\overline{x}_j}$ for each variable $x_j, j \in [n], |J| = 2n$. A due date scenario ξ_i corresponding to clause $C_i = (l_1 \lor l_2 \lor l_3)$ is created in the same way as in the proof of Theorem 3. Additionally, for each variable x_j we create scenario ξ'_j under which the due dates of J_{x_j} and $J_{\overline{x}_j}$ are $2(j-1) + \frac{1}{2}$ and the due dates of the remaining jobs are set to 2n (see Table 5). We first show that the answer to 3- SAT is yes if and only if there is a schedule π such that $\max_{\xi \in U} \sum_{i \in J} T_j(\pi, \xi) \leq 2$.

Assume that the answer to 3-SAT is yes. Consider schedule $\pi = (J_1, J'_1, J_2, J'_2, \dots, J_n, J'_n)$, where $J_j, J'_j \in \{J_{x_j}, J_{\overline{x_j}}\}$. Furthermore, J_{x_j} is processed before $J_{\overline{x_j}}$ if and

only if $x_j = 1$. Since in every clause at least one literal is true, at most two jobs in π are late under each scenario $\xi_i \in \mathcal{U}$. The tardiness of each job in π under any $\xi_i \in \mathcal{U}$ is at most 1. Furthermore, the total tardiness in π under any ξ'_j is exactly 2. In consequence, $\max_{\xi \in \mathcal{U}} \sum_{i \in J} T_j(\pi, \xi) \le 2$.

Assume that there is a schedule π , such that $\max_{\xi \in \mathcal{U}} \sum_{j \in J} T_j(\pi, \xi) \leq 2$. We claim that $\pi = (J_1, J'_1, J_2, J'_2, \dots, J_n, J'_n)$, where $J_j, J'_j \in \{J_{x_j}, J_{\overline{x}_j}\}$. Suppose that this is not the case, and let J_k (J'_k) be the last job in π which is not placed properly. The completion time of J_k (J'_k) is at least 2k + 1. So, its tardiness under ξ'_k is at least $2k + 1 - (2k - 2 + \frac{1}{2}) = 2.5$. Let $x_j = 1$ if and only if J_{x_j} is processed before $J_{\overline{x}_j}$ in π . Since only two jobs can be late under any ξ_i , this assignment satisfies all clauses and the answer to 3- SAT is yes.

In order to prove the lower approximation bound, it is enough to observe that if the answer to 3- SAT is no, then each schedule has the total tardiness 3 under some scenario ξ_i or 2.5 under some scenario ξ'_j , which gives a gap at least $\frac{5}{4}$.

From the fact that $1||\sum T_j$ is weakly NP-hard [see Du and Leung (1990)], we get immediately that more general MIN-EXP $1||\sum T_j$ problem with uncertain due dates is weakly NP-hard as well. The next theorem strengthens this result.

Theorem 5 MIN–EXP 1|| $\sum T_j$ with uncertain due dates is strongly NP-hard.

Proof We will show a polynomial time reduction from the deterministic $1 || \sum w_i T_i$ problem, which is known to be strongly NP-hard (Lawler 1977). Consider an instance of $1 || \sum w_j T_j$. Let $W = \sum_{j \in J} w_j > 0$ and $P = \sum_{j \in J} p_j$. We build an instance of MIN-EXP 1|| $\sum T_j$ with the same set of jobs J and job processing times $p_i, j \in J$. We create K = |J| = n due date scenarios as follows. Under scenario $\xi_i, j \in [n]$, job j has due date equal to d_i and all the remaining jobs have due dates equal to P. We also fix $Pr[\xi_i] = w_i/W$, $i \in [n]$. For any schedule π , we get $\mathbf{E}[\mathbf{F}(\pi)] = \sum_{i \in [K]} \Pr[\xi_i] \sum_{j \in J} T_j(\pi, \xi_i) =$ $\frac{1}{W} \sum_{i \in [n]} w_i \sum_{j \in J} T_j(\pi, \xi_i)$. By the construction, we get $\sum_{i \in J} T_j(\pi, \xi_i) = [C_i(\pi) - d_i]^+$, so $\mathbf{E}[\mathbf{F}(\pi)] = \frac{1}{W} \sum_{i \in [n]} C_i(\pi)$ $w_i[C_i(\pi) - d_i]^+$. In consequence $1 || \sum w_i T_i$ and MIN-EXP 1 $|| \sum T_i$ have the same optimal solutions and the theorem follows.

It was shown in Aissi et al. (2011) that MIN–MAX $1|p_j = 1|\sum U_j$ with uncertain due dates is strongly NP-hard. The following theorem strengthens this result:

Theorem 6 MIN–MAX $1|p_j = 1| \sum U_j$ with uncertain due dates is not approximable within any constant factor unless P=NP.

Proof Consider the following MIN–MAX 0–1 SELECTION problem. We are given a set of items $E = \{e_1, e_2, ..., e_n\}$ and an integer $q \in [n]$. For each item $e_j, j \in [n]$, there is a cost $c_j(\xi_i) \in \{0, 1\}$ under scenario $\xi_i, i \in [K]$. We seek a selection $X \subseteq E$ of exactly q items, |X| = q, which minimizes the maximum cost over all scenarios, i.e., the value of $\max_{i \in [K]} \sum_{e_j \in X} c_j(\xi_i)$. This problem was discussed in Kasperski et al. (2013), where it was shown that it is not approximable within any constant factor $\gamma > 1$. We will show that there is a cost preserving reduction from MIN– MAX 0–1 SELECTION to the considered scheduling problem, which will imply the stated result.

Given an instance of MIN–MAX 0–1 SELECTION, we build the corresponding instance of MIN–MAX $1|p_j = 1| \sum U_j$ as follows. We create a set of jobs J = E, |J| = n, with deterministic unit processing times. For each $i \in [K]$, if $c_j(\xi_i) = 1$ then $d_j(\xi'_i) = n - q$, and if $c_j(\xi_i) = 0$, then $d_j(\xi'_i) = n$. So, we create *K* due date scenarios that correspond to the cost scenarios of MIN–MAX 0–1 SELECTION.

Suppose that there is a solution X to MIN–MAX 0–1 SELECTION such that $\sum_{e_j \in X} c_j(\xi_i) \leq C$ for each $i \in [K]$. Hence, X contains at most C items, $C \leq q$, with the cost equal to 1 under each scenario. In the corresponding schedule π , we first process n - q jobs from $J \setminus X$ and then the jobs in X in any order. It is easily seen that there are at most C late jobs in π under each scenario ξ'_i , and hence, the maximum cost of schedule π is at most C. Conversely, let π be a schedule in which there are at most C late jobs under each scenario ξ'_i . Clearly $C \leq q$ since the first n - q jobs in π must be on-time in all scenarios. Let us form solution X by choosing the items corresponding to the last q jobs in π . Among these jobs at most C under each scenario, and hence, the cost of X is at most C under each scenario ξ_i .

Thus, by Theorem 2, MIN–CVAR $_{\alpha}$ 1 $|p_j = 1| \sum U_j$ is strongly NP-hard for any $\alpha \in (0, 1)$. (Notice that $p_{\text{max}} = 1$ in the proof of Theorem 2 and in the new scenario set \mathcal{U}' still only due dates are uncertain.)

Theorem 7 MIN-EXP 1|| $\sum U_j$ with uncertain due dates is *NP*-hard.

Proof Consider the deterministic $1||\sum w_j U_j$ problem, which is known to be NP-hard (Karp 1972). The reduction from this problem to MIN-EXP $1||\sum U_j$ is the same as the one in the proof of Theorem 5.

It is worth noting that in the proof of Theorem 7 we require an arbitrary probability distribution in the scenario set and we have shown that the problem is only weakly NP-hard. Its complexity for uniform probability distribution is open.

5.2 Uncertain processing times

In this section, we characterize the complexity of the problems under consideration when only processing times are uncertain. It has been shown in Kouvelis and Yu (1997) that MIN-MAX 1|| $\sum C_i$ is strongly NP-hard. Furthermore, this problem is also hard to approximate within $\frac{6}{5} - \epsilon$ for any $\epsilon > 0$ (Mastrolilli et al. 2013). Using Theorem 2, we can immediately conclude that the same negative result holds for MIN–VAR $_{\alpha}$ 1|| $\sum C_i$ for any $\alpha \in (0, 1]$. Also, strong NP-hardness of the min-max problem implies that MIN-CVAR_{α} 1|| $\sum C_i$ is strongly NP-hard for each fixed $\alpha \in (0, 1)$. Observe that the boundary case $\alpha = 0$ (i.e., MIN-EXP 1|| $\sum C_i$) is polynomially solvable, as it easily reduces to the deterministic $1 || \sum C_i$ problem. Since $1 || \sum C_i$ is a special case of $1 || \sum T_i$, with $d_i = 0$ for each $j \in J$, the same negative results are true for the problem with the total tardiness criterion. Observe, however, that MIN-EXP 1|| $\sum T_i$ is also NP-hard, since the deterministic $1 || \sum T_i$ problem is known to be weakly NP-hard (Du and Leung 1990).

It has been shown in Aloulou and Croce (2008) that MIN-MAX $1||\sum U_j$ with uncertain processing times and deterministic due dates is NP-hard. The following theorem strengthens this result:

Theorem 8 MIN–MAX 1|| $\sum U_j$ with uncertain processing times is strongly NP-hard. This assertion remains true even when all the jobs have a common deterministic due date.

Proof We show a polynomial time reduction from the 3- SAT problem (see the proof of Theorem 4). Given an instance of 3- SAT, we create an instance of MIN–MAX 1|| $\sum U_j$ in the following way. For each variable x_i , we create two jobs J_{x_i} and $J_{\overline{x}_i}$, so J contains 2n jobs. The due dates of all these jobs are the same under each scenario and equal 2. For each clause $C_j = (l_1, l_2, l_3)$, we construct processing time scenario ξ_i , under which the jobs $J_{\overline{l}_1}$, $J_{\overline{l}_2}$, $J_{\overline{l}_3}$ have processing time equal to 1 and all the remaining jobs have processing times equal to 0. Then, for each pair of jobs J_{x_i} , $J_{\overline{x}_i}$ we construct scenario ξ'_i under which the processing times of J_{x_i} , $J_{\overline{x}_i}$ are 2 and all the remaining jobs have processing times equal to 0. A sample reduction is shown in Table 6. We will show that the answer to 3- SAT is yes if and only if there is a schedule π such that max $_{\xi \in \mathcal{U}} \sum_{i \in J} U(\pi, \xi) \leq n$.

Assume that the answer to 3- SAT is yes. Then, there exists a truth assignment to the variables which satisfies all the clauses. Let us form schedule π by processing first the jobs corresponding to true literals in any order and processing then the remaining jobs in any order. From the construction of the scenario set, it follows that the completion time of the *n*th job in π under each scenario is not greater than 2. In consequence, at most *n* jobs in π are late under each scenario and $\max_{\xi \in \mathcal{U}} \sum_{j \in J} U(\pi, \xi) \le n$.

Table 6 Processing time scenarios for the formula $(x_1 \vee \overline{x}_2 \vee \overline{x}_3) \wedge (\overline{x}_2 \vee \overline{x}_3 \vee x_4) \wedge (\overline{x}_1 \vee x_2 \vee \overline{x}_4) \wedge (x_1 \vee x_2 \vee x_3) \wedge (x_1 \vee x_3 \vee \overline{x}_4)$. Schedule $\pi = (J_{x_1}, J_{\overline{x}_2}, J_{x_3}, J_{\overline{x}_4}|J_{\overline{x}_1}, J_{x_2}, J_{\overline{x}_3}, J_{x_4})$ corresponds to a satisfying truth assignment

| | ξ1 | ξ2 | ξ3 | ξ4 | ξ5 | ξ_1' | ξ_2' | ξ'_3 | ξ'_4 | d_i |
|----------------------|----|----|----|----|----|----------|----------|----------|----------|-------|
| J_{x_1} | 0 | 0 | 1 | 0 | 0 | 2 | 0 | 0 | 0 | 2 |
| $J_{\overline{x}_1}$ | 1 | 0 | 0 | 1 | 1 | 2 | 0 | 0 | 0 | 2 |
| J_{x_2} | 1 | 1 | 0 | 0 | 0 | 0 | 2 | 0 | 0 | 2 |
| $J_{\overline{x}_2}$ | 0 | 0 | 1 | 1 | 0 | 0 | 2 | 0 | 0 | 2 |
| J_{x_3} | 1 | 1 | 0 | 0 | 0 | 0 | 0 | 2 | 0 | 2 |
| $J_{\overline{x}_3}$ | 0 | 0 | 0 | 1 | 1 | 0 | 0 | 2 | 0 | 2 |
| J_{x_4} | 0 | 0 | 1 | 0 | 1 | 0 | 0 | 0 | 2 | 2 |
| $J_{\overline{x}_4}$ | 0 | 1 | 0 | 0 | 0 | 0 | 0 | 0 | 2 | 2 |

Assume that there is a schedule π such that $\sum_{j \in J} U(\pi, \xi) \leq n$ for each $\xi \in \mathcal{U}$, which means that at most *n* jobs in π are late under each scenario. Observe first that J_{x_i} and $J_{\overline{x}_i}$ cannot appear among the first *n* jobs in π for any $i \in [n]$; otherwise, more than *n* jobs would be late in π under ξ'_i . Hence, the first *n* jobs in π correspond to a truth assignment to the variables x_1, \ldots, x_n , i.e., when J_l is among the first *n* jobs, then the literal *l* is true. Since $f(\pi, \xi_i) \leq n$, the completion time of the *n*-th job in π under ξ_i , so there are at most two false literals for each clause and the answer to 3- SAT is yes.

We thus get from Theorems 2 and 8 that $MIN-VAR_{\alpha} 1||$ $\sum U_j$ is strongly NP-hard for any $\alpha \in (0, 1]$ and $MIN-CVAR_{\alpha} 1|| \sum U_j$ is strongly NP-hard for any $\alpha \in (0, 1)$. The boundary case with $\alpha = 0$ (i.e., MIN-EXP 1|| $\sum U_j$ with uncertain processing times) is an interesting open problem.

6 Positive complexity results

In this section we establish some positive complexity results. Namely, we provide several polynomial and approximation algorithms for particular problems. A summary of the results can be found in Table 3.

6.1 Problems with uncertain due dates

Consider the MIN-EXP $1|p_j = 1| \sum w_j U_j$ problem with uncertain due dates. We introduce variables $x_{ij} \in \{0, 1\}$, $i \in [n], j \in [n]$, where $x_{ij} = 1$ if $j \in [n]$ is the *i*th job in the schedule constructed. The variables satisfy the assignment constraints, i.e., $\sum_{i \in [n]} x_{ij} = 1$ for each $j \in [n]$ and $\sum_{j \in [n]} x_{ij} = 1$ for each $i \in [n]$. If $x_{ij} = 1$, then the completion time of job *j* equals *i*. Define $c_{ijk} = w_j$ if $i > d_j(\xi_k)$ and $c_{ijk} = 0$ otherwise, for each $i, j \in [n]$ and $k \in [K]$. If the variables x_{ij} describe π , then

$$\mathbf{E}[\mathbf{F}(\pi)] = \sum_{k \in [K]} \sum_{i \in [n]} \sum_{j \in [n]} \Pr[\xi_k] c_{ijk} x_{ij} = \sum_{i \in [n]} \sum_{j \in [n]} c_{ij}^* x_{ij},$$

where $c_{ij}^* = \sum_{k \in [K]} \Pr[\xi_k] c_{ijk}$. Hence, the problem is equivalent to the MINIMUM ASSIGNMENT with the cost matrix c_{ij}^* . The same result holds for MIN–EXP $1|p_j = 1|\sum w_j T_j$. It is enough to define $c_{ijk} = w_j[i - d_j(\xi_k)]^+$ for $i, j \in [n]$, $k \in [K]$. We thus get the following results:

Theorem 9 MIN-EXP \mathcal{P} with uncertain due dates is polynomially solvable, when $\mathcal{P} \in \{1|p_j = 1| \sum w_j U_j, 1|p_j = 1| \sum w_j T_j\}$

From Theorems 9 and 1, we immediately get the following approximation result:

Theorem 10 MIN–CVAR_{α} \mathcal{P} with uncertain due dates is approximable within $\rho = \min\{\frac{1}{\Pr_{\min}}, \frac{1}{1-\alpha}\}$, when $\mathcal{P} \in \{1|p_j = 1| \sum w_j U_j, 1|p_j = 1| \sum w_j T_j\}$.

Since MIN-MAX $1|p_j = 1| \sum w_j U_j$ and MIN-MAX $1|p_j = 1| \sum w_j T_j$ are special cases of the min-max version of MINIMUM ASSIGNMENT, which is approximable within *K* [see, e.g., Aissi et al. (2009)], both problems are approximable within *K* as well.

We now study the MIN-EXP 1|| $\sum w_j T_j$ problem with uncertain due dates and deterministic processing times. This problem is strongly NP-hard since $1||\sum w_j T_j$ is strongly NP-hard. The expected cost of π can be rewritten as $\mathbf{E}[F(\pi)] = \sum_{j \in J} \sum_{i \in [K]} \Pr[\xi_i][C_j(\pi) - d_j(\xi_i)]^+$. We thus get a single machine scheduling problem $1||\sum f_j$ with job-dependent cost functions of form $f_j(C_j(\pi)) =$ $\sum_{i \in [K]} \Pr[\xi_i][C_j(\pi) - d_j(\xi_i)]^+$, $j \in J$. Note also that these functions are nonnegative and nondecreasing with respect to $C_j(\pi)$. The same analysis can be done for the MIN-EXP 1|| $\sum w_j U_j$ problem with uncertain due dates and deterministic processing times. Hence and from Cheung et al. (2017), where a $(4 + \epsilon)$ -approximation algorithm, for any $\epsilon > 0$, for this class of problems was provided, we get the following result (see also Theorem 1).

Theorem 11 If $\mathcal{P} \in \{1 || \sum w_j U_j, 1|| \sum w_j T_j\}$, then MIN– EXP \mathcal{P} with uncertain due dates is approximable within $4 + \epsilon$ and MIN–CVAR_{α} \mathcal{P} is approximable within min $\{\frac{4+\epsilon}{P_{rmin}}, \frac{4+\epsilon}{1-\alpha}\}$, $\epsilon > 0$, for any $\epsilon > 0$ and each constant $\alpha \in [0, 1)$.

When the probability distribution in \mathcal{U} is uniform, then the approximation ratio in Theorem 1 can be improved to min{ $(4 + \epsilon)K, \frac{4+\epsilon}{1-\alpha}$ }. Since MIN–MAX \mathcal{P} is a special case of MIN–CVAR_{α} \mathcal{P} with uniform probability distribution and α sufficiently large, we get that MIN–MAX \mathcal{P} , $\mathcal{P} \in \{1 || \sum w_j U_j, 1 || \sum w_j T_j\}$, is approximable within $(4 + \epsilon)K$ for any $\epsilon > 0$.

6.2 The total weighted flow time criterion

In this section, we focus on the problems with the total weighted flow time criterion. We start by recalling a wellknown property [see, e.g., Mastrolilli et al. (2013)], which states that every such a problem with uncertain processing times and deterministic weights can be transformed into an equivalent problem with uncertain weights and deterministic processing times. This transformation goes as follows. For each processing time scenario ξ_i , $i \in [K]$, we invert the role of processing times and weights obtaining the weight scenario ξ'_i . Formally, $p_i = w_i$ and $w_i(\xi'_i) = p_i(\xi_i)$ for each $i \in [K]$. The new scenario set \mathcal{U}' contains scenario ξ'_i with $\Pr[\xi'_i] = \Pr[\xi_i]$ for each $i \in [K]$. We also invert the precedence constraints, i.e., if $i \rightarrow j$ in the original problem, then $j \rightarrow i$ in the new one. Given a feasible schedule $\pi = (\pi(1), ..., \pi(n))$, let $\pi' = (\pi(n), ..., \pi(1))$ be the corresponding inverted schedule. Of course, schedule π' is feasible for the inverted precedence constraints. It is easy to verify that $f(\pi, \xi_i) = f(\pi', \xi'_i)$ for each $i \in [K]$. In consequence $\mathbf{CVaR}_{\alpha}[F(\pi)] = \mathbf{CVar}_{\alpha}[F'(\pi')]$ and $\operatorname{VaR}_{\alpha}[F(\pi)] = \operatorname{VaR}_{\alpha}[F'(\pi')]$, where $F'(\pi')$ is the random cost of π' for scenario set \mathcal{U}' . Hence, the original problem with uncertain processing times and the new one with uncertain weights have the optimal solutions with the same performance measure.

From now on we make the assumption that the jobs have deterministic processing times p_j , $j \in J$ and $w_j(\xi_i)$ is the weight of job j under scenario ξ_i , $i \in [K]$. The value of **CVaR**_{α}[F(π)], for a fixed schedule π , can be computed by solving the following optimization problem (see the formulation (1)b):

$$\min \quad \gamma + \frac{1}{1 - \alpha} \sum_{i \in [K]} \Pr[\xi_k] u_k$$

s.t.
$$\gamma + u_k \ge \sum_{j \in J} w_j(\xi_k) C_j(\pi) \quad k \in [K]$$
$$u_k \ge 0 \qquad \qquad k \in [K]$$
(6)

Let $\delta_{ij} \in \{0, 1\}, i, j \in [n]$, be binary variables such that $\delta_{ij} = 1$ if job *i* is processed before job *j* in a schedule constructed. The vectors of all feasible job completion times (C_1, \ldots, C_n) can be described by the following system of constraints (Potts 1980):

$$VC: C_{j} = p_{j} + \sum_{i \in J \setminus \{j\}} \delta_{ij} p_{i} \ j \in J$$

$$\delta_{ij} + \delta_{ji} = 1 \qquad i, j \in J, i \neq j$$

$$\delta_{ij} + \delta_{jk} + \delta_{ki} \ge 1 \qquad i, j, k \in J \qquad (7)$$

$$\delta_{ij} = 1 \qquad i \rightarrow j$$

$$\delta_{ij} \in \{0, 1\} \qquad i, j \in J$$

Let us denote by VC' the relaxation of VC, in which the constraints $\delta_{ij} \in \{0, 1\}$ are relaxed with $0 \le \delta_{ij} \le 1$. It

has been proved in Schulz (1996) and Hall et al. (1997) that each vector (C_1, \ldots, C_n) that satisfies VC' also satisfies the following inequalities:

$$\sum_{j \in I} p_j C_j \ge \frac{1}{2} \left(\left(\sum_{j \in I} p_j \right)^2 + \sum_{j \in I} p_j^2 \right) \text{ for all } I \subseteq J.$$
(8)

The formulations (7) and (6) lead to the following mixed integer programming model for MIN–CVAR_{α} 1|*prec*| $\sum w_j C_j$ with uncertain weights:

$$\min \quad \gamma + \frac{1}{1 - \alpha} \sum_{i \in [K]} \Pr[\xi_k] u_k$$

s.t.
$$\gamma + u_k \ge \sum_{j \in J} w_j(\xi_k) C_j \quad k \in [K]$$

Constraints VC
$$u_k \ge 0 \qquad \qquad k \in [K]$$
 (9)

We now solve the relaxation of (9), in which VC is replaced with VC'. Let (C_1^*, \ldots, C_n^*) be the relaxed optimal job completion times and z^* be the optimal value of the relaxation. Consider discrete random variable Y, which takes the value $\sum_{j \in J} w_j(\xi_i) C_j^*$ with probability $\Pr[\xi_i]$, $i \in [K]$. The equality $z^* = \mathbf{CVaR}_{\alpha}[Y]$ holds. We relabel the jobs so that $C_1^* \leq C_2^* \leq \cdots \leq C_n^*$ and form schedule $\pi = (1, 2, \ldots, n)$ in nondecreasing order of C_j^* . Since the vector (C_j^*) satisfies VC', it must also satisfy (8). Hence, setting $I = \{1, \ldots, j\}$, we get

$$\sum_{i=1}^{j} p_i C_i^* \ge \frac{1}{2} \left(\left(\sum_{i=1}^{j} p_i \right)^2 + \sum_{i=1}^{j} p_i^2 \right) \ge \frac{1}{2} \left(\left(\sum_{i=1}^{j} p_i \right)^2 \right)$$

Since $C_j^* \ge C_i^*$ for each $i \in \{1 \dots j\}$, we get $C_j^* \sum_{i=1}^j p_i \ge \sum_{i=1}^j p_i C_i^* \ge \frac{1}{2} (\sum_{i=1}^j p_i)^2$ and, finally $C_j = \sum_{i=1}^j p_j \le 2C_j^*$ for each $j \in J$ —this reasoning is the same as in Schulz (1996).

For each scenario $\xi_i \in \mathcal{U}$, the inequality $f(\pi, \xi_i) = \sum_{j \in J} w_j(\xi_i) C_j \leq 2 \sum_{j \in J} w_j(\xi_i) C_j^*$ holds. By Lemma 2, we have $\mathbf{CVaR}_{\alpha}[F(\pi)] \leq 2 \cdot \mathbf{CVaR}_{\alpha}[Y] = 2z^*$. Since z^* is a lower bound on the value of an optimal solution, π is a 2-approximate schedule. Let us summarize the obtained result.

Theorem 12 MIN–CVAR $_{\alpha}$ 1|*prec*| $\sum w_j C_j$ with uncertain processing times is approximable within 2 for each $\alpha \in [0, 1)$.

This result can be refined when the deterministic $1|prec|\sum w_j C_j$ problem is polynomially solvable [for example, when the precedence constraints form an sp-graph, see, e.g.,

Brucker (2007)]. In this case MIN–EXP $1|prec| \sum w_j C_j$ is polynomially solvable, and we can also apply Theorem 1, which leads to the following result.

Theorem 13 If $1|prec| \sum w_j C_j$ is polynomially solvable, then $MIN-CVAR_{\alpha} \ 1|prec| \sum w_j C_j$ with uncertain processing times is approximable within $min\{\frac{1}{1-\alpha}, 2\}$ for each $\alpha \in [0, 1)$.

Observe that $\frac{1}{1-\alpha} < 2$ for each $\alpha < 0.5$. Let us consider MIN–VAR_{α} 1|*prec*| $\sum w_j C_j$ problem. The value of **VaR**_{α}[F(π)], for a fixed schedule π , can be computed by solving the following MIP problem (see (1)a):

min
$$\theta$$

s.t.
$$\sum_{\substack{j \in J \\ k \in [K]}} w_j(\xi_k) C_j(\pi) - \theta \le M_k \beta_k \ k \in [K]$$

$$\sum_{\substack{k \in [K] \\ \beta_k \in \{0, 1\}}} \Pr[\xi_i] \beta_k \le 1 - \alpha$$
(10)

where M_k is an upper bound on the schedule cost under scenario $\xi_k, k \in [K]$. Using the formulation (7) together with (1), we can get a mixed integer programming formulation for MIN–VAR_{α} 1|*prec*| $\sum w_j C_j$. By replacing the constraints *VC* with relaxed *VC'* in the constructed model, we get a mixed integer problem with *K* binary variables. This problem can be solved in polynomial time when *K* is a constant. The same analysis as for MIN–CVAR_{α} 1|*prec*| $\sum w_j C_j$ (we also use Lemma 2) leads to the following result:

Theorem 14 If the number of scenarios is constant, then $MIN-VAR_{\alpha}$ $1|prec|\sum w_j C_j$ with uncertain processing times is approximable within 2 for each $\alpha \in (0, 1]$.

6.3 The bottleneck objective

In this section, we address a class of single machine scheduling problems with a bottleneck objective, i.e., in which $f(\pi) = \max_{i \in J} f_i(C_i(\pi))$, where $f_i(t)$ is the cost of completing job j at time t. An important and well-known example is $1|prec| \max w_i T_i$, in which the maximum weighted tardiness is minimized. This problem can be solved in $O(n^2)$ time by Lawler's algorithm (Lawler 1973). We will use the fact that the minmax versions of the bottleneck problems are polynomially solvable for a wide class of cost functions (Brauner et al. 2016; Kasperski and Zieliński 2016). In particular, the minmax version of $1|prec| \max w_i T_i$ with uncertain processing times and uncertain due dates can be solved in $O(Kn^2)$ time by using the algorithm constructed in Kasperski and Zieliński (2016). In the following, we will assume that $f(\pi, \xi) = \max_{i \in J} w_i T_i(\pi, \xi)$ for a given scenario $\xi \in \mathcal{U}$. We also assume that job processing times and due dates are nonnegative integers under all scenarios and job weights are positive integers. In consequence, the value of $f(\pi, \xi)$ is a nonnegative integer for each ξ .

Let f_{max} be an upper bound on the schedule cost over all scenarios. Let $h : \mathbb{Q}_{+}^{K} \to \mathbb{Q}_{+}$ be a nondecreasing function with respect to \mathbb{Q}_{+}^{K} , i.e., for any $t, t' \in \mathbb{Q}_{+}^{K}$ if $t'_{i} \leq t_{i}$ for each $i \in [K]$, then $h(t'_1, \ldots, t'_K) \leq h(t_1, \ldots, t_K)$. Suppose that h can be evaluated in g(K) time for a given vector $\mathbf{t} = (t_1, \ldots, t_K) \in \mathbb{Z}_+^K$. Consider the corresponding scheduling problem \mathcal{PS} , in which we seek a feasible schedule $\pi \in \Pi$ minimizing $H(\pi) = h(f(\pi, \xi_1), \dots, f(\pi, \xi_K))$. We can find such a schedule by solving a number of the following auxiliary problems: given a vector $t \in \mathbb{Z}_{+}^{K}$, check if $\Pi(t) = \{\pi \in \Pi : f(\pi, \xi_i) \le t_i, i \in [K]\} \text{ is nonempty,}$ and if so, return any schedule $\pi_t \in \Pi(t)$. From the monotonicity of the function h, it follows that for each $\pi \in \Pi(t)$ the inequality $h(f(\pi, \xi_1), \ldots, f(\pi, \xi_K)) \leq h(t)$ is true. Thus, in order to solve the problem \mathcal{PS} , it suffices to enumerate all possible integral vectors $\mathbf{t} = (t_1, \dots, t_K)$, where $t_i \in \{0, 1, \ldots, f_{\max}\}, i \in [K], \text{ and compute } \pi_t \in \Pi(t) \text{ if }$ $\Pi(t)$ is nonempty. The number of such vectors is $(f_{\text{max}}+1)^K$. A schedule π_t with the minimum value of $H(\pi_t)$ is returned.

The crucial step in this method is solving the auxiliary problem. We now show that this can be done in polynomial time for the bottleneck problem with the maximum weighted tardiness criterion. Given any $t \in \mathbb{Z}_{+}^{K}$, we first form scenario set \mathcal{U}' by specifying the following parameters for each $\xi_i \in \mathcal{U}$ and $j \in J$:

$$p_{j}(\xi_{i}') = p_{j}(\xi_{i}), \quad w_{j}' = 1,$$

$$d_{j}(\xi_{i}') = \max\{C \ge 0 : w_{j}(C - d_{j}(\xi_{i})) \le t_{i}\}$$

$$= t_{i}/w_{i} + d_{i}(\xi_{i}).$$

The scenario set \mathcal{U}' can be built in O(Kn) time. We then solve the minmax problem with scenario set \mathcal{U}' , which can be done in $O(Kn^2)$ time by using the algorithm constructed in Kasperski and Zieliński (2016). If the maximum cost of the schedule π returned is 0, then $\pi_t = \pi$; otherwise, $\Pi(t)$ is empty. Since all the risk criteria considered in this paper are nondecreasing functions with respect to schedule costs over scenarios (see Lemma 2 for $\gamma = 1$) and g(K) is negligible in comparison with Kn^2 , we get the following result.

Theorem 15 MIN-EXP \mathcal{P} , MIN-VAR $_{\alpha}\mathcal{P}$ and MIN-CVAR $_{\alpha}$ \mathcal{P} with uncertain processing times and uncertain due dates are solvable in $O((f_{\text{max}} + 1)^K K n^2)$ time, when \mathcal{P} is $1|prec|\max w_j T_j$.

The above running time is pseudopolynomial if *K* is constant. Notice that the special cases, when \mathcal{P} is $1|prec, p_j = 1|\max T_j$ are solvable in $O(K(n + 1)^{K+2})$ time, which is polynomial if *K* is constant (as we can fix $f_{\max} = n$).

We now show that the problems admit an FPTAS if *K* is a constant and $h(\gamma t) \leq \gamma h(t)$, for any $t \in \mathbb{Q}_+^K$, $\gamma \geq 0$.

First we partition the interval $[0, f_{max}]$ into geometrically increasing subintervals: $[0, 1) \cup \bigcup_{\ell \in [\eta]} [(1+\epsilon)^{\ell-1}, (1+\epsilon)^{\ell}),$ where $\eta = \lceil \log_{1+\epsilon} f_{\max} \rceil$ and $\epsilon \in (0, 1)$. Then, we enumerate all possible vectors $\mathbf{t} = (t_1, \ldots, t_K)$, where $t_i \in$ $\{0,1\} \cup \bigcup_{\ell \in [n]} \{(1+\epsilon)^{\ell}\}, i \in [K], \text{ and find } \pi_t \in \Pi(t)$ if $\Pi(t) \neq \emptyset$. Finally, we output a schedule $\pi_{\hat{t}}$ that minimizes value of $H(\pi_t)$ over the nonempty subsets of schedules. Obviously, the running time is $O((\log_{1+\epsilon} f_{\max})^K (Kn^2 +$ $g(K)) = O((\epsilon^{-1} \log f_{\max})^K (Kn^2 + g(K)))$. Let π^* be an optimal schedule. Fix $\ell_i \in \{0, 1, \dots, \eta\}$ for each $i \in [K]$, such that $(1 + \epsilon)^{\ell_i - 1} \leq f(\pi^*, \xi_i) < (1 + \epsilon)^{\ell_i}$, where we assume that $(1 + \epsilon)^{\ell_i - 1} = 0$ for $\ell_i = 0$. This clearly forces $\Pi((1+\epsilon)^{\ell_1},\ldots,(1+\epsilon)^{\ell_K})\neq\emptyset$. Moreover, $(1+\epsilon)^{\ell_i}\leq$ $(1 + \epsilon) f(\pi^*, \xi_i)$ for $\ell_i, i \in [K]$. By the definition of π_i , we get $H(\pi_{\hat{t}}) \leq h((1+\epsilon)^{\ell_1}, \dots, (1+\epsilon)^{\ell_K})$. Since h is a nondecreasing function and $h(\gamma t) \leq \gamma h(t), h((1 + \gamma t))$ $(\epsilon)^{\ell_1}, \ldots, (1+\epsilon)^{\ell_K}) \le (1+\epsilon)h(f(\pi^*, \xi_1), \ldots, f(\pi^*, \xi_K)).$ Hence, $H(\pi_{\hat{t}}) \leq (1 + \epsilon)H(\pi^*)$. By Lemma 2, the risk criteria satisfy the additional assumption on the function h(t). This leads to the following theorem:

Theorem 16 MIN–EXP \mathcal{P} , MIN–VAR $_{\alpha}\mathcal{P}$ and MIN–CVAR $_{\alpha}$ \mathcal{P} with uncertain processing times and uncertain due dates admit an FPTAS, when \mathcal{P} is $1|prec|\max w_jT_j$ and the number of scenarios is constant.

It is worth pointing out that the problems from Theorem 16 are strongly NP-hard when the number of scenarios is a part of the input (see Sect. 5.1). If the number of scenarios is constant, then proving their weak NP-hardness is an interesting open problem.

7 Conclusions and open problems

In this paper, we have discussed a wide class of single machine scheduling problems with uncertain job processing times and due dates. This uncertainty is modeled by a discrete scenario set with a known probability distribution. In order to compute a solution, we have applied the risk criteria, namely the value at risk and conditional value at risk. The expectation and the maximum criteria are special cases of these risk measures. We have provided a number of negative and positive complexity results for problems with basic cost functions. Moreover, we have sharpened some negative ones obtained in Aissi et al. (2011) and Aloulou and Croce (2008). The picture of the complexity is presented in Tables 1, 2 and 3. Obviously, the negative results obtained remain true for more general cases, for instance, for the problems with more than one machine.

There is still a number of interesting open problems on the models discussed. The negative results for uncertain due dates assume that the number of due dates scenarios is a part of input. The complexity status of the problems when the number of due date scenarios is fixed (in particular, equals 2) is open. For uncertain processing times, an interesting open problem is MIN–EXP 1|| $\sum U_j$ (see Table 2). There is still a gap between the positive and negative results; in particular, we conjecture that the negative results for MIN–VAR \mathcal{P} for uncertain processing times (see Table 2) can be strengthened. Now they are just the same as for the MIN–MAX \mathcal{P} .

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