







COVID-19

Risk factors for positive and negative COVID-19 tests: a cautious and in-depth analysis of UK biobank data

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Abstract

Background: The recent COVID-19 outbreak has generated an unprecedented public health crisis, with millions of infections and hundreds of thousands of deaths worldwide. Using hospital-based or mortality data, several COVID-19 risk factors have been identified, but these may be confounded or biased.

Methods: Using SARS-CoV-2 infection test data ($n=4509$ tests; 1325 positive) from Public Health England, linked to the UK Biobank study, we explored the contribution of demographic, social, health risk, medical and environmental factors to COVID-19 risk. We used multivariable and penalized logistic regression models for the risk of (i) being tested, (ii) testing positive/negative in the study population and, adopting a test negative design, (iii) the risk of testing positive within the tested population.

Results: In the fully adjusted model, variables independently associated with the risk of being tested for COVID-19 with odds ratio >1.05 were: male sex; Black ethnicity; social disadvantage (as measured by education, housing and income); occupation (healthcare worker, retired, unemployed); ever smoker; severely obese; comorbidities; and greater exposure to particulate matter (PM) 2.5 absorbance. Of these, only male sex, non-White ethnicity and lower educational attainment, and none of the comorbidities or health risk factors, were associated with testing positive among tested individuals.

Conclusions: We adopted a careful and exhaustive approach within a large population-based cohort, which enabled us to triangulate evidence linking male sex, lower educational attainment and non-White ethnicity with the risk of COVID-19. The elucidation of the joint and independent effects of these factors is a high-priority area for further research to inform on the natural history of COVID-19.

Key words: COVID-19, SARS-CoV-2, prospective cohort, UK Biobank, infection, test data

Key Messages

- We use data from the SARS-CoV-2 infection test ($n = 4509$, including 1325 positive and 3184 negative tests) from Public Health England, linked to the UK Biobank study ($n = 488\,083$).
- Adjusting for potential confounding, male sex; Black ethnicity; social disadvantage (as measured by education, housing and income); occupation (healthcare worker, retired, unemployed); ever smoker; severely obese; comorbidities; and higher exposure to particulate matter (PM) 2.5 absorbance were independently associated with the risk of being tested for COVID-19.
- We found that male sex, non-White ethnicity, and lower educational attainment were independently associated with testing positive among tested individuals.
- None of the health risk factors or comorbidities associated with the risk of being tested were found associated with the risk of testing positive, conditional on being tested.
- We adopted complementary analytical approaches to explore the data and were able to triangulate evidence linking social factors, ethnicity and environmental exposures to COVID-19 risk.
- The elucidation of joint and independent effects of ethnicity, social and environmental factors we report is a high-priority area for further research, which may help clarify the natural history of COVID-19, and suggests possible new avenues for its prevention, diagnosis and treatment.

Background

On 31 December 2019, in Wuhan, China, an outbreak of COVID-19 caused by the novel coronavirus SARS-CoV-2 was first reported. Since then, the infection has spread across continents and is classified as a global pandemic by the World Health Organization.¹ As of 28 May 2020, there have been more than 5.8 million confirmed cases worldwide, and nearly 360 000 deaths; the UK is second only to the USA in number of reported COVID-19 deaths.²

Disease severity appears to be associated with older age, being male and having a range of comorbidities. Severe disease may result in acute respiratory distress syndrome and death.^{3–10} Disease outbreaks have led to rapid saturation of healthcare services, especially intensive care units (ICUs) in regions and conurbations in China, Europe, the USA and elsewhere.^{11,12} In response, many governments implemented quarantine measures¹³ to curtail the spread of infection and limit the number of avoidable deaths.

In the UK, COVID-19 was first documented at the end of January, 2020, although regression-based modelling has inferred probable community spread before detection of first cases in many Western countries.¹⁴ Early in the

epidemic, testing included community cases with typical symptoms or people returning from high-risk areas; this approach was abandoned and testing was then almost exclusively reserved for patients presenting to hospital with high suspicion of COVID-19 based on symptoms and/or clinical/radiological findings.¹⁵ By 23 March, 6650 cases had been reported in the UK and a nationwide lockdown was implemented.

We present here an analysis of UK Biobank data identifying risk factors for testing positive or negative for SARS-CoV-2 infection up to 18 May 2020, as well as those discriminating test positive vs test negative individuals using a test negative design approach.¹⁶

Study and methods

Study population

UK Biobank is a population-based prospective cohort including 502 506 volunteers with active consent, aged 40–69 years at recruitment from 2006 to 2010. At the latest follow-up, 14 423 participants had died leaving

$n = 488\,083$ participants for the present analyses.¹⁷ Each participant provided data on lifestyle, exposures, sociodemographic factors, medical history and medications.

Results of COVID-19 tests from Public Health England's Second Generation Surveillance System microbiology database were linked to UK Biobank participants.¹⁸ These only included 'Pillar 1' data, i.e. swab testing in Public Health England (PHE) labs and NHS hospitals for those with a clinical need, and health and care workers. Samples were analysed using a reverse transcriptase-polymerase chain reaction (RT-PCR) test for SARS-CoV-2. Most of the samples are from combined nose/throat swabs (67%) and upper respiratory tract (25%). In intensive care settings, lower respiratory samples may also be analysed. These included results of 7539 tests from 4509 UK Biobank participants (1824 tested more than once) between 16 March and 18 May 2020 (as available 25 May 2020). Tested participants were classified positive if at least one of their test results was positive, and negative otherwise. Tested participants were considered inpatients if reported as such in the microbiological record for at least one of their tests ($n = 3186$). Inpatient tests arise from specimens collected from an acute (emergency) care provider, an A&E department or an inpatient location. However, microbiology data source has not been linked to admissions data. It can therefore happen that data reported as being from outpatients can arise from an inpatient.

Participant characteristics

Variables were grouped into five categories: demographic, social, health risk, medical factors and environmental exposures (Supplementary Table 1A, available as Supplementary data at *IJE* online).

Demographic variables comprised age calculated as of 31 January 2020, at the time of the first diagnosed UK COVID-19 case, sex and ethnicity, defined as White, Black and Other (South Asian or other ethnic groups).

Social variables included education measured by highest level of qualification attained in three categories: high (College or University degree), intermediate (A/AS levels, O levels/GCSEs, CSEs, NVQ or HND, or equivalent, and other professional qualifications) and low (none of the above). Housing was (i) type of accommodation (house/bungalow or flat), (ii) whether rented or owned, and if owned outright or with a mortgage, and (iii) number of individuals in household. Average household income was included in four categories: <GBP 18 000; GBP 18 000–30 999; GBP 31 000–51 999; >GBP 52 000. Occupation at recruitment was categorized as: retired, employed healthcare workers (including health professionals, health and social welfare associated professionals, healthcare and

related personal services, health and social service managers and hospital porters), employed non-healthcare workers (in paid employment or self-employed) and unemployed (including studying and doing unpaid voluntary work).

Health risk factors were smoking, alcohol drinking (current, former or never), and body mass index (BMI) categorised as: <25, 25–30, 30–40 and >40 kg/m². Comorbidities were derived from self-reported illness at baseline and diagnoses from linked Hospital Episode Statistics data (yes/no): (i) cancer, (ii) cardiovascular disease, (iii) hypertension, (iv) diabetes, (v) respiratory disease, and (vi) autoimmune disease (Supplementary Table 1B, available as Supplementary data at *IJE* online). Number of reported medications at recruitment was categorized as: 0, 1, >1.

Finally, modelled levels of environmental exposure (continuous variables) to nitrogen oxides (NO_x) and particulate matter for particle of diameter smaller than 10 or 2.5 micrometers (PM₁₀, PM_{2.5}), and to soot (PM_{2.5} absorbance), were estimated from residential address in 2010 using land-use regression models at the European level.¹⁹

Statistical analyses

We compared means or proportions for each covariate between tested and non-tested participants (Figure 1); differences between the two populations were assessed using Student's *t*-test (continuous variables: age and exposure levels) and chi-squared test (categorical variables). We compared (i) the tested and non-tested populations at two stages of the UK epidemic (before and after 10 April 2020), (ii) inpatients and outpatients, and (iii) healthcare and non-healthcare workers.

We used univariate logistic regression to model for each covariate the risk of (i) being tested for COVID-19 (tested vs non-tested), (ii) test positive (confirmed case) vs non-tested, (iii) test negative (suspected case) vs non-tested. In order to account for a potential bias in the decision to test and heterogeneity among the participants who tested negative, some of whom may have had illnesses other than COVID-19, we adopted test-negative case-control design,¹⁶ modelling the risk of testing positive conditional on being tested. Such design circumvents some of the selection biases as the analyses are restricted to symptomatic people that were tested. To implement this approach we modelled the risk of testing positive in the tested population only. Continuous covariates (age and environmental exposures) were standardized to ensure comparability and resulting odds ratios (ORs) were expressed as the risk change for a one standard deviation increase in the value of the covariate.

For the four aforementioned analyses, we additionally accounted for correlation across covariates using logistic

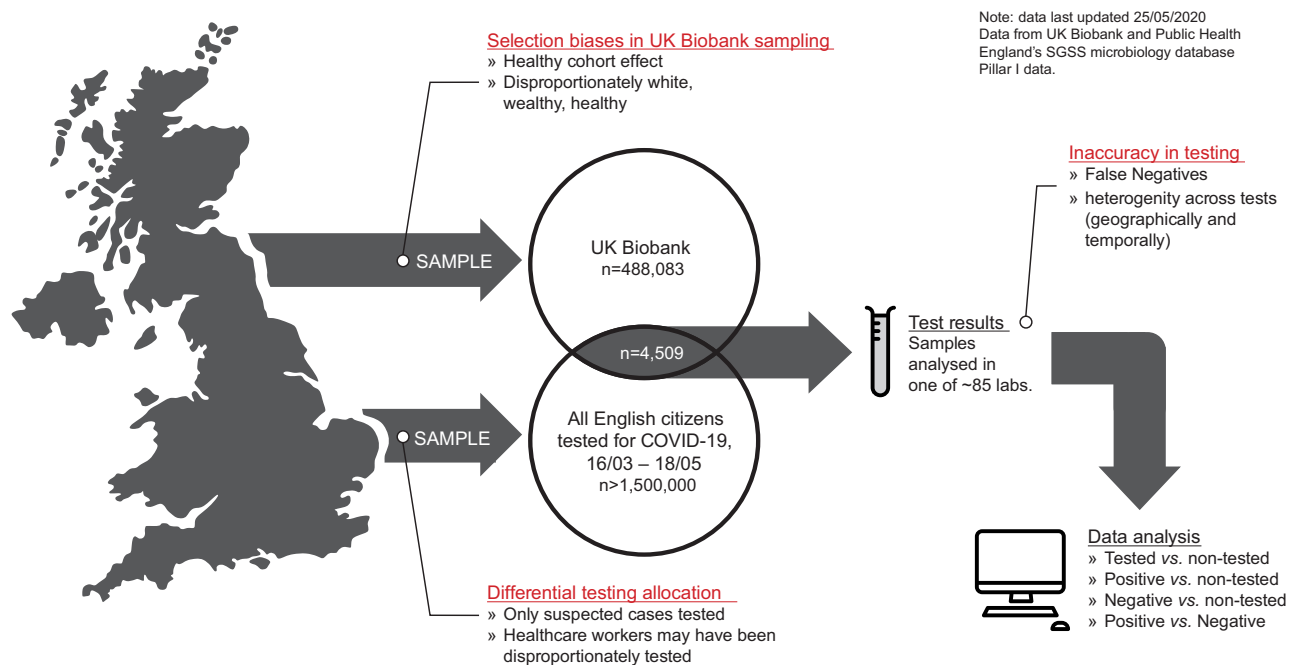


Figure 1 Overview of the data workflow, depicting the synthesis of data from the UK Biobank for COVID-19 testing data. Key biases that are innate to the data gathering processes and test allocation are annotated, as these impact the statistical inferences that can be made from the data.

Least Absolute Shrinkage and Selection Operator (LASSO) regression to model joint effects,²⁰ calibrated using 10-fold cross-validation minimizing the binomial deviance. We investigated stability of the variables selected by fitting logistic LASSO models on ($n=1000$) random 80% subsamples of the full population. As a measure of relevance for each variable, we report its selection proportion.²¹

To account for multiple confounding, we sequentially adjusted for (i) age and sex (base model), (ii) social factors, (iii) health risk factors, (iv) medical variables (comorbidities and number of medications), and (v) environmental factors.

Sensitivity analysis was conducted by separately considering models with healthcare workers and non-healthcare workers.

All analyses were performed in R, version 3.6.3.

Results

Descriptive statistics and univariate analyses

Descriptive statistics comparing the UK Biobank non-tested population ($n=483\,574$) and those tested for COVID-19 ($n=4509$ tested individuals, $n=1325$ positive, and $n=3184$ negative) are shown in [Table 1](#). The distribution of tested participants in relation to the number of tests they underwent is summarized in [Supplementary Table 2](#), available as [Supplementary data](#) at *IJE* online, and shows an excess of men in those who were tested more than

twice. Results of univariate logistic models are shown in [Figure 2](#) and [Supplementary Table 3](#), available as [Supplementary data](#) at *IJE* online.

The probability of being tested was significantly higher in older individuals, among men and people of non-White ethnicity. Tested individuals were more likely to be of lower socio-economic status (SES): having lower educational attainment, living in (i) a flat, (ii) rented accommodation, (iii) a household with an average income <GBP 18 000/year and less likely to be from a household with an average income >GBP 31 000/year. In addition, tested individuals were more likely to have been retired, healthcare workers, unemployed, ever smokers, former or never drinkers, overweight, obese and severely obese. Comorbidities were associated with an increased risk of being tested: cancer, cardiovascular disease, hypertension, diabetes, respiratory disease, autoimmune disease and reporting use of more than one medication. Tested participants were also exposed to higher levels of air pollutants at residence ([Figure 2](#) and [Supplementary Table 3](#), available as [Supplementary data](#) at *IJE* online).

Most of these associations (direction and statistical significance) also held for test positive or test negative separately compared to the non-tested population ([Figure 2](#), [Supplementary Table 3](#), available as [Supplementary data](#) at *IJE* online).

Among the tested population, the risk of testing positive compared with testing negative was (i) higher in men, non-

Table 1 Characteristics of the non-tested and tested populations for COVID-19 in the UK Biobank study. For each variable, the difference between the tested and non-tested populations is evaluated using a Student's *t*-test (for age and environmental exposures) comparing the mean values in the tested and non-tested populations or a chi-squared test (for categorical variables)

| | | Non-tested (n = 483 574) | | Tested (n = 4509) | | P-value (tested vs non-tested) |
|------------------------------|--------------------------------------|--------------------------|----------------------|-------------------|-------------------------|--------------------------------|
| | | n | Mean (SD)/proportion | n | Mean (SD)/proportion | |
| Demographics | Age (years) | 483 574 | 68.06 (8.10) | 4509 | 68.64 (8.88) | 1.66×10^{-5} |
| | Sex | | | | | 7.35×10^{-7} |
| | Female | 265 408 | 54.88% | 2308 | 51.19% | |
| | Male | 218 166 | 45.12% | 2201 | 48.81% | |
| | Ethnicity | | | | | 9.42×10^{-36} |
| | White | 454 766 | 94.56% | 4067 | 90.74% | |
| | Black | 7779 | 1.62% | 167 | 3.73% | |
| | Other | 18 385 | 3.82% | 248 | 5.53% | |
| Social | Education | | | | | 8.39×10^{-31} |
| | High | 156 615 | 33.04% | 1250 | 28.53% | |
| | Intermediate | 237 498 | 50.11% | 2108 | 48.12% | |
| | Low | 79 887 | 16.85% | 1023 | 23.35% | |
| | Type of accommodation | | | | | 4.94×10^{-16} |
| | House | 431 992 | 90.13% | 3826 | 86.46% | |
| | Flat | 47 286 | 9.87% | 599 | 13.54% | |
| | Own or rent accommodation | | | | | 6.71×10^{-70} |
| | Own outright | 248 573 | 52.65% | 2057 | 47.27% | |
| | Own with a mortgage | 178 749 | 37.86% | 1535 | 35.27% | |
| | Rent | 44 782 | 9.49% | 760 | 17.46% | |
| | Number in household | 479 386 | 2.44 (1.32) | 4421 | 2.42 (1.52) | 2.29×10^{-1} |
| | Average household income (GBP) | | | | | 1.25×10^{-46} |
| | <18 000 | 91 289 | 22.27% | 1197 | 32.02% | |
| | 18 000–30 999 | 103 946 | 25.35% | 918 | 24.56% | |
| 31 000–51 999 | 107 804 | 26.29% | 834 | 22.31% | | |
| >52 000 | 106 960 | 26.09% | 789 | 21.11% | | |
| Occupation | | | | | 1.19×10^{-179} | |
| Unemployed | 66 382 | 13.89% | 784 | 17.61% | | |
| Employed (healthcare worker) | 27 789 | 5.81% | 628 | 14.11% | | |
| Employed (other) | 239 832 | 50.17% | 1497 | 33.63% | | |
| Retired | 144 030 | 30.13% | 1542 | 34.64% | | |
| Health risk factors | Smoking status | | | | | 8.89×10^{-21} |
| | Never | 266 023 | 55.33% | 2169 | 48.49% | |
| | Previous | 165 298 | 34.38% | 1718 | 38.41% | |
| | Current | 49 474 | 10.29% | 586 | 13.10% | |
| | Alcohol drinker status | | | | | 1.75×10^{-18} |
| | Never | 21 434 | 4.45% | 278 | 6.20% | |
| | Previous | 16 816 | 3.49% | 241 | 5.37% | |
| | Current | 443 765 | 92.06% | 3968 | 88.43% | |
| | Body mass index (kg/m ²) | | | | | 7.93×10^{-40} |
| | <25 | 159 707 | 33.22% | 1221 | 27.42% | |
| 25–30 | 204 524 | 42.54% | 1847 | 41.48% | | |
| 30–40 | 107 471 | 22.35% | 1212 | 27.22% | | |
| ≥40 | 9076 | 1.89% | 173 | 3.89% | | |
| Medical | Cancer | | | | | 1.19×10^{-19} |
| | No | 405 967 | 83.95% | 3560 | 78.95% | |
| | Yes | 77 607 | 16.05% | 949 | 21.05% | |
| | Cardiovascular | | | | | 2.05×10^{-118} |
| | No | 387 208 | 80.07% | 2985 | 66.20% | |
| Yes | 96,366 | 19.93% | 1524 | 33.80% | | |

(Continued)

Table 1 Continued

| | | Non-tested (n = 483 574) | | Tested (n = 4509) | | P-value (tested vs non-tested) |
|---------------|--------------------------------------|--------------------------|----------------------|-------------------|----------------------|--------------------------------|
| | | n | Mean (SD)/proportion | n | Mean (SD)/proportion | |
| | Hypertension | | | | | 3.00×10^{-67} |
| | No | 326 106 | 67.44% | 2492 | 55.27% | |
| | Yes | 157 468 | 32.56% | 2017 | 44.73% | |
| | Diabetes | | | | | 4.78×10^{-79} |
| | No | 449 911 | 93.04% | 3870 | 85.83% | |
| | Yes | 33 663 | 6.96% | 639 | 14.17% | |
| | Respiratory | | | | | 2.12×10^{-47} |
| | No | 381 403 | 78.87% | 3157 | 70.02% | |
| | Yes | 102 171 | 21.13% | 1352 | 29.98% | |
| | Autoimmune | | | | | 9.22×10^{-30} |
| | No | 415 457 | 85.91% | 3607 | 80.00% | |
| | Yes | 68 117 | 14.09% | 902 | 20.00% | |
| | Number of medications | | | | | 4.43×10^{-37} |
| | 0 | 134 487 | 27.86% | 988 | 21.96% | |
| | 1 | 92 200 | 19.10% | 690 | 15.33% | |
| | >1 | 256 082 | 53.04% | 2822 | 62.71% | |
| Environmental | NO _x (µg/m ³) | 476 556 | 44.06 (15.50) | 4442 | 46.06 (15.93) | 7.76×10^{-17} |
| | PM10 (µg/m ³) | 443 941 | 16.24 (1.90) | 4438 | 16.37 (1.85) | 1.16×10^{-6} |
| | PM2.5 (absorbance/m) | 443 941 | 1.19 (0.27) | 4438 | 1.23 (0.29) | 1.09×10^{-19} |
| | PM2.5 (µg/m ³) | 443 941 | 9.99 (1.06) | 4438 | 10.14 (1.08) | 9.34×10^{-20} |

White individuals, participants with lower educational attainment, renting and owning with a mortgage, living with more people, never drinkers, overweight and obese individuals, and those exposed at residence to higher environmental levels of NO_x and PM2.5, and (ii) slightly lower in older individuals, current smokers, participants previously diagnosed with cancer or autoimmune disease (Supplementary Table 3, available as Supplementary data at *IJE* online).

Comparison of the tested populations in the first and second parts of the UK epidemic (before and from 10 April 2020) showed some differences in that from 10 April, tested participants were younger, included fewer men, had slightly higher SES and fewer comorbidities (Supplementary Table 4, available as Supplementary data at *IJE* online). Compared with outpatients, inpatients were more likely to be older, White, of male sex, of lower SES, to have underlying comorbidities and to be on more than one medication (Supplementary Table 5, available as Supplementary data at *IJE* online).

The healthcare workers tested for COVID-19 differed from the rest of the tested population according to all characteristics except environmental exposures and type of accommodation. In particular, they were younger, more affluent (as measured by income), showed a lower prevalence of all comorbidities, a higher proportion of non-White and lean individuals, of women and of never-smokers (Supplementary Table 6, available as Supplementary data at *IJE* online).

Multivariate and attenuation analyses

We used logistic LASSO models to account for correlation and joint contribution across covariates. Being younger, a man, of non-White background, ever smoker, non-drinker, overweight, obese or severely obese, on more than one medication, of lower SES, exposed to higher environmental levels of NO_x, PM2.5 or PM2.5 absorbance or having comorbidities were all found to jointly contribute to a higher probability of being tested (Figure 3a, blue). Models for the probability of being a confirmed or suspected case in the full UK Biobank population (Figure 3a, beige and green, respectively) selected fewer but consistent sets of predictors ($n = 33$ and 32 respectively). Models for the risk of testing positive, conditional on being tested (Figure 3a, orange), selected 26 covariates notably including age, gender, ethnicity, educational attainment and occupation, obesity, having had a cancer, diabetes, cardiovascular or autoimmune disease, using one or more medication and environmental exposures.

Stability analyses showed that the frequently selected variables to predict being a confirmed case (Figure 3B, $n = 28$ variables with selection proportion >80%) were also selected to predict the probability of being tested (irrespective of the outcome of the test), with a selection proportion close to 100% (except for PM2.5 and household income >GBP 52 000 whose selection proportions were 78 and 85% respectively). Eighteen variables jointly

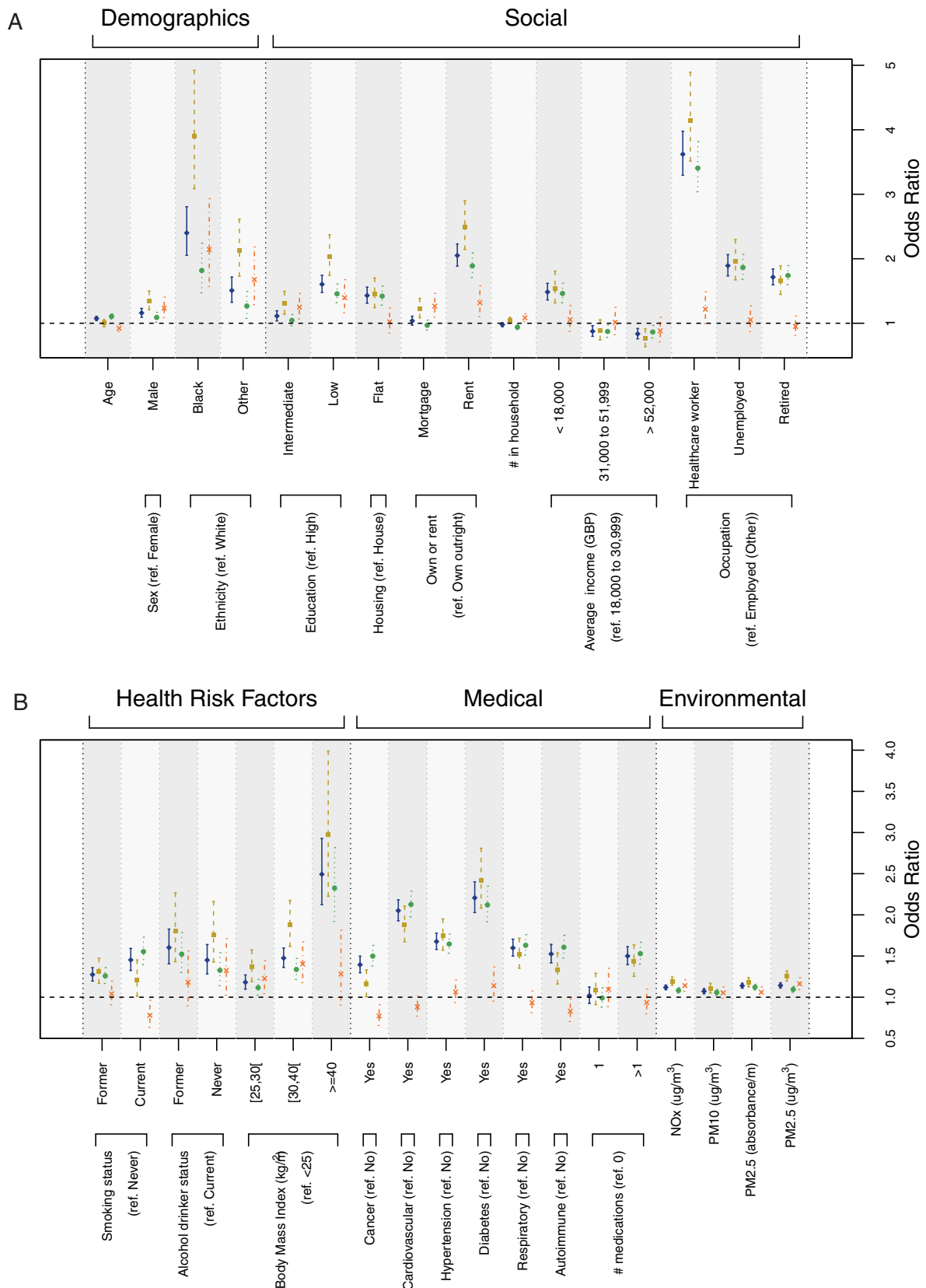


Figure 2 Results from the univariate logistic models predicting from each covariate separately the risk of (i) being tested for COVID-19 (outcome: tested vs non-tested, in blue, plain line), (ii) being tested positive for COVID-19 (outcome: tested positive vs non-tested, in beige dashed line), (iii) being tested negative for COVID-19 (outcome: tested negative vs non-tested, in green, dotted line), and (iv) being tested positive conditional on being tested (outcome: tested positive vs tested negative, in orange dashed/dotted line). Effect size estimates are expressed as odds ratios and are represented for demographic covariates and social factors (A), health risk, medical and environmental factors (B).

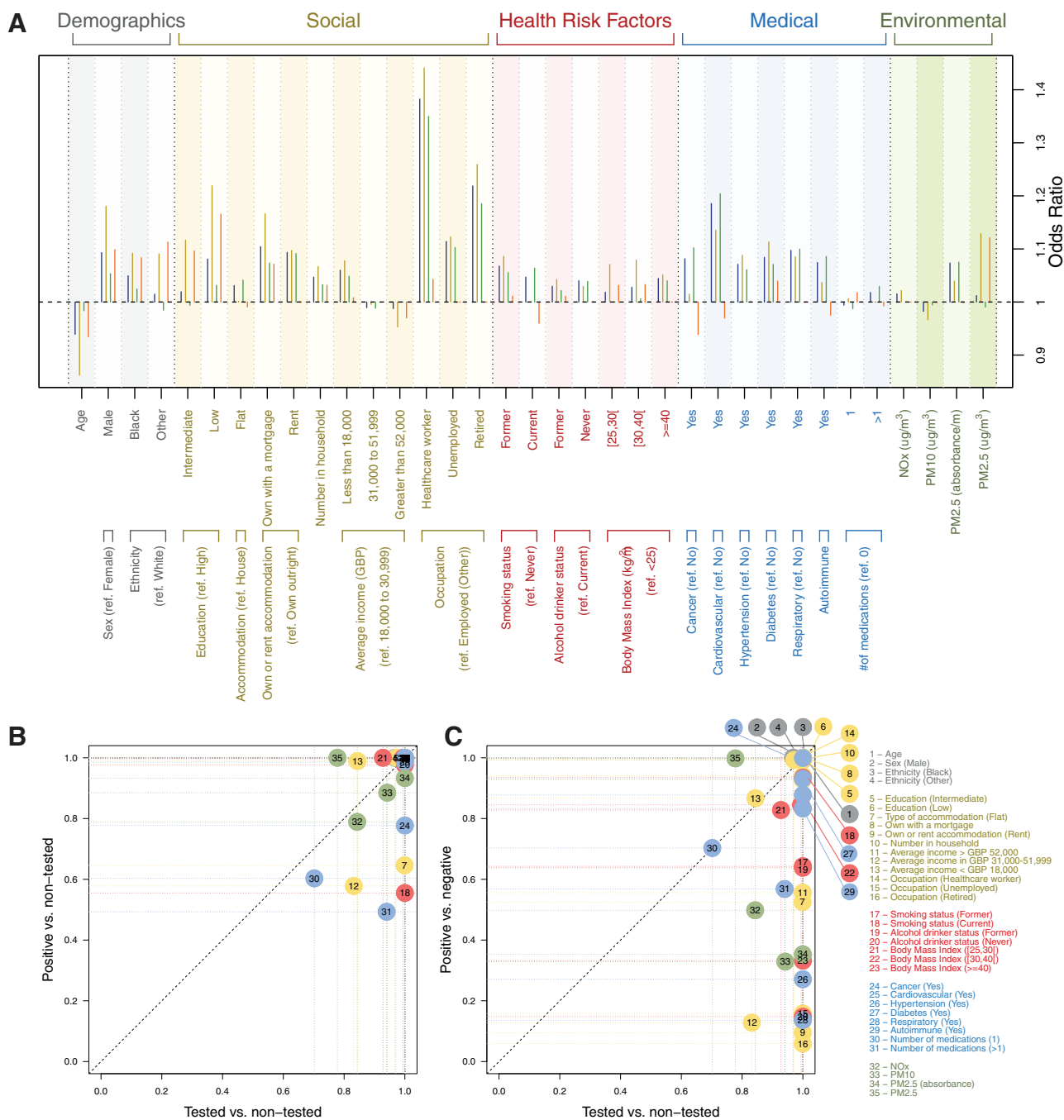


Figure 3 Penalized odds ratios from logistic-LASSO models regressing jointly all predictors against the risk of (i) being tested for COVID-19 (outcome: tested vs non-tested, in blue), (ii) being tested positive for COVID-19 (outcome: tested positive vs non-tested, in beige), (iii) being tested negative for COVID-19 (outcome: tested negative vs non-tested, in green), and (iv) being tested positive conditionally on being tested (outcome: tested positive vs tested negative, in orange) (A). Selection proportion from stability analysis of the LASSO for (i) the risk of testing positive in the full population (B), and (ii) the risk of testing positive for COVID-19 conditionally on being tested (γ -axis) (C) against the selection proportion for the model of the probability of being tested (x -axis). Selection proportions were inferred from 1000 models based on an 80% subsample of the population and are reported for each of the demographics (in grey, $n = 4$), social (brown, $n = 12$), health risk (red, $n = 7$), medical (blue, $n = 8$), and environmental (green, $n = 4$) factors.

differentiated testing positive vs testing negative with selection proportion $>80\%$: being younger, a man, of non-White ethnicity, of lower educational attainment, owning with a mortgage, living with more people, lower income, being a healthcare worker, currently smoking, being

overweight or obese, exposed to higher levels of PM2.5, previously diagnosed with cancer, diabetes, cardiovascular or autoimmune disease.

In the fully adjusted model (Figure 4), variables independently associated with testing positive or negative with

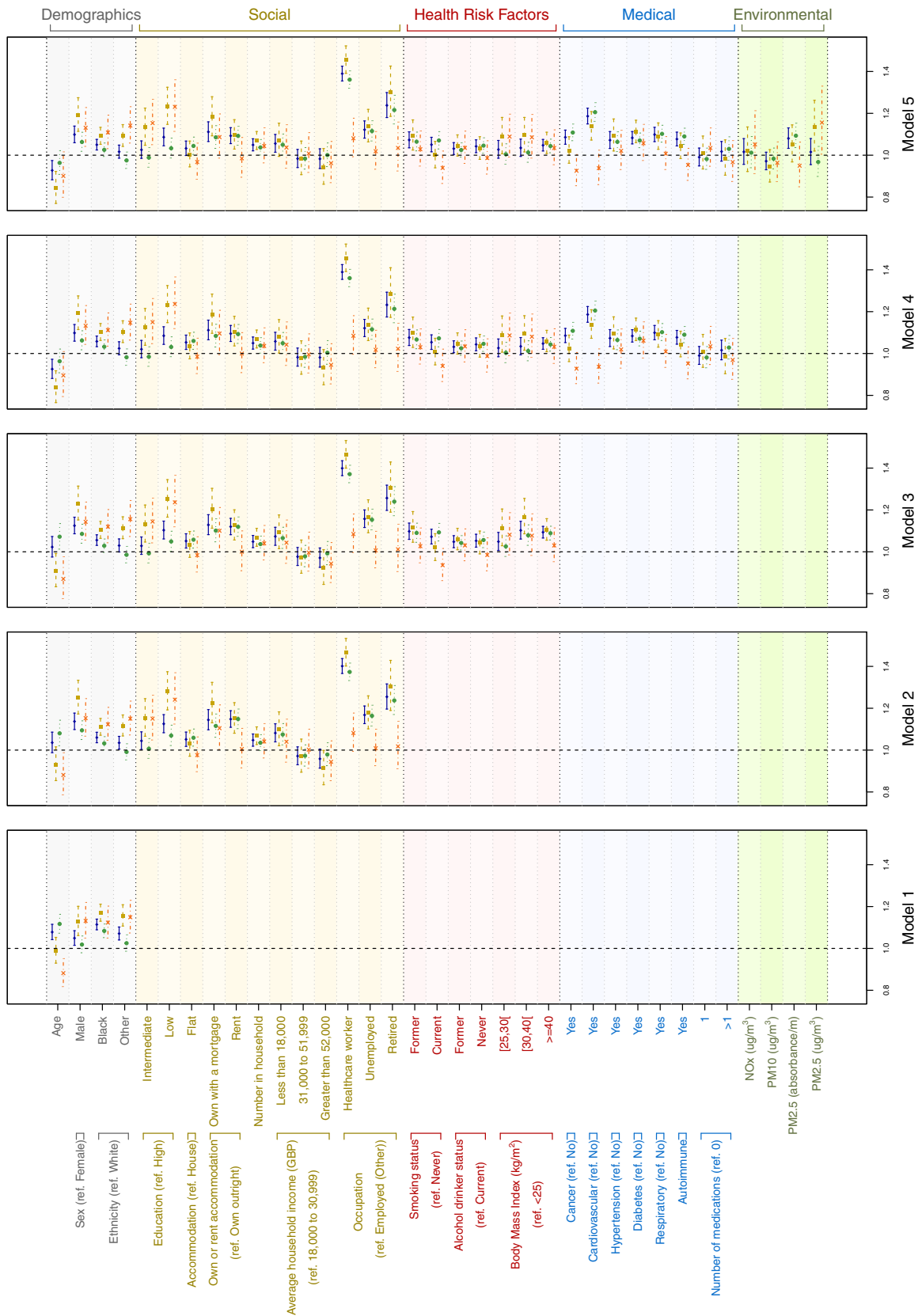


Figure 4 Odds ratio (95% confidence intervals), from the logistic models for the probability of (i) being tested for COVID-19 (blue, plain line), (ii) testing positive (beige, dashed line), (iii) testing negative (green, dotted line), and (iv) testing positive conditionally on being tested (orange, dashed/dotted line). Results are represented for a sequentially adjusted model, where benchmark predictors are defined as demographic descriptors (Model 1), and models are additionally adjusted for social (Model 2), health risk (Model 3), medical (Model 4) and environmental factors (Model 5).

odds ratio (OR) ≥ 1.05 were (Supplementary Table 7A, available as Supplementary data at *IJE* online): male sex (OR = 1.10 [1.06–1.14], $P = 3.43 \times 10^{-7}$); Black vs White ethnicity (OR = 1.05 [1.02–1.08], $P = 7.55 \times 10^{-5}$); low vs high educational attainment (OR = 1.09 [1.05–1.13], $P = 2.84 \times 10^{-5}$); owning a house with a mortgage or renting vs owning outright (OR > 1.09, $P < 10^{-6}$); number in household (OR = 1.05 [1.02–1.08], $P = 6.17 \times 10^{-4}$); earning < GBP 18 000 per year (OR = 1.06 [1.01–1.10], $P = 8.16 \times 10^{-3}$); being a healthcare worker, unemployed or retired vs other employed (OR > 1.12, $P < 10^{-9}$); current or former smoker vs never smoker (OR > 1.05, $P < 10^{-2}$); severely obese vs non-overweight (OR = 1.05 [1.02–1.08], $P = 5.72 \times 10^{-4}$); having had cancer, cardiovascular disease, hypertension, diabetes, respiratory disease, autoimmune disease (OR > 1.07, $P < 10^{-3}$ for all comorbidities). The probability of being tested was inversely associated with age (OR = 0.93 [0.88–0.97], $P = 2.81 \times 10^{-3}$).

The strength of the associations between the probability of being tested and the levels of environmental exposure to NO_x, PM10, PM2.5 were attenuated, and only PM2.5 absorbance was associated in the fully adjusted model (OR = 1.08 [1.03–1.13], $P = 7.99 \times 10^{-4}$, Supplementary Table 7A, available as Supplementary data at *IJE* online).

The fully adjusted models restricted to confirmed or suspected cases (Figure 4, beige and green) gave broadly similar results, except that being younger, of non-White ethnicity, of low educational attainment, a former drinker, which were not associated with risk in suspected cases; and having an average household income < GBP 18 000, being a current smoker, a non-drinker, previously diagnosed with cancer or autoimmune disease, or exposed to higher levels of PM2.5 absorbance were not associated with risk in confirmed cases; and being overweight or obese (OR > 1.09, $P < 0.03$), which were associated with risk in confirmed cases (Figure 4).

Of the associations identified in the univariate analyses for risk of testing positive in the tested population, male sex (OR = 1.13 [1.04–1.23], $P = 3.20 \times 10^{-3}$), lower educational attainment (OR > 1.15, $P < 10^{-2}$), non-White ethnicity (OR > 1.11, $P < 10^{-2}$), and PM2.5 (OR = 1.16 [1.00–1.33], $P = 4.24 \times 10^{-2}$) were weakly associated with test positive vs test negative in the fully adjusted model (Figure 4, Supplementary Table 7B, available as Supplementary data at *IJE* online).

Sensitivity analysis

Excluding healthcare workers from our analyses did not materially affect our results (Supplementary Figure 1A and B, available as Supplementary data at *IJE* online). Results from models restricted to healthcare workers were similar

to those from non-healthcare workers (Supplementary Figure 1C and D, available as Supplementary data at *IJE* online).

Discussion

Main findings

We found important differences between those tested for SARS-CoV-2 infection and the rest of the UK Biobank cohort. Accounting for potential confounding, we found, in a fully adjusted model, that male sex; Black ethnicity; social disadvantage (as measured by education, housing and income); being a healthcare worker, unemployed or retired; a current or former smoker; severely obese; comorbidities (cancer, cardiovascular disease, respiratory or autoimmune diseases); and greater exposure to PM2.5 absorbance were all independently associated with the risk of being tested for COVID-19. We found that the associations linking obesity and the risk of being tested were strongly attenuated while adjusting for comorbidities. We found consistent results when comparing only confirmed COVID-19 cases with the non-tested population. Additionally, comparing data for test positive and test negative individuals within the tested population, we found, in a fully adjusted model, consistent associations linking the risk of testing positive and male sex, lower educational attainment, non-White ethnicity and PM2.5.

Health risk factors and comorbidities were found to be associated with the risk of being tested but not with the risk of testing positive, conditional on being tested in a fully adjusted model. This suggests that these factors may help in predicting the risk of developing COVID-19 symptoms, and therefore the probability of receiving a test. Nevertheless, within the tested population, these factors do not provide any further information that would be relevant to predicting the outcome of the test.

Given the high specificity of the RT-PCR test for SARS-CoV-2, it is likely that all test positive individuals were true cases. Some who tested negative may have had other illnesses with similar clinical presentation and possibly shared risk factors. This is in keeping with the limited and variable sensitivity of the RT-PCR test for SARS-CoV-2 (reportedly $\sim 70\%$)²² and the possibility that those testing negative may have presented to healthcare at a point in the disease course when SARS-CoV-2 RNA may no longer be detectable in the sampled tissues.²³

Nonetheless and despite these limitations we were able to find consistent evidence linking social factors, ethnicity, and marginally, higher levels of environmental exposures, to COVID-19 risk.

Our results add to those from ICU data,²⁴ which showed greater risk of ICU admission for men and individuals from a non-White background. This may explain that we observed an excess of men, who appear to be at higher risk of a severe form of COVID-19, in those who were tested more than twice for the infection. Additionally, severe COVID-19 cases in the ISARIC consortium data²⁵ were more likely to be men and have comorbidities such as cardiovascular or respiratory disease. Furthermore, data from the UK Office for National Statistics show that non-White individuals (specifically from Black, Bangladeshi and Pakistani backgrounds) are 1.6–1.9 times more likely to die from COVID-19 after adjustment for SES.²⁶ Other studies further corroborate the increased risk associated with being male,²⁷ diabetic,²⁷ non-White^{27–29} and of a lower SES.^{27,30,31} Whereas the ICU data included a comparator group with non-COVID-19 viral pneumonia, ISARIC data include COVID-19 cases only. In the present work the cases were drawn from the extensively evaluated UK Biobank cohort; this made it possible to directly compare the characteristics of confirmed or suspected COVID-19 cases with the non-tested population and to evaluate the mutually independent effects of multiple variables on risk.

It has been hypothesized that the higher rate of severe COVID-19 among non-White individuals may reflect lower SES and higher prevalence of comorbidities in Black and minority ethnic groups.³² In the present study, we show that each of these factors independently contributes to the risk of being a confirmed or suspected COVID-19 case. Results from our test negative design approach suggest that non-White ethnicity is associated with increased risk of COVID-19, independently from social factors and comorbidities. Although it is possible that our results may reflect some residual confounding by SES and/or comorbidities, our results are suggestive that non-White individuals have increased risk of severe illness and death from COVID-19. If so, the reasons for these increased risks are unknown and indicate an area for urgent future research.

There has been speculation regarding the extent to which healthcare workers may be at increased risk of COVID-19.³³ Although it is possible that healthcare workers might have been preferentially tested, this was not widespread policy at the time of study,³⁴ and our results did not materially change with or without inclusion of healthcare workers. It has been suggested that the higher risk in healthcare workers may reflect higher or repeated exposure to SARS-CoV-2 and to aerosol-generating procedures^{35,36} as well as reported lack of adequate personal protective equipment at that time.^{35–38}

Comorbidities are also associated with lower SES, but again we found that they were independently and jointly

contributing to increased risk, specifically, cancer, cardiovascular disease, hypertension, diabetes, respiratory and autoimmune diseases. Unlike other studies reporting higher risk in obese individuals,^{39,40} and postulating ACE2 expression in adipose tissue as a potential mechanism for the role of obesity in severe COVID-19,⁴¹ we found that the association of obesity and COVID-19 was attenuated while adjusting for comorbidities. These included cardiovascular diseases, diabetes and respiratory disease, which were the most strongly associated with the risk of developing COVID-19 symptoms, independently of obesity. This points to a potential role for metabolic and pro-inflammatory disorders in the development of COVID-19. Furthermore, we found an excess of ever smokers in the tested group, in agreement with some studies⁶ but not others²⁷—although the former may have been affected by collider bias.⁴²

Although increased risks from higher levels of outdoor air pollutants at the person's residence was seen in the unadjusted analysis, these were attenuated after multiple adjustment and only associations involving PM_{2.5} were (borderline) statistically significant. This suggests that previous reports of an association between outdoor air pollution and COVID-19^{40,43} may have been confounded. Nonetheless, given the effects of outdoor air pollution on respiratory function,⁴⁴ cardiovascular disease⁴⁵ and other infections such as SARS,⁴⁶ this area requires further investigation.

Although age is understood to be a major risk factor for severe disease or death from COVID-19,⁴⁷ age was found to be inversely associated with the risk of being tested but not with the risk of testing positive among the tested population. However, being retired, which remained significant in our adjusted models, is a proxy for older age, and may explain the inverse association with age in the multiply adjusted models. Additionally, older people might have been less likely to have been tested, either due to testing protocols or concerns about attending healthcare settings due to shielding,⁴⁸ and non-referral to hospital, e.g. from care homes where residents were not routinely tested.

Limitations

First, UK Biobank is not representative of the general UK population due to healthy volunteer selection bias and over-representation of White people, participants of higher SES and certain occupations.⁴⁹ In particular, our study population included higher numbers of healthcare workers than in the general population.⁵⁰ This could further be explained by the fact that we adopted a broader definition of healthcare workers (including health and social service manager and care assistants and home carers).

Nonetheless, the range of factors influencing the risk of confirmed or suspected COVID-19 concurs with and extends findings from other studies of hospital-only populations and national mortality data. We also show that our results and conclusions are robust to the exclusion of healthcare workers, which indicates that despite the non-representativeness of the UK Biobank population, our results are not biased by the over-representation of healthcare workers.

Second, mortality data linked to SARS-CoV-2 infection status in UK Biobank is currently unavailable. Future availability of linked hospital outcome and mortality data within UK Biobank will aid in further assessing risk related to SARS-CoV-2 infection.

Despite these limitations, our complementary analytical approaches, including use of the test-negative case-control design, enabled us to triangulate across the different outcomes strengthening the evidence linking a range of exposures to COVID-19 risk.

Conclusions

Linkage of SARS-CoV-2 test results to UK Biobank enabled us to identify independent associations of demographic, social, health risk, medical and environmental factors with risk of testing positive (confirmed) or negative (suspected) for COVID-19. Of these, male sex, lower educational attainment, non-White ethnicity were also found associated with increased risk of testing positive, within the tested population. Elucidation of the joint and independent effects of these factors represents a high-priority area for further research, which may inform on COVID-19 natural history and suggest possible new avenues to pursue for its prevention.

Supplementary data

[Supplementary data](#) are available at *IJE* online.

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Author contributions

M.C.-H., J.E., R.V., P.E., M.K.-I. and C.D. conceived the study and drafted the manuscript. M.C.-H., B.B., M.W., and J.E. performed the statistical analyses. UK Biobank data were extracted, harmonized and analysed by I.T., B.B. and M.W. I.T. and R.V. provided insights into the study design, results interpretation and revised the manuscript. All authors revised the manuscript for important intellectual content and approved the submission of the manuscript. M.C.-H. had full access to the data and takes responsibility for the integrity of the data and the accuracy of the data analysis and for the decision to submit for publication.

Ethics approval

Ethics approval for the nurse visit was obtained from the National Research Ethics Service (Reference: 10/H0604/2). Participants gave written consent for blood sampling.

Conflict of interest

None declared.

References

1. World Health Organization. *WHO Director-General's opening remarks at the media briefing on COVID-19*. Geneva: World Health Organization; 2020.
2. Johns Hopkins University. *New Cases of COVID-19 in World Countries*. Baltimore: Johns Hopkins Coronavirus Resource Center.
3. Guo W, Li M, Dong Y *et al*. Diabetes is a risk factor for the progression and prognosis of COVID-19. *Diabetes Metab Res Rev* 2020;e3319.
4. Huang C, Wang Y, Li X *et al*. Clinical features of patients infected with 2019 novel coronavirus in Wuhan, China. *Lancet* 2020;395:497–506.

5. Shi S, Qin M, Shen B *et al.* Association of cardiac injury with mortality in hospitalized patients with COVID-19 in Wuhan, China. *JAMA Cardiol* 2020.
6. Vardavas CI, Nikitara K. COVID-19 and smoking: a systematic review of the evidence. *Tob Induc Dis* 2020;18:20.
7. Wang D, Hu B, Hu C *et al.* Clinical Characteristics of 138 hospitalized patients with 2019 novel coronavirus-infected pneumonia in Wuhan, China. *JAMA* 2020;323:1061.
8. Wu C, Chen X, Cai Y *et al.* Risk factors associated with acute respiratory distress syndrome and death in patients with coronavirus disease 2019 pneumonia in Wuhan, China. *JAMA Intern Med* 2020;180:934.
9. Yang J, Zheng Y, Gou X *et al.* Prevalence of comorbidities and its effects in patients infected with SARS-CoV-2: a systematic review and meta-analysis. *Int J Infect Dis* 2020;94:91–5.
10. Zhou F, Yu T, Du R *et al.* Clinical course and risk factors for mortality of adult inpatients with COVID-19 in Wuhan, China: a retrospective cohort study. *Lancet* 2020;395:1054–062.
11. Grasselli G, Pesenti A, Cecconi M. Critical care utilization for the COVID-19 outbreak in Lombardy, Italy: early experience and forecast during an emergency response. *JAMA* 2020;323:1545–546.
12. Remuzzi A, Remuzzi G. COVID-19 and Italy: what next? *Lancet* 2020;395:1225–228.
13. Peña S, Cuadrado C, Rivera-Aguirre A *et al.* PoliMap: A taxonomy proposal for mapping and understanding the global policy response to COVID-19. *OSF Preprints*, doi: 10.31219/osf.io/h6mvs, 20 April 2020, preprint: not peer reviewed.
14. Delatorre E, Mir D, Graf T, Bello G. Tracking the onset date of the community spread of SARS-CoV-2 in Western Countries. *medRxiv*, doi: 10.1101/2020.04.20.20073007, 23 April 2020, preprint: not peer reviewed.
15. Public Health England. COVID-19: investigation and initial clinical management of possible cases; 2020.
16. Vandenbroucke J, Brickley E, Vandenbroucke-Grauls C, Pearce N. The test-negative design with additional population controls: a practical approach to rapidly obtain information on the causes of the SARS-CoV-2 epidemic. *arXiv*, arXiv: 2004.06033v2, 14 May 2020, preprint: not peer reviewed.
17. Sudlow C, Gallacher J, Allen N *et al.* UK Biobank: an open access resource for identifying the causes of a wide range of complex diseases of middle and old age. *PLOS Med* 2015;12: e1001779.
18. Armstrong J, Rudkin J, Allen N, *et al.* Dynamic linkage of COVID-19 test results between Public Health England's Second Generation Surveillance System and UK Biobank. *Microb Genom* 2020;6:mgen000397.
19. de Hoogh K, Wang M, Adam M *et al.* Development of land use regression models for particle composition in twenty study areas in Europe. *Environ Sci Technol* 2013;47:5778–786.
20. Friedman J, Hastie T, Tibshirani R. Regularization paths for generalized linear models via coordinate descent. *J Stat Softw* 2010;33:1–22.
21. Vermeulen R, Saberi Hosnizheh F, Bodinier B, on behalf of the EnviroGenoMarkers Consortium Consortium members *et al.* Pre-diagnostic blood immune markers, incidence and progression of B-cell lymphoma and multiple myeloma: univariate and functionally informed multivariate analyses. *Int J Cancer* 2018; 143:1335–347.
22. Fang Y, Zhang H, Xie J *et al.* Sensitivity of chest CT for COVID-19: comparison to RT-PCR. *Radiology* 2020;200432.
23. Zou L, Ruan F, Huang M *et al.* SARS-CoV-2 viral load in upper respiratory specimens of infected patients. *N Engl J Med* 2020; 382:1177–179.
24. Icnarc. ICNARC report on COVID-19 in critical care; 2020.
25. Docherty AB, Harrison EM, Green CA *et al.* Features of 20 133 UK patients in hospital with COVID-19 using the ISARIC WHO Clinical Characterisation Protocol: prospective observational cohort study. *BMJ* 2020;369:m1985.
26. Office for National Statistics. Coronavirus (COVID-19) related deaths by ethnic group, England and Wales: 2 March 2020 to 10 April 2020; 2020. <https://www.ons.gov.uk/peoplepopulationandcommunity/birthsdeathsandmarriages/deaths/articles/coronavirusrelateddeathsbyethnicgroupenglandandwales/2march2020to10april2020> (7 May 2020, date last accessed).
27. Williamson E, Walker AJ, Bhaskaran KJ, *et al.* OpenSAFELY: factors associated with COVID-19-related hospital death in the linked electronic health records of 17 million adult NHS patients. *medRxiv*, doi: 10.1101/2020.05.06.20092999, 7 May 2020, preprint: not peer reviewed.
28. Niedzwiedz CL, O'Donnell CA, Jani BD *et al.* Ethnic and socioeconomic differences in SARS-CoV-2 infection: prospective cohort study using UK Biobank. *BMC Med* 2020;18: 160.
29. Li AY, Hannah TC, Durbin JR *et al.* Multivariate Analysis of Black Race and Environmental Temperature on COVID-19 In the US. *Am J Med Sci* 2020;360:348–56.
30. Alipio M. Do Socio-Economic Indicators Associate with COVID-2019 Cases? Findings from a Philippine Study. *SSRN Electron J* 2020, doi: 10.2139/ssrn.3573353.
31. Zheng Z, Michelle C, Li X. Strong effect of socioeconomic levels on the spread and treatment of the 2019 novel coronavirus (COVID-19) in China. *medRxiv*, doi: 10.1101/2020.04.25.20079400, 7 May 2020, preprint: not peer reviewed.
32. Pareek M, Bangash MN, Pareek N *et al.* Ethnicity and COVID-19: an urgent public health research priority. *Lancet* 2020;395:1421–422.
33. Cheung J-H, Ho LT, Cheng JV, Cham EYK, Lam KN. Staff safety during emergency airway management for COVID-19 in Hong Kong. *Lancet Respir Med* 2020;8:e19.
34. Gov.UK. COVID-19: guidance for healthcare providers who have staff with relevant travel, healthcare or household contact history; 2020. <https://www.gov.uk/government/publications/novel-coronavirus-2019-ncov-guidance-for-healthcare-providers-with-staff-who-have-travelled-to-china> (7 May 2020, date last accessed).
35. Belingheri M, Paladino ME, Riva MA. Beyond the assistance: additional exposure situations to COVID-19 for healthcare workers. *J Hosp Infect* 2020;S0195-6701:30132–0138.
36. Cook TM. Personal protective equipment during the coronavirus disease (COVID) 2019 pandemic: a narrative review. *Anaesthesia* 2020;75:920–27.
37. Iacobucci G. Covid-19: Doctors still at “considerable risk” from lack of PPE, BMA warns. *BMJ* 2020;368:m1316.
38. Wang J, Zhou M, Liu F. Reasons for healthcare workers becoming infected with novel coronavirus disease 2019 (COVID-19) in China. *J Hosp Infect* 2020;105:100–01.
39. Lochlainn MN, Lee KA, Sudre CH *et al.* Key predictors of attending hospital with COVID19: An association study from the

- COVID Symptom Tracker App in 2,618,948 individuals. *medRxiv*, doi: 10.1101/2020.04.25.20079251, 29 April 2020, preprint: not peer reviewed.
40. Wu X, Nethery RC, Sabath BM, Braun D, Dominici F. Exposure to air pollution and COVID-19 mortality in the United States: A nationwide cross-sectional study. *medRxiv*, doi: 10.1101/2020.04.05.20054502, 27 April 2020, preprint: not peer reviewed.
 41. Malavazos AE, Corsi Romanelli MM, Bandera F, Iacobellis G. Targeting the adipose tissue in COVID-19. *Obesity (Silver Spring)* 2020;28:1178–179.
 42. Griffith G, Morris TT, Tudball M *et al.* Collider bias undermines our understanding of COVID-19 disease risk and severity. *medRxiv*, doi: 10.1101/2020.05.04.20090506, 20 May 2020, preprint: not peer reviewed.
 43. Li AY, Hannah TC, Durbin J *et al.* Multivariate analysis of factors affecting COVID-19 case and death rate in U.S. Counties: the significant effects of black race and temperature. *medRxiv*, doi: 10.1101/2020.04.17.20069708, 24 April 2020, preprint: not peer reviewed.
 44. Pope CA, Dockery DW, Schwartz J. Review of epidemiological evidence of health effects of particulate air pollution. *Inhal Toxicol* 1995;7:1–18.
 45. Franklin BA, Brook R, Arden Pope C. Air pollution and cardiovascular disease. *Curr Probl Cardiol* 2015;40:207–38.
 46. Cui Y, Zhang Z-F, Froines J *et al.* Air pollution and case fatality of SARS in the People's Republic of China: an ecologic study. *Environ Health* 2003;2:15.
 47. Jordan RE, Adab P, Cheng KK. Covid-19: risk factors for severe disease and death. *BMJ* 2020;368:m1198.
 48. Public Health England. Disparities in the risk and outcomes of COVID-19. 2020.
 49. Fry A, Littlejohns TJ, Sudlow C *et al.* Comparison of sociodemographic and health-related characteristics of UK Biobank participants with those of the general population. *Am J Epidemiol* 2017;186:1026–034.
 50. Office for National Statistics. International migration and the healthcare workforce. 2019. <https://www.ons.gov.uk/peoplepopulationandcommunity/populationandmigration/internationalmigration/articles/internationalmigrationandthehealthcareworkforce/2019-08-15> (15 August 2019, date last accessed).