## A Search for the Higgs Boson Produced in Association with $Z \rightarrow \ell^{+} \ell^{-}$Using the Matrix Element Method at CDF II

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We present a search for associated production of the standard model (SM) Higgs boson and a $Z$ boson where the $Z$ boson decays to two leptons and the Higgs decays to a pair of $b$ quarks in $p \bar{p}$ collisions at the Fermilab Tevatron. We use event probabilities based on SM matrix elements to construct a likelihood function of the Higgs content of the data sample. In a CDF data sample corresponding to an integrated luminosity of 2.7 $\mathrm{fb}^{-1}$ we see no evidence of a Higgs boson with a mass between $100 \mathrm{GeV} / c^{2}$ and $150 \mathrm{GeV} / c^{2}$. We set $95 \%$ confidence level (C.L.) upper limits on the cross-section for $Z H$ production as a function of the Higgs boson mass $m_{H}$; the limit is 8.2 times the SM prediction at $m_{H}=115 \mathrm{GeV} / c^{2}$.

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In the standard model (SM), the Higgs mechanism is responsible for the observed breaking of the $S U(2)_{L} \otimes U(1)$ symmetry [1, 2], yet the Higgs boson remains the only SM particle that has not been directly observed. Direct searches

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have set a lower limit on the SM Higgs boson mass $m_{H}$ of $114.4 \mathrm{GeV} / c^{2}$ at $95 \%$ C.L. [3], while precision electroweak measurements indirectly constrain its mass to $m_{H}=$ $76_{-24}^{+33} \mathrm{GeV} / c^{2}$ [4]. At hadron colliders the dominant production process for the SM Higgs boson is $g g \rightarrow H$ while its decays are dominated by $H \rightarrow b \bar{b}$ for $m_{H}<140 \mathrm{GeV} / c^{2}$. However, the process $g g \rightarrow H \rightarrow b \bar{b}$ is dwarfed by multi-jet background, necessitating the search for Higgs bosons produced in association with a $W$ or $Z$ boson that decays leptonically. This article reports a search for the process $p \bar{p} \rightarrow$ $Z H \rightarrow \ell^{-} \ell^{+} b \bar{b}(\ell=e, \mu)$ in data with an integrated luminosity of $2.7 \mathrm{fb}^{-1}$ collected with the CDF II detector, nearly 3 times that of the previously reported analysis [5]. The study of Higgs boson production in association with a $W / Z$ gauge boson for low Higgs boson masses is further motivated by the fact that the signal to background ratio is more favorable at the Tevatron compared to the Large Hadron Collider.

For the first time in a $Z H \rightarrow \ell^{-} \ell^{+} b \bar{b}$ search, we utilize a method based on leading-order matrix element calculations [6-8] convoluted with detector resolution functions [9] that form per-event likelihoods. This method, pioneered for use in top quark mass measurements [10, 11], has been recently used in Higgs boson searches in other decay channels [12] by forming a discriminating per-event variable. We extend the technique by expressing the event likelihoods as a function of the $Z H$ signal fraction and maximizing the joint likelihood for the data sample with respect to the signal fraction.

The CDF II detector $[13,14]$ is an azimuthally and forwardbackward symmetric apparatus designed to study $p \bar{p}$ collisions at the Fermilab Tevatron. It consists of a magnetic spectrometer surrounded by calorimeters and muon chambers. The charged particle tracking system, consisting of a silicon detector and drift chamber, is immersed in a 1.4 T magnetic field parallel to the $p$ and $\bar{p}$ beams. Calorimeters segmented in $\eta$ and $\phi$ surround the tracking system and measure the energy of particles detected within them. The electromagnetic and hadronic calorimeters are lead-scintillator and iron-scintillator sampling devices, respectively. Drift chambers located outside the central hadron calorimeters detect muons. The data used in this analysis are collected with an online selection that requires events to have a lepton with $E_{T}>18 \mathrm{GeV}$ (for an electron) or $p_{T}>18 \mathrm{GeV} / c$ (for a muon) [14].

The event selection used in this analysis closely follows that in Ref. [5]. Candidate events are required to have a pair of oppositely charged electrons or muons with invariant mass $76<m_{\text {¢ }}<106 \mathrm{GeV} / c^{2}$. Candidate events are also required to have one jet with $E_{T}>25 \mathrm{GeV}$ and at least one additional jet with $E_{T}>15 \mathrm{GeV}$, both within $|\eta|<2.0$. All jet energies are corrected for non-uniformities in calorimeter response, effects from multiple $p \bar{p}$ interactions and for the hadronic energy scale of the calorimeter [15]. Candidate events are required to have at least one jet with an associated displaced secondary vertex [16] (" $b$-tags", reconstructed using tracks with hits in the silicon detector), thus enriching the $b$-quark content of the sample.

TABLE I: Expected and observed numbers of events with 1 or $2 b$ tagged jets in $2.7 \mathrm{fb}^{-1}$ of data. The $Z H$ expectation is shown for $m_{H}=115 \mathrm{GeV} / c^{2}$ assuming the production cross section at $\sqrt{s}=$ 1.96 TeV for $q \bar{q} \rightarrow Z^{*} \rightarrow Z H$ to be $1.04 \mathrm{pb}[20]$ and the branching ratio $\mathcal{B}(H \rightarrow b \bar{b})$ to be $73 \%$ [21].

| Source | 1 tag | $\geq 2$ tag |
| :--- | :---: | :---: |
| $Z \rightarrow \ell^{+} \ell^{-}+$light partons | $129.6 \pm 24.0$ | $5.5 \pm 0.9$ |
| $Z \rightarrow \ell^{+} \ell^{-}+b \bar{b}, c \bar{c}$ | $107.2 \pm 14.0$ | $19.5 \pm 3.4$ |
| $Z Z, W Z$ | $11.6 \pm 1.3$ | $2.9 \pm 0.4$ |
| $t \bar{t}$ | $13.9 \pm 2.0$ | $7.7 \pm 1.1$ |
| Mis-ID lepton | $15.9 \pm 6.5$ | $0.4 \pm 0.2$ |
| $Z H$ | $1.3 \pm 0.2$ | $0.7 \pm 0.1$ |
| Total expected | $279.5 \pm 28.6$ | $36.3 \pm 3.7$ |
| Data | 258 | 32 |

The backgrounds for this analysis are dominated by events with real $Z$ bosons with additional contributions from $t \bar{t}$ and events where an object, such as a jet, is mis-identified as a lepton. We model the backgrounds with events generated with leading-order event generators, normalized to next-toleading order cross-sections and simulated with a GEANTbased description of the CDF II detector [9]. Z+light-flavor jet contributions are modeled with the ALPGEN [17] simulation code matched with PYTHIA in the MLM scheme [17] for the hadronization and fragmentation. Heavy flavor contributions from $Z+b \bar{b}$ and $Z+c \bar{c}$ are modeled separately with alpgen and combined with the light-flavor jet samples. The $W Z, Z Z$ and $t \bar{t}$ processes are modeled using pythia [18]. Events where a jet is mis-identified as a charged lepton are modeled using jet-enriched data samples [5, 19]. We model the kinematics of $Z H \rightarrow \ell^{+} \ell^{-} b \bar{b}$ events using PYTHIA for $m_{H}$ ranging from $100 \mathrm{GeV} / c^{2}$ to $150 \mathrm{GeV} / c^{2}$. The signal and background contributions expected in $2.7 \mathrm{fb}^{-1}$ and the number of observed events are given in Table I.

We denote the $Z H$ signal probability by $P_{Z H}\left(\mathbf{x}_{\mathbf{i}} \mid m_{H}\right)$ where $m_{H}$ is a parameter and $\mathbf{x}_{\mathbf{i}}$ represents the collection of the measured 4 -vector momenta of the two selected leptons, the two selected jets, and the two components of the missing transverse momentum, in a given event $i$. Similarly we denote the background probability as $P_{b}\left(\mathbf{x}_{\mathbf{i}}\right)$. The per-event likelihood as a function of the signal fraction $s$ for a given event $i$ is

$$
\begin{equation*}
L\left(s, \mathbf{x}_{\mathbf{i}} \mid m_{H}\right)=s P_{Z H}\left(\mathbf{x}_{\mathbf{i}} \mid m_{H}\right)+(1-s) P_{b}\left(\mathbf{x}_{\mathbf{i}}\right) \tag{1}
\end{equation*}
$$

We evaluate $P_{Z H}$ and $P_{b}$ by convoluting the leading-order matrix elements for the process with detector resolution functions and integrating over unmeasured quantities. Thus, $P_{Z H}$ is a probability density in $\mathbf{x}_{\mathbf{i}}$ and can be expressed as

$$
\begin{align*}
P_{Z H}\left(\mathbf{x}_{\mathbf{i}} \mid m_{H}\right)= & \frac{1}{\sigma\left(m_{H}\right)} \int d \Phi\left|\mathcal{M}_{Z H}\left(q, p ; m_{H}\right)\right|^{2} \\
& \times \prod_{j}\left[W\left(p_{j}, \mathbf{x}_{\mathbf{i}}\right)\right] f_{P D F}\left(q_{1}\right) f_{P D F}\left(q_{2}\right) \tag{2}
\end{align*}
$$



FIG. 1: The distribution of $\tan ^{-1} \mathcal{D}$ where the discriminant $\mathcal{D}=$ $\left(P_{Z H}-P_{b}\right) / L$ for expected backgrounds and data for events with one (left) and two (right) $b$-tags. The expected signal ( $\times 10$ ) is overlaid.
where $\mathcal{M}_{Z H}$ is the leading-order matrix element for the process $q \bar{q} \rightarrow Z H \rightarrow \ell^{+} \ell^{-} b \bar{b}$ evaluated for a pair of incoming partons $q$ and outgoing particles $p, W\left(p_{j}, \mathbf{x}_{\mathbf{i}}\right)$ are transfer functions [22] linking the outgoing particle momenta $p_{j}$ to measured quantities $\mathbf{x}_{\mathbf{i}}$ and the $f_{P D F}$ are parton density functions of the incoming partons. The factor $1 / \sigma\left(m_{H}\right)$ ensures that the probability density satisfies the normalization condition, $\int d \mathbf{x}_{\mathbf{i} i} P_{Z H}\left(\mathbf{x}_{\mathbf{i}} \mid m_{H}\right)=1$.

The sample likelihood $\mathcal{L}$ is obtained by taking the product over all events $i$ in the sample

$$
\begin{equation*}
\mathcal{L}\left(s \mid m_{H}\right)=\prod_{i} L\left(s, \mathbf{x}_{\mathbf{i}} \mid m_{H}\right) \tag{3}
\end{equation*}
$$

We enhance our statistical sensitivity by exploiting the expected difference in the rate of signal and background events with two $b$-tagged jets. We replace $P_{Z H}\left(\mathbf{x}_{\mathbf{i}} \mid m_{H}\right)$ by $P_{Z H}\left(\mathbf{x}_{\mathbf{i}}, n \mid m_{H}\right) \equiv P_{Z H}\left(\mathbf{x}_{\mathbf{i}} \mid m_{H}\right) \cdot P_{Z H}\left(n \mid m_{H}\right)$ and $P_{b}\left(\mathbf{x}_{\mathbf{i}}\right)$ by $P_{b}\left(\mathbf{x}_{\mathbf{i}}, n\right) \equiv P_{b}\left(\mathbf{x}_{\mathbf{i}}\right) \cdot P_{b}(n)$, where $P_{Z H}\left(n \mid m_{H}\right)\left[P_{b}(n)\right]$ denotes the probability of tagging signal [background] events with $n$ tags. Table II shows the expected tagging rates for simulated signal and background event samples.

The measured signal fraction $S_{\text {meas }}$ is the value of $s$ which maximizes $\mathcal{L}\left(s \mid m_{H}\right)$. Using Eq. (1), we can define a perevent discriminant $\mathcal{D}_{i} \equiv \partial \ln L / \partial s=\left(P_{Z H}-P_{b}\right) / L$ which increases (decreases) for more signal-like (background-like) events. The maximum-likelihood estimator for the measured signal fraction $S_{\text {meas }}$ corresponds to $\left.\Sigma_{i} \mathcal{D}_{i}\right|_{s=S_{\text {meas }}}=0$. The distribution of $\tan ^{-1} \mathcal{D}_{i}\left(s=S_{\text {meas }}\right)$ for simulated events and data is shown in Fig. 1.

The dominant backgrounds in our data sample are due to $Z+$ jets, $t \bar{t}$ and $Z Z$ processes, in the expected proportions denoted by $\lambda_{Z j j}, \lambda_{t \bar{t}}$ and $\lambda_{Z Z}$ respectively. The background probability in Eq. (1) is given by

TABLE II: Expected single- and double-tag probabilities, $P(n=1)$ and $P(n \geq 2)$, for signal and background events passing our selection. $Z \rightarrow \ell^{+} \ell^{-}+$jets includes jets from both light and heavy quarks.

| Source | $P(n=1)$ | $P(n \geq 2)$ |
| :--- | :---: | :---: |
| $Z \rightarrow \ell^{+} \ell^{-}+$jets | 0.91 | 0.09 |
| $W Z, Z Z$ | 0.80 | 0.20 |
| $t \bar{t}$ | 0.74 | 0.26 |
| $Z H\left(m_{H}=100 \mathrm{GeV} / c^{2}\right)$ | 0.67 | 0.33 |
| $Z H\left(m_{H}=125 \mathrm{GeV} / c^{2}\right)$ | 0.65 | 0.35 |
| $Z H\left(m_{H}=150 \mathrm{GeV} / c^{2}\right)$ | 0.63 | 0.37 |

$P_{b}\left(\mathbf{x}_{\mathbf{i}}, n\right)=\lambda_{Z j j} P_{Z j j}\left(\mathbf{x}_{\mathbf{i}}, n\right)+\lambda_{t t} P_{t \bar{t}}\left(\mathbf{x}_{\mathbf{i}}, n\right)+\lambda_{Z Z} P_{Z Z}\left(\mathbf{x}_{\mathbf{i}}, n\right)$,
where $P_{Z j j}\left(\mathbf{x}_{\mathbf{i}}, n\right), P_{t \bar{t}}\left(\mathbf{x}_{\mathbf{i}}, n\right)$ and $P_{Z Z}\left(\mathbf{x}_{\mathbf{i}}, n\right)$ are the respective probability densities (normalized to unit integral) for the $Z+$ jets, $t \bar{t}$, and $Z Z$ background processes with $n$ tags. Normalization of $P_{b}$ is ensured by requiring $\lambda_{Z j j}+\lambda_{t \bar{t}}+\lambda_{Z Z}=1$.

We construct confidence intervals [23] for the test statistic $R=\mathcal{L}\left(S_{\text {meas }} \mid S_{\text {true }}\right) / \mathcal{L}\left(S_{\text {meas }} \mid S_{\text {true }}^{\text {best }}\right)$ by performing simulated experiments with the expected proportions of background and varying the amounts of signal, such that $S_{\text {true }}$ is the true (input) signal fraction in the simulated experiment. $S_{\text {true }}^{\text {best }}$ is the input signal fraction that has the highest likelihood for a given measured signal fraction, $S_{\text {meas }} . \mathcal{L}\left(S_{\text {meas }} \mid S_{\text {true }}\right)$ is given by Eq. (3) for the simulated experiment with the chosen value of $S_{t r u e}$ and $m_{H}$. Since we are measuring the fractional signal content in the data sample, the number of events in each simulated experiment is held fixed at the value of 290 events observed in the data.

The methodology from Ref. [23] is used to construct confidence intervals in $S_{\text {meas }}$ for each chosen value of $S_{\text {true }}$ and $m_{H}$. This method removes any bias resulting from imperfections in our modeling by relating $S_{\text {meas }}$ to $S_{\text {true }}$. The confidence intervals in $S_{\text {meas }}$ obtained for $m_{H}=115 \mathrm{GeV} / c^{2}$ and $0 \leq S_{\text {true }} \leq 0.25$ are shown in Fig. 2. For a given value of $S_{\text {meas }}$ obtained from the data (or from an independent simulated experiment to evaluate the a priori expectation), we extract the range of $S_{\text {true }}$ for which the confidence intervals contain this value of $S_{\text {meas }}$. A feature of this method is that the resulting range of $S_{\text {true }}$ can be quoted as an upper limit on $S_{\text {true }}$ (if the lower bound is zero) or as a two-sided measurement of $S_{\text {true }}$. As Fig. 2 shows, we obtain an upper limit on $S_{\text {true }}$ given the data, which we convert to the equivalent upper limit on the signal cross section. This procedure is repeated for the range of Higgs boson masses $100 \leq m_{H} \leq 150 \mathrm{GeV} / c^{2}$.

We evaluate systematic uncertainties by varying process rates and kinematic distributions in our simulated experiments. We apply a rate uncertainty of $40 \%$ for $Z$ boson events and of $20 \%$ for diboson and $t \bar{t}$ events. The uncertainty on the rate of heavy flavor production in association with a gauge boson is based on comparisons of data with theoretical predictions [19]. The uncertainty on the diboson and $t \bar{t}$ contribution


FIG. 2: Confidence intervals in the measured signal fraction $S_{\text {meas }}$ (along the $x$-axis) with $68 \%$ C.L., $95 \%$ C.L. and $99.7 \%$ C.L., for a range of true signal fraction $S_{\text {true }}$ values (along the $y$-axis on the left) chosen in the simulated experiments. The signal cross section ratio equivalent to $S_{\text {true }}$ is shown on the $y$-axis on the right. The intervals shown here are computed for a Higgs boson mass of $m_{H}=115 \mathrm{GeV} / c^{2}$, and include statistical and systematic uncertainties. The vertical dashed line indicates the value of $S_{\text {meas }}$ obtained from the data.
includes the uncertainties in the cross sections, selection efficiencies and the top quark mass [5]. A rate uncertainty of $50 \%$ is applied for mis-identified lepton events due to the uncertainty on the lepton misidentification probability [5]. A rate uncertainty of $6 \%$ due to the luminosity uncertainty is applied to all events. The per-jet uncertainty on the $b$-tagging efficiency is $8 \%$ for events with $b$ partons, $16 \%$ for events with $c$ partons and $13 \%$ for events with no heavy flavor [5]. Our analysis is weakly sensitive to uncertainties in the expected total number of events passing our selection, since it relies only on the shapes of measured distributions. Uncertainties in the shapes of kinematic distributions are propagated by varying the amount of QCD radiation in simulated signal events and the jet energy scale in simulated signal and background events within their respective uncertainties [15].

We evaluate confidence intervals for a range of Higgs masses between $100 \mathrm{GeV} / c^{2}$ and $150 \mathrm{GeV} / c^{2}$. We evaluate a priori $95 \%$ C.L. upper limits on the cross section for the process $p \bar{p} \rightarrow Z H \rightarrow \ell^{+} \ell^{-} b \bar{b}$. We express these limits as a ratio with respect to the SM prediction. These expected limits along with those observed in the data are shown in Table III.

In conclusion, we have performed a search for the SM Higgs boson decaying to $b \bar{b}$ produced in association with a $Z$ boson. This is the first analysis performed in this channel with a matrix element method. The data show no excess over expected non-Higgs backgrounds. We set $95 \%$ C.L. upper limits on the cross section of this process for a range of Higgs boson masses. The limit at $m_{H}=115 \mathrm{GeV} / c^{2}$ is 8.2 times greater than the SM prediction. This result improves

TABLE III: Upper limits at $95 \%$ C.L. on the $Z H \rightarrow \ell^{+} \ell^{-} b \bar{b}$ cross section, shown as a ratio to the SM cross section. The column labelled "Expected" shows the median of the limits obtained from simulated experiments containing no signal, and the columns labelled " $\pm 1 \sigma$ " show the range containing $68 \%$ of the expected limits.

| $m_{H}\left[\mathrm{GeV} / c^{2}\right]$ | $-1 \sigma$ <br> $\left[\sigma / \sigma_{S M}\right]$ | Expected <br> $\left[\sigma / \sigma_{S M}\right]$ | $+1 \sigma$ <br> $\left[\sigma / \sigma_{S M}\right]$ | Observed <br> $\left[\sigma / \sigma_{S M}\right]$ |
| :---: | :---: | :---: | :---: | :---: |
| 100 | 6.0 | 8.7 | 12.4 | 7.0 |
| 105 | 6.0 | 8.7 | 12.9 | 6.5 |
| 110 | 7.5 | 11.3 | 16.8 | 7.6 |
| 115 | 8.3 | 12.1 | 18.2 | 8.2 |
| 120 | 9.3 | 13.5 | 20.0 | 9.0 |
| 125 | 13.2 | 18.3 | 27.1 | 13.2 |
| 130 | 17.1 | 24.2 | 35.7 | 17.7 |
| 135 | 21.8 | 31.0 | 44.8 | 22.9 |
| 140 | 31.0 | 44.3 | 65.4 | 32.0 |
| 145 | 42.8 | 61.6 | 89.9 | 43.1 |
| 150 | 73.7 | 104 | 153 | 71.3 |

by a factor of 2 over the previously published result in this channel [5]. We are exploring further improvements in this technique by separating the leading-order and next-to-leading order contributions to the signal and backgrounds, as well as the use of matrix-element-based probabilities in conjunction with other multivariate discriminants.

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