



Original Contribution

Serum Vitamin D and Risk of Prostate Cancer in a Case-Control Analysis Nested Within the European Prospective Investigation into Cancer and Nutrition (EPIC)

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Results from the majority of studies show little association between circulating concentrations of vitamin D and prostate cancer risk, a finding that has not been demonstrated in a wider European population, however. The authors examined whether vitamin D concentrations were associated with prostate cancer risk in a case-control study nested within the European Prospective Investigation into Cancer and Nutrition (1994–2000). Serum concentrations of 25-hydroxyvitamin D were measured in 652 prostate cancer cases matched to 752 controls from 7 European countries after a median follow-up time of 4.1 years. Conditional logistic regression models were used to calculate odds ratios for prostate cancer risk in relation to serum 25-hydroxyvitamin D after standardizing for month of blood collection and adjusting for covariates. No significant association was found between 25-hydroxyvitamin D and risk of prostate cancer (highest vs. lowest quintile: odds ratio = 1.28, 95% confidence interval: 0.88, 1.88; *P* for trend = 0.188). Subgroup analyses showed no significant heterogeneity by cancer stage or grade, age at diagnosis, body mass index, time from blood collection to diagnosis, or calcium intake. In summary, the results of this large nested case-control study provide no evidence in support of a protective effect of circulating concentrations of vitamin D on the risk of prostate cancer.

prostatic neoplasms; serum; vitamin D; 25-hydroxyvitamin D 2

Abbreviations: EPIC, European Prospective Investigation into Cancer and Nutrition; 1,25(OH)₂D, 1,25-dihydroxyvitamin D; 25(OH)D, 25-hydroxyvitamin D.

In 1990, findings of an inverse association between ultraviolet exposure and prostate cancer mortality rates in the United States led to the hypothesis that vitamin D insufficiency may increase prostate cancer risk (1). Results from experimental studies in animal models and cell lines as well as a phase II clinical trial support the hypothesis that vitamin D may play a role in the etiology of prostate cancer (2–4); however, evidence from prospective studies in the United States and Nordic countries

that have assessed the relation between circulating concentrations of 25-hydroxyvitamin D (25(OH)D) and 1,25-dihydroxyvitamin D (1,25(OH)₂D) and risk of prostate cancer has not clearly supported such an association (5–16). Despite the consistent finding of no association between circulating concentrations of vitamin D and the risk of prostate cancer, there is still considerable support for vitamin D insufficiency as a risk factor for prostate cancer (17, 18). To date, we know of no study that has

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investigated this association among men from other European countries.

We examine here the relation between prediagnostic serum concentrations of 25(OH)D and risk of prostate cancer among 652 men with incident prostate cancer and 752 matched controls participating in a large, multicenter cohort study: the European Prospective Investigation into Cancer and Nutrition (EPIC). We also evaluate these associations by stage and grade of disease.

MATERIALS AND METHODS

Participants

The methods of recruitment for EPIC and the design of the study have been described in detail elsewhere (19). The present study includes prostate cancer cases diagnosed after blood collection and individually matched to controls from 7 of the 10 participating countries: Germany, Greece, Italy, the Netherlands, Spain, Sweden, and the United Kingdom. France and Norway, and the regional centers in Utrecht, the Netherlands, and Naples, Italy, were not included in the current study because these cohorts included women only. Centers in Denmark and Malmö, Sweden, did not contribute data to the current study. Each case was matched to 1 control, except in Umeå, Sweden, where each case was matched to 2 controls. Matching criteria were study center, age at enrollment (± 6 months), time of day of blood collection (± 1 hour), and time between blood draw and last consumption of food or drink (<3 , $3-6$, >6 hours; for Umeå <4 , $4-8$, >8 hours). A more detailed description of follow-up for cancer incidence and the selection of cases and controls has been published elsewhere (20).

Diet and lifestyle questionnaires

Dietary intake during the year before recruitment was measured by country-specific validated dietary questionnaires designed to capture local dietary patterns. Information on validation of the food questionnaires has been published previously (19, 21). Calcium intakes were calculated by multiplying the calcium content of each food of a specific portion size or quantity (in grams) by the frequency of consumption as stated on the food questionnaire using country-specific national food tables. Intake from supplemental calcium was not included in this analysis.

Blood collection

A 30-mL blood sample was collected from 139,207 (91%) of the 153,457 men participating in EPIC according to a standardized protocol at the time of recruitment. Filled syringes were kept at 5°C – 10°C , protected from light, and transferred to a local laboratory for further processing and aliquoting, except for participants recruited through the Oxford center. Here, blood samples were collected throughout the United Kingdom and were transported to a laboratory in Norfolk by mail at ambient temperature. Blood for serum was aliquoted into 0.5-mL straws, which were then heat-sealed at both ends and stored in liquid nitrogen tanks at -196°C , except for Umeå, where samples were stored in 1.5-mL plastic tubes at -70°C .

Laboratory assays

Serum 25(OH)D concentration was determined by enzyme immunoassay (OCTEIA 25-Hydroxy Vitamin D kit; Immunodiagnostic Systems, Limited, Bolton, Tyne and Wear, United Kingdom) in the Medical Research Council Human Nutrition Research laboratories in Cambridge by laboratory personnel blinded to the case or control status of participants. Serum samples from each case-control set were assayed within the same batch. The intraassay and interassay coefficients of variation (unblinded) were 14.8% and 12.0%, respectively, at 20.2 nmol/L and 3.9% and 10.8%, respectively, at 72.7 nmol/L. The lowest limit of detection was 5 nmol/L.

Statistical analyses

Differences between cases and controls regarding age at blood collection, height, weight, body mass index, and serum concentration of vitamin D were investigated by using a weighted version of the paired-sample *t* test (22). Differences in the categorical baseline characteristics between cases and controls were compared by using conditional logistic regression models. To assess the influence of month of blood draw on serum concentrations of vitamin D, a simple mathematical model of log-transformed serum vitamin D concentration by month of blood collection (as a categorical variable) was fitted. The “standardized” concentrations of 25(OH)D were calculated by adding the residuals from this model to the overall mean log serum vitamin D value and exponentiating these values, and the resulting serum vitamin D concentrations “standardized for month of blood collection” were used for all analyses. The standardized concentrations were also used to calculate the cutpoints for the quintiles and the “low,” “moderate,” and “high” categories of serum 25(OH)D.

Conditional logistic regression models were used to calculate the odds ratios for prostate cancer by quintiles of 25(OH)D concentration, using cutpoints calculated among controls. Tests for linear trend were obtained by replacing the quintiles with the logarithm of serum concentration of 25(OH)D in the model. The effects of potential confounders—body mass index (weight (kg)/height (m)² in quartiles), smoking status (never, former, current), alcohol intake (<8 g/day, 8–15 g/day, 16–39 g/day, ≥ 40 g/day), education (primary or none, secondary, degree level), marital status (married and/or cohabiting, not married or cohabiting), and physical activity (inactive, moderately inactive, active)—were examined by including these variables in the logistic regression models. Any missing values were assigned to a separate category. Likelihood ratio chi-squared tests were used to examine the heterogeneity of trends in prostate cancer risk with the logarithm of serum 25(OH)D concentration between localized (tumor, node, metastasis staging score of T0 or T1 or T2 and N0 or NX and M0, or stage coded in the recruitment center as localized; $n = 324$) or advanced (T3 or T4 and/or N1+ and/or M1, or stage coded in the recruitment center as metastatic; $n = 122$), low grade (Gleason sum <7 or equivalent (cases coded as well differentiated or moderately differentiated); $n = 361$)

or high grade (Gleason sum ≥ 7 or equivalent (cases coded as poorly differentiated or undifferentiated); $n = 169$), age at diagnosis (< 60 years or ≥ 60 years), body mass index (< 25 kg/m² or ≥ 25 kg/m²), and country of recruitment. We also compared the risk of prostate cancer for men with a diagnosis of prostate cancer less than 4 years ($n = 309$) with that for those men diagnosed 4 or more years ($n = 343$) after blood collection.

These analyses were repeated by using predefined cut-points for concentrations of 25(OH)D of < 50 , 50 – < 75 and ≥ 75 nmol/L to investigate risks associated with low (< 50 nmol/L), moderate (50 – < 75 nmol/L), and high (≥ 75 nmol/L) concentrations of vitamin D (20). We also used the chi-squared test to assess effect modification of a low ($< 1,200$ mg/day) or high ($\geq 1,200$ mg/day) intake of calcium from food sources on the association between low, moderate, and high concentrations of serum vitamin D and risk of prostate cancer.

All statistical analyses were carried out with Stata Statistical Software, release 9 (StataCorp LP, College Station, Texas). All tests of statistical significance were 2-sided, and P values of < 0.05 were considered significant.

RESULTS

When characteristics of cases and controls were compared, the groups did not differ appreciably, except that cases reported slightly less physical activity. The mean concentration of 25(OH)D did not differ significantly between cases and controls ($P = 0.88$) (Table 1).

After we adjusted for age at blood collection, serum concentrations of 25(OH)D varied significantly by month of blood collection among the cases and controls (Figure 1), but there was no significant difference in the mean month of blood collection between the matched sets ($P = 0.643$). Mean concentrations were lowest in cases in January (44.7 nmol/L) and in controls in March (42.9 nmol/L), and they were highest in August (72.4 nmol/L) and September (78.4 nmol/L), respectively. Following standardization of 25(OH)D concentrations for month of blood collection and adjustment for age at blood collection, concentrations varied significantly by country of recruitment ($P < 0.001$) among the controls, with the highest mean concentrations in men in the United Kingdom (57.5 nmol/L) and Sweden (56.5 nmol/L) and the lowest in the Dutch men (45.2 nmol/L) (Figure 2). Following standardization of 25(OH)D concentrations for month of blood collection and adjustment for study center, we found a significant inverse association between 25(OH)D concentrations and age at blood collection; compared with men older than age 70 years, men aged 50–54 years had a vitamin D concentration that was 12.5 nmol/L higher ($P = 0.001$, results not shown).

The odds ratios for prostate cancer by quintiles of 25(OH)D from conditional logistic regression models conditioned on the matching variables and adjusted for lifestyle variables are shown in Table 2. Results showed no significant association between concentrations of 25(OH)D and overall risk of prostate cancer, with or without adjustment for covariates. In the unadjusted results, the risk of prostate cancer for the highest versus the lowest quintile of serum

concentrations of vitamin D was 1.24 (95% confidence interval: 0.87, 1.79; P for trend = 0.265). After additional adjustment for body mass index, smoking status, alcohol intake, education, marital status, and physical activity, the odds ratio for the highest in comparison to the lowest quintile of vitamin D changed very little (odds ratio = 1.28, 95% confidence interval: 0.88, 1.88; P for trend = 0.188). There was no evidence of significant heterogeneity in the trends for the risk of prostate cancer with vitamin D concentration by stage or grade or by the age of men at diagnosis (Table 2). However, when the associations between prostate cancer risk and 25(OH)D were examined separately by age, a borderline significant positive association was observed for men diagnosed at less than age 60 years (odds ratio for the highest compared with the lowest quintile = 2.90, 95% confidence interval: 0.85, 9.88; P for trend = 0.046), and this association was not significant for men aged 60 years or older (P for trend = 0.594); however, neither was the test for heterogeneity between the trends statistically significant (P for heterogeneity = 0.160).

Table 3 results show that the risk of prostate cancer was not significantly different for men with a low (< 50 nmol/L) or a high (≥ 75 nmol/L) concentration of vitamin D in comparison to men with a moderate (50 – < 75 nmol/L) concentration. There was also no association between these predefined categories of vitamin D concentrations and risk of prostate cancer subdivided by stage, grade, or age at diagnosis. In addition, we found no association between the categories of vitamin D and risk of prostate cancer for men with a high ($\geq 1,200$ mg/day) or a low ($< 1,200$ mg/day) intake of calcium (Table 4).

There was no evidence for significant heterogeneity regarding the association between serum 25(OH)D and risk of prostate cancer between men with an early (< 4 years) versus a later (≥ 4 years) diagnosis (P for heterogeneity = 0.857), a high (≥ 25 kg/m²) versus a low (< 25 kg/m²) body mass index (P for heterogeneity = 0.825), or by country (P for heterogeneity = 0.166, results not shown).

DISCUSSION

The findings from this large, prospective study show that the risk of prostate cancer did not vary significantly by serum concentration of 25(OH)D and do not support the hypothesis that circulating 25(OH)D plays a major role in the pathogenesis of prostate cancer. Despite the widespread notion that vitamin D insufficiency is an important risk factor for prostate cancer (17, 18), this theory has not been substantiated by results from the majority of published prospective studies (5, 7, 8, 10–14). Results from 2 studies have shown support for lower concentrations of 25(OH)D and increased risk of prostate cancer. Ahonen et al. (6) demonstrated a greater risk of prostate cancer for Nordic men with a 25(OH)D concentration of ≤ 40 nmol/L in comparison to men with concentrations of > 40 nmol/L; a second Nordic study reported a U-shaped relation, with a higher risk for men with low (≤ 19 nmol/L) and high (≥ 80 nmol/L) concentrations of 25(OH)D in comparison to men with moderate concentrations (15).

Table 1. Baseline Characteristics of Prostate Cancer Cases and Matched Controls in the European Prospective Investigation into Cancer and Nutrition, 1994–2000

Characteristic	Cases (n = 652)	Controls (n = 752)	P Value ^a
	<i>No. (%)^b</i>		
Smoking ^c			
Never	217 (34.5)	243 (34.6)	0.82
Former	283 (45.0)	306 (43.6)	
Current	129 (20.5)	153 (21.8)	
Alcohol consumption ^c			
<8 g/day	260 (41.5)	321 (45.7)	0.15
8–15 g/day	120 (19.2)	148 (21.1)	
16–39 g/day	149 (23.8)	141 (20.1)	
≥40 g/day	97 (15.5)	92 (13.1)	
Physical activity ^c			
Inactive	103 (17.4)	85 (13.2)	0.04
Moderately inactive	212 (33.8)	218 (33.8)	
Active	278 (46.9)	342 (53.0)	
Marital status ^c			
Married and/or cohabiting	480 (88.9)	550 (87.9)	0.59
Not married or cohabiting	60 (11.1)	76 (12.1)	
Educational attainment ^c			
Primary or equivalent	245 (39.8)	290 (41.5)	0.22
Secondary	224 (36.4)	270 (38.6)	
Degree	147 (23.9)	139 (19.9)	
	<i>Mean (SD)</i>		
Age at blood collection, years	61.0 (6.3)	60.5 (6.2)	0.15
Weight, kg ^c	79.0 (11.1)	80.0 (11.2)	0.10
Height, cm ^c	172.3 (6.7)	172.9 (6.9)	0.10
Body mass index, kg/m ^{2c}	26.6 (3.4)	26.8 (3.5)	0.40
Serum 25-hydroxyvitamin D, mean (95% CI), nmol/L ^d	53.6 (52.0, 55.3)	53.5 (51.9, 55.1)	0.88
Year of diagnosis, median (range)	1999 (1994–2005)		
Age at diagnosis, mean (SD), years	65.4 (6.1)		
Months from blood collection to diagnosis, median (range)	49 (0–181)		
Prostate cancer stage, ^e no. (%)			
Localized	326 (50.0)		
Advanced	122 (18.7)		
Unknown	204 (31.3)		
Histologic grade, ^f no. (%)			
Low	362 (55.5)		
High	170 (26.1)		
Unknown	120 (18.4)		

Abbreviations: CI, confidence interval; SD, standard deviation.

^a Weighted tests of mean difference between cases and controls in each matched set or tests of association between category and case-control status using conditional logistic regression, as appropriate.

^b Percentages may not add to 100 because of rounding.

^c Unknown for some subjects; calculations of percentages exclude missing values.

^d Serum concentrations of vitamin D were standardized for month of blood collection.

^e Tumor, node, metastasis (TNM) staging score of T0 or T1 or T2 and N0 or NX and M0 (localized); TNM staging score of T3 or T4 and/or N1+ and/or M1 (advanced).

^f Gleason sum <7 or equivalent or well or moderately differentiated (low grade); Gleason sum ≥7 or poorly differentiated or undifferentiated (high grade).

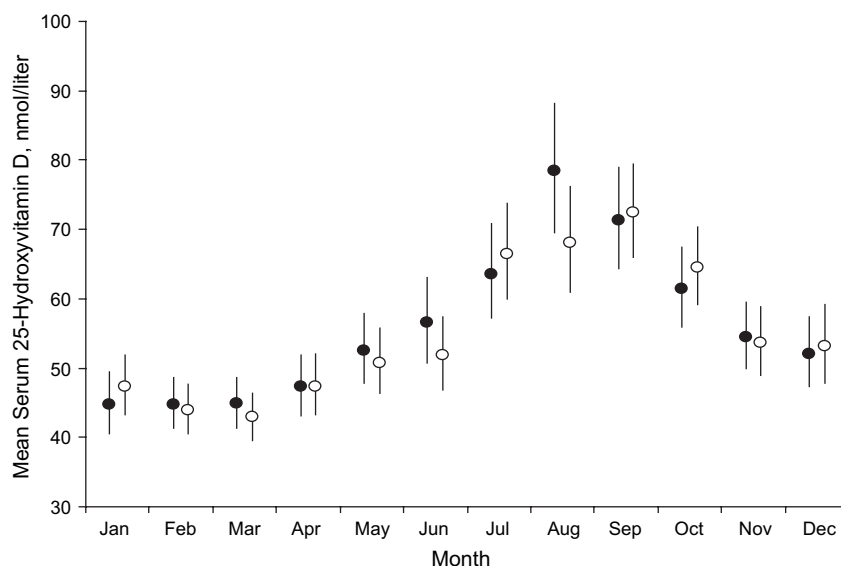


Figure 1. Geometric mean serum 25-hydroxyvitamin D concentrations, by month of blood draw, in cases and controls in the European Prospective Investigation into Cancer and Nutrition, 1994–2000. Means are geometric means and their corresponding 95% confidence intervals adjusted for age at blood draw, study center, and their interaction. Solid circles, means for cases; open circles, means for controls. Jan, January; Feb, February; Mar, March; Apr, April; Jun, June; Jul, July; Aug, August; Sep, September; Oct, October; Nov, November; Dec, December.

Some studies have suggested that the inconsistent results regarding an association between vitamin D and risk of prostate cancer may be due to the variation in vitamin D concentrations between populations. For instance, the proportion of men with low 25(OH)D concentrations (<50 nmol/L) was higher among the Nordic populations (6, 9, 15) than in most of the US cohorts, where sun exposure is

likely to be greater (5, 11, 13), and/or study populations were drawn from health-conscious populations (12, 14) whose intake of vitamin D may be higher. Our data on serum 25(OH)D concentrations suggest a moderate prevalence of vitamin D insufficiency in men across Europe; 39% of the cases and controls had 25(OH)D concentrations of <50 nmol/L. Our finding of a null association was also similar

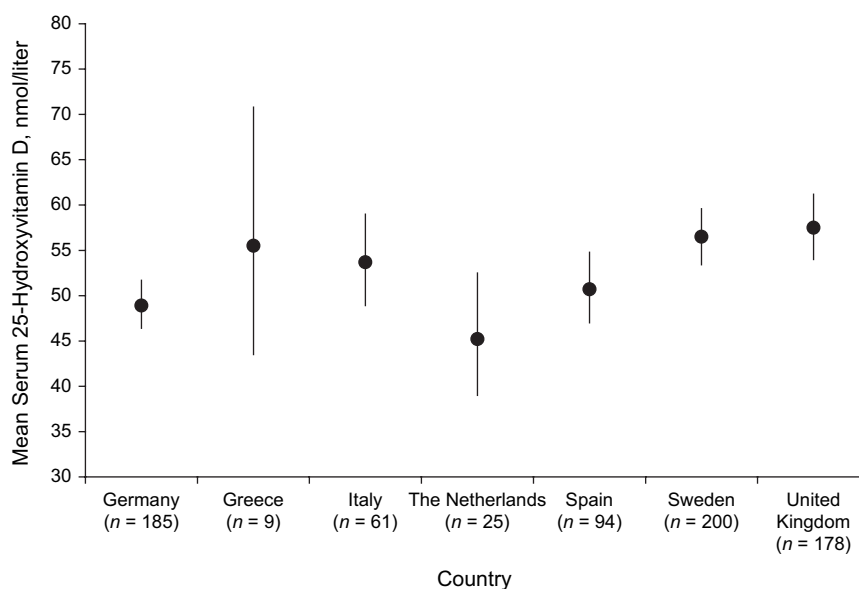


Figure 2. Geometric mean (and 95% confidence interval) serum 25-hydroxyvitamin D concentrations, by country, in controls in the European Prospective Investigation into Cancer and Nutrition, 1994–2000. Means are geometric means and their corresponding 95% confidence intervals standardized for month of blood collection and adjusted for age at blood draw.

Table 2. Odds Ratios (and 95% Confidence Intervals) for Prostate Cancer by Quintile of 25-Hydroxyvitamin D Concentration in the European Prospective Investigation into Cancer and Nutrition, 1994–2000

	Quintile of 25-Hydroxyvitamin D Concentration ^a										Doubling of Concentration		P for Trend ^b	P for Heterogeneity ^c
	1 ^d	2		3		4		5		OR	95% CI			
		OR	95% CI	OR	95% CI	OR	95% CI	OR	95% CI					
Range, nmol/L	2.5–40.4	40.5–50.4		50.5–59.1		59.2–70.8		70.9–163.7						
No. of cases	125	143		128		114		142						
No. of controls	151	150		151		150		150						
Odds ratio ^e	1	1.24	0.88, 1.74	1.15	0.82, 1.62	1.01	0.71, 1.45	1.24	0.87, 1.79	1.13	0.91, 1.41	0.265		
Adjusted odds ratio ^f	1	1.27	0.89, 1.81	1.23	0.85, 1.76	1.06	0.73, 1.55	1.28	0.88, 1.88	1.17	0.93, 1.47	0.188		
Prostate cancer stage ^f														
Localized (<i>n</i> = 326)	1	1.28	0.76, 2.16	1.20	0.72, 1.97	1.09	0.65, 1.83	1.15	0.67, 1.97	1.12	0.80, 1.56	0.508	0.577	
Advanced (<i>n</i> = 122)	1	1.61	0.71, 3.65	1.10	0.40, 2.99	0.83	0.32, 2.12	1.13	0.37, 3.43	1.31	0.73, 2.37	0.364		
Histologic grade ^f														
Low (<i>n</i> = 362)	1	1.45	0.89, 2.37	1.37	0.85, 2.21	1.30	0.79, 2.15	1.52	0.91, 2.56	1.29	0.95, 1.74	0.097	0.108	
High (<i>n</i> = 170)	1	0.94	0.42, 2.10	0.63	0.26, 1.49	0.86	0.36, 2.05	0.83	0.34, 2.07	0.94	0.52, 1.70	0.846		
Age at blood collection ^f														
<60 years (<i>n</i> = 112)	1	2.02	0.59, 6.90	1.87	0.65, 5.37	3.73	1.06, 13.17	2.90	0.85, 9.88	1.90	0.99, 3.64	0.046	0.160	
≥60 years (<i>n</i> = 540)	1	1.30	0.89, 1.90	1.19	0.79, 1.77	0.93	0.62, 1.40	1.19	0.78, 1.80	1.07	0.83, 1.39	0.594		

Abbreviations: CI, confidence interval; OR, odds ratio.

^a Serum concentrations of vitamin D were standardized for month of blood collection.

^b Values were obtained with the logarithm of plasma 25-hydroxyvitamin D replacing the categorical plasma 25-hydroxyvitamin D variable in the model.

^c Values relate to likelihood ratio chi-squared tests of heterogeneity between trends for localized and advanced-stage and low- and high-grade prostate cancer, and for age <60 years and ≥60 years at blood collection.

^d All values of 1 are odds ratios.

^e Conditioned on matching variables by using conditional logistic regression.

^f Conditioned on matching variables and adjusted for body mass index, smoking, alcohol intake, education, marital status, and physical activity by using conditional logistic regression.

to a study in a Finnish population in which 50% of participants had 25(OH)D concentrations of <50 nmol/L (9).

Several studies have evaluated the risk of prostate cancer associated with concentrations of 1,25(OH)₂D, the hormonal biologically active form of vitamin D, which has been shown in experimental studies to reduce the degree of cell proliferation in the prostate (4, 23, 24). In 2 studies, the investigators reported a nonsignificant decreased risk for men with high concentrations of both vitamin D metabolites (8, 13), whereas several others have reported no association (7, 10–12, 14). We chose to assess the association between prostate cancer risk and serum 25(OH)D and not 1,25(OH)₂D because circulating 25(OH)D (from diet, supplementation, and sun exposure) is probably a better marker of an individual's vitamin D exposure than circulating concentrations of 1,25(OH)₂D, which are homeostatically controlled, have a half life of 4 hours, and most probably reflect the production of 1,25(OH)₂D in the kidneys (25). However, a potential limitation of the current study and all previous epidemiologic studies is that it is not clear to what extent circulating 25(OH)D reflects intraprostatic vitamin D concentrations because 1,25(OH)₂D is produced locally in the prostate by cells expressing the enzyme 25(OH)-1 α -hydroxylase (26).

A potential effect modifier of the association between vitamin D and prostate cancer risk is calcium. A high intake

of calcium coupled with low vitamin D status may increase the risk of prostate cancer by reducing the amount of 1,25(OH)₂D synthesized (27). High levels of calcium may suppress the release of parathyroid hormone, and the action of this hormone tightly regulates conversion of 25(OH)D to 1,25(OH)₂D in renal cells (28). Nevertheless, these results showed no evidence that the association between concentrations of 25(OH)D and the risk of prostate cancer varied according to calcium intake, which is consistent with findings reported by other studies that have examined the effect of calcium intake (5, 12, 14) or status (13, 29).

We found no support for a difference in the relation between vitamin D and prostate cancer by stage or histologic grade of the disease, which is consistent with others (7, 8, 12–14) apart from Ahn et al. (5), whose results showed a positive association between 25(OH)D and risk of aggressive disease (Gleason sum ≥7 or clinical stage III or IV). These results also showed no heterogeneity in the association between 25(OH)D and prostate cancer risk by age at diagnosis; however, there was a positive association between serum concentrations of vitamin D and risk of disease for men diagnosed at less than age 60 years. The majority of recent studies have reported no substantial difference in the relation between vitamin D and prostate cancer risk by age (7, 11–14), with the exception of 2 studies (6, 8). Ahonen et al. (6) demonstrated an inverse association between

Table 3. Odds Ratios (and 95% Confidence Intervals) for Prostate Cancer by Predefined Categories of 25-Hydroxyvitamin D Concentration in the European Prospective Investigation into Cancer and Nutrition, 1994–2000

	Category of 25-Hydroxyvitamin D Concentration ^a						P Value ^b
	1		2 ^c	3			
	OR	95% CI		OR	95% CI		
Range, nmol/L	<50		50–74.9	≥75			
No. of cases	258		283	111			
No. of controls	286		353	113			
Odds ratio ^d	1.04	0.82, 1.32	1	1.17	0.85, 1.62		0.265
Adjusted odds ratio ^e	1.00	0.78, 1.28	1	1.14	0.82, 1.58		0.188
Prostate cancer stage ^e							
Localized	1.06	0.75, 1.49	1	1.02	0.64, 1.62		0.508
Advanced	1.22	0.66, 2.25	1	1.17	0.44, 3.13		0.364
Histologic grade ^e							
Low	0.95	0.69, 1.30	1	1.19	0.76, 1.88		0.097
High	1.25	0.71, 2.19	1	1.29	0.62, 2.70		0.846
Age at blood collection ^e							
<60 years	1.08	0.74, 1.58	1	1.00	0.62, 1.62		0.046
≥60 years	0.83	0.58, 1.19	1	1.33	0.81, 2.18		0.594

Abbreviations: CI, confidence interval; OR, odds ratio.

^a Serum concentrations of vitamin D were standardized for month of blood collection.

^b Values were obtained with the logarithm of plasma 25-hydroxyvitamin D replacing the categorical plasma 25-hydroxyvitamin D variable in the model.

^c All values of 1 are odds ratios.

^d Conditioned on matching variables by using conditional logistic regression.

^e Conditioned on matching variables and adjusted for body mass index, smoking, alcohol intake, education, marital status, and physical activity by using conditional logistic regression.

Table 4. Odds Ratios (and 95% Confidence Intervals) for Prostate Cancer by Predefined Categories of 25-Hydroxyvitamin D Concentration and Intake of Dietary Calcium in the European Prospective Investigation into Cancer and Nutrition, 1994–2000

Calcium Intake, mg/day	Category of 25-Hydroxyvitamin D Concentration ^a (Range, nmol/L)												P Value ^b	
	1 (<50)				2 (50–74.9)				3 (≥75)					
	No. of Cases ^c	No. of Controls ^c	OR	95% CI	No. of Cases ^c	No. of Controls ^c	OR	95% CI	No. of Cases ^c	No. of Controls ^c	OR	95% CI		
<1,200	196	186			183	216			75	70				
≥1,200	39	59			63	61			24	24				
Unadjusted ^d														
<1,200			1.19	0.89, 1.60							1.25	0.84, 1.86		
≥1,200			0.73	0.46, 1.16			1.20	0.79, 1.82			1.15	0.63, 2.11		0.305
Adjusted ^e														
<1,200			1.15	0.80, 1.68							1.20	0.80, 1.81		
≥1,200			0.69	0.43, 1.11			1.19	0.78, 1.83			1.09	0.58, 2.05		0.333

Abbreviations: CI, confidence interval; OR, odds ratio.

^a Serum concentrations of vitamin D were standardized for month of blood collection.

^b Values were obtained by using chi-squared tests.

^c For 72 cases and 136 controls, information on dietary calcium intake was missing.

^d Conditioned on matching variables by using conditional logistic regression.

^e Conditioned on matching variables and adjusted for body mass index, smoking, alcohol intake, education, marital status, and physical activity by using conditional logistic regression.

25(OH)D concentrations and prostate cancer risk for men aged than age 52 years, and Corder et al. (8) reported a similar inverse association but only for men older than age 57 years. Notwithstanding the possibility of a differential effect by age, both of these subgroup analyses and our own were based on a small number of cases ($n = 67$, $n = 91$, and $n = 112$ men, respectively), and, because a number of comparisons were made in our analyses, the role of chance cannot be excluded.

Others have also demonstrated that the influence of low concentrations of 25(OH)D on the risk of prostate cancer differed according to several polymorphisms located on the vitamin D receptor gene, including *Cdx2*, *Fok1*, and *Bsm1* (12, 30). Moreover, evidence also suggests that polymorphisms in the vitamin D binding protein affect circulating concentrations of 25(OH)D (31). Given that the genotype of vitamin D receptor or vitamin D binding protein was not determined for the men in this study, our results do not rule out the possibility that low levels of circulating 25(OH)D may be associated with a greater risk of prostate cancer for certain individuals with a specific genotype or haplotype.

Strengths of this study include the large number of cases, prospective design, and inclusion of men from a number of European countries. One of the limitations is that participants were not matched on month of blood collection. Because serum 25(OH)D concentrations varied according to the month in which the blood sample was collected, vitamin D values were standardized for month of blood collection before we assessed the association between vitamin D and risk of prostate cancer so as to minimize the risk of confounding by season of blood collection. The mean preclinical duration of prostate cancer has been estimated to be 10 years (32); in this study, the median time between blood collection and diagnosis of cases was 4 years. Thus, a potential limitation of this analysis was that serum 25(OH)D concentrations in the cases reflected vitamin D status at a time when early preclinical tumors were present rather than prior to initiation of tumor growth. It is possible that any subclinical tumors present at the time of blood collection influenced serum concentrations of 25(OH)D. Notwithstanding these limitations, exclusion of cases diagnosed in the first 4 years of follow-up did not alter the main findings.

In conclusion, the findings from this large prospective study of European men showed no strong association between serum concentrations of 25(OH)D and risk of total prostate cancer. Taken together with the results from other prospective studies, the totality of evidence indicates that low vitamin D concentrations do not have an important role in the etiology of prostate cancer. Further pooled analyses would help to determine whether there are any subgroup associations or an effect of the exposure distribution and range of vitamin D concentrations on the risk of prostate cancer.

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