

Static and Dynamic Path Selection on Expander Graphs: A Random Walk Approach

(Preliminary version)

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Abstract

This paper addresses the problem of virtual circuit switching in bounded degree expander graphs. We study the static and dynamic versions of this problem. Our solutions are based on the rapidly mixing properties of random walks on expander graphs.

In the static version of the problem an algorithm is required to route a path between each of K pairs of vertices so that no edge is used by more than g paths. A natural approach to this problem is through a multi-commodity flow reduction. However, we show that the random walk approach leads to significantly stronger results than those recently obtained by Leighton and Rao [10] using the multi-commodity flow setup.

In the dynamic version of the problem connection requests are continuously injected into the network. Once a connection is established it utilizes a path (a virtual circuit) for a certain time until the communication terminates and the path is deleted. Again each edge in the network should not be used by more than g paths at once.

The dynamic version is a better model for the practical use of communication networks. Our random walk approach gives a simple and fully distributed solution for this problem. We show that if the injection to the network and the duration of connections are both controlled by Poisson processes then our algorithm achieves

a steady state utilization of the network which is similar to the utilization achieved in the static case situation.

1 Introduction

Communication protocols for high-speed high bandwidth networks are based on *virtual circuit switching*. The speed of the network does not allow for on-line routing of individual packets. Instead, upon establishing a connection, bandwidth is allocated along a path connecting the two endpoints for the duration of the connection. These “virtual circuits” are set up on a per-call basis and are disconnected when the call is terminated. Efficient utilization of the network depends on the allocation of virtual circuits between pairs of nodes so that no link is overloaded beyond its capacity.

As in other routing problems we distinguish between a static and a dynamic version. In the static version all the requests are given at once and must be simultaneously satisfied. In the dynamic version requests are continuously generated. A connection once established, continues for a certain amount of time and afterwards its bandwidth can be re-used for other connections.

The static version of this problem translates into the following combinatorial question: Given a network $G = (V, E)$ and a set of K pairs of vertices in V , find for each pair (a_i, b_i) , a path connecting a_i to b_i , such that no edge is used by more than g paths. For arbitrary graphs, the related decision problem is in \mathcal{P} for fixed K - Robertson and Seymour [11], but is \mathcal{NP} -complete if K is part of the input.

In contrast to the negative results for general graphs a significant progress has been made in solving this problem for the interesting class of bounded degree expander graphs. In particular Leighton and Rao [10] have recently obtained a number of constructive and existential results for this problem based on the natural linear relaxation of the circuit switching problem to a multi-commodity flow problem. This paper explores a different approach: we study solutions for this problem based on the rapidly mixing properties of random walks on ex-

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pander graphs. The method we explore here builds on the basic technique developed in [4]. Using the random walk approach we improve the Leighton-Rao results for the static case and obtain the first non-trivial results for the dynamic version of the virtual circuit switching problem.

Our main result for the static case establishes a trade-off between the number of pairs K , and the allowed congestion g .

Theorem 1 *There is an explicit polynomial time algorithm that can connect any set of $K = \alpha(n)n/\ln n$ pairs of vertices on a bounded degree expander so that no edge is used by more than g paths where*

$$g = \begin{cases} O\left(s + \left\lceil \frac{\log \log n}{\log(1/\hat{\alpha})} \right\rceil\right), & \text{for } \alpha < 1/2; \\ O(s + \alpha + \log \log n), & \text{for } \alpha \geq 1/2, \end{cases}$$

$\hat{\alpha} = \min(\alpha, 1/\log \log n)$, and s is the maximal multiplicity of a vertex in the set of pairs.

Since the expected distance between two random vertices on a bounded degree expander is $\Omega(\log n)$ our results are always within an additive factor of $O(\log \log n)$ from optimal. Our theorem improves the results obtained by Leighton and Rao [10] using the multi-commodity flow approach. The case $\alpha(n) = 1/(\ln n)^\epsilon$ improves the bound of Theorem 1 of [10], the case $\alpha(n) = O(1)$ improves the bound of Theorem 2 in [10], and makes the result constructive.

We also note that Kleinberg and Rubinfeld [7] have recently used our result in their analysis of a greedy algorithm for finding *short* disjoint paths on expanders.

The following non-constructive result (See Section 2 for definitions) improves Theorem 3 in [10]:

Theorem 2 *Given a bounded degree (α, β, γ) -expander graph there exists a parameter c that depends on α, β, γ , but not on n , such that any set of less than $cn/(\ln n)^2$ disjoint pairs of vertices can be connected by edge disjoint paths.*

The advantage of the random walks approach is even more significant in the solution of the dynamic virtual circuit switching problems. In practice networks are rarely used in the “batch mode” modeled by the static problem. Real-life network performance is better modeled by a dynamic process whereby requests for connection are continuously arriving at the nodes of the network. A connection has a duration time, and once the communication has terminated its bandwidth can be used for another connection. In section 4 we formulate a model for studying the dynamic virtual circuit switching problem under stochastic assumptions about the injection rate of new requests and the duration of connections.

Using the random walk approach we develop a simple and fully distributed protocol for dynamic path selection on bounded degree expander graphs. For the analysis we adopt the stochastic model assumed in the design of most long-distance telephone networks [9]. Requests arrives according to a Poisson process, and the duration of a connection is exponentially distributed. Our goal is to characterize the relationship between the load and congestion parameters that guarantees system stability (i.e., the expected number of requests in queue is not growing unboundly with time).

Theorem 3 *Let $E(N)$ be the expected number of requests that arrive to the n -node expander network at a given step, let $E(D)$ be the expected duration of a connection. If $E(N)E(D) \leq \alpha(n)n/\ln n$, where $\alpha(n) \leq 1/2$, then there exists a constant c such that for $g \geq c \lceil \ln \ln n / \ln(1/\alpha) \rceil$ the system running our dynamic algorithm is stable, and the expected time a request waits in a queue is $O(n/E(N))$.*

Since $E(N)E(D)$ is the expected number of paths that must be active in the network at the steady state in order to keep the system stable, we see that our dynamic algorithm achieves almost the same edge utilization as that in our static results.

One should note that our approach differs from the work on admission control [2, 6, 3] in that we do not reject requests. All requests are eventually satisfied in our model, but not immediately. In contrast, in the admission control model a request is either immediately satisfied or it is rejected. Our approach better models computer communication, while the admission control approach is a better model for human (telephone) communication.

2 Preliminaries

There are various ways to define expander graphs; here we define them in terms of edge expansion (a weaker property than vertex expansion).

Let $G = (V, E)$ be a graph. For a set of vertices $S \subset V$ let $\text{out}(S)$ be the set of edges with one end-point in S and one end-point in $V \setminus S$, that is

$$\text{out}(S) = \{ \{u, v\} \mid \{u, v\} \in E, u \in S, v \notin S \}.$$

Definition 1 *A graph $G = (V, E)$ is a β -expander, if for every set $S \subset V$, $|S| \leq |V|/2$, we have $|\text{out}(S)| \geq \beta|S|$.*

For the remainder of this paper, when β is not explicitly mentioned we will assume that it is an arbitrary constant greater than 0. For certain results we need expanders that have the property that the expansion of small sets is not too small.

Definition 2 An r -regular graph $G = (V, E)$ is called an (α, β, γ) -expander if for every set $S \subset V$

$$\text{out}(S) \geq \begin{cases} (3r/4 + \alpha)|S| & \text{if } |S| \leq \gamma|V| \\ \beta|S| & \text{if } \gamma|V| < |S| \leq |V|/2 \end{cases}$$

In particular random regular graphs and the (explicitly constructible) Ramanujan graphs of Lubotsky, Phillips and Sarnak [8] are (α, β, γ) -expanders. (See discussion in [4].)

A random walk on an undirected graph $G = (V, E)$ is a Markov chain $\{X_t\} \subseteq V$ associated with a particle that moves from vertex to vertex according to the following rule: the probability of a transition from vertex i , of degree d_i , to vertex j is $1/d_i$ if $\{i, j\} \in E$, and 0 otherwise. (In case of a bi-partite graph we need to assume that we do nothing with probability $1/2$ and move off with probability $1/2$ only. This technicality is ignored for the remainder of the paper.) Its stationary distribution, denoted π , (or π_G) is given by $\pi(v) = d_v/(2|E|)$. Obviously, for regular graphs, the stationary distribution is uniform.

A trajectory W of length τ is a sequence of vertices $\{w_0, w_1, \dots, w_\tau\}$ such that $\{w_t, w_{t+1}\} \in E$. The Markov chain $\{X_t\}$ induces a probability distribution on trajectories, namely the product of the probabilities of the transitions that define the trajectory.

Let P denote the transition probability matrix of the random walk on G , and let $P_{v,w}^{(t)}$ denote the probability that the walk is at w at step t given that it started at v . Let λ be the second largest eigenvalue of P . (All eigenvalues of P are real.) It is known that

$$P_{v,w}^{(t)} = \pi(w) + O(\lambda^t \sqrt{\pi(w)/\pi(v)}). \quad (1)$$

In particular, for regular graphs

$$P_{v,w}^{(t)} = \frac{1}{n} + O(\lambda^t). \quad (2)$$

To ensure rapid convergence we need $\lambda \leq 1 - \epsilon$ for some constant $\epsilon > 0$. This holds for all expanders (Alon [1]). In particular if G is an r -regular β -expander then Sinclair and Jerrum [12] show that

$$\lambda \leq 1 - \frac{1}{2} \left(\frac{\beta}{r} \right)^2$$

It is often useful to consider the separation s of the distribution $P_{v,w}^{(t)}$ from the limit distribution π given by

$$s(t) = \max_{v,w} \frac{\pi(w) - P_{v,w}^{(t)}}{\pi(w)}.$$

Then we can write

$$P_{v,w}^{(t)} = (1 - s(t))\pi + s(t)\sigma$$

where σ is a probability distribution. We can then imagine that the distribution $P_{v,w}^{(t)}$ is producing by choosing either σ with probability $s(t)$ or π with probability $1 - s(t)$. Hence if \mathcal{E} is an event that depends only on the state of the Markov chain we have

$$(1 - s(t)) \Pr(\mathcal{E} \text{ under } \pi) + s(t) \geq \Pr(\mathcal{E} \text{ under } P_{v,w}^{(t)}) \quad (3)$$

and

$$\Pr(\mathcal{E} \text{ under } P_{v,w}^{(t)}) \geq (1 - s(t)) \Pr(\mathcal{E} \text{ under } \pi). \quad (4)$$

We use this in the following scenario:

Experiment A: Choose $u_1 \in V$ uniformly at random and do a random walk W_1 of length τ from u_1 . Let v_1 be the terminal vertex of W_1 .

Experiment B: Choose u_2 and v_2 uniformly and independently from V and do a random walk of length τ from u_2 to v_2 .

Here $\tau = c_0 \log n$ and $s(\tau) \leq N^{-3}K^{-2}$. Since G is regular, u_1 and v_1 and u_2 and v_2 have each the same (uniform) distribution in the two experiments. However v_1 depends on u_1 and therefore the distributions of W_1 and W_2 differ slightly. What we claim though is that for any event \mathcal{E} depending on walks of length τ ,

$$|\Pr((u_1, v_1, W_1) \in \mathcal{E}) - \Pr((u_2, v_2, W_2) \in \mathcal{E})| \leq s(\tau).$$

This follows from the stronger claim that for any $u \in V$ and any event \mathcal{E} depending on walks of length τ

$$\begin{aligned} &|\Pr((u_1, v_1, W_1) \in \mathcal{E} \mid u_1 = u) \\ &\quad - \Pr((u_2, v_2, W_2) \in \mathcal{E} \mid u_2 = u)| \leq s(\tau), \end{aligned}$$

which follows from (3) and (4).

The notation $B(m, p)$ stands for the binomial random variable with parameters $m =$ number of trials, and $p =$ probability of success.

3 Static routing with bounded congestion

In this section we present an algorithm for static routing with bounded congestion. We first use a flow algorithm to randomize the endpoints. We then connect each pair of (new) endpoints by a random path. At this point, most of the edges have a limited congestion but some edges are overloaded. We then remove all paths that use overloaded edges. With high probability the number of disconnected pairs is sufficiently small that we can use an algorithm for finding edge-disjoint paths [4] to reconnect them. More formally our algorithm is:

Algorithm

Input: An r -regular β -expander $G = (V, E)$. A collection of $K = \alpha(n)n/\ln n$ pairs of vertices denoted $\{(a_1, b_1), \dots, (a_K, b_K)\}$ such that no vertex in V participates in more than s pairs.

Output: A set of K paths, $\{P_1, \dots, P_K\}$ such that P_i connects a_i to b_i and the maximum congestion g on any edge is bounded by

$$g = \begin{cases} O\left(s + \left\lceil \frac{\log \log n}{\log(1/\hat{\alpha})} \right\rceil\right), & \text{for } \alpha < 1/2; \\ O(s + \alpha + \log \log n), & \text{for } \alpha \geq 1/2. \end{cases}$$

where $\hat{\alpha} = \min(\alpha, 1/\log \log n)$.

Phase 1. Choose independently (with replacement) uniformly at random, two multisets R_A and R_B of $2K$ vertices each in V .

Phase 2. Select multisets $Q_A \subset R_A$ and $Q_B \subset R_B$ of K vertices, such that every element in $Q_A \cup Q_B$ has multiplicity at most $\max(4K/n, 1)$. If such sets cannot be found, then **stop**. The algorithm has failed.

Phase 3. Let $S_A = \{a_1, \dots, a_K\}$ and $S_B = \{b_1, \dots, b_K\}$. Using a flow algorithm in G twice, connect in an arbitrary manner the vertices of S_A (resp. S_B for the second flow) to the vertices of Q_A (resp. Q_B) by K paths as follows:

- Assume that every edge in G has a capacity equal to $\max(s, 4K/n, 1)/\beta$.
- View each vertex in S_A (resp. S_B) as a source with capacity equal to its multiplicity in S_A (resp. S_B) and similarly every vertex in Q_A (resp. Q_B) as a sink with capacity equal to its multiplicity in Q_A (resp. Q_B).

The expansion properties of G ensure that such flows always exists.

Phase 4. Let \tilde{a}_i (resp. \tilde{b}_i) denote the vertex in Q_A (resp. Q_B) that was connected to the original end-point a_i (resp. b_i). Choose x_1, x_2, \dots, x_K uniformly at random in V and then choose trajectories W_i (resp. W'_i) of length $\tau = c_0 \ln n$ that go from \tilde{a}_i to x_i (resp. \tilde{b}_i to x_i) according to the distribution on trajectories, conditioned on $w_{i,0} = \tilde{a}_i$ and $w_{i,\tau} = x_i$. (The constant c_0 is discussed in the analysis.)

Let $\nu(e)$ be the number of trajectories that use the edge e . Let $g_{\max} = g_{\max}(\alpha)$ be a parameter defined in the analysis below. For every edge e with $\nu(e) > g_{\max}$ delete all the trajectories using it. For all i such that both W_i and W'_i survived, connect \tilde{a}_i to \tilde{b}_i using W_i

followed by W'_i with loops removed. (This will lead to congestion at most g_{\max} .)

If the numbers of disconnected pairs is “too large” (see analysis), then **stop**. The algorithm has failed. Otherwise reconnect them using the algorithm in [4] on G .

The final path from a_i to b_i is the union of the paths from a_i to \tilde{a}_i , and from b_i to \tilde{b}_i found in Phase 3, and the path from \tilde{a}_i to \tilde{b}_i selected here, with loops removed.
End Algorithm

We will show that the algorithm above succeeds with probability greater than $1/3$ for any given input. Thus by repeating it $O(\log n)$ times we prove

Theorem 1 *There is an explicit polynomial time algorithm that with high probability can connect any set of $K = \alpha(n)n/\ln n$ pairs of vertices on a bounded degree expander so that no edge is used by more than g paths where*

$$g = \begin{cases} O\left(s + \left\lceil \frac{\log \log n}{\log(1/\hat{\alpha})} \right\rceil\right), & \text{for } \alpha < 1/2; \\ O(s + \alpha + \log \log n), & \text{for } \alpha \geq 1/2, \end{cases}$$

$\hat{\alpha} = \min(\alpha, 1/\log \log n)$, and s is the maximal multiplicity of a vertex in the set of pairs.

Proof: We leave the discussion of the flow phase to the full paper.

Assume for a moment that we start random walks of length $\tau = c_0 \ln n$ from every element of R_A , without any conditioning on their other endpoint. Let $\nu'(e)$ be the number of walks that use the edge e in this case. Then $\nu'(e) \leq B(2K, 2\tau/(rn))$ in distribution. (Each walk starts at an independently chosen vertex and moves to an independently chosen destination. The steady state of a random walk on G is uniform and so at each stage of a walk, each edge of G is equally likely to be crossed.) Thus for any $t > 0$

$$\begin{aligned} \Pr(\nu'(e) = t) &\leq \binom{2K}{t} \left(\frac{2\tau}{rn}\right)^t \\ &\leq \left(\frac{4Ke\tau}{trn}\right)^t = \left(\frac{4e\alpha(n)c_0}{tr}\right)^t. \end{aligned}$$

Now let $\nu_1(e)$ (resp. $\nu_2(e)$) be the number of trajectories W_i (resp. W'_i) in phase 4 that use e . We choose c_0 such that the separation between the distribution of the endpoint of a walk of length τ and the uniform distribution is less than (say) $1/(n^3 K^2)$. Then in view of the paragraph above, we have

$$\Pr(\nu_1(e) = t) \leq \left(\frac{4e\alpha(n)c_0}{tr}\right)^t + \frac{1}{n^3 K^2}, \quad (5)$$

where the error term comes from the fact that the endpoint of each trajectory is chosen uniformly at random rather than according to the distribution of the endpoint of the walk. Note also that we have to consider ν_1 and ν_2 separately because we want the endpoints to be chosen independently.

The results of [4] imply that given any n vertex, bounded degree, regular β -expander G and given any set of $q \leq n/(\ln n)^\kappa$ disjoint pairs of vertices in G , it is possible to find with high probability a set of paths in G connecting the q pairs, such that each edge in G participates in no more than $1 + 1/\beta$ paths. The parameter κ depends only on the expansion properties of G . Fix it to be the value corresponding to the input graph here. We now consider two cases:

Case 1: $\alpha \leq 1/2$

Define

$$t_0 = \left\lceil 4e^2 c_0 + \frac{3\kappa \ln \ln n}{\ln(1/\hat{\alpha})} \right\rceil$$

$$g_{\max} = 2t_0$$

Observe that

$$\sum_{t=t_0}^{\infty} t \left(\frac{4e\alpha c_0}{tr} \right)^t \leq 2t_0 \left(\frac{4e\alpha c_0}{t_0 r} \right)^{t_0} \leq \left(\frac{4e^2 \alpha c_0}{t_0 r} \right)^{t_0} \quad (6)$$

$$\leq \frac{1}{4r} \left(\frac{\alpha}{t_0} \right)^{t_0/2} \leq \frac{1}{4r(\ln n)^\kappa}.$$

Suppose that we delete all those trajectories W_i or W'_i which use an edge with $\max\{\nu_1(e), \nu_2(e)\} \geq t_0$. The number of pairs of vertices which become disconnected is at most

$$\Delta = \sum_{t_0 \leq t \leq K} 2t |\{e : \nu_1(e) = t\}| + \sum_{t_0 \leq t \leq K} 2t |\{e : \nu_2(e) = t\}|.$$

It follows from (5) and (6) (there are $rn/2$ edges) that $E(\Delta) \leq n/(2(\ln n)^\kappa) + 1/n$ and so with probability almost $1/2$ we find that $\Delta \leq n/(\ln n)^\kappa$. We now re-link these pairs using the algorithm of [4] at an additional congestion cost of $1 + 1/\beta$. Thus the total congestion on an edge is at most

$$\frac{\max(s, 4K/n, 1)}{\beta} + 2t_0 + \left(1 + \frac{1}{\beta}\right) = O\left(s + \frac{\log \log n}{\log(1/\hat{\alpha})}\right).$$

Case 2: $\alpha \geq 1/2$. This time we take

$$t_0 = \lceil 4\alpha e^2 c_0 + 3\kappa \ln \ln n \rceil = O(\alpha + \log \log n),$$

$$g_{\max} = 2t_0$$

and proceed as before. The congestion now is $O(s + \alpha + \log \log n)$. \square

4 Dynamic Selection of Paths

We define a stochastic model for studying a dynamic version of the circuit switching problem. In our model new requests for establishing paths arrive continuously at nodes according to a discrete Poisson process. Requests wait in the processor's queue until the requested path is established. The duration of a path is exponentially distributed.

Our model is characterized by three parameters:

- P_1 is an upper bound on the probability that a new request arrives at a given node at a given step.
- P_2 is the probability that a given existing path is terminated in a given step. A path *lives* from the time it is established until it is terminated.
- g is the maximum congestion allowed on any edge.

We assume that the destinations of path requests are chosen uniformly at random among all the graph vertices.

Our goal is to characterize the relationship between these parameters that ensures stability of the system. (By stability, we mean that the expected length of any queue does not grow unboundedly in time). We also estimate the expected delay incurred by a request in the steady state distribution.

We study a simple and fully distributed algorithm for this problem. In our algorithm each processor at each step becomes active with probability $P'_1 > P_1$ (P'_1 is defined later). An inactive processor does not try to establish a path even if there are requests in its queue. One effect of this is to moderate the injection rate for vertices with large queues. Also, just for the analysis, we imagine that if a process is active but its queue is empty, then it tries to construct a *ghost path* to a random vertex. The lifetime of a ghost path is distributed as that of real paths.

Our algorithm is simply this: Assume that a is active at step t , and the first request in a 's queue is for b . Processor a tries to establish a path to b by choosing a random trajectory of length $\tau = c_0 \ln n$ connecting a to b . If the path does not use any edge with congestion greater than $g - 1$, the path is established, otherwise the request stays in the queue.

Theorem 4 *Let*

$$\Phi = \min \left\{ \frac{1}{\ln(g\tau n)}, \frac{\tau g}{\tau^{(g+1)/g}} \right\}.$$

There exists a constant γ such that if $P_1 \leq \gamma\Phi P_2$, then the system is stable and the expected wait of a request in the queue is $O(1/P_1)$.

Before giving the proof let us see the consequence of this theorem. Let $\mathbf{E}(N) = nP_1$ be the expected number of new requests that arrive at the system at a given step, and let $\mathbf{E}(D) = 1/P_2$ be the expected duration of a connection. For the system to be stable, the expected number of simultaneously active paths in the steady state must be at least $\mathbf{E}(N)\mathbf{E}(D) = nP_1/P_2$. Plugging $g = \ln \ln n / \ln \omega$ for some ω in the range $[1, \ln n]$ in the definition of Φ we get

$$\Phi = \Omega\left(\frac{1}{\omega \log n}\right).$$

Thus the theorem above implies that for such a congestion g , the system remains stable even if we choose P_1 and P_2 such that

$$\mathbf{E}(N)\mathbf{E}(D) = n\frac{P_1}{P_2} = \gamma n\Phi = \Omega\left(\frac{n}{\omega \ln n}\right),$$

in which case the dynamic algorithm utilizes the edges of the network almost as efficiently as the static algorithm analyzed in section 2 (there seems to be an efficiency gap of maximum order $\ln \ln \ln n$ for $\omega \leq \ln \ln n$).

Proof of the theorem: Partition time into intervals of length T (to be determined). Let H_t denote the history of the system during the first t time intervals. Define the event

$$\mathcal{E}(v, t) = \begin{cases} \text{If the queue of processor } v \text{ was not} \\ \text{empty at the beginning of interval } t \\ \text{then } v \text{ served at least one request dur-} \\ \text{ing interval } t \end{cases}$$

Our goal is to show that for all v and t ,

$$\Pr(\mathcal{E}(v, t) \mid H_{t-2}) > \frac{1}{2}.$$

To this end consider the following four events:

• \mathcal{E}_1 defined as

$$\mathcal{E}_1(v, t) = \begin{cases} \text{Processor } v \text{ was not active in any step} \\ \text{of interval } t. \end{cases}$$

Then

$$\Pr(\mathcal{E}_1(v, t) \mid H_{t-2}) \leq (1 - P_1')^T \leq e^{-TP_1'} \leq \frac{1}{10}, \quad (7)$$

provided that

$$TP_1' \geq \ln 10. \quad (8)$$

• \mathcal{E}_2 defined as

$$\mathcal{E}_2(t) = \begin{cases} \text{At the beginning of interval } t \text{ there is} \\ \text{a path in the network that is still alive} \\ \text{although it was established before the} \\ \text{beginning of interval } t - 1. \end{cases}$$

Clearly, at the start of time interval $t - 1$ there are at most grn live paths. So

$$\Pr(\mathcal{E}_2(t) \mid H_{t-2}) \leq grn(1 - P_2)^T \leq grne^{-P_2T} \leq \frac{1}{10}, \quad (9)$$

provided that

$$TP_2 \geq \ln(10grn). \quad (10)$$

• \mathcal{E}_3 defined as

$$\mathcal{E}_3(v, t) = \begin{cases} \text{There are more than } 2nP_1'/P_2 \text{ live} \\ \text{paths that were created within inter-} \\ \text{vals } t - 1 \text{ and } t \text{ when } v \text{ makes its first} \\ \text{attempt in interval } t \end{cases}$$

To evaluate the probability of this event, we overestimate the number of paths in the network when v makes its first attempt at establishing a path in time interval t . We include in the count ghost paths, and paths or ghost path attempts which could not be established because of congestion. The life of such failed paths will also be geometric with parameter P_2 . In this count we exclude paths attempts from before the start of period $t - 1$. Suppose that period $t - 2$ finished j_0 time steps ago. Then our estimate is

$$X = \sum_{w=1}^n \sum_{j=0}^{j_0} X_{w,j}$$

where $X_{w,j}$ is the 0/1 indicator variable for the event

$$\begin{cases} j \text{ time steps ago, a path attempt was} \\ \text{made from } w \text{ and this path is still} \\ \text{alive.} \end{cases}$$

Then $\Pr(X_{w,j} = 1) = P_1'(1 - P_2)^j$ and so

$$\mathbf{E}(X \mid H_{t-2}) = \sum_{w=1}^n \sum_{j=0}^{j_0} P_1'(1 - P_2)^j \leq n\frac{P_1'}{P_2}.$$

An easy argument shows the concentration of X and therefore

$$\Pr(\mathcal{E}_3(v, t) \mid H_{t-2}) = o(1). \quad (11)$$

• \mathcal{E}_4 defined as

$$\mathcal{E}_4(v, t) = \begin{cases} \text{The first path that processor } v \text{ tries in} \\ \text{interval } t \text{ includes an edge used by at} \\ \text{least } g \text{ other paths that were created} \\ \text{within intervals } t - 1 \text{ and } t. \end{cases}$$

We break the ensuing analysis into 2 cases:

Case 1: $2\tau \left(\frac{4e\tau}{rgP_2}\right)^g > 1/10$.

Then

$$\Pr(\mathcal{E}_4(v, t) \mid H_{t-2}, X \leq 2nP_1'/P_2) \quad (12)$$

$$\begin{aligned} &\leq \tau \sum_{h \geq g} \binom{2n \frac{P'_1}{P_2}}{h} \left(\frac{2\tau}{rn} + \frac{1}{n^3 K^2} \right)^h \\ &\leq 2\tau \left(\frac{4e\tau P'_1}{rgP_2} \right)^g \leq \frac{1}{10}, \end{aligned}$$

provided that

$$P'_1 \leq \frac{rgP_2}{4e\tau} \left(\frac{1}{20\tau} \right)^{1/g} \quad (13)$$

In the above calculation we have implicitly used the fact that the X path attempts are a collection of X random walks between randomly chosen pairs. We include unsuccessful path attempts in order to avoid conditioning problems caused by paths blocking paths. Since

$$\bar{\mathcal{E}}(v, t) \subseteq \mathcal{E}_1(v, t) \cup \mathcal{E}_2(t) \cup \mathcal{E}_3(v, t) \cup \mathcal{E}_4(v, t)$$

we see that

$$\Pr(\bar{\mathcal{E}}(v, t) \mid H_{t-2}) \leq \frac{3}{10} + o(1),$$

and we conclude that that in any segment of $2T$ steps processor v is serving at least one request with probability at least $1/2$. The number of new arrivals in this time interval has a Binomial distribution with expectation at most $2TP_1 < 1/2$, provided that

$$TP_1 \leq \frac{1}{4}. \quad (14)$$

By standard analysis the queue under these conditions is stable, and the expected wait in the queue is $O(1/P_1)$.

It remains to check that if γ is small then we can choose T and $P_1 < P'_1 < 1$ to satisfy (8), (10), (13), (14). We take

$$\begin{aligned} T &= \frac{1}{4\gamma P_2 \Phi} \\ P'_1 &= \min \left\{ \frac{rgP_2}{1000\tau^{(g+1)/g}}, 1 \right\}. \end{aligned}$$

This completes the proof for Case 1.

Case 2: $2\tau \left(\frac{4e\tau}{rgP_2} \right)^g \leq 1/10$.

Here we take T as above and $P'_1 = 1$. Here (12) holds by the case hypothesis. \square

5 Existential results regarding edge-disjoint paths

In this section we show how to use the Lovász Local Lemma [5] to prove the *existence* of edge disjoint paths in any r -regular (α, β, γ) -expanders. At the moment we do not see how to make the argument constructive. More precisely we have

Theorem 2 *Given a bounded degree (α, β, γ) -expander graph there exists a parameter c that depends on α, β, γ , but not on n , such that any set of less than $cn/(\ln n)^2$ disjoint pairs of vertices can be connected by edge disjoint paths.*

Proof: Let

$$K = \left\lfloor \frac{rn}{577\tau^2} \right\rfloor,$$

where

$$\tau = \lceil c_2 \ln n \rceil$$

and c_2 is a constant discussed below.

Let $(a_1, b_1), (a_2, b_2), \dots, (a_K, b_K)$ be any set of K disjoint pairs of vertices. We claim that G contains edge disjoint paths joining a_i to b_i for $i = 1, 2, \dots, K$.

Our proof follows the blueprint used in [4]. We start by splitting the original graph $G = (V, E)$ into two disjoint β' -expanders $G_R = (V, E_R)$ and $G_B = (V, E_B)$ exactly as was done in [4]. The salient facts here are: (a) $\beta' > 1$; (b) the construction succeeds with probability $1 - o(1)$, thus such a split always exists; and (c) the maximum degree in G_B is at most $3\tau/4$ and the minimum degree is at least $\tau/4$.

The disjoint paths are constructed in two stages. In the first stage we choose a random set

$$Z = \{\zeta_1, \zeta_2, \dots, \zeta_{2K}\}$$

of $2K$ distinct vertices. We connect the original endpoints to the vertices of Z in an arbitrary fashion via edge-disjoint paths in G_R , such that each Z -vertex is the endpoint of exactly one path. A simple flow argument proves constructively the existence of such edge-disjoint paths on any graph with edge expansion larger than one.

Let \tilde{a}_i (resp. \tilde{b}_i) denote the vertex in Z that was connected to the original end-point a_i (resp. b_i) in the first stage. The core of the proof is to show via the Lovász Local Lemma the existence of edge-disjoint paths in G_B connecting \tilde{a}_i to \tilde{b}_i , for $i = 1, \dots, K$. To this end we choose for each i a random path of length 2τ from \tilde{a}_i to \tilde{b}_i . However the direct application of the Lovász Local Lemma is precluded by the fact that we do not have any control over how the pairing $(\tilde{a}_i, \tilde{b}_i)$ was done, and thus, although the probability of paths sharing an edge is small, the dependency graph is complete. To avoid this calamity, further randomization is necessary, as follows: assume that we rename the vertices of Z at random. Then the pairing induced by the flow becomes a random pairing.

More formally, let ρ be a random permutation of $[2K]$. Let

$$z_i = \zeta_{\rho(i)}, \quad \text{for } 1 \leq i \leq 2K.$$

The flow algorithm defines a pairing $f : Z \rightarrow Z$, that is, a function f such that for all $\zeta \in Z$ we have $f(\zeta) \neq \zeta$ and $f^2(\zeta) = \zeta$. This pairing is defined by $f(\tilde{a}_i) = \tilde{b}_i$;

and $f(\tilde{b}_i) = \tilde{a}_i$. In turn, it induces a pairing ϕ on $[2K]$ via

$$f(\zeta_i) = \zeta_j \leftrightarrow f(z_{\rho^{-1}(i)}) = z_{\rho^{-1}(j)} \leftrightarrow \phi(\rho^{-1}(i)) = \rho^{-1}(j).$$

It is easy to verify that after f is fixed, if ρ is chosen uniformly at random among the permutations of $[2K]$ then ϕ is uniform over the set Φ of all possible pairings of $[2K]$.

Now suppose that for $1 \leq i \leq K$ we choose x_i with distribution

$$\Pr(x_i = v) = \pi_B(v) = \frac{d_B(v)}{2|E_B|}$$

(the steady state distribution of a random walk on G_B) and then choose W'_i and W''_i randomly from all trajectories of length τ which go from \tilde{a}_i to x_i and \tilde{b}_i to x_i respectively. The distribution used for choosing W'_i is that of a random walk of length τ starting at \tilde{a}_i conditional on ending at x_i . Let W_j denote the walk which starts at z_j for $1 \leq j \leq 2K$ and define the event

$$\mathcal{E}_{i,j} = \{W_i \cap W_j \neq \emptyset \text{ and } \phi(i) \neq j\}.$$

Our proof reduces to showing that the event

$$\mathcal{E} = \bigcap_{i < j} \bar{\mathcal{E}}_{i,j}$$

has positive probability.

Now define $\phi_0 \in \Phi$ to be the pairing $\{(1, K+1), (2, K+2), \dots, (K, 2K)\}$. Let \Pr_0 denote probabilities conditional on $\phi = \phi_0$. We will prove that

$$\Pr_0(\mathcal{E}) > 0. \quad (15)$$

Since $\Pr(\phi = \phi_0) > 0$ this suffices to complete the proof. Now if $\phi = \phi_0$ then

$$\mathcal{E} = \bigcap_{|j-i| \neq K} \bar{\mathcal{A}}_{i,j}$$

where

$$\mathcal{A}_{i,j} = \{W_i \cap W_j \neq \emptyset\}.$$

We will apply the local lemma to the events $\mathcal{A}_{i,j}$, conditional on $\phi = \phi_0$.

For each pair (i, j) with $|j - i| \neq K$ we let

$$S_{i,j} = \{\mathcal{A}_{i',j'} : i', j' \notin \{i, j, i+K, j+K\}\}.$$

Fix i, j and let \mathcal{B} denote any event dependent only on the outcome of events in $S_{i,j}$. Then

$$\Pr_0(\mathcal{A}_{i,j} | \mathcal{B}) \quad (16)$$

$$= \sum_{x,y \in V} \Pr_0(W_i \cap W_j \neq \emptyset, z_i = x, z_j = y | \mathcal{B})$$

$$= \sum_{x,y \in V} \Pr_0(W_i \cap W_j \neq \emptyset, z_i = x, z_j = y) \quad (17)$$

$$= \Pr_0(\mathcal{A}_{i,j}). \quad (18)$$

Equation (17) follows from the fact that if we fix z_i and z_j then the occurrence of $\mathcal{A}_{i,j}$ does not depend on the paths $\{W_k : k \notin I_{i,j}\}$. Thus conditional on $\phi = \phi_0$, $\mathcal{A}_{i,j}$ is independent of the events in $S_{i,j}$. Thus the dependency graph has maximum degree d where

$$d \leq 4K. \quad (19)$$

This justifies the complexity of the previous analysis. If no care is taken, the dependence graph will be complete.

Still keeping i, j fixed, let $\Phi_{i,j} = \{\phi \in \Phi : \phi(i) \neq j\}$ and choose an arbitrary $\phi' \in \Phi_{i,j}$. Now

$$\begin{aligned} \Pr_0(\mathcal{A}_{i,j}) &= \sum_{x,y \in V} \Pr_0(\mathcal{A}_{i,j}, z_i = x, z_j = y) \\ &= \sum_{x,y \in V} \Pr(\mathcal{A}_{i,j} | z_i = x, z_j = y, \phi = \phi_0) \\ &\quad \times \Pr(z_i = x, z_j = y | \phi = \phi_0) \\ &= \sum_{x,y \in V} \Pr(\mathcal{A}_{i,j} | z_i = x, z_j = y, \phi = \phi') \\ &\quad \times \Pr(z_i = x, z_j = y | \phi = \phi') \quad (20) \\ &= \Pr(\mathcal{A}_{i,j} | \phi = \phi'). \quad (21) \end{aligned}$$

To justify (20) observe that

- $\Pr(\mathcal{A}_{i,j} | z_i = x, z_j = y, \phi = \phi_0) = \Pr(\mathcal{A}_{i,j} | z_i = x, z_j = y, \phi = \phi')$ since given $z_i = x$ and $z_j = y$ as long as $\phi(i) \neq j$, that is z_i is not paired with z_j , we can decide $\mathcal{A}_{i,j}$ without further reference to ϕ .
- $\Pr(z_i = x, z_j = y | \phi = \phi_0) = \Pr(z_i = x, z_j = y | \phi = \phi')$ since, conditioning on Z and on the pairing induced by the flow phase on Z , if $\{x, y\} \not\subseteq Z$ or x is paired with y , then both sides are 0; otherwise, the LHS is proportional to the number of permutations that induce ϕ_0 and make $z_i = x$ and $z_j = y$. This is clearly the same as the number of permutations that induce ϕ' and make $z_i = x$ and $z_j = y$.

It follows from (21) that

$$\begin{aligned} \Pr_0(\mathcal{A}_{i,j}) &= \frac{1}{|\Phi_{i,j}|} \sum_{\phi' \in \Phi_{i,j}} \Pr(\mathcal{A}_{i,j} | \phi = \phi') \\ &= \frac{|\Phi|}{|\Phi_{i,j}|} \sum_{\phi' \in \Phi_{i,j}} \Pr(\mathcal{A}_{i,j}, \phi = \phi') \\ &= \Pr(\mathcal{A}_{i,j} | \phi(i) \neq j) \end{aligned}$$

and thus

$$\begin{aligned} \Pr_0(\mathcal{A}_{i,j}) &\leq \frac{1}{\Pr(\phi(i) \neq j)} \sum_{x,y \in V} \Pr(\mathcal{A}_{i,j} | z_i = x, z_j = y) \\ &\quad \times \Pr(z_i = x, z_j = y) \quad (22) \end{aligned}$$

Since ϕ is a random pairing,

$$\Pr(\phi(i) \neq j) = \frac{2K-2}{2K-1}. \quad (23)$$

The sum in (22) is the probability of the following event \mathcal{M} : Choose x, y uniformly at random (without replacement). Choose x', y' with probability π_B . Do random walks W, W' of length τ from x to x' and y to y' . The event is $\{W \cap W' \neq \emptyset\}$.

We now prove that

$$\Pr(\mathcal{M}) \leq \frac{(36 + o(1))\tau^2}{rn}. \quad (24)$$

Let $P_{v,\cdot}^{(\tau)}$ be the distribution of a random walk on G_B starting from v after τ steps. Consider the event \mathcal{M}' which differs from \mathcal{M} only in that x and y are chosen independently with distribution π_B , and x' and y' are chosen with the distribution $P_{x,\cdot}^{(\tau)}$ and $P_{y,\cdot}^{(\tau)}$ respectively. (We can now have $x = y$, but this has probability $O(1/n)$ and we will deal with it later.) Let $W = e_1, e_2, \dots, e_\tau$ and $W' = f_1, f_2, \dots, f_\tau$ as edge sequences. Then for $1 \leq \ell, m \leq \tau$ we have

$$\Pr(f_\ell = e_m) = \frac{1}{|E(G_B)|} \leq \frac{4}{rn}.$$

since the edges of each random walk have the uniform distribution over $E(G_B)$ and the start points of these walks are chosen independently. Thus,

$$\Pr(\mathcal{M}') \leq \sum_{\ell=1}^{\tau} \sum_{m=1}^{\tau} \Pr(f_\ell = e_m) \leq \frac{4\tau^2}{rn}.$$

Let now \mathcal{M}'' be the same as \mathcal{M}' except that we choose x' and y' according to π_B . We take c_2 such that the separation $s(\tau)$ is less than $1/n$, and therefore

$$\Pr(\mathcal{M}'') \leq \Pr(\mathcal{M}') + \frac{1}{n}.$$

Finally, given the bounds on the degrees occurring in G_B , when choosing a pair x, y according to π_B we find that no pair occurs with probability more than 9 times any other pair. Thus

$$\Pr(\mathcal{M}) \leq 9 \Pr(\mathcal{M}'' \mid x \neq y) \leq 9 \Pr(\mathcal{M}'') / \Pr(x \neq y)$$

and (24) follows.

We see from (22), (23) and (24) that

$$\Pr_0(\mathcal{A}_{i,j}) \leq \frac{(36 + o(1))\tau^2}{rn}.$$

Using this and (19) in the local lemma yields the theorem. \square

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