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Temporal Multiple Kernel Learning (tMKL) model for predicting resting state FC via characterizing fMRI connectivity dynamics

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Abstract

Over the last decade there has been growing interest in understanding the brain activity in the absence of any task or stimulus captured by the resting-state functional magnetic resonance imaging (rsfMRI). These resting state patterns are not static, but exhibit complex spatio-temporal dynamics. In the recent years substantial effort has been put to characterize different FC configurations while brain states makes transitions over time. The dynamics governing this transitions and their relationship with stationary functional connectivity remains elusive. Over the last years a multitude of methods has been proposed to discover and characterize FC dynamics and one of the most accepted method is sliding window approach. Moreover, as these FC configurations are observed to be cyclically repeating in time there was further motivation to use of a generic clustering scheme to identify latent states of dynamics. We discover the underlying lower-dimensional manifold of the temporal structure which is further parameterized as a set of local density distributions, or latent transient states. We propose an innovative method that learns parameters specific to these latent states using a graph-theoretic model (temporal Multiple Kernel Learning,

tMKL) and finally predicts the grand average functional connectivity (FC) of the unseen subjects by leveraging a state transition Markov model. tMKL thus learns a mapping between the underlying anatomical network and the temporal structure. Training and testing were done using the rs-fMRI data of 46 healthy participants and the results establish the viability of the proposed solution. Parameters of the model are learned via state-specific optimization formulations and yet the model performs at par or better than state-of-the-art models for predicting the grand average FC. Moreover, the model shows sensitivity towards subject-specific anatomy. The proposed model performs significantly better than the established models of predicting resting state functional connectivity based on whole-brain dynamic mean-field model, single diffusion kernel model and another version of multiple kernel learning model. In summary, We provide a novel solution that does not make strong assumption about underlying data and is generally applicable to resting or task data to learn subject specific state transitions and successful characterization of SC-dFC-FC relationship through an unifying framework.

Keywords: rsfMRI, SC, dFC, FC, tMKL

1. Introduction

Since its discovery over two decades ago, there has been a keen interest in investigating the spontaneous intrinsic activity of the human brain. This activity is measured via slow fluctuations in the functional magnetic resonance images (fMRI) when subjects are at rest and not engaged in any task [1]. These fluctuations are highly correlated and discovery of meaningful large-scale functional networks within these correlations led to the use of resting-state fMRI (rsfMRI) to discover human brain function(s) [2, 3]. The resulting matrix of pairwise correlations between regions of interest (ROIs) is termed the functional connectivity (FC) matrix. Many studies of FC have discovered distinct sets of functionally related regions exhibiting temporal correlation in their activities, commonly known as resting state networks (RSNs) [4, 5, 6, 7].

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13 Diffusion tensor imaging (DTI), complementing fMRI, captures the white
14 matter streamlines that form the anatomical pathways along which neural acti-
15 vity spreads over the brain [8, 9]. The topography of the brain anatomy, called
16 the structural connectivity (SC), is estimated by counting the number of stream-
17 lines connecting a pair of ROIs. Over the last decade, understanding the link bet-
18 ween anatomical topology and neural activity has been an important question
19 in neuroscience. How the relatively static SC sculpts the FC over the entire scan
20 duration has been a challenging open research problem in the brain connectome
21 research domain. Initial studies provide evidence that the underlying structu-
22 ral topology largely explains the grand-average functional connectivity [10], the
23 missing link being dynamics. Whole brain computational models aid study and
24 simulation of the temporal dynamics over the structure.

25 Extant whole-brain models advancing our understanding of the SC-FC link
26 can be broadly categorized as follows : (i) models incorporating non-linear dy-
27 namics [11, 12], (ii) graph theoretic models [13, 14, 15], (iii) models at the
28 boundary of biophysics and graph-theoretic abstractions [16, 17, 18]. Becker et
29 al. [15] mapped spectral signatures of the structural and functional topologies
30 based on indirect structural walks of the neural activity. Abdelnour et al. [16]
31 proposed a graph-diffusion framework relating linear diffusion equation of the
32 neural activity over the structural topology to random walks of the activity over
33 the structure. Surampudi et al. [17] proposed abstraction of non-linear diffusion
34 equation into combinations of multi-scale diffusion to map a subject’s SC-FC.

35 Over the last decade, several studies of rsfMRI revealed fluctuating spa-
36 tial patterns which appear and dissolve with time, highlighting the spatiotem-
37 poral repertoire of spontaneous brain activity [19, 20]. Attempts at discover-
38 ing temporal dynamics of rsfMRI can be broadly categorized in the following
39 terms : (i) dynamic functional connectivity (dFC) studies using sliding window
40 approaches providing sequence of windowed FC (wFC) matrices that in turn
41 identify stable transient patterns of functional connectivity fluctuations, called
42 *latent states*, [21, 22, 23, 24], and (ii) Bayesian approaches applied on the time-
43 series themselves [25, 26, 27] which discover latent states in terms of multivariate
44 Gaussian density distributions of the temporal signals. A general perspective is
45 that the neural activity during a task, although being in a high-dimensional

46 space, follows trajectories in a lower-dimensional task-specific manifold during
47 the functional dynamics [28]. This sufficiently motivates the presence of a lower-
48 dimensional manifold for rsfMRI as well.

49 Moreover, the question of how a relatively fixed anatomical structure sup-
50 ports the rich spatiotemporal dynamics is still elusive. Abdelnour et al. [18]
51 have extended their graph-diffusion framework for characterizing SC-dFC rela-
52 tionships. However, theoretical models incorporating principled amalgamation
53 of structural topology and dynamics of rsfMRI are essential. Here, we propose
54 an innovative solution for characterizing the SC-dFC-FC relationship. This is
55 achieved by proposing two novel constructs : (i) discovery of a lower-dimensional
56 manifold that represents the latent structure of the temporal dynamics, (ii) tem-
57 poral multiple kernel learning (*tMKL*) model that learns the SC-dFC mapping,
58 and (iii) generation of latent time series for dFC-FC mapping. The proposed
59 solution estimates grand average FC (gFC or FC) from SC by predicting dFC
60 along with capturing the temporal evolution. Temporal evolution is characteri-
61 zed by using a first-order Markov model between consecutive state transitions.
62 This model is used for generating a long state sequence using the steady state
63 distribution of the Markov random walk. This state sequence is further replaced
64 by sequence of corresponding state-specific FCs generated by the *tMKL* model.
65 Finally, these state-specific FCs are factorized to recover a latent time-series
66 sequence. gFC is then computed on the reconstructed latent time-series and
67 compared with the empirical FC. The proposed model recovers the FCs that
68 are very close to empirical FCs as the state-specific FCs recovered with the
69 *tMKL* model enable realization of subject-specific functional dynamics. Fur-
70 ther, various perturbation experiments demonstrate the robustness and validity
71 of the proposed scheme. This state sequence is further replaced by sequence of
72 corresponding state-specific FCs generated by the *tMKL* model. Finally, these
73 state-specific FCs are factorized to recover a latent time-series sequence. gFC is
74 then computed on the reconstructed time series and compared with the empirical
75 FC. The proposed model recovers the gFCs that are very close to empirical FCs
76 as the state-specific FCs recovered with the *tMKL* model enable realization of
77 subject-specific functional dynamics. Further, various perturbation experiments
78 demonstrate the robustness and validity of the proposed scheme.

79 The specific contributions of the work are the following :

- 80 1. Novel approach for learning the SC-FC mapping through characterizing
81 the dynamic functional connectivity (dFC) over time windows.
- 82 2. Proposal of a novel multiple diffusion kernel model that learns to predict
83 state-specific FCs from SC (tMKL model).
- 84 3. Estimating the latent fMRI time series by using the Markov transition
85 probability matrix in conjunction with the tMKL model.

86 The rest of the paper is organized as follows. In the next section we present the
87 details of the proposed solution. In the subsequent sections we present the de-
88 tails of the neuroimaging data set used, qualitative and quantitative evaluation
89 results along with explanation for the choice of model parameters. Finally, we
90 conclude by pointing out limitations and future research directions.

91 **2. Materials and methods**

92 *2.1. Dataset*

93 Resting state fMRI as well as corresponding diffusion weighted imaging
94 (DWI) data were collected at the Berlin Center for Advanced Imaging, Cha-
95 rité University, Berlin, Germany. The dataset consisted of structural connecti-
96 vity - functional connectivity (SC-FC) pairs of total 46 subjects used in this
97 study. In summary, all the participants underwent resting state functional ima-
98 ging (no task condition) with eyes closed for 22 minutes, using a 3T Siemens
99 Trim Trio scanner and 12 channel siemens head coil (voxel size $3 \times 3 \times 3$ mm).
100 Each fMRI resting state data amount to a total of 661 whole brain scans (time
101 points recorded at TR=2s) were obtained during the resting state functional
102 magnetic resonance imaging (rs-fMRI) session. Thus the blood oxygen level
103 dependent (BOLD) time-series signal available for each participant has 661
104 time points aggregated across 68 regions of interest (ROIs) as per the Desikan-
105 Killiany brain atlas [29]. The diffusion weighted tensors (TR=750 ms, voxel size
106 $2.3 \times 2.3 \times 2.3$ mm) computed from the dwMRI data recorded with 64 gra-
107 dient directions were subjected to probabilistic tractography as implemented
108 in MRTrix [29] in order to obtain subject specific structural connectivity (SC)
109 matrices. Masks derived from high-resolution T1-images were used to determine

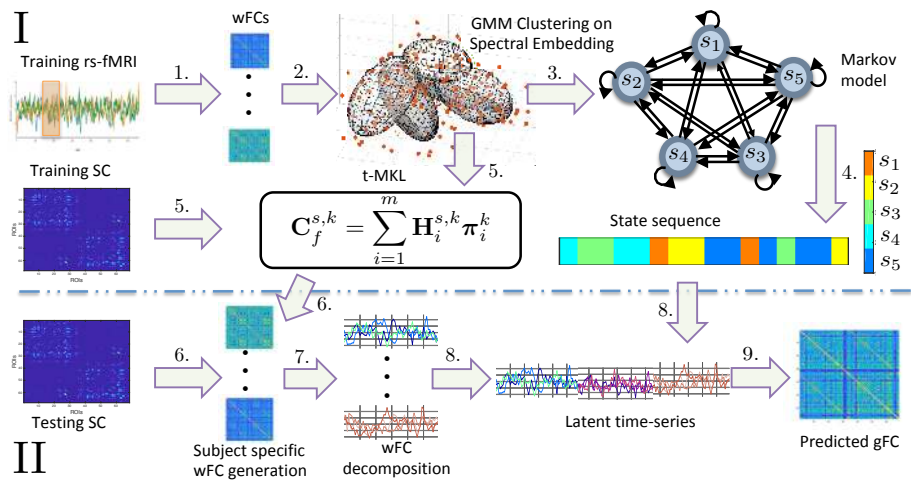


FIGURE 1: **Outline for temporal Multiple Kernel Learning (tMKL) model.** Figure shows the entire pipeline for predicting grand FC for a testing subject. The model incorporates subject specificity along with temporal variation characterization. Part (I) of the model, training phase, consists of three blocks. The first one, learns temporal variations in terms of distinct states via GMM clustering over the underlying manifold of wFCs (steps 1. and 2). The second block utilizes the empirical transitions between these distinct states and captures dynamics in the first order Markov chain (steps 3. and 4). The third block learns subject-specificity by modeling each state by its MKL model [17] (step 5.). Part (II) of the model validates its generalizability on unseen subjects. Importantly, only SC of a testing subject is required (step 6). Each state for the testing subject is characterized in step 7. Each state-specific predicted FC is decomposed into a latent time series which are then concatenated using the steady state distribution of the Markov chain (steps 4. and 8). Finally, grand average FC, static FC, is predicted for that subject (step 9).

110 seed-and target-locations for fibers in the grey/white matter-interface (GWI).
 111 SC matrices contains connection streamlines obtained based on the fiber tra-
 112 cking algorithm with various assumptions based on known limitations imposed
 113 by anatomy, notably the size of the GWI of each region. Further image acqui-
 114 sition, choice of scan parameter details and data pre-processing methodology
 115 adopted are all available in [30].

116 2.2. Proposed model

117 In this section we describe in detail the whole pipeline of the proposed
 118 parametric-model to map the relation between SC and FC using resting state
 119 fMRI data. The proposed model considers the importance of the underlying
 120 anatomical constraints to generate the temporal richness as well as to charac-

121 terize and assess whole-brain FC dynamics. Figure 1 shows a flowchart of the
122 essential elements of the whole pipeline. Proposed model partitions aspects of
123 the whole-brain dynamics essentially into two parts : characterizing temporal
124 dynamics through identification of latent transient states and linking them to
125 the underlying structural geometry. These two aspects are parameterized using
126 a novel combination of unique methods. The model utilizes wFCs (steps 1. – 2.)
127 for identifying states and from the resultant SC-wFC pairs, the relationship
128 between the structure and functional dynamics is learned (steps 3. – 5.). Once
129 these two parts successfully characterize the above mentioned aspects by tuning
130 respective parameters, the model is tested for its generalizability using unseen
131 test subjects (steps 6. – 9.).

132 For identifying latent states within the dynamics, we discover the underlying
133 globally non-linear manifold that spans all the wFCs (step 2.a), thus recovering
134 the lower-dimensional space for meaningful characterization. We employ a pro-
135 babilistic framework for estimating the number of states and the shape of each
136 state in the lower-dimensional space, ensuring soft assignments of wFCs to its
137 neighboring states (step 2.b). These soft assignments are further used to esti-
138 mate the transition dynamics between these states (step 3. – 4.). With respect
139 to second aspect of the model, we adapt the multiple kernel learning (MKL)
140 framework [17] for parameterizing the dependence of SC on wFCs for each state
141 (step 5.). We observe that the parameters to be learned form a non-convex com-
142 bination, necessitating an iterative algorithm. Thus we formulate the learning
143 objective into an optimization formulation and adapt an iterative algorithm for
144 solving this non-convex combinations of parameters.

145 The model predicts state-specific FCs (sFCs) for a test subject (step 6.).
146 These sFCs are decomposed into a latent time-series (step 7.) which is concate-
147 nated using the relative frequency of occurrence of states to generate a global
148 time-series for calculating the static FC of a subject (step 8.). Thus, for a new
149 subject, given the SC, static FC along with its state-specific FCs are predicted
150 by the proposed model (step 9.).

151 In the subsequent subsections we elaborate each part of the proposed model.
152 From now on, let $D = \{\mathbf{F}_w^1, \dots, \mathbf{F}_w^s, \dots, \mathbf{F}_w^p\}$ be the set of all wFC matrices
153 obtained by sliding a window of fixed size ω over the n -dimensional fMRI time-

154 series belonging to all the training subjects.

155 2.2.1. *Spectral Embedding, step 2.a*

156 We propose to soft-cluster these wFC matrices into K states, first by vec-
157 torizing the lower triangular part of a wFC matrix into a column vector of size
158 $\frac{n(n-1)}{2} \times 1$. These wFCs may be sparsely spaced in a higher-dimensional space,
159 but might originally lie on an intrinsic globally non-linear manifold [31]. Spec-
160 tral embedding method is employed to reduce the dimensionality of the data,
161 by finding a mapping to a lower dimensional manifold over which these wFCs
162 reside [32]. The graph constructed over the vectorized wFCs provides a discrete
163 approximation of the continuous manifold. The solution embedding is provided
164 from the eigenmaps (eigenvectors) of the Laplacian operator over the graph,
165 which approximates a natural mapping onto the entire manifold. The Laplacian
166 eigenmaps preserve the local structure in the graph, thus keeping the solution
167 embedding robust to outliers and noise.

168 The spectral embedding method is applied as follows. Firstly, an affinity ma-
169 trix is created by applying a radial basis function over the L1 distance between
170 every pair of wFCs. This matrix captures pairwise relationship between wFCs
171 in a relational graph. Next, we form the corresponding normalized graph Lapla-
172 cian matrix and use the eigenvectors corresponding to its lowest K eigenvalues
173 to define the basis vectors of embedding space [33, 34, 35]. The value of these
174 eigenvectors against each wFC represent its resulting transformation into the
175 embedding space. Finally these K -dimensional embedded wFCs are clustered
176 using Gaussian Mixture Model (GMM), as explained in the next subsection.

177 2.2.2. *GMM Clustering, step 2.b*

178 Following the discovery of an approximation to the continuous lower-dimensional
179 manifold, we now parameterize the local density distribution of wFCs over the
180 manifold using a probabilistic framework, Gaussian mixture model (GMM) [36].
181 Gaussian mixture model is a factor analysis model that represents the proba-
182 bility density of a sample as a weighted combination of component Gaussians.
183 Such a representation facilitates GMM to represent a large class of sample dis-
184 tributions. Specifically, distribution of wFCs over the manifold are modeled as
185 a GMM.

Let the density of \mathbf{F}_w^s be a linear combination of K component Gaussian densities, represented as follows :

$$\begin{aligned} P(\mathbf{F}_w^s; \Theta) &= \sum_{k=1}^K \Psi^k(s) \mathcal{N}(\mathbf{F}_w^s; \mu^k, \Sigma^k) \\ \sum_{k=1}^K \Psi^k(s) &= 1, \forall s = 1, \dots, p \end{aligned} \quad (1)$$

where P denotes the probability density of a wFC. Each component Gaussian is a K -variate Gaussian probability density function of the form :

$$\begin{aligned} \mathcal{N}(\mathbf{F}_w^s; \mu^k, \Sigma^k) \\ = \frac{1}{(2\pi)^{K/2} \det(\Sigma^k)^{1/2}} \exp \left\{ (\mathbf{F}_w^s - \mu^k)^\top \Sigma^{k-1} (\mathbf{F}_w^s - \mu^k) \right\}. \end{aligned}$$

GMM thus represents the manifold as a set of Gaussian densities and parameterizes it in terms of Θ :

$$\Theta = \{ \Psi^k(\cdot), \mu^k, \Sigma^k \}, k = 1, \dots, K. \quad (2)$$

186 As the collection of these component Gaussians forms the manifold, the com-
 187 ponent Gaussians can be interpreted as a *latent transient state* visited by the
 188 brain. Each state is a Gaussian but at different locations and with different
 189 shapes governed by μ^k and Σ^k , respectively in the manifold.

190 2.2.3. State Transition Markov Model, step 3.

191 As described in the previous section, the wFCs are quantized into finite
 192 states $S = \{s_1, \dots, s_K\}$ by GMM clustering. Each wFC sequence now corres-
 193 ponds to a cluster-label (state) sequence and transitions between these states
 194 is representative of the dynamics in the BOLD rsfMRI time series. We assume
 195 first-order dependence among these transitions and learn the Markov transition
 196 probability matrix, $\mathbf{T}_{K \times K}$ by estimating the state transitions from the training
 197 data.

198 Figure 2 shows a depiction of Markov model for $K = 5$ and the corresponding
 199 transition probability matrix. Each edge $t_{i,j}$ captures the probability of transi-
 200 tion from state i to state j . Similarly, self-loop edges $t_{i,i}$ depict the probability
 201 of remaining in the same state. For each state i we compute $t_{i,j}$ by counting
 202 the number of first-order transitions to state j in the state sequence. Finally, we

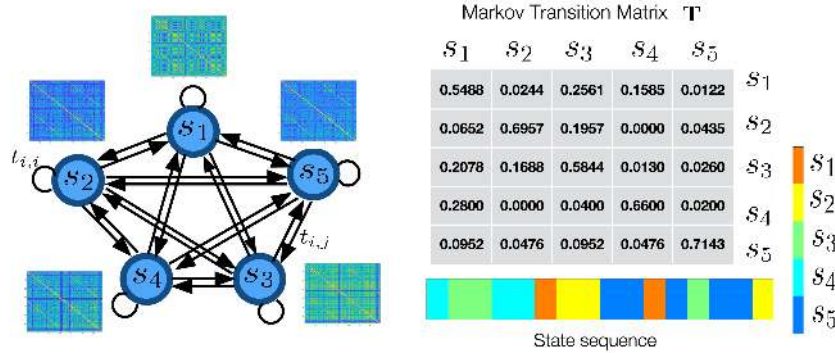


FIGURE 2: **Graphical depiction of proposed Markov state transition model.** An illustration of the first-order Markov chain used as a part of the proposed tMKL model. Each state has its unique distribution of FCs, represented as a Gaussian in the embedding space, from which subject-specific FCs can be sampled. The corresponding transition matrix (for $K=5$) and an example state sequence generated with a Markov random walk over the transition matrix is also depicted.

203 normalize each row of \mathbf{T} to make it a valid transition probability matrix. In the
 204 testing phase, the Markov matrix learned on training wFCs is used to generate a
 205 random state sequence, to eventually construct the latent time-series for testing
 206 subjects. As any Markov chain converges to its steady state distribution with
 207 time regardless of its initial distribution, we find the steady state distribution
 208 over the transition matrix and use this distribution as frequency of occurrence
 209 of states over the time course. This gives us a state transition sequence for a test
 210 subject. Along with the state transition model which captures the dynamics of
 211 the latent states, a model that relates anatomical structure to these states is
 212 required. In the next section, we propose a temporal multiple kernel learning
 213 (tMKL) model that learns this mapping.

214 2.2.4. tMKL Model, step 5.

215 Mean regional activities of all regions are assumed to be in a random walk
 216 over the SC graph. This phenomenon is represented by a linear differential equa-
 217 tion whose analytical solution is the diffusion kernel over the graph defined by
 218 SC which is hypothesized to be representing FC [16]. [14] discovered that physi-

219 cal diffusion over such large scale graphs exhibits multi-scale relationships with
 220 FC, thus a linear combination of multiple diffusion kernels is considered more re-
 221 presentative of FC (this model is referred to as MKL_NIPS from now on). The
 222 linear combination coefficients are scalar values which equally weigh all regional
 223 activities at each diffusion-scale. But it may so happen that activities of non-
 224 physically connected regions may be modulated by other regions. To represent
 225 this phenomenon we introduce the variables $\boldsymbol{\pi}_i$'s of size $n \times n$, that capture the
 226 inter-regional co-activation patterns at diffusion-scale i , $\forall i = 1, \dots, m$, m being
 227 the number of diffusion-scales [17].

Let a diffusion kernel defined at scale i be denoted by \mathbf{H}_i .

$$\mathbf{H}_i = e^{-\tau_i \mathbf{L}} \quad (3)$$

Here, τ_i is the spatio-temporal scale of heat diffusion and \mathbf{L} is the Laplacian matrix corresponding to the SC. We propose that a wFC matrix can be decomposed into a set of diffusion kernels multiplied with their co-activation pattern :

$$\mathbf{C}_f = \sum_{i=1}^m \mathbf{H}_i \boldsymbol{\pi}_i, \quad (4)$$

Here, \mathbf{C}_f denotes predicted wFC. We hypothesize that co-activation patterns are distinctly different for each state and hence we add a superscript index k ($k = 1 \dots K$) to obtain $\boldsymbol{\pi}_i^k$. As the parameters $\boldsymbol{\pi}_i^k$'s are state dependent, state-specific predicted functional connectivity, $\mathbf{C}_f^{s,k}$, will be as follows :

$$\mathbf{C}_f^{s,k} = \sum_{i=1}^m \mathbf{H}_i^{s,k} \boldsymbol{\pi}_i^k = \sum_{i=1}^m e^{-\tau_i^k \mathbf{L}^s} \boldsymbol{\pi}_i^k \quad (5)$$

Here \mathbf{L}^s is the Laplacian matrix of the SC corresponding to wFC^s . This results in the following optimization problem for $\boldsymbol{\Pi}^k$ and $\boldsymbol{\tau}^k$:

$$\begin{aligned} \underset{\boldsymbol{\Pi}^k, \boldsymbol{\tau}^k}{\text{minimize}} \quad & \sum_{s=1}^p \left\| \Psi^k(s) \left(\mathbf{F}_w^s - \mathbf{C}_f^{s,k} \right) \right\|_F^2 \\ & + \lambda_1 \sum_{i=1}^m \|\boldsymbol{\pi}_i^k\|_1 + \lambda_2 \sum_{i=1}^m \|\boldsymbol{\pi}_i^k\|_2 \\ \text{subject to} \quad & \mathbf{C}_f^{s,k} = \sum_{i=1}^m e^{-\tau_i^k \mathbf{L}^s} \boldsymbol{\pi}_i^k \\ & \boldsymbol{\pi}_i^k \in \mathcal{S}_+^n, i = 1, \dots, m, k = 1, \dots, K \\ & \boldsymbol{\tau}^k \succeq \mathbf{0}. \end{aligned} \quad (6)$$

228 Here, \mathcal{S}_+^n is the convex set of positive semi-definite matrices. The objective func-
229 tion takes the form well known in regression analysis as *least absolute shrinkage*
230 *and selection operator* (LASSO) that performs both variable selection and re-
231 gularization. We arrived at the model parameters experimentally, for example,
232 the number of scales m is empirically chosen (see Subsection 3.2).

233 Finally, the model consists of m distinct π_i^k 's which are learned for each of
234 the K states.

235 2.2.5. *Generation of latent time-series for testing subjects, steps 4., 6. –* 236 *9.*

237 As described in the previous section, we predict the state-specific FC matrix
238 for each of the states using the input SC matrix of the testing subject and the
239 learned tMKL model (step 6.). Based on the learned Markov chain state transi-
240 tion matrix, a sequence of states is generated using the steady state distribution
241 of the transition matrix (step 4.). Each of the state-specific FCs in the resulting
242 sequence is factorized into state-specific latent time-series and concatenate to
243 obtain the latent time-series for the testing subject.

In the training phase, wFCs are obtained by computing Pearson correlation coefficients of the windowed BOLD rsfMRI time-series over various regions. We know that Pearson correlation between two time-series A, B is $\rho(A, B) = \frac{\text{cov}(A, B)}{\sigma_A \sigma_B}$. Hence the state-specific wFC matrix works out to be the covariance of its state-specific latent times-series $\hat{Z}_{n \times \omega}$. Thus we can factorize a state-specific wFC as follows :

$$\begin{aligned} \mathbf{C}_f^{s,k} &= \mathbf{U}\mathbf{\Lambda}\mathbf{U}^\top \\ &= (\sqrt{\mathbf{\Lambda}}\mathbf{U}^\top)^\top (\sqrt{\mathbf{\Lambda}}\mathbf{U}^\top) \\ \hat{Z} &= \sqrt{\mathbf{\Lambda}}\mathbf{U}^\top. \end{aligned} \tag{7}$$

244 Thus, using Eq. 7, we recover latent time-series matrix \hat{Z} that can be taken as
245 approximated time-series used for obtaining wFC (step 7.). For a testing sub-
246 ject, each cluster-specific wFC is decomposed into latent time-series and these
247 are concatenated into a grand time-series (step 8.). The latent time series are
248 concatenated by considering the steady state distribution of the Markov chain.
249 Steady state distribution is the probability of being in a state which remains
250 the same throughout transitions. Every random walk over the transition matrix

251 approximates this distribution after infinitely long time. Finally, as Pearson cor-
252 relation is order-agnostic, calculating Pearson correlation matrix of the grand
253 time-series generates the predicted grand-average FC (gFC) for the testing sub-
254 ject (step 9.).

255 **3. Experiments & Results**

256 Performance of the proposed model was evaluated in the following setup. A
257 randomly chosen set of half of the cohort (23 participants) was used for training
258 and the other half (23 participants) for testing. We used Pearson correlation co-
259 efficient between empirical and predicted functional connectivities (FC) as the
260 measure of model performance in order to keep the measure of model perfor-
261 mance consistent with the extant literature. We first compare the performance
262 of the proposed model against several extant methods that provide SC-FC map-
263 ping followed by explaining the rationale behind the choice of optimal model
264 parameters. We also conduct k-fold cross validation results and perturbation
265 experiments, the results of which support generalizability of our model to other
266 data splits. The proposed model predicts state-specific FCs which are thereby
267 used to product the gFC. The quality of the gFC prediction is highly dependent
268 upon the reproducibility of states and their transition patterns across multiple
269 train-test splits. Obtaining different set of states in different splits shall attest
270 the robustness of the proposed model at question. Finally, we analyze the states
271 discovered from our model by observing the state-specificity property of the
272 model and compare it with the states learned using k-means algorithm in Allen
273 et al. [21].

274 *3.1. Grand average FC (gFC) prediction*

275 We compare the performance of the proposed model with several existing
276 approaches : single diffusion kernel (SDK) model [16], the non-linear dynamic
277 mean field (DMF) model [12] and multiple kernel learning (MKL) model [17].
278 To our knowledge, ours is the only model that incorporates structural infor-
279 mation along with temporal dynamics for predicting grand average FC. DMF
280 and SDK models do not incorporate learning in their formulation and tune the

281 parameters for each subject separately. DMF model inherently captures non-
282 stationarity, therefore it is directly used for gFC prediction without computing
283 wFCs. We estimated the optimal parameters of the DMF and SDK models from
284 the training wFCs and predicted the gFCs of testing subjects using these op-
285 timal parameters. The mode of the performance distribution histogram for the
286 training set was used to select the optimal model parameters. Figure 3 shows
287 that tMKL has superior performance compared to the others.

288 To validate the generalizability of the tMKL model over unseen testing
289 data, we performed k-fold cross-validation experiment whose results are listed
290 in Table 1. These results suggest that performance of our solution is consistent
291 across various splits, hence supporting our claim of generalizability of our model
on unseen data.

k	fold-1	fold-2	fold-3	fold-4	fold-5	mean
2	0.757	0.732	-	-	-	0.745
3	0.771	0.811	0.778	-	-	0.787
5	0.785	0.809	0.813	0.809	0.808	0.805

TABLE 1: **Cross-validation experiments suggesting generalizability of tMKL model.** Mean k -fold cross-validation results for $k = 2, 3, 5$ are shown in the corresponding rows for k -values. As the number of training samples increases with the number of folds, the mean performance also increases suggesting that the model is learning well with increased samples and is able to replicate the same for testing subjects.

292

293 Now, the choice of various model parameters is explained in the next sub-
294 section.

295 3.2. *Parameter Selection*

296 1. **Choice of size of sliding-window, ω** : Within the extant literature, the
297 choice of a suitable sliding window size is an open problem with respect
298 to the analysis of temporal dynamics in rs-fMRI [20]. The sliding window
299 size should be small enough so as not to miss the state transitions and
300 should be large enough to capture the state transitions reliably. Keeping
301 this in mind, we followed Allen et al. [21] by using a sliding window of
302 $\omega = 22$ TRs. The window was tapered at the ends by convolving it with



FIGURE 3: **Model performance comparison between tMKL and existing models.** Pearson correlation between the empirical and predicted gFCs for all the testing subjects is shown for all models. As can be seen, MKL model outperforms other two models, and tMKL model is at par or better than MKL for all but one testing subjects. Even though there is marginal gain in the overall prediction quality, tMKL provides rich insights into the temporal dynamics thus gaining its superiority over extant models.

303 a Gaussian of $\sigma = 3$ TRs and was slid with a stride of 5 TRs to create
 304 wFCs.

305 2. **Choice of GMM parameters :** Each *latent transient state* in which
 306 the wFCs lie is represented using a component Gaussian of the GMM. In
 307 order to choose the optimal number of these states, K , we selected the
 308 GMM model corresponding to a minimum BIC score. Bayesian informa-
 309 tion criterion (BIC) is a statistical measure based on the log-likelihood
 310 function used for selecting a model amongst a finite set of alternatives,
 311 where the model corresponding to the lowest BIC score is chosen. The
 312 plot in Figure 4 shows BIC scores for the models obtained by fitting GMM
 313 for a large range of K (2 to 19), where the minimum value was obtained
 314 for $K = 12$. For each K , we ran GMM 100 times and noted the minimum
 315 BIC score, these BIC scores were used in the figure. To retain generality
 316 of the component Gaussians, we ran our experiments by considering a
 317 unique full covariance matrix for each component Gaussian.

318 3. **Choice of number of diffusion scales for tMKL, m :** The scale
 319 values were sorted in ascending order, where lower values correspond
 320 to local diffusion phenomenon and higher values correspond to global
 321 diffusion phenomenon. Scale values lying in-between correspond to in-
 322 termediate diffusion phenomena. We ran several experiments by varying

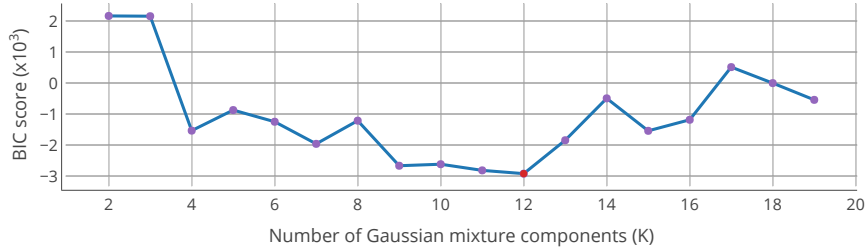


FIGURE 4: **Bayesian information criterion (BIC) score for selecting the number of components in Gaussian mixture model.** The GMM is fit over the training wFCs lying in the lower dimensional manifold. The BIC score is reported by varying the number of component Gaussians (K) from 2 till 19. The Gaussian mixture model corresponding to $K = 12$ (shown in red) has the lowest BIC score among others and is therefore preferred. The plot shows local minima at $K = 4, 7,$ and 15 which may mislead the user while selecting the optimum model. This local minima suggests the choice of number of components in Allen et al. [21].

323 m at powers of 2 from 2 to 32. While the performance for all the scales
324 was reasonable, however in order to carry out comparative analysis with
325 the MKL model, we chose the number of scales as $m = 16$ for all the
326 experiments.

327 3.3. *Robustness of the model*

328 In order to validate the robustness of our model we performed various experi-
329 ments to assess whether our solution overfits the training data and also whether
330 the prediction of the grand average FC is agnostic to the particular SC matrix.

331 1. **Reproducibility of states** : As mentioned in Section 2.2.2, GMM yields
332 K soft assignment vectors for the training wFCs. We validated reproduc-
333 ibility of this clustering by ensuring replication of the same for wFCs of
334 the testing subjects. We generated wFCs for all the testing subjects using
335 the sliding window approach. Soft assignment vectors were generated for
336 these testing wFCs using the GMM employed on the training data, which
337 is then used to compute the Markov transition matrix and the correspon-
338 ding steady state distribution. Figure 5 shows an example of the steady
339 state distribution for our proposed method. We evaluated the similarity

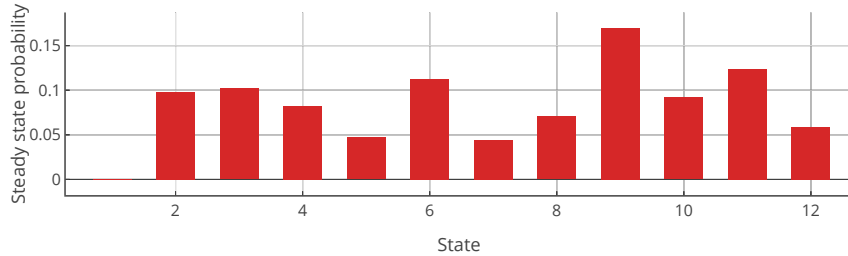


FIGURE 5: **Markov chain steady state distribution.** After the states are retrieved using GMM, the Markov chain transition matrix was learned over the resulting state-sequences of the wFCs of the training subjects. The figure shows the steady state distribution of the transition matrix, which represents the probability distribution of occurrence of a state after infinite amount of time.

340 between the Markov transition matrix and steady state distributions of
341 the training and testing wFCs by finding the Pearson correlation coef-
342 ficients. Table 2 shows that the states are highly replicable for multiple
343 train-test splits of the data.

344 2. **Perturbation experiments** : Each testing subject SC was perturbed
345 $N = 150$ times and using the learned model we predict the grand average
346 FC. We perturbed every SC by randomly generating it from the power law
347 distribution followed by its elements. The generated state-specific wFCs
348 may have non-positive eigenvalues. Here we considered only the real part
349 of the generated time series in order to estimate (predict) grand average
350 FC. Figure 6 shows this observation over all the 23 testing subjects. Box
351 plots for each subject depict the range of correlation values for random
352 SCs. Here we observe less correlations between empirical and predicted
353 gFCs using the perturbed SC, validating that our model respects the
354 topology of input SC. This suggests that the model is not overfitting the
355 data and is sensitive to perturbation in SC.

356 3.4. *State-specificity of the tMKL model*

357 In the previous section, we successfully investigated whether the estimated
358 state transition matrix is general enough in the sense of being reproducible with
359 several train-test splits of the data (refer Table. 2). Other critical questions are

Run Index	ρ_{TM}	ρ_{SSD}	e_{TM}	e_{SSD}
1	0.9947	0.9509	0.1337	0.0564
2	0.8683	0.8546	0.7379	0.2703
3	0.9440	0.8839	0.5120	0.1433
4	0.9035	0.9809	0.7154	0.1004
5	0.8624	0.9604	0.7094	0.1332
6	0.9665	0.8337	0.3824	0.1119
7	0.9131	0.8563	0.6263	0.1107
8	0.9746	0.6824	0.3381	0.1521
9	0.9275	0.8691	0.6671	0.0950
10	0.8623	0.9599	0.7299	0.1608
11	0.9777	0.9596	0.3250	0.0482
mean	0.9301	0.8692	0.5068	0.1358
stdev	0.0501	0.1155	0.2093	0.0636

TABLE 2: Comparison of Markov chain transition matrix (TM) and its steady state distribution (SSD) between training and testing subjects. Comparison is done computing the Pearson correlation coefficient (ρ) and the L2 distance (e) between the training-TM, testing-TM and training-SSD, testing-SSD respectively. This experiment is repeated for 11 train-test splits of the data. Consistent high values of ρ and low values of e across multiple splits show similarity of the states and their transition behavior across train-test subjects, therefore establishing the reproducibility of states.

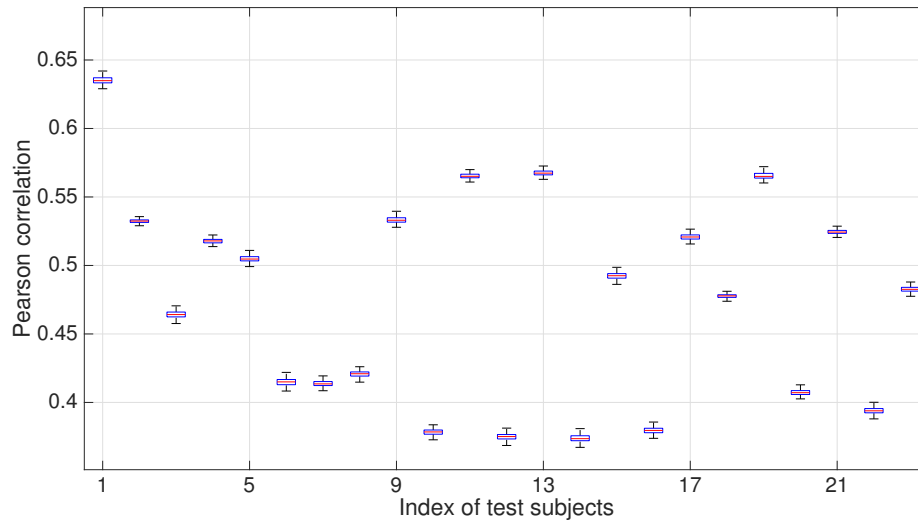


FIGURE 6: **Effect on performance of tMKL model due to perturbation of SC matrices of test subjects.** Shown here are the box plots (blue) of Pearson correlation between empirical and predicted grand FCs when SCs are perturbed for all the testing subjects.

360 whether different model components such as the Π^k 's as well as the predicted
 361 FCs are distinct for different states or not. If they are not distinct, the resulting
 362 MKL models for different states become redundant. In order to verify the state-
 363 specificity of the model, we performed three simulation experiments : i) to show
 364 that the learned state-specific model parameters on the training data are distinct
 365 for different states, ii) to show that the predicted state-specific FCs during
 366 testing phase are also distinct from one another and, iii) to evaluate the accuracy
 367 of state-specific assignments of the model prediction using precision and recall
 368 measures.

369 As summarized in Section 2.2.4, the full model consists of estimating $m = 16$
 370 π_i^k 's for all the $K = 12$ states (i ranging from 1 to 16 and k ranging from 1
 371 to 12). We perform a comparison experiment to see whether, for a fixed i , π_i^k 's
 372 are dissimilar from each other. The results of the first experiment are depicted
 373 as $m = 16$ similarity matrices in Figure 7. It appears that the learned π_i^k 's
 374 are indeed different for different states, especially in the similarity matrices for
 375 global scales (see the top row of Figure 7).

376 In the second experiment, we verified whether the predicted state-specific
 377 FCs during the testing phase are distinct from one another. For a test subject,

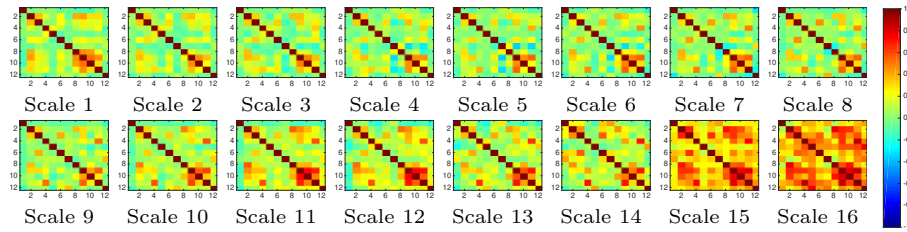


FIGURE 7: **Distinctness of π_i^k 's.** After state-specific MKL models are learned, we check the distinctness of π_i^k 's for every scale value ranging from $i = 1, \dots, m (= 16)$ using Pearson correlation coefficient between every pair of states. Each of these m matrices is a $K (= 12) \times K$ similarity matrix. Distinctness of π_i^k 's would ideally result in a $K \times K$ identity matrix and any deviation would indicate lack of distinctness of the learned π_i^k 's. As observed in most of these m similarity matrices, majority of the off-diagonal entries in these pairwise correlation matrices are zero, indicating the distinctness π_i^k 's. They are significantly distinct for global scales (scale indices $i = 1, \dots, 8$ in the top row) in comparison to local scales (scale indices $i = 9, \dots, 16$ in the bottom row), where they appear to be similar.

378 there are K state-specific FCs predicted based on the SC of the subject (see
 379 step 6. of Figure 1). In the previous experiment, as the Π^k 's have been demons-
 380 trated to be distinct from each other, given a fixed test SC, the state-specific
 381 predictions are also expected to be distinct. Consequently, we computed pair-
 382 wise correlations between the K predicted FCs leading to a $K \times K$ similarity
 383 matrix for each test subject. We then calculated the element-wise mean (8(a))
 384 and standard deviation (8(b)) across the 23 similarity matrices. As shown in
 385 the figure, the dominant identity matrix pattern observed in the mean matrix
 386 combined with low values in the standard deviation matrix, verifies that the
 387 predictions are indeed distinct from one another.

388 In the third experiment, we evaluated the accuracy of the predicted state-
 389 specific assignments of the proposed model. A wFC in the training phase is
 390 assigned to a state which it belongs to with the maximum probability of be-
 391 longingness as described in 1 in section 2.2.2. The cluster assignments of wFCs
 392 to states should obey the principle of maximum intra-state similarity as well
 393 as maximum inter-state dissimilarity. Therefore the predicted FC for a state
 394 in the testing phase should have maximum similarity with the training wFCs
 395 belonging to the same state and also minimum similarity with training wFCs
 396 belonging to other states. For each predicted state-specific FC, a set of trai-

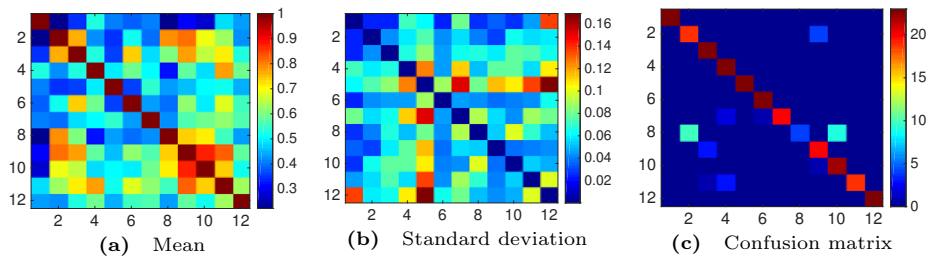


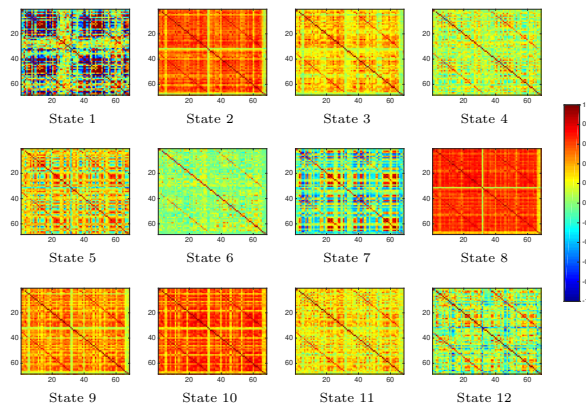
FIGURE 8: **Quantitative state specificity of the model.** Pearson correlation between all the possible pairs of the K state-specific FCs for a subject was calculated and stored in a $K \times K$ matrix. Element-wise mean and standard deviation across all the subject-specific matrices is shown in (a) and (b) respectively. The confusion matrix in (c) can be used for measuring the overall accuracy of state-specific predictions, and precision and recall for each state. For each testing subject there are K number of state-specific FCs. The state label against which a state-specific FC is predicted from tMKL model serves as the ground truth for this experiment. Empirically, each state-specific FC must be nearer to the training wFCs belonging to that state than from the training wFCs belonging to other states, thus attesting the accuracy of model prediction. As the manifold was constructed based on $L1$ similarity, we found the neighbors of the predicted FCs in the original space (of size $n(n-1)/2$). For this purpose, we searched for 25 nearest neighbors for each state-specific prediction and voted for the empirical state-belongingness. Rows (columns) in the confusion matrix depict the actual (predicted) states. Overall accuracy of tMKL model prediction for all the test subjects is 87.68%. It can be seen that non-zero off-diagonal entries result in reduced accuracy. To get a subject-specific measure of the state-specificity, we ran the same experiment for all the testing subjects independently. Noticeably, mean matrix is similar to the confusion matrix with very less standard deviation.

397 ning wFCs is computed which lie in its proximity in the original space (of size
 398 $n(n-1)/2$). The mode of the wFC state-labels of this set of neighbouring train-
 399 ing wFCs would indicate the estimated state label for the predicted FC. Recall
 400 from step 6. of Figure 1 that the tMKL model implicitly assigns a state-label to
 401 the predicted FCs. In this experiment, our aim is to compare the implicit label
 402 with the estimated label. The concurrence is measured through a confusion ma-
 403 trix aggregated over all the test subjects (see Figure 8(c)). The accuracy of the
 404 predicted state-specific assignment measures the number of instances for which
 405 this estimated state label matches the implicit state label. The confusion ma-
 406 trix (Figure 8(c)) has a dominant main diagonal and low off-diagonal elements,
 407 indicating that the implicit assignments seem to be valid. Overall accuracy of
 408 the state-specific assignments for the test subjects works out to be 87.68%.

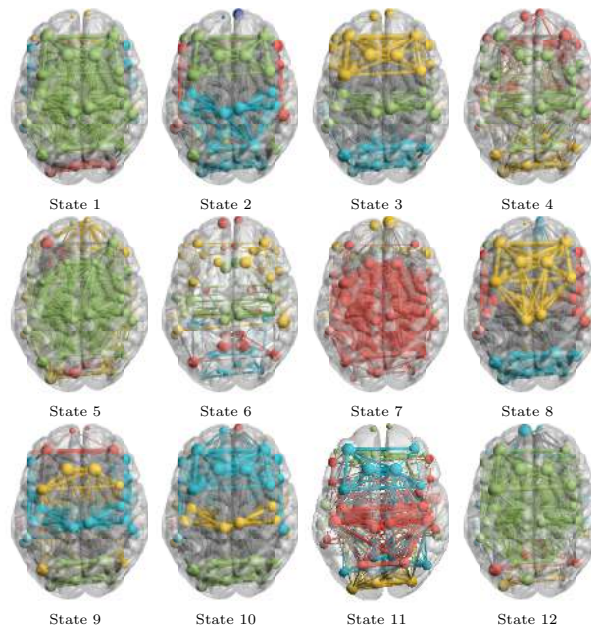
409 In summary, the above-mentioned experiments establish the ‘distinctness’ of
410 states and their corresponding predictions.

411 4. Discussion

412 Besides understanding the relationship between the anatomical architecture
413 and the functional dependencies, over the last decade, characterization of the
414 temporal richness of the resting state functional MRI signal also has been a
415 major trend in the field of cognitive neuroscience. Several approaches have
416 been proposed to understand the inherent richness observed in the sponta-
417 neous spatio-temporal BOLD activity. Operator-based formulations of neural
418 dynamics [11, 12] propose a generative model to predict functional connecti-
419 vity from the structural connectivity via incorporating temporal dynamics into
420 the model. Another class of techniques introducing spectral graph theoretic me-
421 thods [16, 14, 18, 17] primarily focus on mapping the eigen-spectrum of SC and
422 FC of individual subjects, but with minimal focus on the temporal richness.
423 Here, we have proposed an innovative method which combines both anatomical
424 constraints as well as incorporating temporal richness present in the endoge-
425 nous activity. More specifically, our proposed model learns parameters specific
426 to these latent states using a temporal Multiple Kernel Learning (tMKL) and
427 finally predicts the grand average functional connectivity (FC) of the unseen
428 subjects by employing a state transition Markov model. One of the interesting
429 proposal in the framework is that tMKL learns a mapping between the under-
430 lying anatomical network and the temporal structure present in the empirical
431 data to quantify gFC. Further, we have introduced a learning framework to find
432 model specific parameters via state-specific optimization formulations and yet
433 the model performs at par or better than state-of-the-art models for predicting
434 the gFC. Moreover, our proposed model shows sensitivity towards individual
435 subject’s SC as we have clearly demonstrated here with perturbation experi-
436 ments. However, before we can clearly appreciate the novelty in the proposed
437 techniques we need to understand the existing methods of relating underlying
438 SC with windowed FC.



(a) State-specific FCs



(b) Visualization of states

FIGURE 9: **Qualitative state specificity** State specific FCs predicted for every subject are averaged across all testing subjects. (a) Visually distinct FC matrices are shown for all the 12 states. (b) Communities are identified for these mean FCs using Louvian algorithm available in brain-connectivity-toolbox [37] and Brain-net-viewer [38] was used for visualization of these communities. The distinct community structures clearly suggest that transient states are modeling different brain dynamics.

439 *4.1. Relating underlying structural connectivity to Windowed FC*

440 A different line of study by Allen et al. [21] and subsequent works by these
441 and other authors have focused on the temporal structure of the windowed
442 FCs (wFCs) and were able to successfully characterize state transitions, but
443 without specifically relating the temporal dynamics to the underlying structure
444 [39, 40, 41, 20]. Allen et al. collected all the wFCs from the training subjects and
445 learned a k-means clustering model to discover distinct latent states (common
446 to the cohort) that are visited by the brain. Our preliminary attempts at fusing
447 SC using MKL model [17] with the temporal structure learned with k-means
448 did not give satisfactory results (summarized in Figure S1 in Supplementary
449 materials and the description therein). The proposed temporal multiple kernel
450 learning (tMKL) model in this paper belongs to the class of spectral graph theo-
451 retic methods [16, 14, 18, 17]. The model attempts both at improving upon the
452 quality of SC-FC mapping and also aims to characterize the temporal richness
453 of the signal while incorporating the structural information in a principled way.
454 The proposed temporal multiple kernel learning (tMKL) model is an attempt
455 towards generating BOLD time series of a subject using only the SC. As sum-
456 marized in Figure 1, the proposed pipeline partitions the BOLD time series of a
457 subject into windows yielding wFCs. The underlying structure of the wFCs was
458 learned via a manifold whose structure was further parameterized using GMM.
459 The GMM components were hypothesized to be the states whose temporal evo-
460 lution was succinctly captured in a Markov chain transition matrix. For each
461 state, MKL model was learned to capture the SC-dFC relationship. The learned
462 model was utilized to predict state-specific FCs for a test subject. These pre-
463 dicted FCs were further factored into latent time series and concatenated using
464 the steady-state properties of the transition matrix. Pearson correlation of this
465 final time series generates the predicted FC for a subject.

466 *4.2. Rationale for t-MKL pipeline to discover latent temporal struc-* 467 *ture*

468 In the following we will explain the rationale for various steps in the proposed
469 pipeline. It appears that while modeling dFC using unsupervised techniques
470 for clustering wFCs into states, one faces the curse of dimensionality problem

471 head-on. During clustering, wFCs ought to be assigned to the same state as
472 that of their neighbors because they are temporally contiguous and might share
473 similarities. As we can see, wFCs lie in a high dimensional space, but based
474 on their similarity with respect to their neighbors, they may lie on an intrinsic
475 lower-dimensional manifold. This lower dimensional manifold becomes the space
476 over which temporal structure could be precisely identified. Spectral embedding
477 techniques utilize the similarity between the neighboring wFCs to discover the
478 underlying manifold. After representing the temporal structure as a manifold,
479 the next task is to parameterize the lower-dimensional structure. Once we obtain
480 a lower-dimensional embedding, we need to cluster the wFCs to discover the
481 discrete state space. Unsupervised approaches such as K-means clustering would
482 yield spherical clusters, limiting the shape and size of states, whereas GMM
483 clustering is a generalized clustering scheme. We parameterize the local density-
484 distribution of wFCs over the manifold to a factor analysis model that further
485 represents the manifold as a set of component Gaussians at various locations
486 whose shape, orientation, and size depend on the local densities of the wFCs.

487 The proposed model is cohort-based and hence the underlying assumption
488 is of the generalizability of the model to unseen test subjects. We have learned
489 the Markov transition probability matrix on training wFCs and used this to
490 generate long sequences of time series for test subjects eventually yielding a
491 good approximation of the grand average FC with a maximum of 0.8 (see
492 Figure 3).

493 *4.3. Reproducibility of latent states and FC configurations*

494 After presenting the rationale behind designing the proposed model that is
495 shown to be successful at mapping SC-dFC-FC tripartite relationship, several
496 expectations arise such as reproducibility of discovered states and their corres-
497 ponding predicted FCs, sensitivity of the model to the underlying anatomical
498 structure, state-specificity of the tMKL model and importantly verifying that
499 the model does not overfit the training data. Several experiments were conducted
500 in order to verify that the model satisfies these claims and the results presented
501 in Section 3 point to the robustness of the performance of the model. Further,
502 in order to verify whether the state-specific FCs predicted for a subject are
503 distinct, we performed community detection over the mean state-specific FCs

504 of the test cohort (see Figure 9). As can be clearly seen in the figure, regions
505 in each state show distinct interaction patterns among themselves. The states
506 seem to characterize the transient relationship among the ROIs which appear
507 and disappear across the duration of the resting state scan. The markov chain
508 state transition model further allows the characterization of the temporal fluc-
509 tuations of the states that approximates latent temporal structure. Significantly,
510 the MKL models were learned only over the individual states without any glo-
511 bal error measure governing the learning process. Yet, the grand average FC
512 prediction is at par or better than that of the MKL model and superior to the
513 other competing approaches.

514 *4.4. Conclusion*

515 As part of future work, it will be interesting to explore the biophysical mea-
516 ning of the model parameters. Other direction could be to characterize the
517 dynamics better by predicting the time series itself rather than working with
518 correlation matrices. An immediate investigation would be to explore the rela-
519 tionship between the latent time-series and the actual BOLD time-series. Ano-
520 ther line of work would be to apply the proposed model to characterize dFC
521 in various conditions such as neurodegenerative and psychiatric disease, healthy
522 and pathological aging etc.

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