

The Bayesian Reader: Explaining word recognition as an optimal Bayesian decision process.

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Abstract

This paper presents a theory of visual word recognition that assumes that, in the tasks of word identification, lexical decision and semantic categorization, human readers behave as optimal Bayesian decision-makers. This leads to the development of a computational model of word recognition, the Bayesian Reader. The Bayesian Reader successfully simulates some of the most significant data on human reading. The model accounts for the nature of the function relating word-frequency to reaction time and identification threshold, the effects of neighborhood density and its interaction with frequency, and the variation in the pattern of neighborhood density effects seen in different experimental tasks. Both the general behavior of the model, and the way the model predicts different patterns of results in different tasks, follow entirely from the assumption that human readers approximate optimal Bayesian decision-makers.

Introduction

Words that appear frequently in the language are recognized more easily than words that appear less frequently. This fact is perhaps the single most robust finding in the whole literature on visual word recognition. The basic result holds across the entire range of laboratory tasks used to investigate reading. For example, frequency effects are seen in lexical decision (Forster & Chambers, 1973; Murray & Forster, 2004) in naming (Balota & Chumbley, 1985; Monsell, Doyle, & Haggard, 1989), semantic classification (Forster & Hector, 2002; Forster & Shen, 1996), perceptual identification (Howes & Solomon, 1951; King-Ellison & Jenkins, 1954) and eye fixation times (Inhoff & Rayner, 1986; Just & Carpenter, 1980; Rayner & Duffy, 1986; Rayner, Sereno, & Raney, 1996; Schilling, Rayner, & Chumbley, 1998). Frequency effects are also seen reliably in spoken word recognition (Connine, Mullennix, Shernoff, & Yelen, 1990; Dahan, Magnuson, & Tanenhaus, 2001; Howes, 1954; Luce, 1986; Marslen-Wilson, 1987; Pollack, Rubenstein, & Decker, 1960; Savin, 1963; Taft & Hambly, 1986) and therefore appear to be a central feature of word recognition in general.

The fact that high frequency words are easier to recognize than low frequency words seems intuitively obvious. However, possibly because the result seems so obvious, very little attention has been given to explaining why it is that high frequency words should be easier to recognize than low frequency words. All models of word recognition contain some mechanism to ensure that high frequency words are identified more easily than low frequency words but, in many models, high frequency words are easier by virtue of some arbitrary parameter setting. For example, in the E-Z Reader model (Reichle, Pollatsek, Fisher, & Rayner, 1998; Reichle, Rayner, & Pollatsek, 1999, 2003), lexical access is specified to be a function of log frequency, where the slope of the frequency function is a model parameter. In the logogen model (Morton, 1969) resting levels of logogens for high frequency words are set to be higher than resting levels for low frequency words.

However, resting levels or thresholds could equally well be set to make low frequency words easier than high. The only factor preventing this move is that it would conflict with the data. More generally, even when models contain a mechanism that necessarily produces a frequency effect (e.g. Forster's, 1976, search model) one might still ask why there should be a frequency effect at all. That is, wouldn't it be better if all words were equally easy to recognize? The present paper attempts to answer this question by presenting a rational analysis (Anderson, 1990) of the task of recognising written words. This analysis assumes that people behave as optimal Bayesian recognizers. This assumption leads to an explanation of why it is that high frequency words ought to be easier to recognize than low frequency words. Furthermore, it explains why the function relating frequency to reaction time (Whaley, 1978) and perceptual identification threshold (Howes & Solomon, 1951; King-Ellison & Jenkins, 1954), is approximately logarithmic (although for further qualification see Balota, Cortese, Sergent-Marshall, Spieler, & Yap, 2004, and Murray & Forster, 2004). It also explains why neighborhood density influences word recognition, and why neighborhood density can have a different influence on tasks such as lexical decision, word identification, and semantic categorization.

Explaining the word-frequency effect

Explanations of the word frequency effect fall into three main categories: First, frequency could influence processing efficiency or sensitivity (e.g. Solomon & Postman, 1952). That is, perceptual processing of high frequency words might be more effective than perceptual processing of low frequency words. Second, frequency might alter the bias or threshold for recognition (e.g. Broadbent, 1967; Grainger & Jacobs, 1996; McClelland & Rumelhart, 1981; Morton, 1969; Norris, 1986; Pollack et al., 1960; Savin, 1963). That is, frequency might not alter the effectiveness of perceptual processing, but would simply make readers more prepared to recognize a high frequency word on the basis of less evidence than would be required to identify a low frequency word.

Finally, lexical access could involve a frequency ordered search through some or all of the words in the lexicon (Becker, 1976; Forster, 1976; Glanzer & Ehrenreich, 1979; Murray & Forster, 2004; Paap, McDonald, Schvaneveldt, & Noel, 1987; Paap, Newsome, McDonald, & Schvaneveldt, 1982; Rubenstein, Garfield, & Millikan, 1970).

Frequency as sensitivity

The idea that frequency might have a direct effect on the efficiency or sensitivity of perceptual processing was first proposed by Solomon and Postman. However, there has been very little direct evidence for this view, and there have been no explicit accounts of how changes in word frequency might alter the nature of perceptual processing, beyond what might be involved in the initial learning of a new word. Broadbent (1967) directly compared bias and sensitivity accounts of frequency in spoken word perception and concluded that the evidence was entirely consistent with a response bias account.

However, the debate over whether frequency effects are due to changes in sensitivity or bias has been reengaged recently by Wagenmakers, Zeelenberg, and Raaijmakers (2000), Wagenmakers, Zeelenberg, Schooler, and Raaijmakers (2000), and Ratcliff and McKoon (2000). The debate centers on data from a two-alternative forced-choice tachistoscopic identification task. In this task, participants see a briefly presented word followed by a display consisting of two words, one of which is the briefly presented word. The participant's task was to decide which of these two words had actually been displayed. Wagenmakers, Zeelenberg, and Raaijmakers showed that participants perform better when the alternatives are both high-frequency words than when they are both low in frequency. This is exactly what one would expect if frequency increased perceptual sensitivity rather than bias. A simple response bias account would predict that, with words of equal frequency,

the biases would cancel out, and performance on high and low frequency words would be identical. However, Wagenmakers, Zeelenberg, Schooler, and Raaijmakers suggest that this result can be explained by assuming that participants sometimes make a random guess when words have not reached a sufficiently high level of activation. As participants will be more likely to guess high-frequency words, high-frequency words will tend to be identified more accurately. This suggestion is similar to Treisman's (1978a, 1978b) Perceptual Identification model. Treisman assumed that once perception had narrowed its search to some subvolume of perceptual space, words within that subvolume were chosen at random, with a bias proportional to frequency, regardless of how near they were to the centre of the subvolume. As Wagenmakers, Zeelenberg, Schooler, and Raaijmakers note, this is not an optimal strategy, as some potentially useful information is discarded.

However, it is also possible that this task may make it difficult for participants to behave optimally. The best strategy would be to consider only the two alternative words presented, and to ignore the rest of the lexicon. With words of equal frequency, participants should then choose the alternative that best matches the target. However, because of the delay (300ms) between presentation of the target and the alternatives, participants may begin to identify the target word in the normal way, such that all of the words in the lexicon are potential candidates. If this were to happen, low frequency words would be less likely to be identified correctly, and participants might often misidentify low frequency targets as a word other than one of the two alternatives. When the two alternatives are presented, participants would then have to make a random choice between them. There would therefore be more random guesses for low than high-frequency words. The important point here is that participants can only behave optimally if they can completely disregard other words in the lexicon.

Response Bias Theories

The most familiar example of a response bias account of word frequency is Morton's (1969) logogen model. In the logogen model, each word in the lexicon has a dedicated feature counter, or logogen. As perceptual features arrive, they increment the counts in matching logogens. A word is recognized when its feature count exceeds some threshold. Frequency effects are explained by assuming that high-frequency words have higher resting levels (or equivalently, a lower threshold) than low frequency words. High frequency words can therefore be identified on the basis of fewer perceptual features than low-frequency words. In such a model, it might seem that word identification would be most efficient when thresholds are set as low as possible, consistent with each logogen responding to only the corresponding word, and not to any other word. However, as Forster (1976) pointed out, if this were so, then increasing the resting level of high frequency words beyond that point would often cause a logogen for a high frequency word to respond in error when a similar word of lower frequency was presented. To avoid this, and allow headroom for frequency to modulate recognition, thresholds must initially be set at a conservative level where many more than the minimum number of features is required for recognition. So, if the baseline setting is that words need N more features than are actually required for reliable recognition, the resting levels of high-frequency words can be raised by up to N features before causing errors. However, all other words will now be harder to recognize than they would have been using the original threshold. That is, in order to be able to safely raise the resting levels of high-frequency words, all other words in the lexicon have had to have their thresholds set higher than necessary. The overall effect of making a logogen system sensitive to frequency is therefore to make word recognition harder. This seems to be a quite fundamental problem with the account of frequency given by the logogen model: Incorporating frequency into resting levels decreases the efficiency of word recognition. However, as will be shown later, this is not a problem with criterion bias models in general, but

rather with the way the logogen model combines frequency and perceptual information in a completely additive fashion.

Search Models

The other main class of theory is search models. The most fully developed search model is that of Forster (1976). In Forster's search model, a partial analysis of a word directs the lexical processor to perform a serial search through a frequency ordered subset of the lexicon (a 'bin'). The straightforward prediction of this model is that the time to identify a word should be a linear function of the rank position of that word in a frequency ordered lexicon. Recently, Murray and Forster have presented evidence that rank frequency does actually give a slightly better account of the relation between RT and frequency in a lexical decision task than does log frequency. Although the difference between rank frequency and log frequency correlations is small, it is quite possible that the choice between alternative models will ultimately hinge on such subtle differences. However, although the search model does give a principled explanation of the form of the relation between frequency and RT, this is the only prediction that follows directly from the assumption of a search process. For example, Murray and Forster (2004) showed that the function relating frequency and errors was also closely approximated by a rank function. However, to explain this in a search model, they had to suggest that, as each lexical entry was searched, there was some probability that the search might get side-tracked and move to the wrong subset of the lexicon. A search of the wrong bin should always lead to a 'no' response. Murray and Forster pointed out that if the probability of a side-tracking error is sufficiently small, error rate will be approximately proportional to the number of lexical entries that must be searched before encountering the correct word, i.e. rank frequency.

Note that the search model also incorporates a degree of 'direct access'. The initial analysis of a word directs the search to a bin containing only a subset of the words in the lexicon which share some common orthographic characteristics. By reducing the bin size to 1, the search model would become a direct access model. The fewer words there are to be searched in each bin, the faster recognition would become. In other words, search is a suboptimal process. Direct access would be more efficient, but then there would be no word frequency effect.

Frequency and learning

A deficiency in all of these explanations is that they simply indicate how a word frequency effect might arise. None offers a convincing explanation for why word recognition should be influenced by word frequency at all. Monsell (1991) has directly considered the question of why there should be frequency effects. He argued that frequency effects follow naturally from a consideration of how words are learned. His arguments were cast in the framework of connectionist learning models, and he suggested that it was an inevitable consequence of such models that word recognition would improve the more often a word was encountered. Indeed, all models of word recognition in the parallel-distributed processing framework do show a word frequency effect (e.g. Plaut, 1997; Plaut, McClelland, Seidenberg, & Patterson, 1996; Seidenberg & McClelland, 1989). Depending on the exact details of the network, a connectionist model could predict that frequency would influence bias, sensitivity or both. However, two main factors undermine the learning argument somewhat. The first is that not all connectionist learning networks need extensive experience to learn. In particular, localist models are capable of single trial learning (see Page, 2000, for a discussion of the merits of localist connectionist models). Human readers are also capable of very rapid and long-lasting learning of new words (Salasoo, Shiffrin, & Feustel, 1985). The second is that readers will have extensive experience with words from all but the lowest end of the frequency range. Why should these words not end up being learned as well as high frequency words? In particular, one

would expect that readers with more experience would show a smaller frequency effect. Over time, performance on high frequency words should approach asymptote, while low frequency words would continue to improve. There is some support for this prediction. Tainturier, Tremblay, and Lecours (1992) found that individuals with more formal education (18 years vs 11 years) showed a smaller frequency effect. As Murray and Forster (2004) point out, the word frequency effect should also get smaller with age. However, the frequency effect seems to either remain constant with age (Tainturier, Tremblay, & Lecours, 1989) or to increase (Balota, Cortese, & Pilotti, 1999; Balota et al., 2004; Spieler & Balota, 2000). This would seem to imply that learning is not very effective. That is, a learning-based explanation of the word frequency effect seems to be predicated on the assumption that the learning mechanism in the human brain fails to learn words properly even after thousands of exposures. In effect, this form of explanation amounts to a claim that sensitivity to word frequency is an unfortunate and maladaptive consequence of an inadequate learning system.

Search models are open to a similar criticism. A frequency ordered search process will make low frequency words take longer to identify than high frequency words, but a parallel access system would eliminate the disadvantage suffered by low frequency words completely. Once again, the word frequency effect is explained as an undesirable side-effect of a suboptimal mechanism.

A Bayesian recognizer

One answer to the question of why word recognition should be influenced by frequency is provided by asking how an ideal observer should make optimal use of the available information. The concept of an ideal observer has a long history in vision research (Geisler, 1989; Hecht, Shalae, & Pirenne, 1942; Weiss, Simoncelli, & Adelson, 2002), but has less commonly been applied to higher cognitive processes. However, recently, the ideal observer analysis has also been applied to object perception (Kersten, Mamassian, & Yuille, 2004), eye movements in reading (Legge, Klitz, &

Tjan, 1997), and letter and word recognition (Pelli, Farell, & Moore, 2003). The basic concept of the ideal observer is very simple: given some form of perceptual input, and a clear specification of the task to be performed, what is the behavior of a system that performs the task optimally?¹.

In cases where the perceptual input is completely unambiguous, an ideal observer would simply select the lexical entry whose representation matches the input. The actual implementation of the matching process might be very complex, but the optimal strategy is just to select the matching word from the lexicon. An optimally designed system would be able to match the input against all words in the lexicon in parallel, and there would be no effect of word-frequency. However, if the input is ambiguous, the requirements are different. With ambiguous input it is no longer sufficient simply to select the best matching lexical entry. Under these circumstances an ideal observer must also take the prior probabilities of the words into account. That is, an ideal observer should be influenced by the frequency with which words occur in the language.

The behavior of an ideal observer is critically dependent on the precise specification of the task or goal. In what follows, the goal of the observer will be taken to be that of performing the experimental task as fast as possible, consistent with achieving a specified level of accuracy. In practice, goals are much more complex than this. Even in a lexical decision experiment, accuracy has to be traded off against speed. Responses have to be made before some deadline imposed by the experimenter. There are costs in making errors, and a cost in not responding in time. There may be different costs in making errors on words and nonwords. In an ideal observer analysis this trade-off between different constraints is expressed in the form of a utility function (or, conversely, a loss function). The task of the ideal observer then becomes one of maximizing the value of the utility function. However, the form of the utility function is likely to vary between experiments, and may even be influenced by factors such as the interpersonal interaction between experimenter and participant. In practice, participants have no way of knowing how to maximize the utility function

until they have had some experience with the experimental stimuli. For example, if the goal is to respond as quickly as possible consistent with making about 5% errors, the response threshold will depend on the specific experimental stimuli used. Are the items easy or hard, or a mixture of the two? Some suggestions as to how participants might adapt to the characteristics of the stimuli in a particular experiment are given in Mozer, Colagrosso and Huber (2002) and Mozer, Kinoshita and Davis (2004). However, in the present context, the goal will be taken simply as performing the experimental task with certain accuracy (95% correct).

Ideal observers almost always use Bayesian inference (cf. Geisler, 2003; Knill, Kersten, & Yuille, 1996) as this is the optimal way to combine perceptual information with knowledge of prior probability. Bayes' theorem is shown in Equation 1.

$$P(H | E) = P(H) \times P(E | H) / \sum_{i=0}^{i=n} (P(H_i) \times P(E | H_i)) \quad (1)$$

Starting from knowledge of the prior probabilities with which events or hypotheses (H) occur, Bayes' theorem indicates how those probabilities should be revised in the light of new evidence (E). Given the prior probabilities of the possible hypotheses $P(H_i)$, and the likelihood that the evidence is consistent with each of those hypotheses $P(E/H_i)$, Bayes' theorem can be used to calculate $P(H_i/E)$, the revised, or posterior probability of each hypothesis, given the evidence.

$$P(\text{Word} | \text{Input}) = P(W) \times P(I | W) / \sum_{i=0}^{i=n} (P(W_i) \times P(I | W_i)) \quad (2)$$

$$P(\text{Input}) = \sum_{i=0}^{i=n} (P(W_i) \times P(I | W_i)) \quad (3)$$

In the case of word recognition, the hypotheses will correspond to words, and $P(H)$ is given by the frequency of the word. As shown in Equation 2, the probability that the input corresponds to a particular word is then given by the probability that the input was generated by that word, divided by the probability of observing that particular input. Any particular input could potentially be produced by many different words. The probability of generating a particular input is therefore obtained by summing the probabilities that each word might have generated the input (Equation 3). As will be discussed later, this definition of $P(\text{Input})$ carries the implication that the input was indeed generated by one of the words in the lexicon. This will not always be the case. Bayesian methods are commonly employed in automatic pattern recognition systems (cf. Jelinek, 1997). For example, in tasks such as automatic speech recognition, the process of matching a given input to words in the lexicon is hardly ever perfect. Whether because of ambiguity in the signal itself, or because of limitations on the ability of the recognizer, there is always some residual ambiguity in the match. Even when the input appears to match a particular word very well, there remains some probability that the input to the recognizer could have been generated by another word. In the absence of any knowledge of the prior probabilities of the words, the best strategy will always simply be to choose the word that best matches the input. However, if prior probabilities are available, these should be taken into account by the application of Bayes' theorem.

It might appear that in many laboratory word recognition tasks, such as lexical decision, there is no ambiguity in the input, so prior probability, or frequency, need not be considered. Words are generally presented clearly under circumstances where the participant will have no difficulty in identifying the stimulus accurately. However, the critical question is not whether the stimulus itself could potentially be identified unambiguously, but whether there is ambiguity at the point where a decision is made. Participants in a lexical decision experiment are encouraged to respond as quickly as possible. Indeed, participants generally respond so quickly that they make errors. If they respond before they have reached a completely definitive interpretation of the input, there will necessarily

be some residual ambiguity at the point where the decision is made. Under these circumstances, frequency should still influence responses, even though the stimulus itself is presented clearly.

In order to appreciate the importance of including prior probability in the Bayes' equation, it is helpful to consider a simple example. In common with most expositions of Bayes' theorem I will consider a simple case of medical diagnosis and use Bayes' theorem to determine the probability that a patient has disease D , given that they test positive for that disease (Equation 4). Assume that the test can correctly identify the disease 95% of the time when the disease is present, and has a false alarm rate of 10%. That is, on 10% of occasions when the disease is not present, the test incorrectly produces a positive result.

Equations 5 and 6 show the probability that a patient testing positive for the disease really does have the disease when the prevalence of the disease is 1 in 5, or 1 in 1000. $P(\text{Disease} | \text{Test Positive})$ is much lower when the prevalence of the disease is also low. The reason is very simple: Because there is a substantial probability that the test will produce a false positive result, the rarer the disease, the more likely it is that a positive result will be due to a false alarm. In a sense, we can think of Bayes' theorem as providing a correction for false alarms.

$$P(\text{Disease} | \text{Test Positive}) = P(D) \times P(\text{TestPos} | D) / ((P(D) \times P(\text{TestPos} | D)) + (P(\text{Not}D) \times P(\text{TestPos} | \text{Not}D))) \quad (4)$$

$$P(\text{Disease} | \text{Test Positive}) = 0.2 \times 0.95 / ((0.2 \times 0.95) + (0.8 \times 0.1)) \quad (5)$$

$$P(\text{Disease} | \text{Test Positive}) = 0.70$$

$$P(\text{Disease} | \text{Test Positive}) = 0.001 \times 0.95 / ((0.001 \times 0.95) + (0.999 \times 0.1)) \quad (6)$$

$$P(\text{Disease} | \text{Test Positive}) = 0.009$$

Exactly the same reasoning applies to word recognition. If the recognizer is presented with a noisy representation of a word (whether because the stimulus itself is noisy, or because there is noise in the perceptual system) there will be some probability that the word that most closely matches the input will not be the correct word. If the input most closely matches a low frequency word (tests positive for that word), there is some probability that the input was actually produced by another word (i.e. the fact that the low frequency word is the best match is a false alarm). If the other word is much more frequent than the low frequency word, it may be more likely that the input was produced by the less well matching high frequency word than the more closely matching low frequency word. That is, information about word frequency effectively alters the weighting of perceptual evidence. Note that frequency is not the only factor that can influence the prior probability of a word. Under natural reading circumstances a variety of sources of contextual information will also alter the expected probability of words. For example, syntactic knowledge, sentence level semantic interpretation, discourse representations, and parafoveal preview (e.g. Rayner, 1998) might all alter the prior probabilities. Frisson, Rayner and Pickering (2005) and McDonald and Shillcock (2003a; 2003b) have shown that eye movements are influenced predictability or transition probability. McDonald and Shillcock (2003a) interpreted their data in terms of Bayes' theorem, and they suggest that transition probability and frequency act on the same stage of reading. In automatic speech recognition, all of these probabilities are usually derived from what is referred to as a language model. In fact, word frequency is simply a unigram language model.

The general form of Bayes' theorem shares a family resemblance to the Luce (1959) choice rule and Massaro's Fuzzy Logical Model of Perception (FLMP Massaro, 1987; Oden & Massaro, 1978). In common with the Luce choice rule, the probability of a word is given by the ratio of the evidence

for that word, divided by the evidence for all other words. What is perhaps unfamiliar in Bayes' theorem is the fact that the measure of the evidence for a word contains the term $P(I|W)$, that is, the probability of observing the perceptual input I , given that the word W has been presented. $P(I|W)$ could be determined by experience. For example, each time a word is encountered it will produce some representation I_x at the input of the recognizer. The recognizer could then (assuming it had some source of knowledge about the true identity of the word) learn the probability density function (pdf) representing the probability of receiving a particular input, given that a particular word was presented. For any new input, the system would then be able to look up the probability that that input came from presentation of a particular word ($P(I|W)$). This is effectively how some automatic speech recognition systems are trained to recognize words using hidden Markov models. In fact, $P(I|W)$ could be computed from the products of $P(I|\text{Letter})$ over letter positions, so there is no need for the system to have extensive experience of every word in the lexicon. Experience of letters should suffice. However, as will be explained below, here I will take a rather different approach to computing $P(I|W)$ and estimate it from the input.

In effect, frequency in the Bayesian approach acts as a bias. For example, in a perceptual identification task, a Bayesian recognizer should respond with the word with the largest posterior probability. Other things being equal, high frequency words would therefore tend to be generated as responses more often than low frequency words. However, frequency would not improve the perceptual sensitivity of the system in terms of its ability to discriminate between a pair of words in a forced-choice task. This is consistent with Broadbent's (1967) data (for speech recognition) that suggested that word frequency is best characterized as a response bias, and with Wagenmakers, Zeelenberg, Schooler, and Raaijmakers' (2000) interpretation of forced-choice perceptual identification data.

It is important to note that the frequency bias in a Bayesian recognizer (and also in the Luce choice theorem and FLMP) acts quite differently from the response bias in, for example, Morton's logogen model, or the criterion bias model described by Broadbent (1967). In the logogen model, the response bias is independent of the amount of perceptual information available. High frequency words just require fewer features to exceed threshold than do low frequency words. As explained earlier, this leads to a problem whereby incorporating frequency makes recognition less rather than more efficient. However, the bias in Bayes' theorem trades off against perceptual evidence. If $P(I|\text{correct word})$ is 1.0, and all other $P(I|W)$ are 0.0, then frequency $P(W)$ cancels out. The better the perceptual evidence, the smaller will be the influence of frequency. Frequency can never override reliable perceptual evidence. This is clearly a desirable property. No matter how large the frequency difference between two similar words, a Bayesian decision process will always select the correct word when the input is unambiguous.

Having established that an optimal word recognizer should take frequency into account, the next section of the paper will develop a simple Bayesian model of visual word recognition (the Bayesian Reader). The aim here is to produce a computational account of word recognition that makes as few assumptions as possible. The model can then be used to simulate word identification and lexical decision tasks, and to investigate how the model behaves when presented with words varying along dimensions such as frequency and orthographic neighborhood density. In effect, this is an exercise in what Anderson (1990) terms 'rational analysis'. Beginning with an account of how an optimal word recognizer should work, we can ask whether the characteristics of human word recognition can be explained by assuming that we approximate optimal recognizers. Remarkably, it turns out that this simple model shows striking parallels with human behavior.

Rational analysis tries to explain cognition in terms of an optimal adaptation to the environment. In this framework, a theory should be more than just an account of how behavior is produced. A good theory should also explain why it is that people behave as they do. This is the link with Bayesian, or ideal observer approaches. Ideal observers behave the way they do because this is the optimal way to perform the task. This contrasts with the alternative accounts of the word-frequency effect reviewed here. All current word recognition models propose some mechanism that can produce a word-frequency effect, but none explains why there should be a word-frequency effect. For a concise summary of the rational analysis position, see Chater and Oaksford (1999), or Schooler (2001).

Rational analysis attaches very little importance to mechanisms or to considerations of neural implementation. The burden of explanation is carried by what Marr (1982) refers to as the computational level. Theories at the computational level specify both what functions a particular information processing system must compute, and why those computations are required to achieve the goals of the system. For any given computational-level theory there will be many different algorithms that could compute the necessary functions, and many different ways those algorithms might be implemented in the brain. However, in the absence of additional constraining data (e.g. from neurobiology) there is typically little basis for choosing one implementation over another, so the computational level provides the most appropriate characterization of the theory. The model presented here is therefore a computational-level theory, although some suggestions will be offered as to how the model might be reformulated as a connectionist network, and how some functions of the model might be implemented. Norris (2005) discusses the relation between computational models and theories in more depth.

It is important to note that neither the rational analysis nor ideal observer approaches imply that human perception or cognition is perfect. The ideal observer is generally assumed to operate on a restricted input where some information is either noisy, or missing completely. There may also be processing limitations, such as limited memory capacity. The ideal observer then makes optimal use of the information and resources available. Differences between the performance of an ideal observer and human perception can therefore help reveal the nature of any information loss in the perceptual system. Information loss might come about as a result of internal noise, or from recoding perceptual representations in a way that throws some information away. The ideal observer analysis can therefore provide significant insights into the nature of the input available to a particular component of the perceptual system. Because the ideal observer only has access to limited information, it will often produce errors. In fact, some of the main benefits of the approach have been in explaining patterns of errors. For example, rational analysis has been used to explain errors in reasoning (McKenzie, 1994; Oaksford & Chater, 1994), syntactic parsing (Chater, Crocker, & Pickering, 1998; Crocker, 1999; Jurafsky, 1996), and memory (Anderson & Milson, 1989). In recent work in vision, Weiss, et al. (2002) have shown how an ideal observer analysis can explain errors in motion perception.

The remainder of this paper is devoted to developing the Bayesian Reader, an ideal observer model of human word recognition. The goal is to establish whether some of the most important features of visual word recognition can be explained simply by assuming that readers approximate ideal observers operating on a noisy perceptual input. All of the important features of the model follow from this simple assumption. The model has quite deliberately been kept as simple as possible to ensure that, to the extent that the model does simulate human data, this is solely a consequence of the assumption that the recognition process is optimal. Like all theory development, rational

analysis is an iterative procedure, whereby theories are refined in the light of discrepancies between theory and data. The Bayesian Reader is the first step in this process.

The Bayesian Reader

As a first step in assessing the behavior of an optimal word recognizer, it is necessary to have an estimate of $P(I|W)$. Although $P(I|W)$ could be learned, for the purposes of the simulations presented here, we will take a rather different approach and assume that $P(I|W)$ can be estimated from the current input. This depends on three assumptions:

1. All words can be represented as points in a multidimensional perceptual space.
2. Perceptual evidence is accumulated over time by successively sampling from a distribution centered on the true perceptual co-ordinates of the input with samples being accumulated at a constant rate.
3. $P(I|W)$ for all words can be computed from an estimate of variance of the input distribution

The first assumption follows simply from the fact that some words are more easily confused than others. The notion of words as being represented in a multidimensional perceptual space has been most clearly articulated by Treisman (1978a; 1978b) in his Perceptual Identification model, and has subsequently been incorporated in other models of word recognition (Luce & Pisoni, 1998; Norris, 1986). In fact, almost all computational models of word recognition either explicitly or implicitly share this assumption. For example, the letter features used in Rumelhart and McClelland's (1982) interactive activation model ensure that words are located in a multidimensional space.

The second assumption should also be fairly uncontroversial, and is similar to the feature accumulation process in the model of lexical decision presented by Wagenmakers, Steyvers, Raaijmakers, Shiffrin, van Rijn, and Zeelenberg (2004) although, in their model, the rate of feature accumulation is not linear with time. It also has parallels with random walk or diffusion models (e.g. Ratcliff, 1978; Ratcliff, Gomez, & McKoon, 2004). This assumption is necessary for the present exercise because it provides a way of investigating the behavior of an optimal recognizer under varying degrees of perceptual uncertainty. As noted above, if perceptual information were completely unambiguous, word frequency should have no influence on recognition.

The third assumption avoids the need to learn the probability density function of $P(I|W)$ for individual words. It also helps to keep the model simple and general, and avoids making any arbitrary assumptions about learning. As successive samples arrive, it is possible to compute the mean location and the standard error of the mean (SEM, Equation 7) of all samples received so far.

$$\sigma_M = \sigma / \sqrt{N} \quad (7)$$

The mean represents a point in perceptual space (the co-ordinates of the best guess as to the perceptual form of the input), and the SEM is computed from the distances between each sample and the sample mean. The SEM is measured in units corresponding to Euclidean distance in perceptual space.

INSERT FIGURE 1 ABOUT HERE

The calculations in the model are illustrated graphically in Figure 1. The figure illustrates the case of a lexicon with only two words, where those words differ along a single perceptual dimension. Each curve represents the probability density function $f(I|W)$ for a given SEM. Note that in Equation 3 the input I is implicitly assumed to be one of a discrete set of values over which the probability $P(I|W)$ is distributed. Given the form of input being assumed here, I is a continuous valued variable whose probability distribution is then correctly represented as a density function $f(I|W)$. Under these assumptions the equivalent Bayes' equation is:

$$P(W | I) = P(W) \times f(I | W) / \sum_{i=0}^{i=n} (P(W_i) \times f(I | W_i)) \quad (8)$$

where $f(I|W)$ corresponds to the height of the pdf at I . For a given I , $f(I|W)$ is called the likelihood function of W . When comparing different candidate W 's on the basis of input I , it is the ratio of the likelihoods that influence the revision of the prior probabilities. The full set of equations governing the calculation of $P(W|I)$ is given in *Appendix A*.

In Figure 1 the current estimate of the mean of the input distribution falls in between the two words, although closer to word 1 than word 2. The lower two curves represent some point early in processing where the SEM is large, and the upper two curves represent a later point, where the SEM is much smaller. As more samples arrive, the SEM will decrease. As shown in Equation 7, the SEM is proportional to the inverse square root of the number of samples. Words that are far away from the mean of the input distribution will tend to become less and less likely as more samples are accumulated. Consequently, $P(W|I)$ of the word actually presented will tend to increase, while the $P(W|I)$ of all other words will decrease. One noteworthy feature of this kind of sampling model is that, given enough samples, there is no limit as to how small the SEM can become. In the absence of any restrictions on the amount of data available (i.e. number of samples), the $P(W|I)$ of a clearly presented word will always approach 1.0 in the limit.

For clarity, the mean of the input distribution is assumed to be the same at both the early and late processing times. In practice, the estimate of the input mean would change over time. The relative heights of the curves for the two words at the point corresponding to the mean of the input distribution reflect the relative likelihoods of the two words. At the early time point the heights of the two curves at the input mean are fairly similar, indicating that the input distribution is only slightly more likely to have been generated by word 1 than word 2. Later in time there is a very large difference in the heights. This is as much due to the fact that $f(I|\text{word 2})$ is now very low as due to $f(I|\text{word 1})$ being high. In effect, word 2 has stopped competing with word 1. The two curves centered on word 1 are identical to those centered on word 2. Because of this, we would get the same calculated heights from a curve centered on the input mean. However, this will only be the case under the assumption that the SEM of the input distribution can be used to estimate $f(I|W)$ for all words. If the pdf of $f(I|W)$ were learned, then each word is likely to have a different pdf.

Although Figure 1 has been described as representing recognition of words, the same principles would apply to recognition of any object in perceptual space. For example, letter recognition from a noisy input would operate in exactly the same way. Indeed, another version of the model has been constructed that operates by computing letter probabilities and then computing word probabilities from the letter probabilities. $P(I|W)$ is then simply computed from the product of the $P(\text{Letter}|\text{Input})$ for each letter in the word. Including a letter level in this way does not alter the behavior of the model.

For some purposes it may be possible to derive $P(I|W)$, or $P(\text{Input} | \text{Letter})$ from perceptual confusion matrices, rather than estimating them from the input. This has been done for visual word identification by Pelli, Farrel and Moore (2003), who used perceptual confusion data to simulate

performance in the Reicher-Wheeler task (Reicher, 1969; Wheeler, 1970) in a Bayesian framework. However, a limitation of this technique is that construction of confusion matrices requires a large amount of data, and each confusion matrix can only characterize the information available at a single point in time. Pelli et al. were only able to perform simulations for three different exposure durations in a perceptual identification task. Their data and simulations will be discussed in more detail later in the paper.

Note that as more perceptual information arrives, $P(W)$ will have less and less influence on $P(W|I)$. In the limit, $P(I|W)$ for all but the word actually presented will approach 0, and $P(W)$ will have no effect whatsoever. However, in general, as $P(W)$ gets lower, the number of samples required to reach a given $P(W|I)$ will increase. That is, high frequency words will be identified more quickly than low frequency words.

It is important to bear in mind that the posterior probabilities being calculated here are the probabilities that the input is a particular word, given that the input really is a word. Because of the properties of the normal distribution, the closest word to the input mean will always have a probability approaching 1.0 in the limit, even if the input does not correspond exactly to any particular word. The decision being made is: given that the input is a word, which word is it? Even an unknown word will produce a high $P(W|I)$ for one existing word in the lexicon. When simulating identification of known words, this limitation is not a problem. However, consideration of how to handle unknown words will become important later when modeling lexical decision.

The representation of word frequency

Equation 1 implies that readers have access to information about each word's prior probability of occurrence in the language. Ideally this would be an estimate of the expected probability of

encountering each word in the current context. However, here I will simply assume that this can be approximated by the measure of word frequency recorded in CELEX (Baayen, Piepenbrock, & Gulikers, 1995). It is important to note that behavior will only be optimal if $P(W)$ is a true probability based on the absolute frequency counts, and not on log frequency. This is an important contrast between the present model and almost all psychological accounts of frequency. Even Rumelhart and Siple (1974), who incorporated a Bayesian decision rule in their model of word recognition, used log frequency.

It is clearly possible that the psychological representation of prior probability might be modulated by factors other than frequency itself. McDonald and Shillcock (2001) suggested word recognition is strongly influenced by the number of different contexts a word can appear in. It has also been argued that age of acquisition (Juhasz, 2005; Morrison & Ellis, 1995; Morrison, Ellis, & Quinlan, 1992) or cumulative frequency (Zevin & Seidenberg, 2004) are more powerful determinants of recognition than overall frequency of occurrence. However, this is still an active area of debate (Juhasz, ; Stadthagen-Gonzalez, Bowers, & Damian, 2004) and, for present purposes, we can think of these factors as simply influencing the psychological estimate of a word's prior probability.

In many psychological experiments the distribution of word frequencies is not at all the same as the distribution of word frequencies in the language. For example, in some experiments (e.g. Forster, 1981; Glanzer & Ehrenreich, 1979) participants may see only high-frequency words, or only low-frequency words. An ideal Bayesian decision process should adapt to these local probabilities (cf. Mozer et al., 2002). However, it is worth bearing in mind that, even if participants are aware that an experiment contains predominantly low frequency words, low frequency words will have lower effective probabilities than would high frequency words in an experiment containing only high-frequency words. Because there are more low-frequency than high-frequency words (Zipf's law),

the probability of encountering any particular low-frequency word will still be less than the probability of encountering a particular high-frequency word.

Possible mechanisms

Although the Bayesian Reader has been developed within the rational analysis framework, and has therefore been presented purely at a computational level, it should be relatively straightforward to implement the theory as a connectionist network. Following from the work of MacKay (1992) there has been considerable interest in developing connectionist networks to perform Bayesian inference and classification. A detailed account of how to construct connectionist networks to perform Bayesian inference is given by McClelland (1998). Possible neural mechanisms to compute likelihood ratios are discussed in Gold and Shadlen (2001), while Rao (2004) shows how a recurrent network architecture can implement Bayesian inference for an arbitrary hidden Markov model. This latter paper is particularly interesting in the present context as the network has to learn probability density functions. Further discussion of the neural mechanisms that might perform Bayesian computations are discussed in Burgi, Yuille and Grzywacz (2000) and Kersten, Mamassian and Yuille (2004).

From the rational analysis perspective a connectionist implementation of the Bayesian Reader would not increase the explanatory value of the theory. However, an illustration of how the model might be implemented as a network might make it easier to appreciate the relationship between the Bayesian Reader and connectionist models of word recognition like MROM (Grainger & Jacobs, 1996) and DRC (Coltheart, Rastle, Perry, Langdon, & Ziegler, 2001). A possible architecture is shown in Figure 2. The topology of this network is exactly the same as that presented in figure 2.1 of McClelland (1998). Although this network assumes a direct mapping between input features and

words, it would be trivial to extend it to have an additional level of representation mapping from input features to words via letters.

INSERT FIGURE 2 ABOUT HERE

The input units in the network would each correspond to an element in a feature vector corresponding to the mean value of that element in the input samples. So, for a 4-letter word there would be 4×26 input units. The input units therefore code the vector representing the location of the current input in perceptual space, and are exactly equivalent to the input to the model as already described. Layer 1 consists of Gaussian word units whose output is a Gaussian function of the distance between the stored representation of the word and the input words. These units produce a larger output the closer the input vector is to the representation stored in the weights of the word unit. These units calculate the heights of the pdfs. The sharpness of the Gaussian tuning function is modulated by the current SEM. The smaller the SEM the more finely tuned (peaked) the Gaussian function will be. These units therefore calculate the likelihoods of the words. The Layer 1 word units also need to have a bias representing the frequency of the word so that the likelihoods can be multiplied by $P(W)$. The sigma unit sums the activation of all layer 1 word units, and passes that to the output units which calculate $P(W|I)$ by dividing their input from Layer 1 by the output of the sigma unit. Responses can therefore be controlled by setting a threshold on the activation of the output units. In common with other connectionist models in psychology, it is likely to be the case that the functions computed by the units in the network might be best computed by a neuronal assembly consisting of a set of simpler elements.

The first point to note about this network is that, by design, it is guaranteed to produce exactly the same output as the Bayesian Reader. In fact, the network is simply a redescription of the

computations performed in the Bayesian Reader. For example, the computer code required to simulate a Gaussian word unit is exactly the same as the code required to compute the heights of pdfs in the model. Indeed, the program currently implementing the Bayesian Reader is exactly the program one would write to implement the network. Only the names given to functions and objects might change. The only difficulty in transforming the Bayesian Reader into a network is really in deciding which of the many possible implementations one might choose. In the absence of additional constraining data (e.g. from neurobiology) there is little basis for choosing one implementation over another, therefore the computational level provides the most appropriate characterization of the theory. The network in Figure 2 is not the theory (see Norris (2005), for further discussion of the relationship between networks, models, and theories). The network representation of the model highlights an important similarity, and difference, between the Bayesian Reader and MROM and DRC. Both MROM and DRC rely on a measure of global activation in order to perform lexical decisions. The Bayesian Reader requires a measure of global activation computed by the sigma unit in order to perform word recognition at all. Although normal word recognition depends solely on individual $P(W|I)$, the system must always compute global activation too.

The following section presents a simulation to illustrate the behavior of the Bayesian Reader when identifying words of different frequencies.

Simulation of the word-frequency effect in identification tasks

In all of the simulations presented here, each input letter is represented as a 26-element vector. Each letter of the alphabet is coded by setting a single element to 1.0 and all other elements are set to zero. Words are represented by a concatenation of position-specific letter vectors. Every word, or letter string, can therefore be considered to be represented by a point in multidimensional space

(104-space for 4-letter words). Because this coding scheme does not lend itself to representing words of different length in a single multidimensional space, all simulations use only words of a single length. This particular coding scheme was chosen because of its simplicity rather than because of a theoretical commitment to this form of input representation. When a letter string is presented to the Bayesian decision process a series of samples is generated from a distribution centered on the corresponding point in space. Each sample is generated by adding zero-mean gaussian noise to all dimensions. The exact value of the variance of the noise can be considered to be a scaling parameter. The more noise is added, the longer it will take to reduce the SEM to a point where decisions can be made with a high degree of certainty.

After each new sample is received, the mean location of the input distribution is estimated, and the SEM is calculated. The SEM is then used to calculate the $P(I|W)$ for all words (see Figure 1). The $P(I|W)$ are then entered into Equation 2. This allows $P(W|I)$ to be calculated for all words. If the probability of a word exceeds some threshold, then that word is generated as a response. This procedure is very closely related to the M-ary Sequential Probability Ratio Test (MSPRT) described by Baum and Veeravalli (1994), which is an extension of Wald's (1947) Sequential Probability Ratio Test (SPRT). Given that one sample is accumulated per unit time, the number of samples required to reach a given probability threshold will be linearly related to both RT and identification threshold.

It is also possible to set a time threshold rather than a probability threshold, and then to read out probabilities after a given amount of time (cf. Van Rijn & Anderson, 2003; Wagenmakers et al., 2004). Because the model has to cycle through the lexicon computing probabilities for all words, and response times have to be averaged over many runs, it is very computationally expensive. All

simulations were run on a network of Windows PCs running the Condor system (Litzkow, Livny, & Mutka, 1988).

The first set of simulations of word frequency use a lexicon containing 4106 5-letter words from the CELEX database. One hundred and thirty words were selected at random from this lexicon as test stimuli. These words were divided into 13 sets of 10. Each set of 10 words was assigned to one of 13 different frequency values (1,2,5,10,20,50,100,200,500,1000,2000,5000,10000), while the remaining words were given their normal CELEX written frequency values³. In order to ensure that these initial simulations would reflect word frequency rather than properties of individual words, there were 13 different lists of the 130 words, with assignment of frequency to words over lists determined by a latin square. Identification time for all words in each frequency value was averaged over 50 simulation runs i.e. each mean is the results of 50 x 13 x 10 runs. Figure 3 plots log word frequency against the number of steps required to reach a response threshold of $P(W|I) = 0.95$ for a single word. The function is very nearly a straight line, and the best fitting log function accounts for 0.99 of the variance. This simple model therefore automatically produces the logarithmic relation between frequency and both RT and identification threshold observed in the literature. Changes in the variance, or the response threshold, do not alter the form of the function. This result is consistent with Baum and Veeravalli's (1994) analysis of the MSPRT which showed that the expected time to reach a probability threshold should be a decreasing linear function of the logarithm of the prior probability of the hypothesis, with the slope depending inversely on the information (Kullback-Leibler) distance to the nearest competitor hypothesis.

INSERT FIGURE 3 ABOUT HERE

The data that these simulations most directly map onto are the identification threshold data from Howes and Solomon (1951) or King-Ellison and Jenkins (1954). There are no comparable data available from speeded response tasks, such as progressive demasking, that might be considered to be pure measures of the time taken to identify a word. Almost all of the available data comes from lexical decision or naming. It might appear that the model could be compared with data from word naming, such as the Balota and Spieler (1999) data set. However, as is widely recognized, naming generally involves contributions from both lexical information, and knowledge of spelling to sound correspondences. Lexical effects in naming are therefore likely to be greatly attenuated by the contribution of sublexical information. Indeed, unless special measures are taken to prevent the use of non-lexical pronunciation procedures, naming usually produces a much smaller frequency effect than tasks like lexical decision (Monsell et al., 1989). In fact, it is not even clear that word naming always involves identification of a unique word in the lexicon. In cascaded systems like the DRC, and in the multiple-levels model (Norris, 1994), sublexical information allows responses to be generated earlier than would be possible if responses had to await unique identification of the word. So, while the Bayesian Reader could be used to simulate the word recognition processes underlying naming (e.g. the lexical component of the multiple-levels model of Norris, 1994), it is difficult to extend the rational analysis or ideal observer approaches to all of the operations involved in reading aloud. The first problem is that it is hard to specify exactly what decision an ideal observer would need to make in order to pronounce a word. For example, one might specify that all of the phonemes in the pronunciation of a word had to be determined with a given level of accuracy before pronunciation. However, perhaps there is left-to-right processing (as in the DRC) such that pronunciation can be initiated as soon as the first one or two phonemes have been determined, even if later phonemes were still uncertain. Of course, it is quite possible that the phonological representations underlying pronunciation might not be phonemes at all. Finally, even if the nature of the phonological representations were known, the performance of an ideal observer would

depend on the relative time courses with which lexical and sublexical sources of phonological information become available. However, despite the difficulty of applying the Bayesian approach to reading aloud, it can readily be extended to the lexical decision task.

Lexical decision

So far, I have described the optimal procedure for word identification. This assumes that the task of the recognizer is to identify which word has been presented. Importantly, I have assumed that the input is always a word in the model's lexicon. This is clearly not the case in a lexical decision task. Generally, half of the stimuli in a lexical decision experiment will be unfamiliar nonwords. The task is no longer to identify which specific word has been presented (although human readers may well do this under some circumstances) but to decide whether the input is a known word or not. The primary interest is not in calculating $P(W|I)$, but $P(\text{a word rather than a nonword} | I)$, as in Equation 9.

$$P(\text{A Word} | \text{Input}) = \frac{P(\text{A Word}) \times P(I | \text{A Word})}{(P(\text{A Word}) \times P(I | \text{A Word})) + (P(\text{A Non-Word}) \times P(I | \text{A Non-Word}))} \quad (9)$$

It might seem that the simplest procedure for lexical decision would be to respond 'Yes' as soon as any word in the lexicon exceeds the recognition threshold. Although simple, such a procedure has a number of drawbacks. The first is that this procedure has no way of making a 'No' response to a nonword. One word will always reach the recognition threshold, even when a nonword is presented. As noted earlier, because of the properties of the normal distribution, eventually the probability of one word will always approach 1.0. The Bayesian $P(W|I)$ does not actually depend on a high value of $P(I|W)$, so long as $P(I|W)$ for all other words is sufficiently low. $P(W|I)$ reflects the relative probability of words in the lexicon, under the assumption that the input is indeed a word. In lexical decision, this is clearly not a valid assumption. An alternative procedure might be

to use the normal recognition mechanism to identify the best matching word, and then to perform a spelling check to determine whether the input really does correspond to that word (cf. Balota & Chumbley, 1984). This is a strategy that participants might adopt if they are being very cautious; however, is it not the most efficient way to use the available information.

Although lexical decision seems like an artificial laboratory task, the ability to distinguish between familiar words and unfamiliar letter strings is also an essential part of the normal reading process (Chaffin, Morris, & Seely, 2001). Readers will often encounter unfamiliar words, and these should be recognized as being unknown words rather than simply being identified as a token of the nearest known word. It turns out that the problems of identifying unknown words and performing lexical decision have the same solution. That is, lexical decision can be based on information that has to be computed anyway in the course of normal reading.

Some of the critical differences between identification and lexical decision can be appreciated by considering how lexical decision might be performed if somehow there was a fixed set of nonwords with known locations that could appear in an experiment. Assume that the nonwords in the experiment are equiprobable (i.e. all $P(NW_i)$ are the same). Because words and nonwords are equally likely to appear in a lexical decision experiment, word and nonword probabilities must each sum to 0.5. For any input I , we can then calculate the likelihood of that input, given that the input was produced by a word (Equation 10), and the likelihood that the input was produced by a

$$f(I | word) = \sum_{i=0}^{i=n} (P(W_i) \times f(I | W_i)) \quad (10)$$

$$f(I | nonword) = \sum_{i=0}^{i=n} (P(NW_i) \times f(I | NW_i)) \quad (11)$$

$$f(I) = f(I | word) + f(I | non - word) \quad (12)$$

$$P(word | I) = f(I | word) / f(I) \quad (13)$$

$$P(nonword | I) = f(I | nonword) / f(I) \quad (14)$$

nonword (Equation 11).

Equation 12 gives the likelihood of observing this particular input. Equation 13 gives the probability that the input was generated by presentation of a word, and Equation 14 gives the probability that the input was generated by a nonword. Informally, these equations indicate that, of all of the ways that the input I could have been produced, some proportion comes from words, and some from nonwords. If most of the ways that the input could have been produced are from words, then the input is likely to have been a word. The important feature to note about these equations is that $P(\text{word}|I)$ depends on summing the contribution of the likelihoods of all individual words, and does not depend of a high value of $P(W_i | I)$ for any particular word. That is, the appropriate procedure for performing lexical decision does not depend on first recognising any specific word.

This example illustrates how to discriminate between two sets of inputs (words and nonwords) where the items in the sets are known in advance. However, the set of nonwords used in a real lexical decision experiment is generally not known in advance. In lexical decision experiments the nonwords are usually constructed according to some quite general rule. For example, nonwords might be one letter different from words, be pronounceable letter strings, pseudohomophones, or be letter strings matched on bigram frequencies. The exact form of the nonwords generated by an experimenter might be influenced by implicit rules that even they themselves are unaware of. Nevertheless, a participant (or ideal observer) needs to have some way of representing members of both the word and nonword categories in order to construct a decision rule to discriminate between them. Part of the participants' task in lexical decision is therefore to construct a model of the nonwords used in the experiment. However, even an ideal observer will be unable to anticipate the exact set of nonwords used in an experiment. A model of the nonwords will therefore inevitably be based on a heuristic approximation. Fortunately, it is possible to achieve high levels of

performance in lexical decision simply by estimating how close words and nonwords are to each other in lexical space.

If nonwords are formed by changing a single letter in a real word (e.g. BLINT), they will be very close in space to real words, whereas consonant strings (e.g. CVXR), will be quite far from the nearest word. Assume for the moment that all nonwords in an experiment differ from real words by only a single letter. Effectively, the participant must decide whether the input is a real word, or a letter string that differs only slightly from a real word. The simplest way to do this is just to approximate $f(I|nonword)$ by a single point in space, where the point corresponds to a nonword one letter different from the nearest word. At least when using high response thresholds, the main influence on lexical decision will be from words and possible nonwords located very close to the input. When the SEM is small, most of the words and nonwords (or empty spaces in the lexicon) will have almost no influence on $P(I|word)$. Because of this, lexical decision can be performed simply by assuming that there is a single nonword positioned close to the input. But, where exactly should that nonword be placed?⁴

With the input representation used here, any two letter-strings that differ by a single letter will always be the same fixed distance apart in perceptual space (in fact, 1.41, i.e. $\sqrt{2}$). The estimate of how close nonwords are to words in an experiment will be referred to as the nonword distance, ND. In trying to discriminate between words and nonwords, the question then becomes one of whether the input is more likely to have been generated by a word, or by a nonword located near the input, but at least ND from the nearest word. That is, no nonword in the experiment should be closer than ND to a word. In general, if the input is estimated to be much closer than ND to a word, then the input is likely to have been generated by that word. If the input is much further away than ND from the nearest word, the input is more likely to have been produced by a nonword. The further the

input is from the nearest word, the less likely it is that the input is a word. In practice, an estimate of $f(I|\text{nonword})$ can be computed simply by assuming that there is a single virtual nonword located on a line connecting the input with the nearest word. This virtual nonword is placed as near to the input as possible on that line, but never closer than ND to the word. The value of ND should reflect the difficulty of the word-nonword discrimination in the experiment. If nonwords are very similar to words, ND should be small. If nonwords are very dissimilar to words (e.g. consonant strings) ND should be large. If the virtual nonword is set to have a frequency that is approximately the same as the average frequency of the words, it is now possible to calculate $f(I|\text{nonword})$ in exactly the same way as for a word (Equation 15). $P(\text{virtual nonword}|I)$ can then be calculated and used as a basis for lexical decision. If $P(\text{virtual nonword}|I)$ is $\gg 0.5$, respond 'No', if $P(\text{virtual nonword}|I)$ is $\ll 0.5$, respond 'Yes'.

$$P(\text{virtual nonword} | I) = f(I | \text{virtual nonword}) / (f(I | \text{word}) + f(I | \text{virtual nonword})) \quad (15)$$

In other words, if the input is more likely to have been generated by a word than the virtual nonword, respond 'Yes', otherwise respond 'No'. This procedure works because, late on in processing, the virtual nonword is usually only competing with one or two nearby words. Setting the nonword frequency to the average word frequency makes the competition between words and the virtual nonword roughly even. Although this procedure can reliably perform lexical decision, and also gives a logarithmic relation between frequency and RT, it is not ideal. The main problem is that it takes no account of the contribution that other potential nonwords (i.e. nonwords of the sort that might appear in the experiment) should make to $f(I|\text{nonword})$. This causes the model to have a strong bias to respond 'Yes' if it is required to make decisions based on very little evidence (e.g. speeded responding, or the response-signal procedure studied by Wagenmakers et al., 2004).

When the SEM is large, many words in a large subvolume of perceptual space will all make some contribution to $f(I|\text{word})$. The contribution of each word will be multiplied by its frequency, and the combined influence of all of these words can overwhelm $f(I|\text{nonword})$. As the model is forced to respond faster $P(\text{nonword}|I)$ decreases. That is, with fast responses the model has a bias to respond 'Yes'. What is needed is some way of representing the influence of the whole range of potential nonwords that could have appeared in an experiment. I will call these nonwords *background nonwords*. There is no need for these background nonwords to be explicitly enumerated, or added to the lexicon. All that is required is to make allowance for the fact that the experiment might contain nonwords distributed throughout lexical space, in addition to the virtual nonword. All that we need to know about background nonwords is their distance from the input. For example, if we wish to consider the influence of a nonword some distance away from the current estimate of the mean co-ordinates of the input, it really makes little difference exactly where in space that nonword is located (e.g. at a position that would correspond to a pronounceable nonword, or perhaps a string of consonants). The only information that goes into Bayes' equation is the distance between the nonword and the current estimate of the input co-ordinates. Furthermore, most of the contribution to $P(I|\text{a nonword})$ is from locations very close to the input.

Because the exact location of background nonwords does not matter, it is possible to make the simplifying assumption that nonwords are distributed homogeneously in lexical space. This ignores the fact that, if nonwords are chosen to be pronounceable, both words and nonwords will tend to be clumped together in small regions. However, this simplification has little effect on the behavior of the model, as these possible nonwords only exist to balance the probabilities of words and nonwords. Their exact location in perceptual space is of little consequence because they simply represent the fact that there is a possibility that the experiment might contain nonwords that are located in perceptual space some distance from the current estimate of the input co-ordinates.

Given these assumptions, we can calculate the likelihoods of background nonwords at various distances from the input, and use these to model the contribution of the entire set of possible nonwords. The details of the procedure for estimating the influence of background nonwords will depend on the form of input representations used. The present simulations simply take advantage of the fact that all letter strings differing from any particular letter string by a given number of letters will lie at the same distance, and therefore all have the same $f(I|NW)$. For example, for a 4-letter word we can calculate the number of strings differing by 1,2,3, and 4 letters, and multiply those numbers by the $f(I|NW)$ at the corresponding distance. The probability of these nonwords is then set to sum to $0.5 - P(\text{virtual nonword})$. Multiplying this probability by the sum of the $f(I|NW)$, then gives an estimate of the contribution of the background nonwords. The calculations are given in *Appendix B*, along with a description of an alternative procedure that can be used if there is an intermediate letter level.

As already noted, this simple model of background nonwords takes no account of factors such as whether or not nonwords might be pronounceable. Also, all regions of space have the same nonword density, regardless of how many words are in that region. Furthermore, because the distances to background nonwords are measured relative to the input, there is no guarantee that they will correspond exactly to real letter strings. Although it is possible to develop a more elaborate model of background nonwords, the method described here does what is required to roughly balance the overall word and nonword probabilities early in processing when there is still ambiguity in the input. Indeed, simulations using slightly different models of background nonwords produce the same pattern of results. When the SEM is large, probabilities will now be influenced both by words that are some way away from the input word, and also by distant nonwords. If a very high probability threshold is used, it makes little difference whether the background nonwords are included. With high thresholds, only words and nonwords very close to the input have any

influence on the probabilities. The addition of the background nonwords adds hardly anything to the computational complexity of the model as it only adds one term for each letter in the input string. Most of the computation still involves calculation of word probabilities.

The procedure for modeling $f(I|\text{nonword})$ is exactly what is required to deal with unknown words in normal reading. We need to have some idea of what the input from an unknown word would look like. As in the case of lexical decision, the most important information is that an unknown word will differ from any known word by at least one letter, and that unknown words could be located almost anywhere in perceptual space. An important difference is clearly that the probability of encountering an unknown word in normal text is going to be very much lower than in a lexical decision experiment. However, the main difference between normal reading and lexical decision is in how the available information is used. In normal reading, the critical probabilities for word identification are the $P(W_i|I)$. $P(\text{nonword}|I)$ is the probability that the input is a novel word. In lexical decision, 'No' responses can also be based on $P(\text{nonword}|I)$, but 'Yes' responses should be based on $P(\text{a word}|I)$ (or, equivalently, $1-P(\text{nonword}|I)$), and not on the probabilities of individual words ($P(W_i|I)$). So, although this method of making lexical decisions might seem complicated, something very much like it is essential for normal reading. Furthermore, because most of the work is being done by the virtual nonword nearest to the input, there is no need to have a very precise model of $f(I|\text{nonword})$. The interesting implication of this analysis is that lexical decision does not require any specialized information or procedures that are not already required for normal reading.

Simulation of the frequency effect in lexical decision

In all simulations ND is set to 2.0, which is the distance between two letter strings differing by 2 letters. The inclusion of background nonwords actually means that there is some small contribution from background words closer than 2.0, so the effective ND is actually slightly smaller than 2.0,

meaning that nonwords 1 letter different from words will tend to be correctly classified as nonwords. This gave good levels of accuracy for both words and nonwords in all of the simulations reported. In practice, an increase in response threshold can have the same effect as a decrease in ND. However, if ND bears no relation to the difficulty of the discrimination, the resulting value of $P(\text{a word})$ will no longer be a good estimate of $P(\text{a word})$.

Other than when demonstrating the effect of varying the response threshold, all simulations use a threshold of $P(\text{word}) > 0.95$, for 'yes' responses, and $P(\text{word}) < 0.05$ for 'no' responses. All simulation results represent the average of 50 simulation runs of each stimulus. The simulations therefore use no free parameters. However, in some instances an additional set of figures is also reported where two additional parameters are used to fit the model response times to actual RTs. One parameter represents time that human participants spend performing response execution and other processes outside of the scope of the model, and the other parameter specifies the correspondence between model time-steps and real time. These numbers are reported simply to make it easier to appreciate the exact relationship between the simulations and the data.

Figure 4 shows the results of a simulation of lexical decision using exactly the same materials as the earlier simulation of identification. As in the case of identification, the function relating RT to log frequency is close to linear, with an R^2 of 0.99.

INSERT FIGURE 4 ABOUT HERE

In the next simulation, the model is compared to real data from the Balota, Cortese and Pilotti (1999) study of lexical decision which included almost all of the monosyllabic words in English. Table 1 shows the results of simulations of their 4-letter words for which CELEX frequencies were

available, using the data from their younger group of participants. Simulations of their 5-letter words produce similar results. The table shows correlations between model RT, data, both log and rank frequency, and neighborhood density (N). Correlations are shown for the model with recognition thresholds of both 0.95 and 0.99. As would be expected from the previous simulations, there is a very high correlation between simulated RT and both log and rank frequency. All three of these measures correlate to a similar degree with the data, although, consistent with Murray and Forster's (2004) data, rank frequency has a slightly higher correlation than the others. It is interesting to note that the correlation between model RT and N increases as the response threshold is decreased. This is a quite general property of the model, and is also seen in a simulation of data from Forster and Shen (1996) to be presented later.

INSERT TABLE 1 ABOUT HERE

The analyses performed by Balota et al. (2004) suggest that the CELEX database may not be the best predictor of lexical decision latency. However, for the present purposes we are comparing CELEX frequency measures (and rank frequency derived from CELEX) with the predictions of the model, which also incorporates CELEX frequencies. Using a different frequency measure might alter the correlations overall, but should not alter the intercorrelations between the different frequency-based predictor variables.

Neighborhood effects

So far we have seen that the Bayesian Reader gives a good explanation of the relationship between frequency and RT (or exposure duration) in both identification and lexical decision. The form of the frequency function, in both the model and the data, is the same in the two tasks. However, one factor, namely neighborhood density (usually defined in terms of Coltheart, Davelaar, Jonasson and

Besner's, N (1977), the number of words differing from a target word by exactly 1 letter), has generally been found to have opposite effects in lexical decision and identification (for reviews see Andrews, 1997 and Perea & Rosa, 2000). In identification tasks, words in dense neighborhoods are usually harder to identify (Grainger & Segui, 1990; Havens & Foote, 1963; Perea, Carreiras, & Grainger, in press; Perea & Pollatsek, 1998; Pollatsek, Perea, & Binder, 1999), whereas in lexical decision, such words are generally easier to classify (e.g. Andrews, 1989, 1992; Forster & Shen, 1996). Furthermore, in lexical decision, high N nonwords are harder to classify than are low N nonwords (e.g. Coltheart et al. 1977; Forster and Shen, 1996). Note that there are some exceptions to the general finding of facilitatory neighborhood effects on words in lexical decision. For example, the original Coltheart et al. study only found an effect of N on nonwords, but not on words. Johnson and Pugh (1994) found facilitatory effects of N on words when the nonwords were unpronounceable, but inhibitory effects when they were pronounceable. However, a facilitatory effect of N is by far the dominant pattern, especially in English (see Andrews, 1997). Some studies of perceptual identification have also reported facilitatory effects of density. For example, Sears, Lupker and Hino (1999) found facilitatory effects in perceptual identification when performance levels were low.

However, the most convincing evidence that having many neighbors produces an inhibitory effect on identification comes from a study by Pollatsek, Perea and Binder (1999). They demonstrated that words that showed a facilitatory effect of neighborhood density in lexical decision produced an inhibitory effect in reading, as measured by eye tracking. As eye tracking during reading probably represents the most 'ecologically valid' measure of word identification, this result cannot readily be attributed to unusual task demands or strategies.

It might seem that data from word naming (reading aloud) might also provide a useful source of information on neighborhood effects. Indeed, nonwords in dense neighborhoods have been found to be easier to name (McCann & Besner, 1987). However, this may well be due to phonological factors rather than the process of word identification itself. Words in dense neighborhoods are more likely to have common spelling to sound correspondences than words in sparse neighborhoods. For example, the multiple-levels model of reading aloud described by Norris (1994) produces facilitatory effects of neighborhood density on naming, even though lexical activation in that model is not directly influenced by orthographic similarity. Norris presented simulations of Andrews' (1989) lexical decision data and showed that the naming latencies produced by the model closely paralleled the data.

The fact that neighborhood density has the opposite effect in identification and lexical decision poses problems for existing models of word recognition. For example, the basic mechanism in interactive activation models predicts that words in dense neighborhoods should be harder to identify than words in sparse neighborhoods. Inhibition between similar words should make words with many neighbors harder to identify. The search model predicts that there should be an inhibitory neighborhood effect on nonwords, but no effect for words. This is because close matches are initially simply flagged for later verification. This means that responses to words are unaffected by the presence of neighbors. However, the flagged entries are checked before 'No' decisions are made, and this produces an inhibitory effect of neighborhood density on nonwords. In contrast, Sears, Hino and Lupker (1999) showed that both the Plaut et al. (1996) and Seidenberg and McClelland (1989) models predict that neighborhood effects should be facilitatory, at least for low frequency words. None of these models can readily account for the fact that effects of neighborhood density are inhibitory in tasks requiring perceptual identification, but facilitatory in lexical decision. In a further complication, words that show facilitatory effects of density in lexical

decision have been claimed to show no density effect in a semantic classification task (Forster & Shen, 1996; Forster & Hector, 2002), although Sears, Lupker and Hino (1999) reported facilitatory effects of density in semantic classification.

This variability poses problems for any simple notion that all tasks tap directly into the same underlying word recognition process. One way of accommodating these task differences is to assume that one kind of task is a true reflection of the underlying recognition process, while other tasks are contaminated by task-specific processes. In her review, Andrews (1997) concluded that it was the facilitatory effects in lexical decision that were driven by the recognition system, while identification tasks were contaminated by guessing. On the other hand, Jacobs and Grainger (1992) concluded that it was the inhibitory effect that was fundamental. In order to explain the facilitatory effect in lexical decision, Grainger and Jacobs (1996) developed the Multiple Read-Out Model (MROM). In MROM, lexical decision responses are based both on the activation of individual words (the M criterion) and on the \sum criterion, which is sensitive to the total amount of activity in the lexicon. Similar assumptions are made in the DRC model (Coltheart et al., 2001). Total activity will be greater for words in dense than in sparse neighborhoods. To the extent that responses are influenced by this criterion, there should be a facilitatory effect of neighborhood density in lexical decision. Note that lexical decision responses can not be based solely on the \sum criterion as high N nonwords can produce more lexical activation than low N words. Therefore, the \sum criterion cannot discriminate reliably between words and nonwords. Although MROM can simulate most of the data on neighborhood effects, Grainger and Jacobs offer no independent motivation for adding the \sum criterion to the model. The model does not explain why lexical decision should be based on multiple response criteria rather than simply on the activation of individual words. However, as we shall see, the Bayesian Reader necessarily predicts that neighborhood effects are inhibitory in identification, and facilitatory in lexical decision.

It should be apparent that, as applied to perceptual identification, the Bayesian Reader has to predict that words with many lexical neighbors will be harder to identify than words with few neighbors. For a given SEM, a nearby neighbor will have a much larger $P(I|W)$ than a distant neighbor. Close neighbors add to the denominator of Equation 2 and decrease the $P(W|I)$ of the target word. Words in dense neighborhoods will therefore take longer to identify than words in sparse neighborhoods. Although the model does not incorporate explicit inhibition between neighbors, neighbors do effectively compete with each other. For any input, the probability of the target word will be lower if it has more neighbors. The probability will be lower still if those neighbors are also high in frequency.

In lexical decision, on the other hand, the neighborhood density effect will be reversed. Words in dense neighborhoods will be classified more easily than words in sparse neighborhoods. This is because all words in the neighborhood of the input contribute to the likelihood that the target is a word. This clearly makes intuitive sense if one thinks of the lexicon as consisting of a space populated by words, and by 'holes' (to use the terminology of Treisman, 1978a, 1978b) that could be nonwords in an experiment. If the estimate of the input mean is located in a very dense neighborhood, the neighborhood will contain more words than holes, and the input is most likely to have been produced by a word. In a sparse neighborhood, the ratio of words to holes will be much smaller, and the input is more likely to correspond to a hole (nonword) (see Figure 5). Note that the bias to respond 'word' in dense neighborhoods will apply to both words and nonwords. Nonwords in dense neighborhoods will therefore be classified more slowly, as more perceptual evidence will be required to overcome the bias to respond 'word'.

INSERT FIGURE 5 ABOUT HERE

The behavior of the Bayesian Reader in an identification task follows inevitably from the assumptions about the form of the input available to the system. Although the procedure for making lexical decisions is more complex, its behavior also follows directly from the demands of performing the task in an optimal fashion. The change from an inhibitory effect of neighborhood density in identification, to a facilitatory one in lexical decision, is therefore a direct consequence of the requirement to perform the tasks optimally. If readers showed an inhibitory effect of neighborhood density in lexical decision, this would indicate that they were not performing the task in optimally. For example, a strategy of identifying the closest word to the input and then performing a spelling check would be an inefficient way to perform lexical decision, and would produce an inhibitory effect of neighborhood density. Unlike existing accounts of neighborhood effects, there is no need for *ad hoc* assumptions simply to make it fit the data: this is simply how an optimal decision-maker has to behave. The notion of task requirements here therefore differs somewhat from standard discussions of 'task demands' in word recognition. Task-specific processes are often invoked to explain the different patterns of results found in different tasks. For example, according to Grainger and Jacobs (1996), lexical decision imposes additional task demands that turn an underlying inhibitory effect of neighborhood density into a facilitatory one. According to Andrews (1997), the natural state of affairs is for neighborhood density to have a facilitatory effect and this is counteracted by the demands of perceptual identification tasks. On the grounds that they found no neighborhood effects on 'No' responses in a semantic classification task ("is this word an animal?"), Forster and Shen (1996) concluded that all neighborhood effects are due to task specific strategies. The common feature of all of these accounts is that they assume that there is an underlying recognition system with quite fixed properties (e.g. inhibitory vs. facilitatory effects of neighbors). In response to the demands of different tasks, the output of this recognition system can be used in different ways that can alter, or even reverse, the behavior of the system. The problem

with these appeals to task demands is that the task-specific mechanisms exist solely to explain data that is at variance with the predictions of the underlying model. In the Bayesian procedures developed here, there is no underlying default mechanism. All perceptual decisions require some task-specific computations. In the absence of such computations (and a clear specification of what decision is to be made), there simply is no behavior. Therefore neither inhibition nor facilitation are general properties of the system. This can be further illustrated by considering how other tasks should be performed.

Semantic classification

Forster and Shen (1996) reported that there were no systematic effects of neighborhood density on 'no' responses to words in a semantic classification task, even though the same words showed a facilitatory effect of N in lexical decision (although there was, in fact, a 20ms facilitatory effect of density in semantic classification in their experiment 5). Similar results were obtained by Forster and Hector (2002). Despite this result, there remains some uncertainty as to whether there is a genuine effect of neighborhood density on 'no' decisions in semantic classification. Sears, Lupker and Hino (1999) found facilitatory effects of N in both perceptual identification and semantic classification. Nevertheless, it is instructive to consider what pattern of data we would expect to see if participants in such tasks behaved optimally.

Let us examine how an ideal observer might perform a semantic classification task. Note that an ideal observer would be able to perform all necessary computations instantly, and would be limited only by the quality of the available data. In order to decide that 'bear' is an animal, it is necessary to distinguish 'bear' from neighboring words like 'pear'. Failure to discriminate between these alternatives could lead to an error because a 'pear' is not an animal. Now consider what should happen when the task is to decide whether 'bear' is a member of a different category, such as a

vehicle. At some stage in processing, when 'bear' and 'pear' are still hard to distinguish, it would be possible to determine that neither is a vehicle. In psychological terms, it may be possible to determine whether the words have appropriate semantic features. This would enable a 'no' response to be made without fully analysing the input. If the ideal observer is able to make this semantic assessment before word identification has been fully resolved, neighborhood effects in semantic categorization should be much smaller for 'no' responses than for 'yes' responses. Note that one would still expect to see effects of frequency, as frequency will still alter the probability of all candidate words. Indeed, Forster and Shen (1996) did find frequency effects in semantic categorization.

Forster and Hector (2002) have shown that neighborhood effects on nonwords also vary as a function of task demands. For example, whether or not a nonword has a close neighbor does not influence 'no' responses in a semantic categorization task, unless one of the neighbors is a member of the relevant category (e.g. 'turtle' when the category is 'animal'). This result would seem to have exactly the same explanation as Forster and Shen's (1996) data. If none of the neighbors is a category member, then those nonwords can be rejected without needing to perform accurate identification (i.e. to establish definitively that they are not specific words). If a nonword has a neighbor that is a category member, the nonword must be processed further to establish that the stimulus really is a nonword and not its neighbor. The explanation being proposed here has points in common with Forster and Hector's own account. They suggest that, each word is linked to broad semantic fields. When a nonword like 'turtle' is presented, the search process locates the entry for 'turtle'. As 'turtle' is linked to the semantic field for animal, 'turtle' is subject to a more detailed form check that takes extra time.

The idea that these data should be explained by assuming that semantic information becomes available before words are uniquely identified receives support from a recent study by Pecher, Zeelenberg and Wagenmakers (2005) using a semantic categorization task. When participants were required to judge whether a word like 'cat' was animate or not, they responded more quickly if they had previously seen an orthographically similar word from the same category ('rat') than a similar word from a different category ('mat'). Relative to a neutral baseline of words whose neighbors had not been seen, both the facilitatory effect of having seen a same-category word, and the inhibitory effect on having seen a different category word were significant. In another experiment, participants were faster at categorizing words with orthographic neighbors belonging to the same category than words whose neighbors were not in that category. In a related experiment, Rodd (2004) found that words like 'leopard' that had an animal neighbor ('leopard') were responded to more slowly than words like 'toffee' which had a neighbor that was not an animal ('coffee').

Depending on the task, neighborhood effects can be facilitatory, inhibitory, or non-existent. If one starts from the assumption that there is an underlying word recognition system with fixed properties, this pattern of data is difficult to explain without invoking special task-specific mechanisms. In general, task-specific mechanisms serve little purpose other than to alter the default behavior of the model to fit the data. For example, the extra decision criteria in Grainger and Jacobs' (1996) MROM model seem to have no adaptive function other than to produce facilitatory effects of neighborhood density in lexical decision. In contrast, the Bayesian decision procedures developed here are motivated solely by the requirement to perform the task in an optimal fashion. These procedures follow entirely from a rational analysis of the tasks, and have not been adapted to fit the data.

One might object that the rational analysis approach permits an unconstrained proliferation of task-specific decision rules. In part this is true, but the rules are determined by the task and not by the data. It seems likely that the empirical results themselves follow as a direct consequence of the adaptability of human decision making. There is effectively no limit on the number of different tasks that humans can perform, and each task must necessarily use slightly different decision rules. Given that human behavior in these word recognition tasks is in line with the predictions of the rational analysis, this tells us something very important about the flexibility of human decision making. Even with minimal training, humans can perform all manner of complex tasks. For example, participants in a lexical decision experiment are unlikely to have any prior experience in making word-nonword decisions, yet they adapt readily to the task. Furthermore, as argued here, and by Wagenmakers et al. (2004), they behave in a manner that is close to optimal. This ability to reconfigure cognitive processes adaptively in response to environmental demands is an essential part of human intelligence (Duncan, 2001). However, as was noted when developing Bayesian Reader lexical decision procedure, lexical decision should be based on the kind of information that must already be available to detect unknown words in normal reading. All that really changes between tasks is that, whereas responses in identification are based on $P(W|I)$, responses in lexical decision are based on $P(\text{nonword}|I)$. A similar case can be made that the optimal procedure for semantic categorization also capitalizes on information already required for normal reading. Norris (1986) suggested that, as a word is being recognized, it should be evaluated for its compatibility with the context, and that this was essential for disambiguation. Access to the semantic representation of a word before it has been uniquely identified is exactly what is required for the semantic classification strategy proposed here. Viewed in this way, it is perhaps not so surprising that readers can perform these tasks using what appears to be a near optimal strategy: all of the information required is already available from the processes involved in normal reading,

The next section investigates the behavior of the Bayesian Reader further by presenting simulations of a number of experiments studying the influence of word frequency and number of orthographic neighbors in the lexical decision task.

Simulations of neighborhood density effects.

A number of studies have investigated the detailed pattern of neighborhood effects seen in the lexical decision task. For example, Andrews (1989,1992) showed that there was an interaction between density and frequency, with a larger effect of density being observed for low-frequency than for high-frequency words. Forster and Shen (1996) parametrically manipulated number of neighbors (N) for both words and nonwords. They found that while increasing neighborhood density facilitated responses to words, it inhibited responses to nonwords (see also Coltheart et al., 1977). Siakaluk, Sears and Lupker (2002) reported a series of experiments where they examined the neighborhood effects for high and low frequency words while manipulating nonword N, so as to alter the difficulty of the word-nonword discrimination. They found a tendency for density to have less of an effect the more similar the nonwords were to words. In their experiment with the most word-like nonwords (1D) they found that high frequency words in sparse neighborhoods were actually responded to faster than high frequency words in dense neighborhoods.

In order to illustrate the similarity between the behavior of the recognizer and human data the remainder of this section reports simulations of the main lexical decision experiments in all four of these studies. Each simulation uses the model exactly as described above, apart from the fact that simulations of 4, 5 and 6 letter words had to be performed separately with lexicons consisting only of words of that length. When simulating experiments where different length words are used, the reported means are always unweighted means of the mean number of steps for words of each length, as this eliminates any confounds due to the model responding differently to different length

words. The simulations all use exactly the same materials as in the original experiments, apart from words that are not in the CELEX database.

Andrews (1989,1992)

The first experiments in both the Andrews (1989) and Andrews (1992) studies looked at lexical decision performance on words of high and low frequency that had either a large or small number of neighbors. Andrews found that there was an interaction between frequency and N such that the facilitatory effect of N was larger for low- than high-frequency words. The original data and the simulations are shown in Tables 2 and 3. The tables also show the adjusted model output where the raw output is transformed as described earlier. These numbers are presented to make it easier for the reader to appreciate the similarity between the pattern of results produced by the model and the human data. In both simulations, the model shows a facilitatory effect of neighborhood density and an interaction between frequency and density. The main divergence between the model and the data is that the model produces too big an effect of N on high frequency words in the 1989 simulations, but too small a effect in the 1992 data. Snodgrass and Mintzer (1993) used the Andrews (1989) items in a perceptual identification task. They found that the effect of N was inhibitory. Table 4 shows the result of a simulation of identification time rather than lexical decision time for the words from Andrews (1989). In line with Snodgrass and Mintzer's results, the facilitatory effect of N seen in lexical decision becomes an inhibitory effect when the model simulates identification.

INSERT TABLES 3 & 4 ABOUT HERE

Forster and Shen (1996)

Experiment 2 of Forster and Shen (1996) manipulated N for both words and nonwords in a lexical decision task. Simulations of this experiment are shown in Figure 6. Note that the simulation omits

the item 'flyer' as this does not appear in CELEX. In line with the data, the model produces a facilitatory effect of N on words, and an inhibitory effect on nonwords. The main discrepancy between the model and the data is that the model responds equally quickly to words and nonwords, whereas the data show the standard finding of slower responses to nonwords. Forster and Shen's participants had quite high error rates (up to 23.4% for nonwords with 3 or 4 neighbors). With the 0.95 threshold the model makes a maximum of only 2% errors in any one condition. Therefore Figure 5 also shows the model performance using a lower threshold of 0.75. In the case of the materials used in this experiment the maximum error rate of the model is 11% errors to the nonwords with 3 or 4 neighbors. With this lower threshold (and higher error rate) the model output looks more like the data, as the effect of neighborhood density for words and nonwords becomes more similar. As noted earlier, this tendency of the simulations to show larger effects of N as the response threshold is reduced is a very general characteristic of the model. With a threshold of 0.99 the model becomes almost completely insensitive to N.

The fact that word and nonword responses are equally fast is not an inevitable property of the model. As Figure 6 shows, reducing the threshold for nonwords from 0.95 to 0.75 would make word responses faster than nonword responses. If the nonword threshold is reduced too much, this could sometimes cause low frequency words to be misclassified. However, the maximum increase in word error rate from reducing the nonword threshold from 0.95 to 0.75 is less than 3% in any one condition. In many experiments it may not be appropriate to set word and nonword thresholds to the same value. In general, if the Bayesian Reader were set up to have a perfect model of the distribution of possible nonwords in an experiment, and all words had the same frequency, one would expect reaction times for words and nonwords to be the same if the same threshold was used for both. Words and nonwords would be equivalent sets of items contributing equally to $P(\text{Word})$. However, in practice, there will inevitably be some differences between words and nonwords.

Words will each have their own individual frequency, whereas nonwords will be assigned to a single frequency corresponding to the mean frequency of the words. Because roughly 2% of word types account for about 50% of all word tokens, the mean word frequency will be very much larger than the frequency of most words. As the model is currently set up, most words are effectively competing with a virtual nonword that is higher in frequency. For low frequency words, $P(\text{Word})$ can therefore drop very low before rising toward 1.0 as more samples arrive. If the nonword threshold is set too liberally, many low frequency words will be misclassified as nonwords. In contrast, a low word threshold does not lead to so many errors with nonwords. An ideal observer should therefore adapt the word and nonword thresholds to the conditions of each specific experiment to track the selected accuracy level. In practice, nonword thresholds are likely to have to be higher than word thresholds.

INSERT FIGURE 6 ABOUT HERE

Siakaluk, Sears and Lupker (2002)

Siakaluk, Sears and Lupker (2002) carried out a number of experiments investigating neighborhood effects, and the way they vary as a function of the nonwords used in the experiment. Experiments 1A, 1B, 1C, and 1D all used the same set of words, but with progressively more word-like nonwords. Nonwords in 1A had no neighbors, while those in 1D had large numbers of neighbors. Their data for all four experiments and the model simulations are shown in Table 5. The table shows both the raw model output and also the output adjusted to fit the data from Experiment 1A. The general pattern of the data is that N effects are marginal for high-frequency words, apart from in 1D where they were actually inhibitory. For low frequency words there is a facilitatory effect of N for all but the words in 1D with high-frequency neighbors. Overall, words with a high-frequency

neighbor tend to be faster than words with no high-frequency neighbor. Once again the model output corresponds closely to the experimental data, and captures the interaction between N and word frequency and the effect of high-frequency neighbors. The most problematic aspect of these data for the Bayesian Reader is that there is an 11.5ms inhibitory effect of neighborhood density for high-frequency words in experiment 1D. This difference was significant by subjects, but not items. Other things being equal, the Bayesian Reader should always produce a facilitatory effect of neighborhood density in the lexical decision task. However, whether or not people produce this pattern of data depends, of course, on whether they behave as optimal Bayesian recognizers. As noted earlier, it is quite possible that people might adopt a suboptimal strategy such as identifying the nearest word and then performing a spelling check. If this were the case (and identification were based on the procedure described here) then we would expect to see the same inhibitory effect on N found in identification. Interestingly, this possibility might provide an explanation for the different effect of N seen in older and younger individuals. Balota et al. (2004) found that N had a smaller effect in older than younger participants, Perhaps older participants are more likely to adopt the suboptimal strategy of first identifying the word. This is consistent with the suggestion of Balota et al. (2004) that older participants may be less likely to engage the specific task demands of the lexical decision task.

INSERT TABLES 5 & 6 ABOUT HERE

Experiment 2 of the Siakaluk et al. (2002) paper examined the effect of having a single high-frequency neighbor. Investigations of the effect of the frequency of lexical neighbors have produced inconsistent results (see the reviews by Andrews, 1997, and by Perea & Rosa, 2000). While some studies have found facilitatory effects of the presence of high frequency neighbors (Forster & Shen, 1996; Sears, Hino, & Lupker, 1995), others have found that high frequency

neighbors inhibit lexical decision (Carreiras, Perea, & Grainger, 1997; Grainger, O'Regan, Jacobs, & Segui, 1992; Grainger & Segui, 1990; Huntsman & Lima, 1996; Perea & Pollatsek, 1998).

While simulations cannot resolve the empirical debate, it is informative to simulate a study of neighborhood frequency effects to illustrate the behavior of the model. One merit of the Siakaluk et al. (2002) study is that they manipulated both the presence of high-frequency neighbors, and N. Given that the major theoretical issue is whether, as predicted by MROM, high-frequency neighbors and N have different effects, it is hard to draw any conclusions where there is no clear effect of N in an experiment.

Experiment 2A in Siakaluk et al. (2002) used words and nonwords with small neighborhoods, and Experiment 2B used stimuli with large neighborhoods. The data and simulations are shown in Table 5. In line with the data, the model shows that the effect of having a single high-frequency neighbor is facilitatory. Note that, whereas the simulations produce faster RTs in Experiment 2B than in Experiment 2A, the human RTs were faster in Experiment 2A. This is most likely because the nonwords in Experiment 2B also had more neighbors. This is likely to require a more conservative threshold to produce similar error rates in the two experiments. Siakaluk et al. attempted to simulate this data in MROM. They report that they were unable to simulate the data from Experiment 2A, and the only way they could produce a facilitatory neighborhood effect in Experiment 2B was by using parameters that produced a nonword error rate of 21.6%, compared to 6% in the data. When the probability threshold for a 'No' response is set to 0.95, the Bayesian Reader makes no errors to any of the nonwords used in the Siakaluk et al. experiments. However, errors do start to appear as the threshold is lowered beyond this.

This illustrates an interesting difference between the Bayesian Reader and MROM. In MROM, a single high-frequency neighbor will have a strong inhibitory effect on the target word, and it is this inhibitory effect that dominates the model's behavior, rather than the increase in global activation that produces N effects. In the Bayesian Reader there is no inhibition between words, and lexical decision is facilitated by the presence of neighbors, regardless of their frequency.

Summary of simulations

The simulations reported here are intended to give an indication of the behavior of the Bayesian Reader when words vary in frequency and in similarity to other words. The first simulations showed that both lexical decision and identification time in the model are an approximately logarithmic function of word frequency. Simulations of the experiments by Andrews (1989, 1992), Forster and Shen (1996), and Siakaluk et al. (2002) all showed that the model produces the expected facilitatory effect of N on lexical decision. As shown in Table 4, the facilitatory effect of N in lexical decision becomes an inhibitory effect when the model is required to identify the same items. Simulations of the Andrews and Siakaluk et al. experiments also showed that the effect of N is greater for low- than for high-frequency words. The model also successfully simulates the inhibitory effect of N on nonwords in a lexical decision task (Forster & Shen). Finally, the model correctly predicts that the presence of a single high-frequency neighbor should have a facilitatory effect on words in lexical decision (Siakaluk et al.). Overall, the performance of the model shows a remarkable parallel to that of human readers in all of the experiments simulated, even when considering only those simulations using the standard threshold of 0.95. There has been no need to perform an extensive search through parameter space to get the model to fit the data. In part this is because the only parameters available to manipulate are the nonword distance (ND) and the response threshold. Furthermore, these two parameters have very similar effects. Decreasing ND or increasing the threshold both reduce the model's sensitivity to N. This can be seen in the

simulations of the data from Balota, et al. (1999), and from Forster and Shen. However, the main reason for the model's insensitivity to parameter settings is that the behavior of the model observed in these simulations follows automatically from the demands of optimal decision making.

Interpretation of Bayesian probabilities

Although Bayes' theorem delivers an estimate of the posterior probability for each word, these probabilities need to be interpreted with some caution. The probabilities will only be accurate to the extent that the underlying assumptions are satisfied. For example, in the characterisation of the identification task, the probabilities depend on the input corresponding to a known word, and on the frequency of the word being accurately represented. In the lexical decision simulations, probabilities also depend critically on the accuracy of the parameter representing the similarity of words and nonwords, and on assumptions about distribution of words and nonwords in perceptual space. As already pointed out, prior estimates of word frequency may not accurately reflect the probabilities of the words actually presented in an experiment. In practice then, experimental participants are likely to have to modify their response thresholds during the course of an experiment in order to achieve the desired speed-accuracy trade-off. In the context of the present simulations, this means that setting the response threshold to say, 0.95, will be unlikely to produce an error rate of exactly 5%. Of course, in contrast to human participants, the Bayesian Reader never presses the wrong button by accident.

Limitations of the model

The simulations reported here are intended to demonstrate the general characteristics of a Bayesian recognizer. No attempt has been made to incorporate extra assumptions or mechanisms into the model to help make the model give a better fit to the data. Primarily this is because the goal of the enterprise has been to establish whether many of the standard empirical findings can be explained

purely on the basis of a rational analysis of the tasks involved. Any move to incorporate additional assumptions would increase the complexity of the model and make it harder to determine the true cause of the model's behavior. However, having established that the model does give a principled explanation of a wide range of data, it is possible to consider how it might be developed further. One of the major limitations of the model is in the form of the input representations. The assumptions that all letters are equally confusable, and that letter representations are position-specific, are unlikely to be an accurate characterisation of human perception (Chambers, 1979; O'Connor & Forster, 1981; Perea & Lupker, 2003, 2004). However, experimenting with alternative input representations that take some account of letter similarity made little difference to the behavior of the model. Even incorporating a separate level of letter representations does not alter the behaviour of the model. Arguably, discovering the perceptual representation of orthographic form is likely to remain one of the main goals of reading research for some time to come.

As it stands, the Bayesian Reader has access to only two representations, the written forms of words, and their frequencies. Human readers have access to a great wealth of knowledge about words. For example, they know their pronunciation, their meaning, and whether they are likely in the current context. All of these factors influence ease of word recognition (see Balota et al., 2004 for a recent review, and also Balota, 1990). These factors could also potentially influence an optimal Bayesian recognizer. Although lexical decision can only be performed reliably on the basis of orthographic information, recognition could be influenced by phonological information. If a phonological representation corresponding to a word becomes available during recognition, this may bias the recognizer toward a 'Yes' response. The extent of any such bias might depend, for example, on whether or not the nonwords in an experiment were pseudohomophones or not (Coltheart et al., 1977; Rubenstein, Lewis, & Rubenstein, 1971). Finally, semantic information may correlate with the kind of contexts words appear in (cf. McDonald & Shillcock, 2001) which, in

turn, may alter the expectation that different classes of word (e.g. abstract/concrete) may appear in isolation in a psychological experiment.

As already noted, $P(W)$ should really be an estimate of the probability that a word will appear in the current context, so one would expect recognition to be influenced by contextual factors. The Bayesian account is neutral as to the exact mechanism responsible for calculating prior probabilities. For example, contextual probabilities could be modulated by spreading activation (Collins & Loftus, 1975; McNamara, 1992) compound cueing (Ratcliff & McKoon, 1988), or plausibility checking (Norris, 1986). Although context must be mediated by its influence on the Bayesian prior probabilities, this does not mean that the priors have to be determined ‘prior’ to the presentation of the target word. Priors simply need to be available at the point when the Bayesian computations are performed. In the Bayesian Reader these computations are being performed continuously so, for example, the priors could be updated as a result of a plausibility checking process taking place during the course of recognition. In contrast to spreading activation models, probabilities should go down as well as up. That is, words that are unlikely in a particular context should have $P(W)$ reduced, and therefore should be inhibited (for review of context effects in reading see Neely, 1991). According to the Bayesian analysis, inhibition, like frequency, is an important factor in helping recognition. Reducing the probability of contextually improbable words will benefit the recognition of contextually probable words by decreasing the effective competition. Importantly, facilitation and inhibition (increasing and decreasing prior probabilities) should act as an integrated process. This would seem to be more compatible with a unitary account like plausibility checking than two-process accounts involving facilitation by spreading activation, combined with attention-based inhibition.

The simulations presented here have all focussed on mean RT. In recent years, a number of researchers have also performed analyses of RT distributions in lexical decision and naming (Andrews & Heathcote, 2001; Balota & Spieler, 1999; Grainger & Jacobs, 1996; Ratcliff et al., 2004). Because the simulations presented here include noise in the sampling process, we can also measure the distributions of the model RTs. However, a consistent feature of these simulations is that RTs have very small variance. For example, in a simulation comparing two sets of 20 words, one with a Kucera and Francis (1967) frequency of approximately 100, and one with a frequency of approximately 5 (the same frequency range used by Balota and Spieler (1999), the model produces RT distributions that have almost no overlap. Given that the model simulates an ideal observer, and that the only source of noise is in the sampling process, this is perhaps not too surprising. Even if the human perceptual system did approximate an ideal observer, it is highly implausible that there would be no additional sources of noise in decision or response execution processes. For example, there might be trial-to-trial variability in the variance of the input noise, the threshold value, or ND. Note that Wagenmakers et al. (2004) had to add noise to their model in the form of trial-to-trial variation in the effective start of feature activation, in order to simulate the detailed pattern of speed-accuracy trade off that they observed. Variability in the decision criteria also plays an important role in MROM.

I have not yet investigated the effects of adding noise elsewhere. Partly this is because the model is already computationally expensive, and collecting sufficient data to examine changes in response distributions while varying several noise parameters would be an exceedingly slow process.

Although I have argued that the assumption that human readers approximate optimal Bayesian decision-makers goes a long way to explaining the general character of human visual word recognition, it would seem unlikely that humans can ever be completely optimal in their behavior.

As noted above, there are almost certainly significant uncontrolled sources of noise in the human perceptual system. Furthermore, not all participants in an experiment will be guaranteed to home in on an optimal strategy. For example, Balota et al. (2004) suggested that their older participants might not engage fully with the demands of the lexical decision task. Some readers might adopt a cautious strategy of first identifying the closest word, and then checking to see whether the word is spelled correctly. The issue of optimality in the early perceptual stages of reading has been investigated by Pelli, et al. (2003). They presented evidence that, under some circumstances at least, word recognition is in fact suboptimal. They investigated whether readers were able to make full use of all of the visual information in a word, or were dependent on identifying individual letters in the word. Using a perceptual identification task they found that readers were not able to integrate all of the available visual information. An ideal observer should be able to recognize all words equally well so long as they have the same overall contrast energy. That is, longer words should require less contrast energy per letter than shorter words. However, Pelli et al. found that identification performance was determined by the contrast energy of the individual letters. Readers therefore behaved as though each letter had to be identified separately, rather than being able to identify words as wholes.

Pelli et al. (2003) simulated their data using a two-stage model in which the first stage outputs a single hypothesis as to the identity of each letter. The second stage then uses knowledge of the confusions produced at the first stage to produce a best guess as to the identity of the word. Their simulations were based on empirically derived letter-confusion matrices at each of three contrast levels. For each input letter the confusion matrix gives the probability that the letter will be identified as any particular output letter. They assume that, given a particular letter as input, the first stage of recognition emits an output letter in accordance with the probabilities from the confusion matrices. The second stage has knowledge of these confusion probabilities. That is, for

each letter it knows $P(\text{letter produced by first stage} \mid \text{true identity of input letter})$. The probabilities $P(\text{Word} \mid \text{Input})$ can be computed from the products of the letter probabilities. Combined with knowledge of $P(\text{Word})$, Bayes' theorem can then be used to compute $P(\text{Word} \mid \text{Input})$. This is exactly how the Bayesian Reader operates when it is implemented with an intermediate level of letter representations.

A focus of some concern in the word recognition literature has been the effect of list composition. The main empirical finding is that responses are faster in lists consisting of only easy words, than when lists contain both easy and hard words together. Responses to hard words tend to be faster in mixed lists than in pure lists of hard words. This effect is found both for naming (Lupker, Brown, & Colombo, 1997; Rastle, Kinoshita, Lupker, & Coltheart, 2003; Taylor & Lupker, 2001) and lexical decision (Dorfman & Glanzer, 1988; Glanzer & Ehrenreich, 1979; Gordon, 1983; Perea et al., in press; Rastle et al.) and also for tasks not involving word recognition (Strayer & Kraymer, 1994). The standard explanations for list composition effects are in terms of adjustments in response criteria (Grainger & Jacobs, 1996; Lupker et al., 1997; Perea et al., ; Taylor & Lupker). For example, criterion bias models would explain these data by assuming that participants can set a lower threshold when lists contain only high-frequency words. When there are no low-frequency words, a more liberal threshold can be set for responding 'No', leading to faster 'No' responses in pure high-frequency lists. Both of these options are clearly available to the Bayesian Reader, and an ideal observer would be expected to adjust its decision criteria to optimize performance. However, Mozer et al. (2004) reported that they were unable to simulate the full pattern of list composition effects by a simple manipulation of response thresholds. Instead, they developed a Bayesian explanation for how participants adapt to list composition, which is compatible with the Bayesian Reader. They assume that a participant's placement of a response threshold is influenced by what they refer to as the *Historical Accuracy Trace*. In terms of the Bayesian Reader, this can be

thought of as the average function relating $P(\text{a word})$ to time over recent trials. This is combined with the *Current Activity Trace* (CAT, equivalent to the growth of $P(\text{a word})$ over time in the Bayesian Reader) to generate a weighted *Mean Activity Trace* (MAT). Responses are determined by a threshold placed on the MAT rather than the CAT itself, and this is what causes the list composition effect.

If a high-frequency word appears in a list of mixed high and low-frequency words, the MAT for the word will under-estimate the accuracy expected for a given level of $P(\text{a word})$, and responses will be slower than they would have been in a pure high-frequency list. Conversely, if a low-frequency word appears in a mixed list, the MAT will over-estimate the accuracy expected for a given level of $P(\text{a word})$ and responses will speed up relative to what would be expected in a list of pure high-frequency words.

Comparison with other models

REM-LD

The Bayesian heritage of the model places it in the same lineage as the REM-LD model of Wagenmakers et al. (2004). Both models depend on accumulating evidence over time, and on making a Bayesian decision as to whether the input is a word or a nonword. However, the models differ in the form of the input representations, the details of the evidence accumulation, and their accounts of frequency and neighborhood density. Furthermore, whereas the Bayesian Reader was

developed to establish the general characteristics of a Bayesian recognizer in a range of tasks, REM-LD was designed to provide a detailed fit to results from experiments using the response-signal lexical decision task. Consequently, REM-LD requires many more parameters than the Bayesian Reader.

In the response-signal task, participants are required to make a response at a particular point in time, rather than when they are confident in making a correct decision. As Wagenmakers et al. (2004) note, performance in this task is difficult to explain with models such as MROM where there are multiple response criteria (see below). However, Bayesian models are ideally suited to this kind of task as they generate a single output representing the probability that the stimulus is a word. When participants are signalled to respond they simply need to determine whether $P(\text{a word})$ is greater or less than 0.5.

A significant difference between the models is in the form that the input takes. In REM-LD, words are coded in terms of a set of 30 discrete features that are sampled without replacement. Each input feature has some probability of matching the corresponding feature of a word, given that the word contains that feature (*beta 1*), and some probability that the probe will match a word, given that the word does not contain that feature (*beta 2*). These probabilities place a fixed limit on the amount of perceptual evidence that can be extracted from the input. The vectors that represent words in their simulations are generated according to these probabilities, and do not represent the actual form of real words. *beta 1* and *beta 2* are free parameters in the model. They effectively determine how reliably the input features help to discriminate between alternative words and, in that respect, they are broadly analogous to the distances between words in the present model. In REM-LD the probabilities are set for the entire lexicon, meaning that all words in any single simulation effectively have the same neighborhood density. REM-LD is therefore only used to simulate the

behavior of sets of words with general properties, and does not simulate responses to specific words.

Wagenmakers et al. (2004) reported simulations of a response-signal experiment using nonwords varying in word-likeness. They simulated responses to 4-letter nonwords that differed from words either by a single letter, or by 2 letters. Although nonwords differing from words by two letters clearly must have an N of 0, nonwords differing from words by a single letter could have any value of N between 1 and 25, so this manipulation is rather different from standard manipulations of N . Word-likeness was simulated by lowering the probability that a feature in a word would match the input given that the feature was not actually in the input for less word-like nonwords. In effect, this makes less word-like nonwords easier to discriminate from words.

However, this way of simulating word-likeness seems difficult to motivate given that both words and nonwords are comprised of the same set of features. This implies that a given feature has one probability of incorrectly matching a word when the feature comes from presentation of a word-like nonword or a word, and a different probability when the feature is in a less word-like nonword. In other words, identical features must be marked according to which type of stimulus they belong to. Another significant difference between the models is that in REM-LD the rate of accumulation is not-linear, but is a negatively accelerated function that is determined by a rate parameter, b .

Whereas the central focus of the present work is on the proper treatment of word frequency, Wagenmakers et al. (2004) do not commit themselves to any particular explanation of the word frequency effect. For the purposes of their simulations, they assume that frequency alters probability of a match between a probe (input) feature and a feature in a word, given that the word and probe feature are the same. The main advantage of the current model over REM-LD is

therefore that the current model gives a natural and completely Bayesian account of word frequency. Furthermore, by using a form of input that can represent individual words, the Bayesian account automatically explains neighborhood effects without any additional assumptions or parameters.

ACT-R

Van Rijn and Anderson (2003) presented a model of speeded lexical decision in the ACT-R (Anderson, 1993) framework. They simulated the same response-signal data as Wagenmakers et al. (2004). This model has the advantage that it uses a standard lexicon, and frequency is a function of each word's CELEX frequency. However, the input to the model is log frequency rather than linear frequency. The growth of word activation over time in the model is handled by a "competitive latency mechanism" whereby time to retrieve chunk depends on its own activation relative to the activation of other chunks. This clearly has parallels with Bayes' equation, the Luce choice rule, and the FLMP. However, the basic principles of the model do not explain why the more word-like nonwords should be harder to classify than the less word-like nonwords (as they were in the Wagenmakers et al., 2004 data). In order to get the model to simulate this aspect of the data, high word-like nonwords are assumed to retrieve more words that are similar to the nonwords than are less word-like nonwords.

In common with REM-LD then, the ACT-R account of lexical decision has no principled reason for the particular choice of word frequency mechanism. Although both ACT-R and REM-LD have an intuitively reasonable account of the influence of word-likeness on nonword responses, the account does not follow necessarily from the basic assumptions of the models.

Search models

The search model (Forster, 1976; Murray & Forster, 2004) shares only one feature with the Bayesian Reader: both predict that word recognition RT should be a roughly logarithmic function of frequency. Murray and Forster have argued that rank position in a frequency ordered lexicon is a better predictor of RT than is log frequency. However, as can be seen from the correlations presented in Table 1 the differences between the two measures are very small.

A limitation of the search model is that the function relating frequency to RT and error rate is the only prediction that follows directly from the basic search model. All other predictions require additional assumptions. For example, facilitatory effects of neighborhood density in the lexical decision task are attributed to an unspecified bias in the decision process (Forster & Shen, 1996). One might argue that the Bayesian Reader also has a bias to respond positively to more word-like stimuli. However, this bias is an integral and unavoidable part of the process of making optimal decisions.

Ratcliff, Gomez and McKoon (2004)

Ratcliff, et al. (2004) have recently applied Ratcliff's (1978) diffusion model to the lexical decision task. In their account of lexical decision, a diffusion process is driven by an input corresponding to a value on a word-nonword continuum. The drift rate of the diffusion process is determined by the value on the word-nonword continuum. On the basis of an analysis of RT distributions, Ratcliff et al. concluded that the effect of word frequency in lexical decision can be characterized solely by a change in the drift rate of the diffusion process. This has intriguing implications for the nature of the word recognition process. It implies that word frequency has no effect at all on the duration of word recognition itself. Apart from some variation due to noise, the word recognition process must

always take the same amount of time to complete, regardless of word frequency. Once a word has been recognized, the recognizer outputs a value on the word-nonword continuum as a signal to commence the diffusion process. This ensures that, although the word recognition system is able to identify high and low frequency words equally quickly, the diffusion process is told to respond more slowly to low frequency words. In effect, the explanation for the word frequency effect in lexical decision is that, for some unspecified reason, low frequency words are responded to more slowly than high frequency words. In the absence of any reason why frequency should modulate drift rate, the diffusion account suffers from the same problem as the other models reviewed here. Having a lower drift rate for low-frequency words is suboptimal, because recognition would be faster if low-frequency words had the same drift rate as high-frequency words. Part of the problem with the diffusion model may be that it assumes a constant drift rate. The probabilities in the Bayesian Reader can also be considered to be a random walk, with probabilities starting near $P(\text{Word}) = 0.5$, and drifting toward an upper boundary determined by the threshold for a word response, and a lower boundary determined by the threshold for a nonword response. However, in the Bayesian Reader the drift rate (expected change in probability per step) changes over time. More importantly, the random walk begins at the very beginning of the word recognition process, not the end. The change in drift rate both over time, and as a function of frequency, reflects the dynamics of the word recognition process itself, and is not simply an uninteresting side-effect of a decision process.

Jacobs and Grainger's MROM

MROM probably simulates the broadest range of data from perceptual identification and lexical decision tasks of any current model. The underlying mechanism of MROM is an interactive activation model similar to that of Rumelhart and McClelland (1982). However, in order to simulate lexical decision data, the interactive activation model has to be supplemented with extra

decision criteria (the Σ and T criteria). MROM simulations use three response criteria. The M criterion is sensitive to the activation in individual word units, the Σ criterion is sensitive to the total level of activation in the lexicon, and the T criterion is a time deadline for making ‘No’ responses. During processing, these criteria can be altered on the basis of the overall lexical activation. If overall activation is high after seven processing cycles, T is increased and Σ is decreased. Because overall activation will tend to be higher for high N words and nonwords, this will tend to make responses to high N words faster than to low N words, while making responses to high N nonwords slower than to low N nonwords. By adding these two new criteria to the basic interactive activation model, MROM is able to simulate both the facilitatory effect of neighborhood density on words in the lexical decision task and the inhibitory effect of density on nonwords. The DRC model of (Coltheart et al., 2001) explains lexical decision in the same way as MROM. However, what is lacking in either MROM or DRC is an independent motivation for why lexical decision should be influenced by overall lexical activation. As Grainger and Jacobs themselves note (Grainger & Jacobs, 1996, pp 559 & 565), the total level of lexical activation is not a reliable cue to lexical decision. This means that placing too much reliance on overall lexical activation can make the model produce a very high error rate. For example, Siakaluk et al. (Siakaluk et al., 2002) found that MROM could not simulate their neighborhood data from words without producing far too many errors on nonwords (Siakaluk et al., , p 677). The Bayesian Reader was able to simulate these data while still classifying nonwords accurately. Grainger and Jacobs motivate MROM as extending the scope of interactive activation models, and providing an integrated account of identification and lexical decision. However, the central problem with the model is that the only reason for using the Σ criteria on summed activation is to transform the inhibitory effect of N seen in identification into the facilitatory effect found in lexical decision. Interestingly, the Bayesian Reader has to compute a measure of summed ‘activation’ ($P(I)$) in order to identify words at all. This reason for this can be seen both in the basic Bayes’ equation (Equation 2) and in Figure 1. In

fact, if decisions are made very slowly it is possible to make completely accurate lexical decisions based solely on $P(I)$. After a sufficiently large number of samples, $P(I)$ should approach 1.0 for words, and 0.0 for nonwords. However, using a threshold on $P(I)$ is a very suboptimal decision rule because $P(I)$ can be quite large early on in processing, decrease, and then increase again. Although the Bayesian approach might seem to provide some justification for the use of summed activation in MROM, it could only be fully justified if MROM were also modified to deliver an estimate of $P(I)$ based on the $P(I|W)$ s. As $P(I|W)$ is not available from the interactive activation network in MROM, the entire network would have to be replaced with a network able to compute the relevant probabilities. This would effectively transform MROM into a notational variant of the Bayesian Reader.

Wagenmakers et al. (2004) raised a number of criticisms of MROM with regard to its ability to simulate data from the response-signal paradigm. For example, it is not clear how the model could be modified to generate a response before any of its response criteria were exceeded. In fact, this is a problem for all activation models. How can an activation model decide whether a stimulus is a word or nonword in a response-signal task where the criteria cannot be reliably set before stimulus presentation? From the onset of stimulus presentation both the activation of individual words and the overall lexical activation will grow. If a response is required before the criteria are reached, how is a decision to be made? The only principled way to make a decision is to determine whether the amount of activation is more or less than would be expected at that time if the target were a word. This implies that the model needs to know how the values of the response criteria should change over time. In contrast, in both REM-LD and the Bayesian Reader, there is a single probability measure that can be continuously monitored. At any point in time it is possible to make a lexical decision based on whether that single value of $P(\text{a word})$ is above or below 0.5. For the same

reason, a Bayesian mechanism doesn't need a response deadline criterion for making 'no' responses, as all that is required is to monitor $P(\text{a word})$.

The Neighborhood Activation Model (Luce & Pisoni, 1998)

Luce and Pisoni's Neighborhood Activation Model (NAM) of spoken word recognition has some important similarities with the Bayesian Reader. In their model, estimates of the confusibility of words are derived empirically by measuring identification of words in noise. This allows them to derive an estimate of the probability that listeners will respond with a given word, given a particular phoneme sequence as input. Using the Luce (1959) choice rule, the product of these probabilities and word frequency is then used to derive an estimate of the activation of each word. As already noted, the Luce choice rule and Bayes' theorem have a similar form. Luce and Pisoni found that NAM was a better predictor of lexical decision latency than was log frequency. Interestingly, in contrast to the standard finding with visual lexical decision, the effect of neighborhood density in their study was inhibitory. This difference between visual and spoken word recognition could well be attributable to the serial nature of speech processing. However, setting aside the fact that it is a model of spoken rather than visual word recognition, NAM has two main limitations. The first is that it only gives a single static estimate of word activation at one point in time, and does not simulate RT directly. The second is that both lexical decision and identification are modelled in exactly the same way.

Discussion

The present paper has developed a model of word recognition based on a rational analysis of the tasks of visual word identification and lexical decision. The rational analysis rests on some very simple assumptions about the nature of lexical representations of word form, and the kind of input

available to a recognizer. In fact, the assumptions are so general that exactly the same principles could be used to model recognition of other stimuli such as faces or objects, and a similar pattern of results should emerge. Words are assumed to be represented as points in a multidimensional space. Input is assumed to accumulate over time by sampling from a noisy distribution centered on the point in space corresponding to the string of letters presented to the recognizer. These assumptions then entirely determine the behavior of an ideal observer performing word identification or lexical decision. The Bayesian Reader was shown to simulate some of the most important empirical findings in the literature on visual word recognition. The recognizer produces a logarithmic function relating word frequency to recognition time, and accounts for the different patterns of orthographic neighborhood effects seen in identification and lexical decision tasks.

The major achievement of the Bayesian Reader is that it explains why a recognition system should be influenced by word frequency and neighborhood density. As noted in the *Introduction*, although existing models can simulate the effects of word frequency, what they fail to do is offer an explanation of why there should be an effect of word frequency at all. Similarly, while some models can simulate the change in the neighborhood density effect between identification and lexical decision tasks, only the Bayesian Reader explains why it is that performance in these tasks should differ in this way.

The rational analysis approach adopted here started by developing an analysis of the task, which effectively involves construction of what (Marr, 1982) terms a "computational theory". A computational theory concerns itself with the nature of the computations that must be performed to carry out the task of interest, and does not necessarily make any claims about human behavior. In contrast, the starting point for most psychological models of visual word recognition is the choice of a mechanism, such as search, logogens, or interactive activation networks. Having

selected a mechanism, the model is then modified as required to make it fit the data. This is exemplified quite clearly in the development of MROM. By beginning with an interactive activation network that produces inhibitory effects of neighborhood density, Granger and Jacobs (1996) were forced to modify their model by adding extra decision criteria that would make it consistent with fact that density effects are facilitatory in lexical decision. These moves enable the model to simulate the data but, unlike the Bayesian Reader, they don't explain why the pattern of neighborhood effects should vary according to the task. The accounts of word frequency offered by most theories suffer from a similar problem. As was argued in the *Introduction*, sensitivity to word frequency generally emerges as a consequence of some arbitrary or maladaptive property of the recognition system. For example, frequency effects might be due to frequency dependent recognition thresholds or resting levels, as in the logogen model or interactive-activation models. However, thresholds could equally well be set to produce a reverse frequency effect. If frequency effects are a consequence of learning in connectionist networks, as suggested by Monsell (1991), then they are an unfortunate side effect of an inefficient learning mechanism. In contrast, according to the Bayesian account offered here, sensitivity to word frequency is an entirely beneficial and adaptive property of the recognition system.

One of the most important features of the Bayesian Reader is that its behavior follows entirely from the requirement to make optimal decisions based on the available information. Given the initial assumptions, and a specification of the decision to be made, there is only one way the model can behave. The model has not been influenced or changed in any way by consideration of the data. Although it may seem strange to extol the virtues of a model that has not been influenced by the data, the strength of this approach is that it offers an explanation for why people behave the way they do. To the extent that people behave in line with the predictions of the Bayesian Reader, this suggests that people are recognising words in an optimal manner. The explanation for why people

are sensitive to word frequency and neighborhood density is therefore that this is exactly how they should behave if they are to perform recognition optimally.

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Footnotes

1. See Geisler & Kersten (2002) , for a brief introduction to the concept of the ideal observer, or Geisler (2003), for a slightly more formal treatment.
2. The program used to run the simulations (Windows version only) is available to download from <http://www.mrc-cbu.cam.ac.uk/~dennis/BayesianReader>
3. CELEX frequencies were obtained by collapsing over all occurrences of each orthographic form in the CELEX database.
4. Under some circumstances it would also be possible perform lexical decision by assuming that the lexicon contained a single ‘nonword’ for which $f(I/nonword)$ had the same small value regardless of the perceptual input. A lexical decision response could be based on the corresponding $P(nonword/I)$. If the input is a non-word, all $f(I/W_i)$ will approach zero, and $P(nonword/I)$ will therefore approach 1.0. If the input is a word, $f(I/nonword)$ will be small relative to the sum of $P(W_i/input) \times P(W_i)$, and $P(nonword/I)$ will approach zero.

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Appendix A

Calculation of word probabilities in the Bayesian Reader

Representation of words

In the representations used in the simulations reported here, each word is represented by a vector of 26 L, where L is the number of letters in the word. In the sub-vectors corresponding to each letter one coordinate is set to 1 and the remainder to 0. Each word vector therefore corresponds to a point in multidimensional perceptual space.

Note that, in principle, the coordinates of the vector can take any real value and can be altered to reflect, for example, letter similarity.

The input letter-string is also represented by a vector of the same form. The input vector therefore also corresponds to a point in perceptual space.

Sampling

Each sample presented to the model is derived by adding zero-mean Gaussian noise, with standard-deviation σ , to each coordinate of the vector corresponding to the input letter string. After each new sample s_i is received, the mean μ of the sample vectors is updated. This is simply calculated by computing the average value of the sample vectors for each coordinate.

The standard error of the mean (SEM) is then computed. The SEM is calculated on the basis of the distance between the location of each sample vector and the location of the mean of the sample vectors. Distance between two vectors (the sample mean μ and a sample vector s_i) is given by

$$d_i = \|\mu - s_i\| = \sqrt{\sum_{i=1}^{i=n} (\mu_i - s_{ii})^2} \quad (16)$$

where μ is the vector corresponding to the mean, and s_i is the sample vector and n is the number of coordinates of the vector (number of dimensions/features) which, in the present case is 26 L.

The SEM is calculated in the usual way, where SEM is

$$\sigma_M = \sigma / \sqrt{N} \quad (17)$$

And sigma is given by :

$$\sigma^2 = \sum_{i=1}^{i=N} d_i^2 / nN \quad (18)$$

where N is the sample size.

Calculating the likelihood of each word:

$$f(\mu | W_i) = (2\pi\sigma_M^2)^{-n/2} e^{-D_i^2/2\sigma_M^2} \quad (19)$$

The likelihood of each word is then calculated on the basis of the distance, $D_i = \|\mu - W_i\|$ between each word and the mean of the input sample. The likelihood, $f(\mu|W_i)$, is then given by

Note that as all the calculations are based on likelihood ratios, the first term of equation (19) can be ignored, as it is common to all likelihoods, giving:

$$f(\mu | W_i) = e^{-D_i^2/2\sigma_M^2} \quad (20)$$

This, in conjunction with word frequency ($P(W_x)$), provides the values used to compute $P(W_x/\mu)$ for every word,

$$P(W_x | \mu) = P(W_x) \times f(\mu | W_x) / \sum_{i=1}^{i=m} (P(W_i) \times f(\mu | W_i)) \quad (21)$$

where m is the number of words in the lexicon. Equation 21 is equivalent to Equation 8 in the main text.

In lexical decision the same equations govern the calculation of nonword probabilities. When the task is identification, a response is generated when $P(W/I)$ for one word exceeds the specified response threshold.

Appendix B

Calculation of Background Nonwords

The likelihood or probability of background nonwords can be calculated, or approximated, in different ways, depending on the nature of the input representations. The current simulations take advantage of the fact that, given the form of the letter representations used, all letter strings must fall at one of a set of discrete distances from the closest letter string to the mean, as determined by the number of letters they share with that string. For any given letter string there will be background nonwords that differ by one letter, by two letters, and so on up to the length of the letter string. For any word or letter string of a given length we can calculate how many other letter strings can be constructed that differ by 1, 2 or more letters. For example, for 4-letter strings the figures are 100, 3750, 62500 and 390625, for 1, 2, 3 and 4 letters. We also know how far these letter strings are from a given letter-string. These distances are 1.41, 2.0, 2.45 and 2.83 respectively (i.e. $\sqrt{2}$, $\sqrt{4}$, $\sqrt{6}$, $\sqrt{8}$). So, for any input I, we can use this information to estimate the contribution of the background nonwords to $P(I | \text{a nonword})$, and hence calculate $P(\text{a word} | I)$, as in Equation 8. Note that the frequency of the virtual nonword is still set to be the same as the average word frequency, and the full set of background nonwords plus the virtual nonword are set to have the same total frequency as the words. That is, overall, both the words and the nonwords have a summed probability of 0.5. The relevant calculations are shown in Equations 22-25.

$$f(I | \text{background nonwords}) = f(I | \text{BNW}) = \sum_{i=1}^{i=L} f(I | D_i) \times N_i \quad (22)$$

$$P(\text{virtual nonword}) = P(\text{VNW}) = 0.5 / m \quad (23)$$

$$P(\text{background nonword}) = P(\text{BNW}) = 0.5 / (26^L - 1) \times (m - 1) / m \quad (24)$$

$$P(a \text{ word} | I) = f(I | a \text{ word}) / (f(I | a \text{ word}) + P(\text{VNW}) \times f(I | \text{VNW}) + P(\text{BNW}) \times f(I | \text{BNW})) \quad (25)$$

Where m is the number of words in the lexicon, 26 is the number of letters in the alphabet, L is the word length, D_i is the distance from the input mean of nonwords differing by i letters, and N_i is the number of nonwords at that distance. As already noted, this simple model of background nonwords takes no account of factors such as whether or not nonwords might be pronounceable. Also, all regions of space have the same nonword density, regardless of how many words are in that region. In fact, because the number of possible letter strings is so much greater than the number of actual words, no correction is made for the fact that some very small proportion of possible strings will correspond to words. Furthermore, because the distances to background nonwords are measured relative to the input, there is no guarantee that they will correspond exactly to real letter strings. Although it is possible to develop a more elaborate model of background nonwords, the method described here does what is required to roughly balance the overall word and nonword probabilities early in processing when there is still ambiguity in the input. Indeed, simulations using slightly different models of background nonwords produce the same pattern of results. When the SEM is large, probabilities will now be influenced both by words that are some way away from the input word, and also by distant nonwords. If a very high probability threshold is used, it makes little difference whether the background words are included. With high thresholds, only words and nonwords very close to the input have any influence on the probabilities. The addition of the background nonwords adds hardly anything to the computational complexity of the model as it only

adds one term for each letter in the input string. Most of the computation still involves calculation of word probabilities.

If words are recognized via letter representations (as in Pelli et al., 2003), the background likelihood can be calculated in a different way. $P(I|W)$ is given by the product of $P(I|\text{letter})$ for each letter in the word. An estimate of the likelihood of the background nonwords can then be

based on $1 - \sum_{i=0}^{i=n} P(I | Letter_i)$. This is effectively the summed likelihood of each string of letters

that is not one of the words. The virtual nonword itself is simply the most likely sequence of letters that is not a word (this is similar to setting ND to a distance corresponding to a letter). A version of the model using an intermediate letter level has been implemented and produces results very similar to the simulations reported here.

The procedure used for letters illustrates that, even when not using a letter level, the influence of background nonwords can be estimated by working in the domain of probabilities rather than likelihoods. In the same way that the probability of all letter strings must sum to 1, the area under the pdf of the normal distribution (as in Figure 1) must also sum to 1. Combined with the assumption that no nonword can be closer than ND to the nearest word, this provides another way of obtaining an approximate correction for background nonwords. If the distribution is centered on the word closest to the mean, there is a slice $2 \times ND$ wide in the center of the pdf that cannot contain a nonword. With a large SEM this slice will correspond to only a small proportion of the total probability. If the input is a word, and the SEM is very small, this slice will encompass the entire pdf. The contribution of background nonwords will therefore be proportional to the area in the tails of the normal distribution beyond $\pm ND$, which depends only on the SEM. It would also

be possible to take this further and subtract out an estimate of the probability of the words derived from their likelihoods.

Figure Captions

Figure 1

Figure 1 shows the input distributions centered on two words at times early and late in processing. The figure is based on the assumption that there is a lexicon of only two words, and that they vary on a single perceptual dimension. Each curve represents the probability density function of $f(I|W_i)$ for a given SEM. Calculations are based on likelihoods, that is, on the heights of the probability density function of each word at the point corresponding to the mean of the input distribution. The lower two curves correspond to a time early in processing where the SEM is large. The ratio of the heights of these two curves at the position corresponding to the mean of the input distribution is close to one, indicating that both words have similar probabilities. For the upper curves, corresponding to later in processing, the SEM is much smaller. As $P(\text{input}|\text{word 2})$ is now almost zero, $P(\text{word 1}|\text{input})$ will be close to 1.0. Note that, at any given time, the pdfs for all words are the same as they are all determined by the SEM of the input distribution.

Figure 2

Possible network implementation of the Bayesian Reader. Not all connections shown. See text for details.

Figure 3

Simulation of the effect of word-frequency in word identification. Model RT (in samples) is plotted against log word frequency. The figure also shows the regression line for RT against log frequency. This line accounts for 99% of the variance.

Figure 4

Simulation of the effect of word-frequency in lexical decision. Model RT (in samples) is plotted against log word frequency. The figure also shows the regression line for RT against log frequency. This line accounts for 99% of the variance.

Figure 5

Illustration of areas of high- and low neighborhood density in perceptual space. Each dot represents a word, and 'High' and 'Low' correspond to the centers of input distributions in high and low density areas of lexical space respectively.

Figure 6 a-c

Data and simulations of Forster and Shen showing data (in Milliseconds) and simulated lexical decision times to words and nonwords as a function of number of neighbors. Panel *a* shows the data, Panel *b* shows simulated RT with a response threshold of 0.95, and Panel *c* shows simulations with a threshold of 0.75.

Figure 1

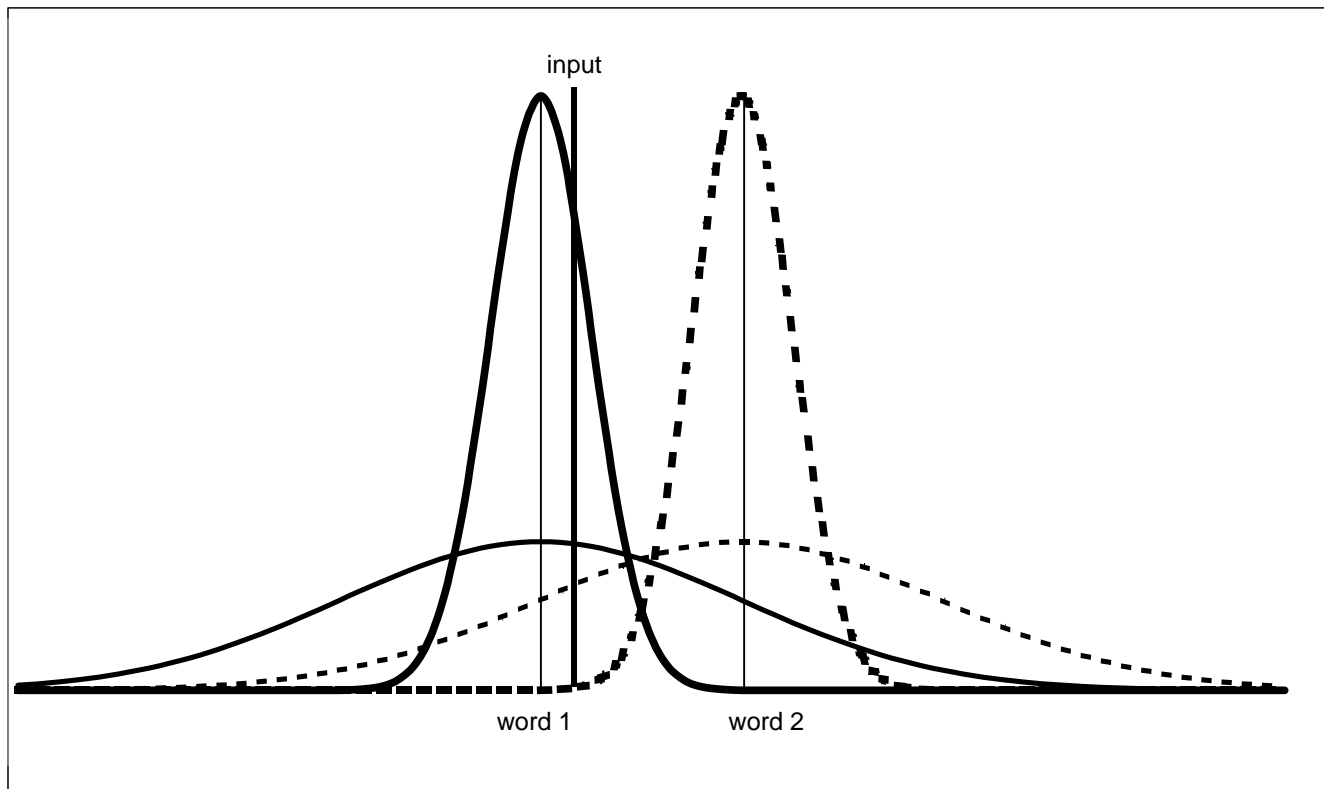


Figure 2.

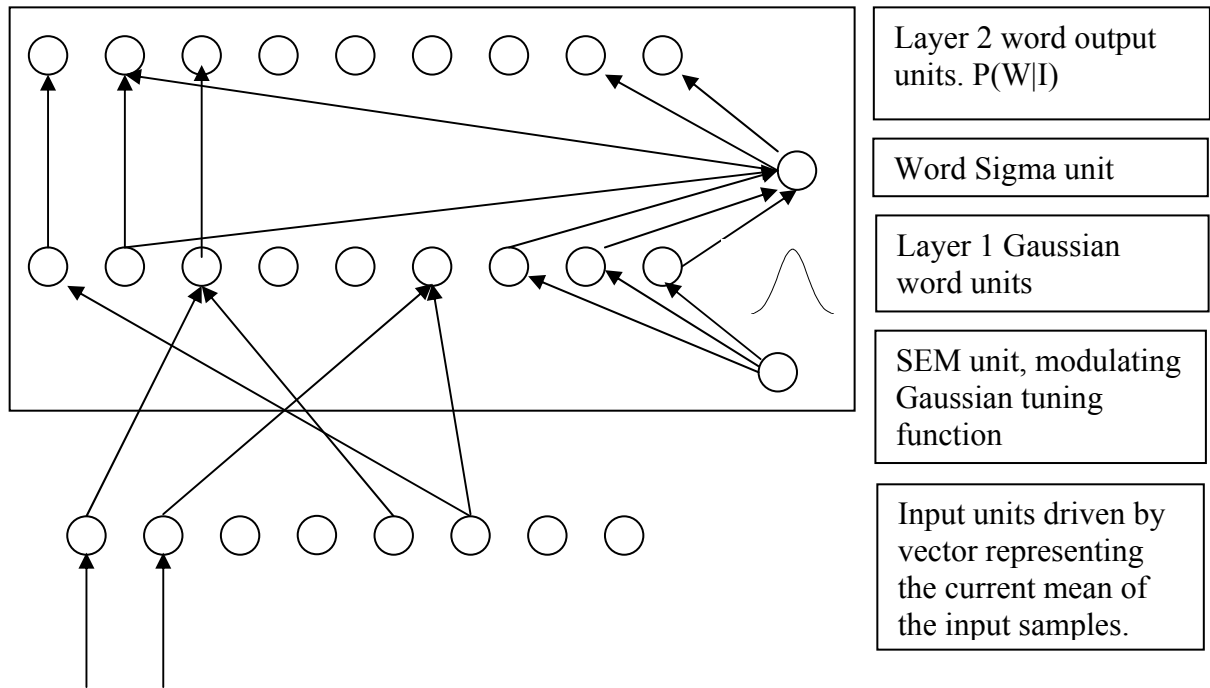


Figure 3

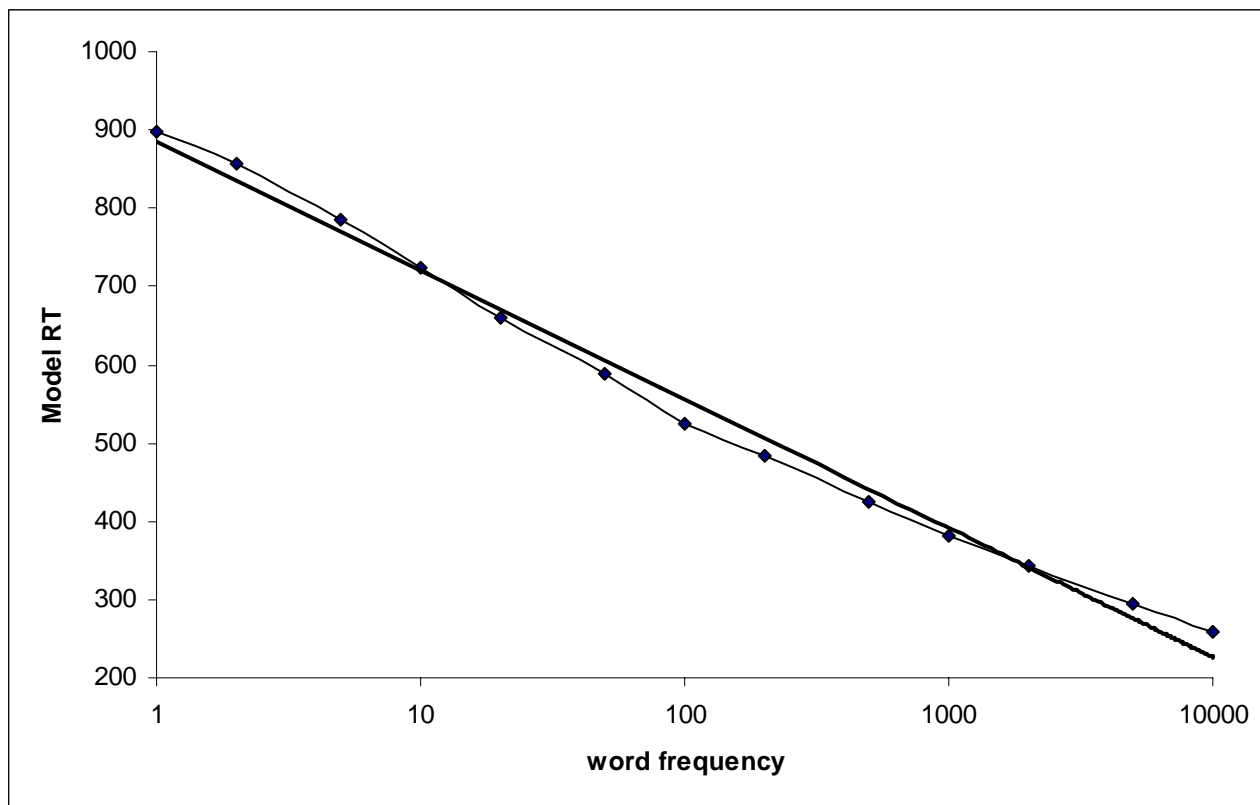


Figure 4

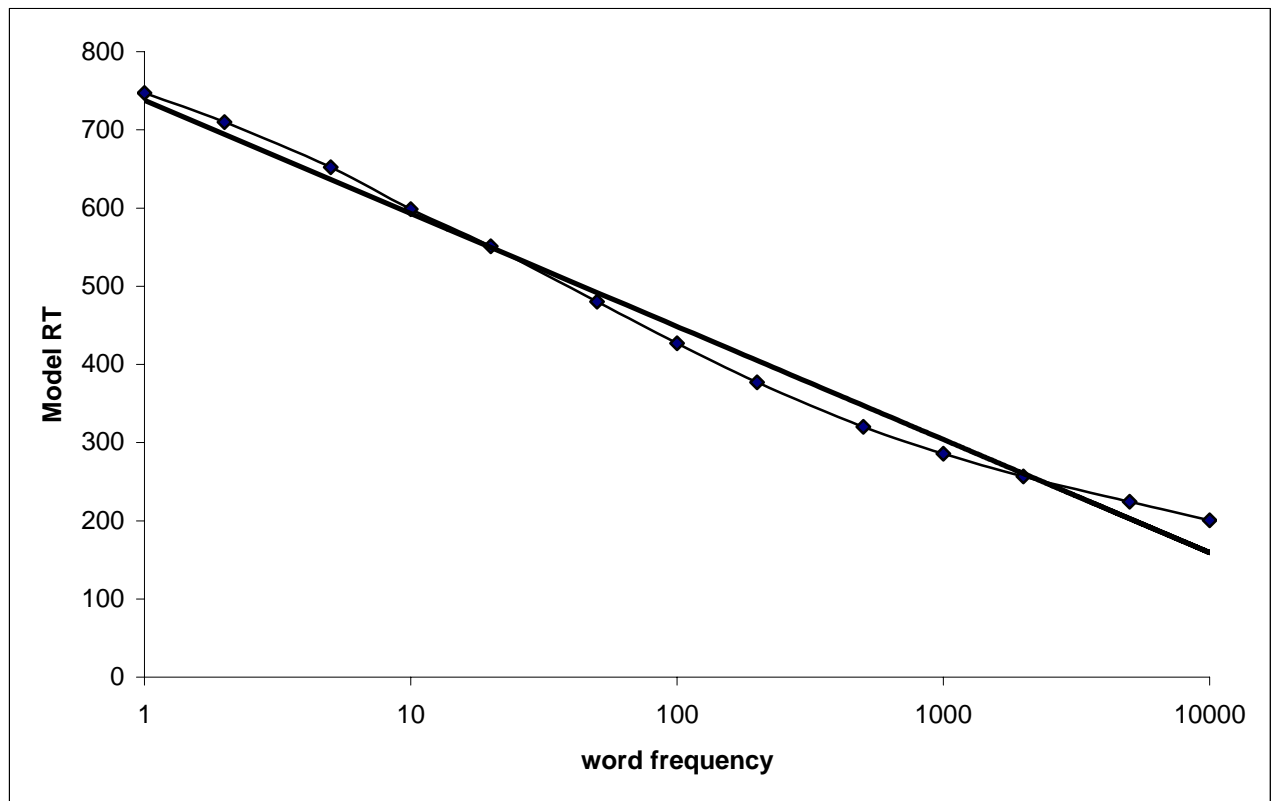


Figure 5

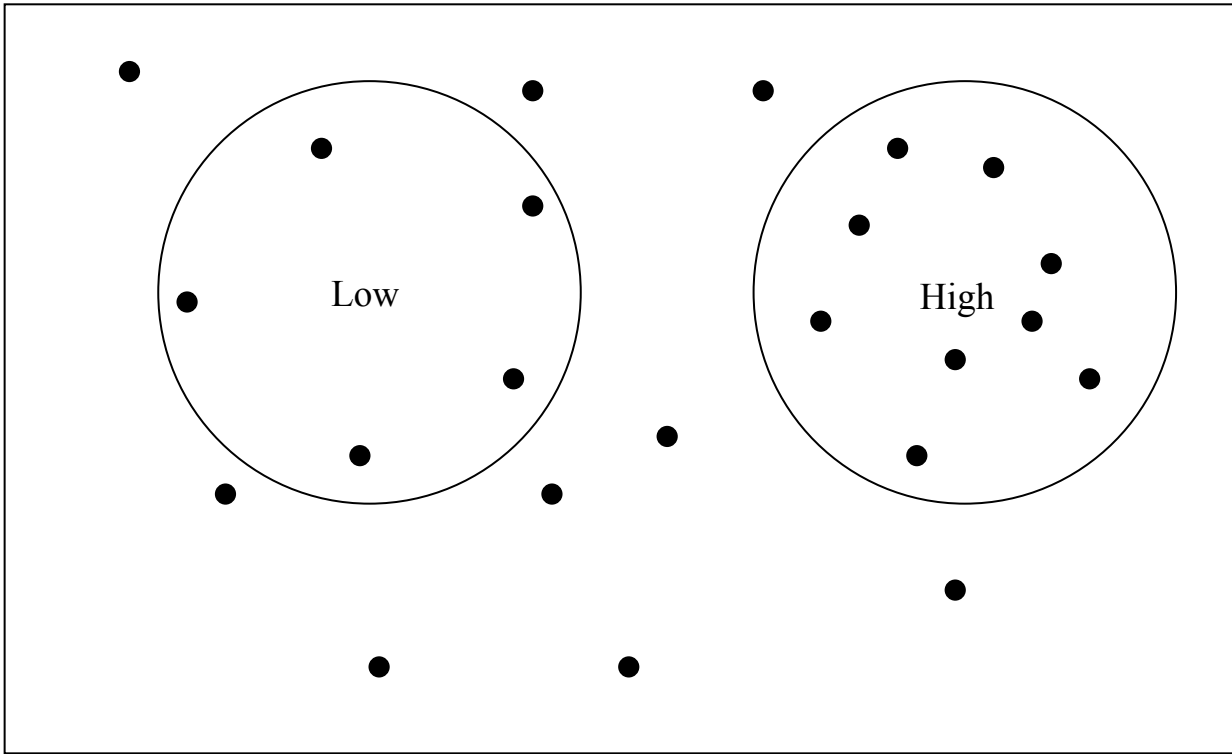


Figure 6 a-c

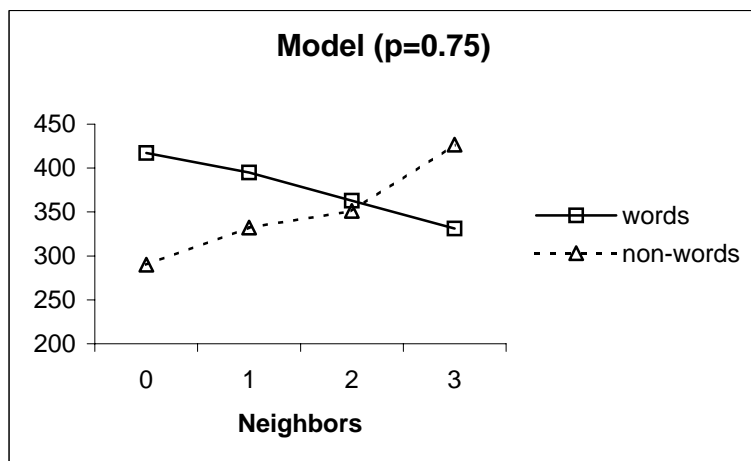
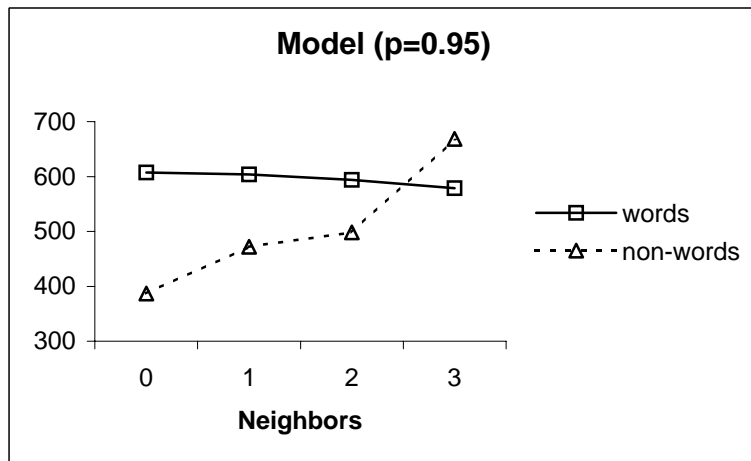
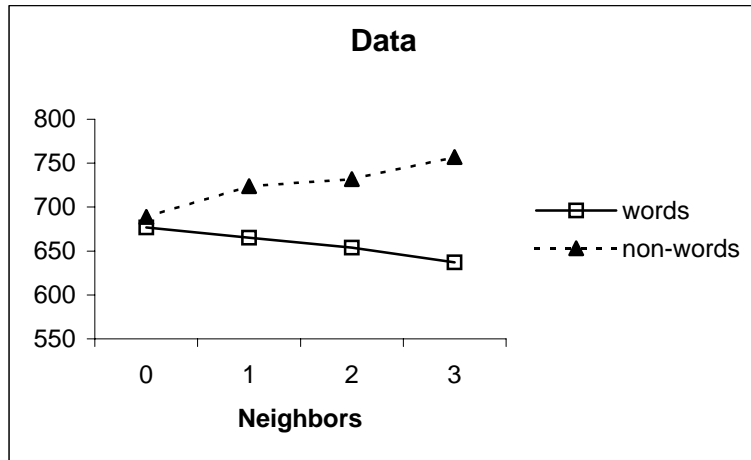


Table 1

| | | | | |
|--------------|-------|----------|-------|-------|
| RT | 1.0 | | | |
| Log freq | -0.60 | 1.0 | | |
| Rank freq | -0.63 | -0.89 | 1.0 | |
| N | -0.11 | 0.10 | -0.10 | 1.0 |
| Model (0.99) | 0.61 | -0.99 | 0.91 | -0.14 |
| Model (0.95) | 0.56 | -0.89 | 0.82 | -0.25 |
| | RT | Log freq | Rank | N |

Table 1.

Correlations of model with data from Balota, Cortese and Pilotti (1999). Correlations are based on 1215 4-letter words. Model correlations are shown with thresholds of both 0.99 and 0.95.

Table 2. Andrews (1989) Experiment 1

| | Data | | Model | | Adjusted Model | |
|--------|--------|-------|--------|-------|----------------|-------|
| | High N | Low N | High N | Low N | High N | Low N |
| High F | 602 | 608 | 371 | 404 | 597 | 613 |
| Low F | 693 | 733 | 570 | 651 | 693 | 732 |

Table 2 shows mean lexical decision latencies (in Milliseconds) from Experiment 1 of Andrews (1989) (Data), the unadjusted results of the model simulations (Model), and the model output fitted to the data by regressing the model against the data.

Table 3. Andrews (1992) Experiment 1

| | Data | | Model | | Adjusted Model | |
|--------|--------|-------|--------|-------|----------------|-------|
| | High N | Low N | High N | Low N | High N | Low N |
| High F | 586 | 570 | 323 | 326 | 577 | 578 |
| Low F | 714 | 757 | 602 | 698 | 712 | 759 |

Table 3 shows mean lexical decision latencies from Experiment 1 of Andrews 1992 (Data), the unadjusted results of the model simulations (Model), and the model output fitted to the data by regressing the model against the data.

Table 4

Simulation of identification times (raw model) for the items from Andrews (1989), Experiment 1.

| | Model | |
|--------|--------|-------|
| | High N | Low N |
| High F | 586 | 419 |
| Low F | 889 | 643 |

Table 5. Simulations of Siakaluk, Sears and Lupker (2002), Experiment 1

| Source of data | 1A | 1B | 1C | 1D | Raw Model | Adjusted Model |
|------------------|-----|-----|-----|-----|-----------|----------------|
| Condition | | | | | | |
| hf largeN hfN | 476 | 521 | 541 | 564 | 353 | 478 |
| hf smallN hfN | 481 | 519 | 543 | 552 | 360 | 479 |
| hf largeN no hfN | 480 | 519 | 542 | 568 | 360 | 479 |
| hf smallN no hfN | 486 | 525 | 544 | 557 | 372 | 482 |
| lf largeN hfN | 494 | 548 | 579 | 602 | 457 | 502 |
| lf smallN hfN | 507 | 564 | 598 | 600 | 472 | 505 |
| lf largeN no hfN | 509 | 562 | 593 | 609 | 485 | 508 |
| lf smallN no hfN | 517 | 575 | 613 | 615 | 524 | 517 |

Table 5.

Lexical decision data (in Milliseconds) from experiments 1A-1D of Siakaluk et al. (2002), along with model simulations. The table shows both Raw Model output (in samples) and Adjusted Model output where the model has been fitted to the data from 1A.

Note.

Hf: high frequency word, lf: low frequency word, largeN: large neighborhood, smallN: small neighborhood, hfN: high frequency neighbor.

Table 6. Simulations of Siakaluk, Sears and Lupker (2002), Experiment 2

| Experiment | Condition | Data | Model |
|------------|----------------|------|-------|
| 2A | smallN no hfN | 561 | 524 |
| | smallN one hfN | 545 | 473 |
| 2B | largeN no hfN | 603 | 492 |
| | largeN one hfN | 588 | 470 |

Table 6

Mean lexical decision latencies (in Milliseconds) and model simulations (samples) for the data from Siakaluk, Sears and Lupker (2002) Experiment 2A (small N words) and Experiment 2B (large N words). Note that the Model RTs here represent the raw output of the model, and that 2A and 2B used different participants.