The effect of *CYP1A1* polymorphisms on the risk of lung cancer: a global meta-analysis based on 71 case–control studies

Zhiwei Chen[†], Ziming Li[†], Xiaomin Niu, Xiangyun Ye, Yongfeng Yu, Shun Lu* and Zhen Chen^{1,*}

Shanghai Lung Tumour Clinical Medical Centre, Shanghai Chest Hospital Affiliated to Shanghai Jiaotong University, Shanghai 200030, China and ¹Department of Pathology, Shidong Hospital, Shiguang Road 999, Shanghai 200438, People's Republic of China

[†]These two authors contributed equally to this work.

*To whom correspondence should be addressed. Shanghai Lung Tumour Clinical Medical Centre, Shanghai Chest Hospital Affiliated to Shanghai Jiaotong University, Huaihai West Road 241, Shanghai 200030, People's Republic of China. Tel: +86 21 62821990; Fax: +86 21 62804970; Email: supershun_lu@hotmail.com or Department of Pathology, Shidong Hospital, Shiguang Road 999, Shanghai 200438, People's Republic of China. Tel: +86 21 65882999; Fax: +86 21 65881977; Email: superchenzhen@126 .com

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The cytochrome P450 1A1 (CYP1A1) is a phase I enzyme involved in many oxidative reactions that has attracted considerable attention as a candidate gene for lung cancer susceptibility based on its function as a key factor required for bioactivation of carcinogenic polycyclic aromatic hydrocarbons and catechol oestrogen formation. In the past decade, the relationship between CYP1A1 and lung cancer has been reported in various ethnic groups; however, these studies have yielded contradictory results. To investigate this inconsistency, we performed a meta-analysis of 71 studies involving a total of 30 368 subjects for the MspI and Ile-Val polymorphism of the CYP1A1 gene to evaluate the effect of CYP1A1 on genetic susceptibility for lung cancer. In a combined analysis, the summary per-allele odds ratios for lung cancer of the MspI and Ile-Val polymorphism were 1.19 [95% confidence interval (CI): 1.11-1.28] and 1.20 (95% CI: 1.08–1.33), respectively. Significant results were also observed using dominant or recessive genetic model. In the subgroup analysis by ethnicity, significantly increased risks were found for the MspI and Ile-Val polymorphism among East Asians in almost all genetic models. However, only marginal significant associations were detected for the MspI polymorphism among Caucasians and other population, while no significant associations were observed for the Ile-Val polymorphism in Caucasians and other population. This meta-analysis demonstrated that the MspI and Ile-Val polymorphism of CYP1A1 is a risk factor associated with increased lung cancer susceptibility, but these associations vary in different ethnic populations.

Introduction

Lung cancer is the most common malignancy and the leading cause of cancer death world wide, with an estimated 5-year survival rate of 15% (1). Despite much investigation, the

causes are not yet fully understood. Epidemiological evidence suggests that exposure to tobacco-associated carcinogens is clearly implicated in its aetiology (2). However, it has also been reported that only <20% of smokers develop lung cancer, suggesting that genetic variations and other environmental factors also play important roles in determining individual differences in lung cancer susceptibility (3,4).

The cytochrome p450 (CYP) enzymes play an important role in the metabolism of many endogenous substrates such as steroids and eicosanoids and exogenous carcinogens such as polycyclic aromatic hydrocarbons (PAHs), aromatic and heterocyclic amines that can then covalently bind to DNA forming DNA adducts, thereby initiating the carcinogenic process (5). Within the CYP system, cytochrome p450 1A1 (CYP1A1) is key to the metabolic activation of PAHs found in tobacco smoke and is mainly expressed in extrahepatic tissues, including the lung (6). Several important single nucleotide polymorphisms have been identified in the CYP1A1 gene. Among them, two genetically liked polymorphisms in CYP1A1, the 3801T > C substitution (rs4646903) creating a MspI restriction site in the 3' non-coding region and the 2455A > G substitution (rs1048943) resulting in an amino acid change (Ile462Val), have been wildly studied for their association with lung cancer susceptibility.

Over the past few years, considerable efforts have been devoted to exploring the relationships between the *CYP1A1* polymorphisms and lung cancer risk among various populations. However, existing studies have yielded inconsistent results. These disparate findings may be due partly to insufficient power, false-positive results and publication biases. The interpretation of these studies has been further complicated by the use of different populations or different control source. To help clarify the inconsistent findings, we therefore conducted a comprehensive meta-analysis to quantify the overall risk of *CYP1A1* polymorphisms on developing lung cancer.

Materials and methods

Search strategy and selection criteria

Eligible literatures published before the end of September 2010 were identified by a search of PUBMED, EMBASE, Web of science and China National Knowledge Infrastructure databases using combinations of the following keywords: 'cytochrome p450 1A1', 'cyp1a1', 'polymorphism' or 'variant' and 'lung cancer' or 'lung carcinoma' without restriction on language. All the searched studies were retrieved, and their references including relevant reviews were checked as well for other relevant publications.

Eligible studies had to meet all the following criteria: (i) original papers containing independent data, (ii) identification of lung cancer cases was confirmed histologically or pathologically, (iii) case–control or cohort studies and (iv) genotype distribution information in cases and controls or odds ratio (OR) with its 95% confidence interval (CI) and *P*-value. The major reasons for exclusion of studies were (i) overlapping data and (ii) case-only studies, family-based studies and review articles.

Data extraction

The following information was independently extracted from each report by two participants in the meta-analysis: the first author, publication year, ethnicity of the study population, number of cases and controls, gender, age of cases and controls, cigarette smoking status, source of control, genotyping method and

the study population, number of cases and controls, gender, ntrols, cigarette smoking status, source of control, genotyp ronmental Mutagen Society. number of genotypes in cases and controls. Disagreement was resolved by discussion between authors. Where essential information was not presented in articles, every effort was made to contact the authors.

Statistical analysis

The strength of association between MspI and Ile-Val polymorphisms of *CYP1A1* and lung cancer risk was assessed by OR with the corresponding 95% CI. We first used the chi-square test to check if there was significant deviation from Hardy–Weinberg equilibrium (HWE) among the control subjects in each study. If controls of studies were found not to be in HWE, sensitivity analyses were performed with and without these studies to test the robustness of the findings. The per-allele OR of the risk allele was compared between cases and controls. Then we examined the association between risk genotype of these polymorphisms and the lung cancer susceptibility using dominant and recessive genetic models.

Heterogeneity across individual studies was calculated using the Cochran chi-square Q test followed by subsidiary analysis or by random-effects regression models with restricted maximum likelihood estimation (7-9). Random-effects and fixed-effect summary measures were calculated as inverse variance-weighted average of the log OR. The results of random-effects summary were reported in the text because it takes into account the variation between studies. In addition, we investigated potential sources of identified heterogeneity among studies by stratification according to the number of cancer cases (<300 and \geq 300), ethnic group, source of control (population or hospital), status of HWE (yes or no) and pathological type [lung squamous carcinoma (SCC), adenocarcinoma (AC), small cell lung cancer (SCLC) and large cell lung cancer (LCLC)]. Ethnic group was defined as East Asians, Caucasians (i.e. people of European origin) and others (e.g. Indian and African-American). The Z test was used to determine the significance of the pooled OR. Gender distribution in cases and controls, genotyping method and mean age of cases and controls were analysed as covariates in meta-regression.

We assessed publication bias by using an ancillary procedure attributed to Egger *et al.* (10), which uses a linear regression approach to measure funnel plot asymmetry on the natural logarithm of the OR. The larger the deviation from the funnel curve of each study the more pronounced the asymmetry. The results from small studies tend to scatter widely at the bottom of the graph, with the spread narrowing among larger studies. The significance of the intercept is evaluated using the *t* test. Sensitivity analysis was performed by removing each individual study in turn from the total and re-analysing the remainder. This procedure was used to ensure that no individual study was entirely responsible for the combined results. All statistical analyses were carried out with the Stata software version 10.0 (Stata Corporation, College Station, TX, USA). The type I error rate was set at 0.05. All the *P*-values were for two-sided analysis.

Results

Characteristics of the included studies

The combined search yielded 247 references. Study selection process was shown in Figure 1. A total of 71 studies were finally included with 12 815 patients and 17 553 controls (11–81). The detailed characteristics of the studies included in this meta-analysis are shown in Table I. For the MspI polymorphism, 56 studies were available, including a total of 9688 cases and 12 670 controls. For the Ile-Val polymorphism, 47 studies involved a total of 8920 cases and 12 540 controls. These two polymorphisms were found to occur in frequencies consistent with HWE in the control populations of the vast majority of the published studies. Of the cases, 32% were East Asians, 44% were Caucasians and 24% were of other ethnic origins.

Meta-analysis results

Association of MspI variant with lung cancer. Using randomeffect model, the per-allele overall OR of the C variant for lung cancer was 1.19 [95% CI: 1.11–1.28; $P(Z) < 10^{-5}$; $P(Q) < 10^{-4}$], with corresponding results under dominant and recessive genetic models of 1.22 [95% CI: 1.11–1.34; $P(Z) < 10^{-4}$; $P(Q) < 10^{-5}$; Figure 2] and 1.35 [95% CI: 1.18–1.54; $P(Z) < 10^{-4}$; P(Q) = 0.03], respectively.

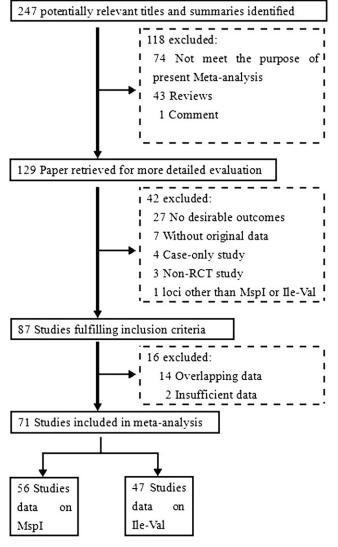


Fig. 1. Flow chart of literature search for studies examining *CYP1A1* gene polymorphism and risk of lung cancer.

When studies were stratified for ethnicity, significant risks were found among East Asians in all genetic models (C allele: OR = 1.19, 95% CI: 1.07–1.32; dominant model: OR = 1.24, 95% CI: 1.07–1.44; recessive model: OR = 1.36, 95% CI: 1.13-1.64). Similar results were also found in the Caucasians and other populations with a per-allele OR of 1.15 (95% CI: 1.01-1.32) and 1.23 (95% CI: 1.05-1.43), respectively. Subsidiary analyses of control source yielded a per-allele OR for hospital-based controls of 1.22 (95% CI: 1.09-1.67) and for population-based controls of 1.17 (95% CI: 1.06-1.28). However, in the stratified analysis according to sample size, no significant association was detected among five large studies. In the subgroup analyses by pathological type, significant associations were found for lung SCC in all genetic modes. However, we failed to detect any association between lung AC, SCLC or LCLC and the polymorphism for all genetic models (Table II).

In meta-regression analysis, mean age of cases (P = 0.96) and controls (P = 0.73), genotyping method (P = 0.89) and gender distribution in controls (P = 0.10) did not significantly explain such heterogeneity. By contrast, the gender distribution

Reference	Year	Ethnicity	No. of cases/controls	Age of cases/controls	Gender distribution in cases/controls (male %)	Source of control	Genotyping method
Kawajiri (11)	1990	Japanese	68/104	NA/NA	NA/NA	Population	RFLP
Tefre (12)	1991	Norwegain	221/212	63.0/NA	79.6/89.7	Population	RFLP
Hirvonen (13)	1992	Finnish	87/121	63.0/52.0	80.5/87.6	Population	RFLP
Nakachi (14)	1993	Japanese	85/170	65.2/65.2	NA/NA	Population	RFLP
Shields (15)	1993	American	56/48	65.1/62.1	96.4/93.8	Hospital	RFLP
Drakoulis (16)	1994	German	142/171	65.0/59.0	80.9/57.9	Hospital	RFLP
Sugimura (17)	1994	Brazilian	110/112	61.4/61.9	80.0/80.4	Hospital	RFLP
Kihara (18)	1995	Japanese	97/258	NA/NA	86.4/77.9	Population	RFLP
Hamada (19)	1995 1996	Brazilian	99/108 44/81	62.9/61.8 62.0/NA	76.7/81.1 NA/NA	Hospital	RFLP RFLP
Jacquet (20) Ishibe (21)	1990 1997	Belgian Mexican– and African–Americans	171/295	62.1/61.9	NA/NA NA/NA	Population Hospital	RFLP
García-Closas (22)	1997	American	416/446	67.0/64.0	53.6/46.9	Population	RFLP
Bouchardy (23)	1997	French	150/171	58.4/55.0	NA/NA	Hospital	RFLP
Taioli (24)	1998	African–Americans	289/190	58.9/41.1	76.2/51.5	Population	RFLP
Hong (25)	1998	Korean	85/63	NA/NA	NA/NA	Hospital	RFLP
Marchand (26)	1998	Japanese, American	339/454	64.5/65.1	61.0/65.1	Population	RFLP
Sugimura (27)	1998	Japanese	247/185	67.8/62.1	100/100	Hospital	RFLP
Persson (28)	1999	Chinese	76/119	NA/NA	55.3/NA	Population	RFLP
Hu (29)	1999	Chinese	59/132	56.0/55.1	72.9/64.4	Population, Hospital	RFLP
Dresler (30)	2000	American	158/149	65.3/60.6	52.0/36.0	Population	RFLP
London (31)	2000	Chinese	214/669	63.0/63.0	100/100	Population	RFLP
Lin (32) Dolzan (33)	2000 2000	Chinese Slovenian	132/259 195/100	63.0/58.0 NA/NA	69.7/61.4 NA/NA	Population Population	RFLP RFLP, allele-specific PCR
Quiñones (34)	2001	Chilean	60/140	62.8/50.3	83.3/60.6	Population	RFLP
Chen (35)	2001	Chinese	106/106	56.9/53.6	82.1/82.1	Population	Allele-specific PCR
Song (36)	2001	Chinese	217/404	57.8/58.3	80.2/83.7	Population	RFLP
Gsur (37)	2001	Australian	134/134	63.9/64.1	100/100	Hospital	RFLP
Ratnasinghe (38)	2001	Finn	282/324	60.0/59.0	100/100	Population	RFLP
Zhang (39)	2002	Chinese	65/60	59.4/55.6	86.2/90.0	Hospital	RFLP
Zhou (40)	2002	Chinese	92/98	57.5/55.6	70.7/69.4	Hospital	RFLP
Oztürk (41)	2003	Turkish	55/65	57.8/56.3	72.1/56.9	Population	RFLP
Kiyohara (42)	2003	Japanese	158/259	NA/NA	0/0	Hospital	RFLP
Taioli (43)	2003	Italian	110/707	34.8/40.7	66.7/68.7	Population, hospital	RFLP
Wang (44)	2003	Chinese	162/181	55.7/54.8	64.6/66.9	Population	RFLP
Dialyna (45)	2003	Greek	122/178	64.6/58.8	86.0/83.1	Population	RFLP
Li (46)	2004	Chinese	217/200	NA/NA	NA/NA	Hospital	RFLP
Dong (47) Sobti (48)	2004 2004	Chinese Indian	82/91 100/76	NA/NA 55.5/50.9	NA/NA 95.0/96.1	Hospital Population	RFLP, allele-specific PCI RFLP
Liang (49)	2004	Chinese	152/152	60.9/60.5	70.4/70.4	Hospital	RFLP
Yang (50)	2004	Chinese	197/144	56.0/56.0	0/0	Population	TaqMan
Alexandrie (51)	2004	Swedes	524/530	66.0/44.0	65.6/77.9	Population	RFLP, allele-specific PCR
Demir (52)	2005	Turkish	32/37	55.0/34.0	93.5/64.9	Population	RFLP
Wrensch (53)	2005	Latinos, African–Americans	370/944	64.0/63.5	47.1/47.8	Population	RFLP
Li (54)	2005	Chinese	103/138	58.9/58.9	77.5/69.8	Population	RFLP
Ng (55)	2005	Chinese	126/162	62.0/63.4	0/0	Hospital	RFLP
Adonis (56)	2005	Chilean	57/103	63.2/53.3	70.2/54.4	Population	RFLP
Sreeja (57) Warglaff (58)	2005	Indian	146/146	58.2/56.1	91.1/87.7	Hospital	RFLP
Wenzlaff (58)	2005	American	159/179	63.2/54.9	41.3/49.2	Population	RFLP
Belogubova (59) Li (60)	2006 2006	Russian Chinese	141/450 150/152	60.6/59.5 57.1/54.8	87.9/48.4 76.7/77.0	Population, hospital Population	RFLP Allele-specific PCR
Chang (61)	2006	Chinese	163/163	57.1/54.8 NA/NA	56.4/56.4	Hospital	RFLP
Wang (62)	2006	Chinese	91/91	NA/NA NA/NA	65.9/65.9	Hospital	RFLP
Qian (63)	2000	Chinese	108/108	59.4/57.6	61.1/61.1	Population	RFLP
Yang (64)	2000	Korean	314/349	55.4/48.3	67.6/61.8	Population	RFLP
Gu (65)	2007	Chinese	279/684	NA/NA	NA/NA	Population	RFLP
Cote (66)	2007	American	354/440	42.5/41.4	49.6/44.3	Population	RFLP
Pisani (67)	2007	Thai	168/287	NA/NA	66.4/61.1	Population, hospital	RFLP
Yoon (68)	2008	Korean	213/213	57.0/57.0	0/0	Population	RFLP
Gallegos-Arreola (69)	2008	Mexican	222/248	60.7/46.2	60.8/53.2	Population	RFLP
Shah (70)	2008	Indian	200/200	56.0/43.0	100/100	Population	RFLP
Xia (71)	2008	Chinese	58/116	57.0/57.0	89.7/89.7	Hospital	RFLP
Liu (72)	2008	Chinese	110/125	58.1/59.2	63.6/65.6	Population	Oligonucleotide chip
	2000	American	502/523	59.7/59.0	0/0	Population	RFLP, TaqMan
Cote (73) Honma (74)	2009 2009	Brazilian	200/264	64.0/NA	72.0/60.6	Population	RFLP, Taqinan RFLP

Table I. Studies investigating the association between the CYP1A1 polymorphisms and lung cancer risk

Table 1. Continued							
Reference	Year	Ethnicity	No. of cases/controls	Age of cases/controls	Gender distribution in cases/controls (male %)	Source of control	Genotyping method
Kumar (75)	2009	Indian	93/253	42.6/39.8	86.0/80.2	Population	RFLP
Klinchid (76)	2009	Thai	85/82	59.8/57.8	64.0/46.0	Population	RFLP
Timofeeva (77)	2009	German	619/1264	45.1/45.0	63.6/62.9	Population	Mass array
Shaffi (78)	2009	Indian	109/163	52.7/53.2	79.8/41.7	Hospital	RFLP
San Jose (79)	2010	Spanish	104/265	66.0/63.9	NA/NA	Hospital	RFLP
Jin (80)	2010	Chinese	124/154	59.3/54.5	NA/NA	Hospital	RFLP
Wright (81)	2010	Australian	1040/784	65.7/63.1	68.0/68.0	Population, hospital	RFLP

NA, not available; RFLP, restriction fragment length polymorphism; PCR, polymerase chain reaction.

in cases (P = 0.015) was significantly correlated with the magnitude of the genetic effect, explaining 36% of the heterogeneity.

Association of Ile-Val variant with lung cancer. Overall, the per-allele OR of the 462Val variant for lung cancer was 1.20 [95% CI: 1.08–1.33; P(Z) = 0.0007; $P(Q) < 10^{-5}$], with corresponding results under dominant and recessive genetic models of 1.21 [95% CI: 1.05–1.38; P(Z) = 0.007; P(Q) < 10^{-5}] and 1.57 [95% CI: 1.30–1.90; $P(Z) < 10^{-5}$; P(O) =0.15], respectively (Table III).

In the stratified analysis by ethnicity, significant associations were detected among East Asians in all genetic models (G allele: OR = 1.20, 95% CI: 1.06–1.35; recessive model: OR = 1.54, 95% CI: 1.24–1.91). Unfortunately, we failed to detect any association to cancer risk for Caucasians and other populations in all genetic models (Figure 3). Subsidiary analyses of control source yielded a per-allele OR for hospital-based controls of 1.30 (95% CI: 1.06-1.59) and for population-based controls of 1.14 (95% CI: 1.01-1.29). In sample size subgroup, this association became nonsignificant when the meta-analysis was restricted to larger studies.

Significant heterogeneity was present among the 47 studies of the Ile-Val polymorphism $[P(Q) < 10^{-5}]$. However, mean age of cases (P = 0.76) and controls (P = 0.52), genotyping method (P = 0.17) and gender component of cases (P = 0.05) and controls (P = 0.91) only explained little heterogeneity.

Gene-environment interaction. The data on genotypes of the MspI and Ile-Val polymorphism among cases stratified by smoking status were available in 18 (including 2451 cancer cases) and 11 (including 1837 cancer cases) studies, respectively. Using dominant genetic model, smokers with the risk allele of MspI or Ile-Val variant had no significantly increased lung cancer risk compared to non-smoker cancer cases with an OR of 1.41 [95% CI: 0.77–2.02; $P(Q) < 10^{-5}$] and 1.08 [95% CI: 0.67–1.74; P(Q) = 0.0004], respectively.

Sensitivity analyses and publication bias

Sensitivity analysis indicated that no single study influenced the pooled OR qualitatively, suggesting that the results of this meta-analysis are stable (data not shown).

The shape of the funnel plots was symmetrical (Figures 4 and 5). The statistical results still did not show publication bias in these studies for MspI (Begg test, P = 0.25; Egger test, P = 0.29) and Ile-Val polymorphism (Begg test, P =0.32; Egger test, P = 0.18).

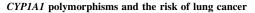
Discussion

Large sample and unbiased epidemiological studies of predisposition gene polymorphisms could provide insight into the in vivo relationship between candidate genes and diseases. This is the most comprehensive meta-analysis that examined the *CYP1A1* polymorphisms and the relationship to susceptibility for lung cancer. Its strength was based on the accumulation of published data giving greater information to detect significant differences. In total, the meta-analysis involved 71 studies for lung cancer, which provided 12 815 cases and 17 553 controls.

Our results demonstrated that the MspI and Ile-Val polymorphism of CYP1A1 is a risk factor for developing lung cancer. In the stratified analysis by ethnicity, significant associations were found in East Asians for the two polymorphisms in almost all genetic models. However, only marginal significant associations were detected among Caucasians and other population for MspI polymorphism of CYP1A1, while no associations were found in Caucasians and other population for Ile-Val variant. In fact, the distribution of the less common Val allele varies extensively between different races, with a prevalence of $\sim 25\%$ among East Asians, \sim 5% among Caucasians and \sim 15% among other population. Thus, failing to identify any significant association in Caucasians and other populations could be due to substantially lower statistical power caused by the relatively lower prevalence of Val allele of CYP1A1. Therefore, additional studies are warranted to further validate ethnic difference in the effect of this functional polymorphism on lung cancer risk. In addition, studies using different source of controls also get consistent positive results.

When subgroup analyses by pathological types were considered, the C allele of MspI polymorphism has an increased risk factor in lung SCC subgroup but not in lung AC, SCLC or LCLC subgroup. Our results support the hypothesis by Marchand et al. (26) that genetic susceptibility to PAHs predominantly causes squamous cell carcinoma. The reason for the observed tumour-specific difference in the risk conferred by the CYP1A1 MspI polymorphism is unknown. However, different carcinogenic processes may be involved in the genesis of various tumour types because of the presence of functionally different CYP1A1 MspI polymorphism. Hence, future studies should use homogeneous cancer patients.

Compared with the previous meta-analysis (82,83), the present study is much larger, with almost four times as many cases as the earlier meta-analysis. In addition, we also investigated the interaction between CYP1A1 gene polymorphisms, smoking status and lung cancer susceptibility. Furthermore, we explored



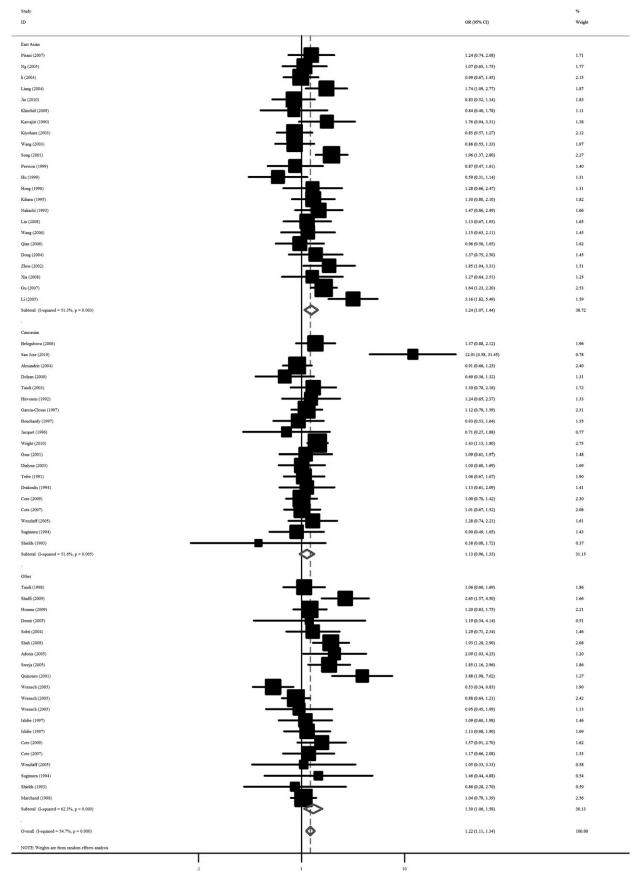


Fig. 2. Forest plot (random-effects model) of lung cancer risk associated with CYP1A1 MspI polymorphism using dominant genetic model.

Subgroup analysis	No. of cases/controls	C allele versus T allele			Dominant model			Recessive model		
		OR (95% CI)	P(Z)	P(Q)	OR (95% CI)	P(Z)	P(Q)	OR (95% CI)	P(Z)	P(Q)
Total Ethnicity	9688/12 670	1.19 (1.11–1.28)	$< 10^{-5}$	$< 10^{-4}$	1.22 (1.11–1.34)	$< 10^{-4}$	$< 10^{-5}$	1.35 (1.18–1.54)	$< 10^{-4}$	0.03
East Asians	3095/4569	1.19 (1.07-1.32)	0.001	0.01	1.24 (1.07-1.44)	0.005	0.003	1.36 (1.13-1.64)	0.001	0.009
Caucasians	4237/4865	1.15 (1.01 - 1.32)	0.04	0.04	1.13 (0.96–1.33)	0.15	0.005	1.58 (1.08–2.30)	0.02	0.78
Others	2356/3236	1.23 (1.05–1.43)	0.01	0.0002	1.30 (1.06–1.58)	0.01	0.0001	1.25 (0.97 - 1.62)	0.09	0.11
Control source		(,					
Hospital	6087/8460	1.22 (1.09-1.37)	0.0008	0.003	1.26 (1.08-1.48)	0.004	0.001	1.41 (1.19-1.67)	$< 10^{-4}$	0.37
Population	3801/4210	1.17 (1.06–1.28)	0.0009	0.001	1.18 (1.05–1.34)	0.007	$< 10^{-4}$	1.33 (1.10–1.61)	0.003	0.02
Sample size		· · · · ·			· · · · · ·			· · · · ·		
Small	6984/10 057	1.20 (1.10-1.30)	$< 10^{-4}$	$< 10^{-4}$	1.24 (1.11-1.38)	0.0001	$< 10^{-5}$	1.34 (1.16-1.55)	$< 10^{-4}$	0.02
Big	2704/2613	1.16 (0.99–1.37)	0.07	0.20	1.11 (0.94–1.32)	0.22	0.17	1.32 (0.92-1.89)	0.13	0.55
HWĔ		, í								
Yes	8690/11 309	1.19 (1.10-1.28)	$< 10^{-4}$	$< 10^{-4}$	1.22 (1.10-1.35)	$< 10^{-4}$	$< 10^{-5}$	1.34 (1.15-1.55)	0.0002	0.07
No	998/1361	1.21 (0.96-1.52)	0.11	0.02	1.23 (0.85-1.80)	0.28	0.002	1.25 (0.94-1.65)	0.13	0.24
Pathological type										
SCC	1483/4867	1.59 (1.25-2.01)	0.0001	$< 10^{-5}$	1.61 (1.30-2.01)	$< 10^{-4}$	0.0001	1.80 (1.22-2.65)	0.003	0.05
AC	912/4296	1.02 (0.87-1.19)	0.81	0.57	0.98 (0.83-1.16)	0.85	0.58	1.09 (0.78–1.52)	0.62	0.95
SCLC	421/3689	1.07 (0.86–1.34)	0.52	0.88	0.88 (0.69–1.11)	0.26	0.53	1.65 (1.02-2.68)	0.04	0.60
LCLC	59/864	1.34 (0.79–2.29)	0.28	0.81	1.36 (0.72–2.58)	0.34	0.51	2.69 (0.72–10.04)	0.14	0.87

Table II. Main results of pooled ORs with CI for association of the MspI polymorphism and lung cancer risk in the meta-analysis

Table III. Main results of pooled ORs with CI for association of the Ile-Val polymorphism and lung cancer risk in the meta-analysis

Subgroup analysis	No. of cases/controls	G allele versus A allele			Dominant model			Recessive model		
		OR (95% CI)	P(Z)	P(Q)	OR (95% CI)	P(Z)	P(Q)	OR (95% CI)	P(Z)	P(Q)
Total	8920/12 540	1.20 (1.08–1.33)	0.0007	$< 10^{-5}$	1.21 (1.05–1.38)	0.007	$< 10^{-5}$	1.57 (1.30–1.90)	$< 10^{-5}$	0.15
Ethnicity		· · · · ·			· · · · ·			· · · · ·		
East Asians	2780/3863	1.20 (1.06-1.35)	0.003	0.003	1.20 (0.99-1.47)	0.06	$< 10^{-5}$	1.54 (1.24–1.91)	$< 10^{-4}$	0.43
Caucasians	4093/5312	1.19 (0.92–1.54)	0.17	0.0005	1.16 (0.88–1.53)	0.29	0.0002	1.52 (0.80-2.88)	0.20	0.76
Others	2047/3365	1.19 (0.93–1.51)	0.16	$< 10^{-4}$	1.26 (0.96-1.65)	0.10	0.0002	1.50 (0.91-2.48)	0.11	0.01
Control source										
Hospital	6113/9623	1.30 (1.06-1.59)	0.01	0.0003	1.38 (1.07-1.77)	0.01	0.0006	1.71 (1.18-2.48)	0.004	0.18
Population	2866/2917	1.14 (1.01-1.29)	0.03	$< 10^{-4}$	1.16 (0.99-1.36)	0.06	$< 10^{-5}$	1.50 (1.20-1.88)	0.0003	0.24
Sample size										
Small	6223/9288	1.21 (1.08-1.35)	0.0007	$< 10^{-5}$	1.22 (1.05-1.41)	0.01	$< 10^{-5}$	1.69 (1.38-2.06)	$< 10^{-5}$	0.23
Big	2697/3252	1.10 (0.80-1.52)	0.56	0.002	1.14 (0.81–1.59)	0.45	0.003	0.91 (0.56-1.49)	0.71	0.56
HWE										
Yes	5987/8467	1.20 (1.05-1.36)	0.007	$< 10^{-5}$	1.20 (1.03-1.39)	0.02	$< 10^{-5}$	1.53 (1.17-1.99)	0.002	0.05
No	2933/4073	1.19 (0.99–1.42)	0.06	0.005	1.16 (0.83-1.61)	0.38	$< 10^{-5}$	1.68 (1.25-2.25)	0.0006	0.79
Pathological type										
SCC	1286/3294	1.20 (0.96-1.50)	0.11	0.10	1.32 (0.97-1.81)	0.08	0.001	1.66 (0.95-2.92)	0.08	0.18
AC	918/3474	0.91 (0.74-1.13)	0.39	0.21	0.92 (0.70-1.20)	0.52	0.10	1.67 (0.69-4.02)	0.25	0.005
SCLC	331/2181	1.12 (0.67–1.86)	0.66	0.002	0.77 (0.43–1.37)	0.37	0.03	3.21 (1.62-6.34)	0.0008	0.32
LCLC	60/649	1.25 (0.37-4.31)	0.72	0.23	1.23 (0.29-5.25)	0.78	0.22	NA	NA	NA

NA, not available.

potential sources of heterogeneity across studies. Besides, our results suggest an overestimation of the true genetic association by small studies, consistent with the phenomenon known as 'winner's curse' (84,85).

A number of factors predict lung cancer; however, detailed pathogenesis mechanisms of lung cancer remain a matter of speculation. *CYP1A1* is a key phase I enzyme that converts many lung carcinogens into DNA-binding metabolites with carcinogenic potential (5). Evidence from *in vitro* expression studies indicates that the C allele of MspI polymorphism of *CYP1A1* could affect the transcriptional control elements involved in enzyme inducibility (86,87). The second polymorphism is an $A \rightarrow G$ transition in exon7 of *CYP1A1* that results in substitution of isoleucine to valine in the hemebinding region, which increases microsomal enzyme activity (6,88). Thus, the risk allele may therefore have more ability to metabolise mutagens and pro-carcinogens, hence more liable to metabolically activate mutagens and carcinogens.

In interpreting the results, some limitations of this metaanalysis should be addressed. First, in the subgroup analyses, different ethnicities were pooled in other population, which may bring in some heterogeneity. As studies among the Indians and Africans are currently limited, further studies including a wider spectrum of subjects should be carried to investigate the role of these variants in different populations. Second, the two subgroup meta-analyses considering interactions between *CYP1A1* polymorphisms and cigarette smoking, as well as between different pathological types of lung cancer were performed on the basis of a fraction of all the possible data to be pooled, so selection bias may have occurred and our results may be overinflated. Nevertheless, the total number of subjects included in this part of the analysis comprises the largest

CYP1A1 polymorphisms and the risk of lung cancer

D State Stat	OR (95% CI)	% Weight
East Asian	001/055.110	2.65
Yoon (2008) Yang (2007)	0.81 (0.55, 1.19) 1.24 (0.91, 1.69)	2.85
rang (2007)	1.24 (0.91, 1.69)	2.66
Ng (2005)	0.91 (0.57, 1.46)	2.39
Yang (2004)	2.53 (1.62, 3.97)	2.46
Chen (2001)	1.61 (0.92, 2.80)	2.45
Song (2001)	1.93 (1.37, 2.71)	2.79
London (2000)	0.92 (0.63, 1.33)	2.70
Persson (1999)	0.95 (0.52, 1.75)	2.02
Ни (1999)	0.40 (0.19, 0.82)	1.73
Hong (1998)	0.14 (0.03, 0.64)	0.65
Sugimura (1998)	1.01 (0.69, 1.48)	2.66
Kihara (1995)	0.78 (0.48, 1.26)	2.36
Nakachi (1993)	1.28 (0.75, 2.19)	2.21
Liu (2008)	1.99 (1.05, 3.76)	1.93
Li (2006)	1.27 (0.75, 2.17)	2.21
Dong (2004)	2.25 (1.22, 4.16)	1.99
Zhang (2002)	1.95 (0.96, 3.99)	1.74
Zhou (2002)	1.95 (0.96, 3.97)	1.75
Klinchid (2009)	1.29 (0.69, 2.41)	1.96
Subtotal (I-squared = 67.6%, p = 0.000)	1.20 (0.99, 1.47)	43.89
	T	
Caucasian		
Timofeeva (2009)	0.94 (0.64, 1.37)	2.66
Alexandrie (2004)	0.76 (0.39, 1.47)	1.88
Dresler (2000)	2.37 (1.27, 4.44)	1.96
Dolzan (2000)	0.72 (0.32, 1.62)	1.53
Taioli (2003)	1.61 (0.91, 2.86)	2.11
San Jose (2010)	4.05 (1.11, 14.70)	0.83
Bouchardy (1997)	0.74 (0.32, 1.71)	1.48
Wright (2010)	2.06 (1.43, 2.97)	2.71
Gsur (2001)	1.09 (0.49, 2.41)	1.56
Ratnasinghe (2001)	0.86 (0.55, 1.35)	2.45
Drakoulis (1994)	1.98 (0.89, 4.37)	1.56
Cote (2009)	0.97 (0.57, 1.65)	2.21
Cote (2007)	0.54 (0.28, 1.04)	1.89
Wenzlaff (2005)	0.39 (0.14, 1.10)	1.11
Hamada (1995)	2.22 (1.13, 4.36)	1.84
Subtotal (I-squared = 66.1%, p = 0.000)	1.16 (0.88, 1.53)	27.78
	1	
Other	1 <u> </u>	
Quinones (2001)	2.45 (1.24, 4.87)	1.81
Gallegos-Arreola (2008)	1.66 (1.15, 2.40)	2.71
Kumar (2009)	1.34 (0.74, 2.42)	2.05
Shah (2008)	1.79 (1.14, 2.79)	2.47
Oztürk (2003)	0.57 (0.27, 1.20)	1.67
Taioli (1998)	1.45 (0.63, 3.33)	1.48
Sobti (2004)	2.82 (0.82, 9.75)	0.88
Shaffi (2009)	2.25 (1.36, 3.73)	2.29
Wrensch (2005)	0.53 (0.33, 0.84)	2.39
Wrensch (2005)	0.89 (0.39, 2.02)	1.50
Wrensch (2005)	1.13 (0.54, 2.33)	1.71
Ishibe (1997)	0.92 (0.50, 1.70)	2.00
Ishibe (1997)	1.00 (0.26, 3.81)	0.78
Cote (2009)	2.79 (0.85, 9.15)	0.94
Cote (2007)	2.16 (0.50, 9.25)	0.69
Wenzlaff (2005)	0.32 (0.01, 8.24)	0.17
Marchand (1998)	0.80 (0.57, 1.11)	2.80
Subtotal (I-squared = 63.8%, p = 0.000)	1.26 (0.96, 1.65)	28.33
Overall (I-squared = 64.6%, p = 0.000)	1.21 (1.05, 1.38)	100.00
NOTE: Weights are from random effects analysis	1	

Fig. 3. Forest plot (random-effects model) of lung cancer risk associated with CYP1A1 Ile-Val polymorphism using dominant genetic model.

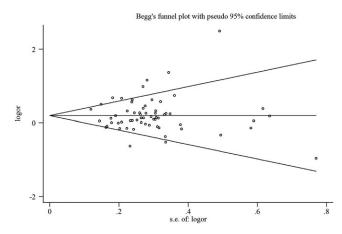


Fig. 4. Begg's funnel plot of *CYP1A1* MspI polymorphism and lung cancer risk; Egger's test was also performed to investigate the symmetry of the funnel plot (P = 0.29).

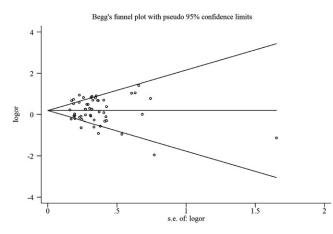


Fig. 5. Begg's funnel plot of *CYP1A1* lle-Val polymorphism and lung cancer risk; Egger's test was also performed to investigate the symmetry of the funnel plot (P = 0.18).

sample size so far. Third, only published studies were included in this meta-analysis. Therefore, publication bias may have occurred, even though the use of a statistical test did not show it. In spite of these, our present meta-analysis also had some advantages. First, substantial number of cases and controls were pooled from different studies, which greatly increased statistical power of the analysis. Second, no publication biases were detected, indicating that the whole pooled results may be unbiased.

To conclude, this meta-analysis suggests that the MspI and Ile-Val polymorphism of *CYP1A1* is associated with lung cancer susceptibility. For future association studies, strict selection of patients, well-matched controls and larger sample size will be required. More studies should also be carried out to examine the impact of *CYP1A1* on lung cancer. Moreover, gene–gene and gene–environment interactions should also be considered in future studies.

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