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The elevated preoperative platelet-to-lymphocyte ratio predicts poor prognosis in breast cancer patients

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Background: The elevation of the platelet-to-lymphocyte ratio (PLR), an easily applicable blood test based on platelet and lymphocyte counts has been associated with poor prognosis in patients with different types of cancer. The present study was aimed to investigate the prognostic significance of the preoperative PLR in a large cohort of breast cancer patients.

Methods: Data from 793 consecutive non-metastatic breast cancer patients, treated between 1999 and 2004, were evaluated retrospectively. The optimal cutoff values for the PLR were calculated using receiver operating curve analysis. Cancer-specific survival (CSS), overall survival (OS) as well as distant metastasis-free survival (DMFS) were assessed using the Kaplan–Meier method. To evaluate the independent prognostic significance of PLR, multivariable Cox regression models were applied for all three different end points.

Results: Univariable analysis revealed a significant association between the elevated preoperative PLR and CSS (hazard ratio (HR): 2.75, 95% confidence interval (CI): 1.57–4.83, $P < 0.001$) that remained statistically significant in multivariable analysis (HR: 2.03, 95% CI: 1.03–4.02, $P = 0.042$). An increased PLR was also significantly associated with decreased OS in univariable (HR: 2.45, 95% CI: 1.43–4.20, $P = 0.001$) and in multivariable analysis (HR: 1.92, 95% CI: 1.01–3.67, $P = 0.047$). Furthermore, univariable analysis showed a significant impact of increased PLR on DMFS (HR: 2.02, 95% CI: 1.18–3.44, $P = 0.010$). Subgroup analysis revealed significant associations of the elevated PLR on the primary end point CSS for all breast cancer subtypes. This association retained its significance in multivariable analysis in patients with luminal B tumours (HR: 2.538, 95% CI: 1.043–6.177, $P = 0.040$).

Conclusions: In this study, we identified the preoperative PLR as an independent prognostic marker for survival in breast cancer patients. Independent validation of our findings is needed.

Breast cancer is the most commonly diagnosed cancer in woman worldwide and its incidence is increasing owing to mammographic screening and an ageing population. Postmenopausal hormone replacement therapy, Western-style diet, obesity, and consumption

of alcohol and tobacco contribute to the rising incidence of breast cancer. Because of improved treatment and earlier detection, the mortality rate has decreased in recent years in most Western countries; however, breast cancer is still the third leading cause of

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cancer-related death in Europe and the United States (Ferlay *et al*, 2012).

The prognosis of breast cancer patients depends on biological characteristics of the cancer and on patient and treatment characteristics. Established prognostic and predictive factors are age, tumour size, grade, nodal status, and endocrine/Her2neu receptor status. These parameters can be integrated into prognostic models that allow a relatively accurate estimation of the probability of recurrence and death from breast cancer (Kraeima *et al*, 2013).

In addition, a large number of translational research studies have revealed an association of various molecular biomarkers with clinical outcome in breast cancer. Molecular predictors of prognosis (for example, the Amsterdam 70 gene signature, Mammprint, or the 16-gene Recurrence Score, Oncotype Dx) may be used to gain additional prognostic and/or predictive information, but their relevance for integration into current clinical practice is still under investigation (Albain *et al*, 2010; Saghatchian *et al*, 2012; Drukker *et al*, 2013). In addition, high costs, lack of standardisation and regional availability limit their application in routine clinical practice.

Tumour progression and metastatisation comprise a cascade of steps that involve the interaction between the tumour and the host-derived stromal microenvironment, which includes factors that support angiogenesis and inflammation (Balkwill and Mantovani, 2001; Coussens and Werb, 2002). Recent data have expanded the concept that the systemic inflammatory response that is usually measured by surrogate blood-based parameters, such as C-reactive protein, neutrophil, or platelet count, has an important role in the progression of several solid tumours (Coussens and Werb, 2002; Mantovani *et al*, 2008; Roxburgh and McMillan, 2012; Dethlefsen *et al*, 2013; Szkandera *et al*, 2013).

Systemic inflammation is associated with the release of several pro-inflammatory mediators such as interleukin (IL)-1, IL-3, and IL-6 that are known to stimulate megakaryocyte proliferation leading to thrombocytosis. Platelet aggregation and degranulation along with the consequent release of platelet-derived pro-angiogenic mediators have been suggested as the important determinant of tumour growth (Klinger and Jelkmann, 2002). In addition, a number of immunologic mediators, most notably IL-10 and transforming growth factor- β are released, which can result in a significant immunosuppressive effect with consequent impaired lymphocyte function and reduced lymphocyte counts (Salazar-Onfray *et al*, 2007).

Pre-therapeutic indices of systemic inflammation have been suggested to provide prognostic information in various cancer entities. Among these inflammatory parameters, the platelet-to-lymphocyte ratio (PLR) has been proposed as an easily accessible and reliable marker to predict cancer prognosis (Smith *et al*, 2008; Koch *et al*, 2009; Smith *et al*, 2009; Asher *et al*, 2011; Proctor *et al*, 2011; Carruthers *et al*, 2012; Kwon *et al*, 2012; Raungkaewmanee *et al*, 2012; Roxburgh and McMillan, 2012; Liu *et al*, 2013).

Cumulating evidence suggests that a high PLR might represent an independent adverse prognostic factor in ovarian cancer (Asher *et al*, 2011; Raungkaewmanee *et al*, 2012), colorectal cancer (Kwon *et al*, 2012; He *et al*, 2013), advanced gastric cancer (Lee *et al*, 2012) as well as in pancreatic cancer (Smith *et al*, 2009). However, data regarding the prognostic significance of the PLR in breast cancer are sparse, and currently available data from small scale studies provide conflicting results about its prognostic significance with regard to different clinical end points in breast cancer patients (Seretis *et al*, 2012; Azab *et al*, 2013).

The purpose of our study was to evaluate the effect of the preoperative PLR on cancer-specific survival (CSS), overall survival (OS), and distant metastases-free survival (DMFS) in a large cohort of non-metastatic breast cancer patients in order to further clarify the prognostic significance of the PLR for breast cancer.

MATERIALS AND METHODS

Subjects. A total of 793 patients with histologically confirmed breast cancer were included in this retrospective study. The data set included all female patients with non-metastasised breast cancer who had been treated and/or had been included in the breast cancer surveillance programme at the Division of Clinical Oncology and the Department of Therapeutic Radiology and Oncology, Medical University of Graz from December 1999 to June 2004. Clinical and histopathological features were retrospectively obtained from the patient's history. In 81 patients (10%), a modified radical mastectomy was performed; the remaining patients were treated with breast conserving surgery and adjuvant radiotherapy. The preoperative blood cell count was obtained within 3 days before surgery. Adjuvant chemotherapy was administered in 93 patients (12%), adjuvant hormonal treatment in 378 patients (48%), and 202 patients (26%) were treated with adjuvant chemotherapy and hormonal therapy.

Follow-up investigations were performed in regular intervals (3-month intervals in years 1–3, 6-month intervals in years 4–5, and 12-month intervals in years 6–10 after diagnosis) and included clinical examination, laboratory tests including CEA and carbohydrate antigen 15-3, radiological assessment (mammography and breast ultrasound, liver scan or ultrasound and chest X-ray), and gynaecological examination. Loss to follow-up rate was 4.3% after 3 years and 7.1% after 5 years. This study has been approved by the Institutional Review Board of the Medical University of Graz. All participants were Caucasians.

Statistical analysis. The primary end point was cancer-specific survival (CSS) defined as the time from breast cancer diagnosis to the date of patients' breast cancer-related death. The secondary end points included OS calculated from the time of diagnoses to death of any cause, and DMFS calculated from the date of diagnosis to the date of breast cancer metastases.

The PLR was calculated as the absolute platelet count measured in $\times 10^9 l^{-1}$ divided by the absolute lymphocyte count measured in $\times 10^9 l^{-1}$. The neutrophil-to-lymphocyte ratio (NLR) was calculated as the absolute neutrophil count measured in $\times 10^9 l^{-1}$ divided by the absolute lymphocyte count measured in $\times 10^9 l^{-1}$. The ideal cutoff value for the PLR and NLR was determined applying receiver operating curve analysis as previously published (Absenger *et al*, 2013).

The relationship between the PLR and other clinico-pathological features was studied by χ^2 -test. Patients' clinical end points were calculated using the Kaplan–Meier method and compared by the log-rank test. Backward stepwise multivariable Cox proportion analysis was performed to determine the influence of potential confounders such as age, pathologic tumour stage, grade, lymph node involvement, oestrogen receptor (ER), progesterone receptor (PR), and HER2 status on CSS, OS, and DMFS. Hazard ratios (HRs) estimated from the Cox analysis were reported as relative risks with corresponding 95% confidence intervals (CIs). All statistical analyses were performed using the Statistical Package for Social Sciences version 20.0 (SPSS Inc., Chicago, IL, USA). A two-sided $P < 0.05$ was considered statistically significant.

RESULTS

A total of 793 female patients with histologically confirmed breast cancer were included in the present analysis. The median age at the time of diagnosis was 57.9 ± 12.2 years.

Baseline patient characteristics are shown in Table 1. Patients have been classified into the subtypes luminal A, luminal B, Her2, and basal like according to the St Gallen recommendations

Table 1. Patient characteristics

	Patient characteristics, n (%)
n	793
Age at diagnosis (years)	
≤60	410 (51.7%)
>60	383 (48.3%)
Tumour stage	
T1	450 (56.7%)
T2	178 (22.4%)
T3	102 (12.9%)
T4	38 (4.8%)
Missing data	25 (3.2%)
Lymph node involvement	
N0	454 (57.3%)
N1	207 (26.1%)
N2	79 (10.0%)
N3	45 (5.7%)
Missing data	8 (1%)
AJCC stage	
IA	316 (39.8%)
IB	1 (0.1%)
IIA	164 (20.7%)
IIB	113 (14.2%)
IIIA	98 (12.4%)
IIIB	33 (4.2%)
IIIC	43 (5.4%)
Missing data	24 (3%)
Tumour grade	
G1	57 (7.2%)
G2	383 (48.3%)
G3 + G4	350 (43.9%)
Missing data	3 (0.4%)
Molecular subtype	
Luminal A	369 (46.5%)
Luminal B	251 (31.7%)
Basal like	70 (8.8%)
Her2	30 (3.8%)
Missing data	73 (9.2%)

Abbreviations: AJCC stage = American Joint Committee on Cancer stage; n = number of patients.

published by Goldhirsch *et al* (2011). A total of 369 patients (46.5%) have been classified into luminal A, 251 patients (31.7%) into luminal B, 30 patients (3.8%) into the HER2-positive subtype, and 70 patients (8.8%) into the basal-like subtype. Classification was possible in 720 patients (91%). Determination of ER, PR, and HER2 receptor status revealed ER expression in 623 cases (78.6%), PR expression in 483 cases (60.9%), and Her2 overexpression in 85 cases (10.7%). HER2 status was generally determined using the HerceptTest. A negative assay was reported with 0 and 1+ staining, whereas 3+ was reported as positive; in case of a 2+ level of staining, a confirmatory testing by fluorescence *in situ* hybridisation was performed.

The mean platelet count was 271.2 ± 69.6 , the mean lymphocyte count was 1.7 ± 0.6 , and the mean PLR was 181.1 ± 131.0 . In 747 patients (94.2%), the preoperative PLR was available.

Applying the criteria mentioned above, we determined a cutoff value of 292 for the PLR to be optimal to discriminate between

patients' CSS that prompted us to select 292 as the optimal cutoff value for all subsequent analyses to differentiate between low (<292) and high (≥ 292) PLR. Overall, there were 699 patients with a low PLR and 48 patients with a high PLR.

A high PLR significantly correlated with lymph node involvement, high tumour grade, and ER-negative tumours (all $P < 0.05$), but not with age, advanced T stage, PR status, or Her2 overexpression.

Median follow-up time was 98 ± 29.2 months. Of the 793 breast cancer patients, 167 (21.1%) developed metastatic disease, 116 (14.6%) died due to disease progression. Overall, 136 patients (17.2%) died for any cause. Figures 1, 2 and 3 show the Kaplan–Meier curves for CSS, OS, and DMFS and reveal that a high PLR is a consistent factor for poor prognosis in breast cancer patients ($P < 0.01$ for all three tested end points, log-rank test).

Univariable analysis revealed a significant impact of a high PLR on CSS. In our multivariable analysis that included age, tumour stage, lymph node involvement, oestrogen and PR status, Her2 status, and PLR, the PLR was identified as an independent prognosticator of poor outcome for patients' CSS (HR: 2.03, 95% CI: 1.03–4.02; $P = 0.042$, Table 1). Furthermore, multivariable analysis showed a significant association of advanced T stage and N stage with poor CSS (Table 1). Regarding OS, a high PLR was identified as an independent prognostic factor for poor survival (HR: 1.92, 95% CI: 1.01–3.67, $P = 0.047$, Table 2). In addition, tumour stage and lymph node stage were independently associated with OS.

Univariable analysis showed a significant association between increased PLR and the occurrence of distant metastases (HR: 2.02, 95% CI: 1.18–3.44, $P = 0.010$). In multivariable analysis, the PLR did not remain a significant prognosticator of DMFS, whereas a significant impact of advanced T stage and lymph node involvement on poor DMFS was found (Table 3).

Internal validation has been performed by randomly splitting the data set into two subsets. In subgroup 1 ($n = 391$ patients), the HR for CSS was 2.588 (95% CI: 1.103–6.069) and in subgroup 2 ($n = 402$ patients), the HR for CSS was 2.867 (95% CI: 1.356–6.061).

Subgroup analysis revealed significant associations of the elevated PLR on the primary end point CSS for all breast cancer subtypes (Tables 3 and 4). This association retained its significance in multivariable analysis in patients with luminal B tumours (HR: 2.538, 95% CI: 1.043–6.177, $P = 0.040$). In this breast cancer subgroup, the PLR turned out to be superior in terms of the prognostic impact compared with age, T status, and N1 and N1 status.

In the present cohort, a total of 120 patients (15%) did not receive any adjuvant systemic therapy. In univariable analysis, the PLR was significantly associated with CSS among these patients (HR: 4.449, 95% CI: 1.612–12.27, $P = 0.004$). However, multivariable analysis did not show a significant effect of an elevated PLR on CSS (HR: 2.745, 95% CI: 0.728–10.34, $P = 0.136$).

Applying the criteria mentioned above, the cutoff value for the NLR to be optimal to discriminate between patients' CSS was 2.5. Overall, there were 299 patients (37.7%) with a low NLR (<2.5) and 448 patients (56.5%) with a high NLR (≥ 2.5). In univariable analysis, we detected a significant association between an elevated NLR and cancer-specific survival (HR: 1.514, 95% CI: 1.008–2.275, $P = 0.046$). In multivariable analysis, a significant impact of NLR on CSS was not detected (HR: 1.247, 95% CI: 0.811–1.918, $P = 0.315$).

DISCUSSION

Despite recent progress in the identification of genetic, epigenetic, and common molecular alterations, the routine diagnostic and

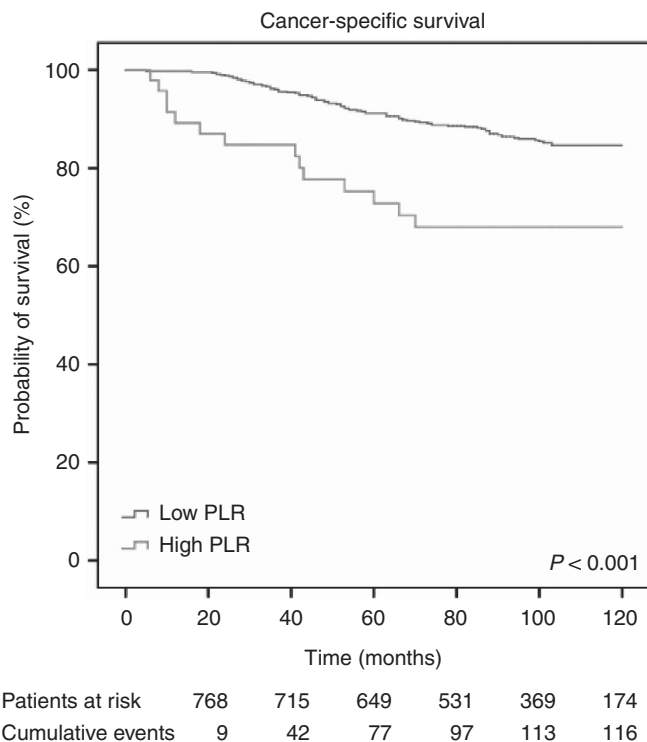


Figure 1. Kaplan–Meier curves for CSS of breast cancer patients categorised by the PLR.

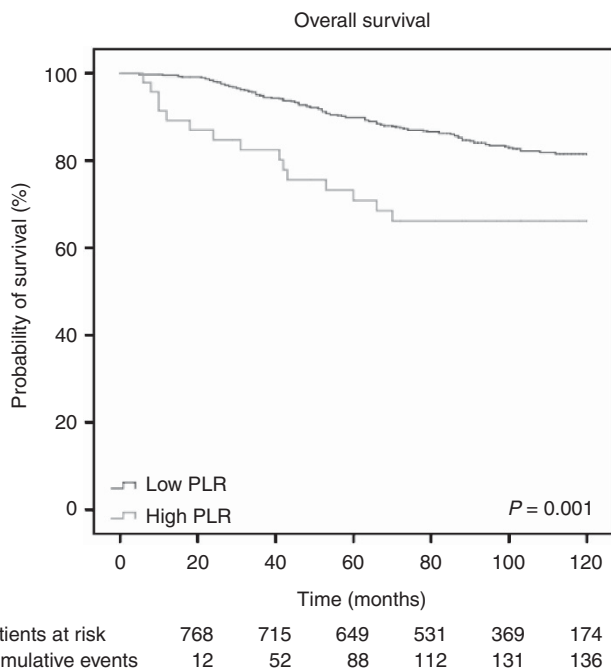


Figure 2. Kaplan–Meier curves for OS of breast cancer patients categorised by the PLR.

prognostic assessment of breast cancer currently relies on pathological tissue examination and traditional clinico-pathological prognostic variables (Blamey *et al*, 2007). Regularly used blood-based parameters, such as platelet or lymphocyte counts, are relatively easy to assess without additional laborious efforts, making them attractive parameters for an improved individualised risk assessment.

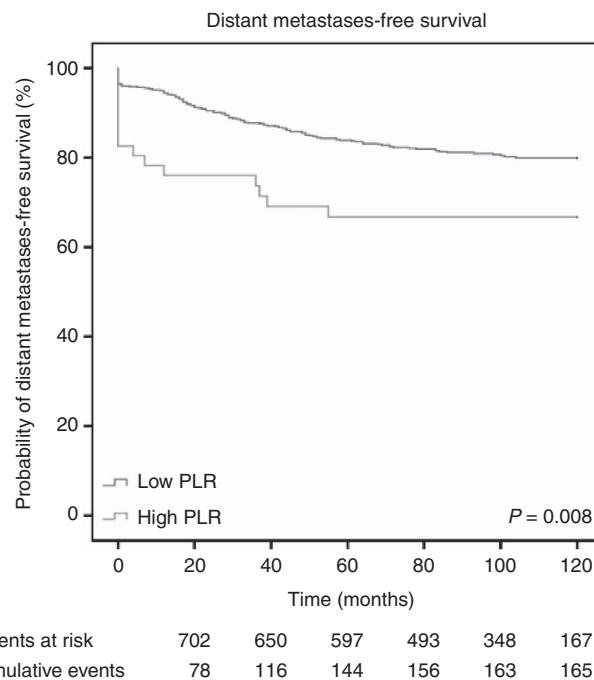


Figure 3. Kaplan–Meier curves for DMFS of breast cancer patients categorised by the PLR.

Recent experimental and clinical data indicate that the activation of platelets is crucial for cancer progression promoting angiogenesis, degradation of the extracellular matrix, and release of adhesion molecules and growth factors (Egan *et al*, 2011; Kono *et al*, 2012).

Previous studies have shown a direct correlation between the number of circulating platelets and the level of serum VEGF-A, a major factor in tumour-induced angiogenesis (Benoy *et al*, 2002). Furthermore, it has been shown that platelets promote invasiveness of tumour cells by enhanced metalloproteinase-9 (MMP-9) secretion and that the addition of anti-platelet drugs may inhibit invasiveness of tumour cells because of decreased MMP-9 secretion (Suzuki *et al*, 2004).

In addition, numerous studies have suggested that pro-inflammatory mediators, such as IL-1, IL-3, and IL-6, that are released in various cancer entities stimulate the proliferation of megakaryocytes, the platelet progenitor cells (Klinger and Jelkmann, 2002; Alexandrakis *et al*, 2003). An increased blood platelet count might therefore also reflect an index of systemic inflammation induced by the tumour.

Lymphocytes have a major role in cancer immune-surveillance, which inhibits tumour cell proliferation and metastasisation (Ownby *et al*, 1983). Tumour-infiltrating lymphocytes have been linked with improved outcomes in different types of cancer (Ropponen *et al*, 1997; Mahmoud *et al*, 2011; West *et al*, 2011). The peritumoural inflammatory response is thought to reflect the interaction between the tumour and the host. In previous studies, a high lymphocytic infiltrate has been linked with improved survival, independent of clinico-pathological characteristics in primary operable ductal invasive breast cancer (Mohammed *et al*, 2012). High levels of tumour-infiltrating lymphocytes have also been associated with pathological complete response to primary systemic therapy irrespective of breast cancer subtype (Seo *et al*, 2013; Dieci *et al*, 2014). However, data about the correlation of intra-tumoural immune cells and blood-based cells constituting the systemic inflammatory response are sparse.

Previous studies have demonstrated an association between lower peripheral blood lymphocyte count and poor survival in

Table 2. Univariable and multivariable analysis of clinico-pathological parameters for the prediction of cancer-specific survival in patients with breast cancer (n = 793)

	Univariable analysis		Multivariable analysis	
	HR (95% CI)	P-value	HR (95% CI)	P-value
Age at diagnosis				
<60	1 (Referent)		1 (Referent)	
≥60	0.97 (0.68–1.41)	0.89	0.97 (0.62–1.50)	0.876
Tumour stage				
T1	1 (Referent)		1 (Referent)	
T2	3.31 (2.07–5.30)	<0.001	3.06 (1.79–5.22)	<0.001
T3	3.31 (1.92–5.70)	<0.001	2.35 (1.29–4.27)	0.005
T4	6.76 (3.48–13.14)	<0.001	3.96 (1.88–8.34)	<0.001
Lymph node involvement				
N0	1 (Referent)		1 (Referent)	
N1	1.97 (1.26–3.10)	0.003	1.92 (1.12–3.28)	0.017
N2	2.66 (1.53–4.62)	0.001	1.92 (1.02–3.63)	0.043
N3	5.87 (3.41–10.12)	<0.001	5.21 (2.71–10.02)	<0.001
Tumour grade				
G1	1 (Referent)		1 (Referent)	
G2	1.99 (0.62–6.42)	0.251	2.06 (0.49–8.74)	0.328
G3 + 4	4.28 (1.35–13.60)	0.014	3.45 (0.82–14.51)	0.090
Estrogen receptor				
Negative	1 (Referent)		1 (Referent)	
Positive	0.65 (0.43–0.98)	0.038	0.90 (0.50–1.61)	0.725
Progesterone receptor				
Negative	1 (Referent)		1 (Referent)	
Positive	0.74 (0.51–1.07)	0.110	0.73 (0.45–1.19)	0.205
Her2 status				
Negative	1 (Referent)		1 (Referent)	
Positive	1.24 (0.75–2.06)	0.402	0.79 (0.44–1.42)	0.426
Platelet-to-lymphocyte ratio				
<292	1 (Referent)		1 (Referent)	
≥292	2.75 (1.57–4.83)	<0.001	2.03 (1.03–4.02)	0.042

Abbreviations: CI = confidence interval; HR = hazard ratio.

different types of cancer (Fogar *et al*, 2006; Ray-Coquard *et al*, 2009). Azab *et al* (2013) showed an improved survival in breast cancer patients with elevated lymphocyte counts compared with those with lower lymphocyte counts.

Furthermore, previous studies demonstrated that normalisation of an initial lymphocytopenia resulted in an improved clinical outcome in breast cancer patients treated with chemotherapy (Nieto *et al*, 2004; West *et al*, 2011). An elevated lymphocyte count has also significantly been associated with prolonged OS in patients with multiple myeloma and metastatic colorectal cancer (Leitch *et al*, 2007; Ege *et al*, 2008).

A combined index using platelet and lymphocyte counts has been investigated as prognostic marker for different cancer entities. In advanced gastric cancer patients, Lee *et al* (2012) demonstrated that an elevated PLR is associated with worse OS. A poor prognostic impact of an elevated PLR has also been demonstrated in ovarian cancer and pancreatic cancer (Smith *et al*, 2009;

Table 3. Univariable and multivariable analysis of the platelet-to-lymphocyte ratio (PLR) for the prediction of cancer-specific survival in different breast cancer subtypes (n = 720)

	Univariable analysis		Multivariable analysis ^a	
	HR (95% CI)	P-value	HR (95% CI)	P-value
Luminal A				
PLR ≤292	1 (Referent)		1 (Referent)	
PLR ≥292	1.14 (0.27–4.71)	0.86	1.44 (0.32–6.43)	0.65
Luminal B				
PLR ≤292	1 (Referent)		1 (Referent)	
PLR ≥292	2.78 (1.31–5.91)	0.008	2.54 (1.04–6.18)	0.04
Basal like				
PLR ≤292	1 (Referent)		1 (Referent)	
PLR ≥292	8.49 (2.22–32.47)	0.002	3.90 (0.62–24.32)	0.15
Her2				
PLR ≤292	1 (Referent)		1 (Referent)	
PLR ≥292	NA	0.665	NA	0.977

Abbreviations: CI = confidence interval; HR = hazard ratio; NA = not available.
^aAdjustment for age at diagnosis, tumour stage, and lymph node involvement.

Table 4. Univariable and multivariable analysis of clinico-pathological parameters for the prediction of cancer-specific survival in luminal B breast cancer patients (n = 251)

	Univariable analysis		Multivariable analysis ^a	
	HR (95% CI)	P-value	HR (95% CI)	P-value
Age at diagnosis				
<60	1 (Referent)		1 (Referent)	
≥60	1.120 (0.661–1.897)	0.675	1.212 (0.647–2.271)	0.549
Tumour stage				
T1	1 (Referent)		1 (Referent)	
T2	2.221 (1.154–4.275)	0.017	1.749 (0.841–3.637)	0.134
T3	1.840 (0.842–4.019)	0.126	1.420 (0.619–3.258)	0.408
T4	3.929 (1.544–9.997)	0.004	2.825 (0.994–8.031)	0.051
Lymph node involvement				
N0	1 (Referent)		1 (Referent)	
N1	2.199 (1.117–4.329)	0.023	2.069 (0.950–4.509)	0.067
N2	1.626 (0.689–3.836)	0.267	1.632 (0.652–4.084)	0.295
N3	8.228 (3.835–17.653)	<0.001	7.969 (3.371–18.837)	<0.001
PLR				
PLR ≤292	1 (Referent)		1 (Referent)	
PLR ≥292	2.780 (1.308–5.910)	0.008	2.538 (1.043–6.177)	0.040

Abbreviations: CI = confidence interval; HR = hazard ratio; PLR = platelet-to-lymphocyte ratio.
^aAdjustment for all factors listed in the table.

Asher *et al*, 2011; Raungkaewmanee *et al*, 2012). Furthermore, recent studies have shown a significant association between an elevated PLR and decreased OS after curative resection of colorectal cancer (Kwon *et al*, 2012; Liu *et al*, 2013).

In breast cancer, Seretis *et al* (2012) found a correlation between an elevated PLR and the number of infiltrated lymph nodes. More recently, Azab *et al* (2013) studied the impact of the PLR on OS in 437 breast cancer patients. The authors categorised the included patients according to PLR quartiles and found that patients in the highest PLR quartile had a significant higher 5-year mortality rate (Azab *et al*, 2013).

In breast cancer patients, elevated inflammatory markers such as C-reactive protein, serum amyloid A, and serum IL-6 have also been associated with reduced survival (Bachelot *et al*, 2003; Pierce *et al*, 2009). Furthermore, the measurement of the systemic inflammatory response using a selective combination of C-reactive protein and albumin (termed the modified Glasgow prognostic score) has been shown to have prognostic value in different cancer entities including breast cancer (Proctor *et al*, 2011). Besides PLR, other pre-therapeutic indices of systemic inflammation such as NLR and the lymphocyte-to-monocyte ratio have been suggested to provide prognostic information in different cancer entities (Azab *et al*, 2013; Stotz *et al*, 2014).

Azab *et al* (2013) have also shown a significant association of the pretreatment NLR with mortality that was superior to the prognostic effect of PLR. However, the authors have only analysed OS but not CSS, which might be influenced by numerous other factors including non-cancer-related deaths. In the present study, we did not detect a significant impact of an elevated NLR on CSS in multivariable analysis. Our findings are consistent with the data previously reported by Smith *et al* (2009) who found that PLR, not NLR, was a predictor of mortality in pancreatic cancer. Furthermore, Asher *et al* (2011) demonstrated that elevated PLR, not NLR, was a predictor of poor survival among patients with ovarian cancer. In a study by Kwon *et al* (2012), the elevated PLR, not NLR, has been shown to be a significant predictor of mortality in 200 colorectal cancer patients.

In the present study, an elevated preoperative PLR was significantly associated with decreased CSS and OS in breast cancer patients. These statistical associations retained their significance after adjusting for other potential predictors of patients' outcome and were independent of patient age, T- and N-stage, tumour grade, and ER, PR, and Her2 status.

Our findings indicate that an elevated PLR might be a poor prognostic factor in breast cancer patients. Patients with an elevated PLR might be considered as candidates for additional, more aggressive treatment approaches or more stringent follow-up schedules. The results of subgroup analysis indicate that the preoperative PLR carried the most significant prognostic information in patients with luminal B tumours. In this subgroup, the impact of the PLR on CSS turned out to be superior when compared with age, T status, and N1 and N2 status, and might contribute to the identification of patients who would benefit from a more aggressive treatment approach. If the present findings are replicated in future studies, determination of the PLR may help to obtain a more precise individual risk profile for breast cancer mortality and contribute to a tailored treatment of breast cancer patients.

However, our data have to be regarded as preliminary and have to be interpreted cautiously until validated by additional studies. Confirmation of the present results in an independent data set is imperative to draw firm conclusions about the role of PLR for breast cancer prognosis.

Major strengths of our study include the relatively large sample size and the narrow time frame for blood collection within 3 days before surgery excluding possible clinically significant infections.

However, as with all retrospective studies, limitations of our study are inherent to its design including the retrospective data collection. In addition, it has to be taken into account that lymphocyte and platelet counts may be influenced by several other conditions or factors. Increased platelet counts are often caused by inflammatory diseases, such as connective tissue disorders,

medication, bacterial diseases, or viral infections. The most common cause of low lymphocyte counts is a recent infection. In addition, lymphocytopenia may also be caused by infections, malnutrition, connective tissue disorders, severe stress, intense physical exercise, and medical treatments. Nevertheless, even considering these limitations, our data indicate that an increased preoperative PLR might represent an independent prognostic factor for CSS and OS in non-metastatic breast cancer patients.

In conclusion, an increased preoperative PLR seems to significantly affect CSS and OS in non-metastatic breast cancer patients and may support oncological therapy decisions. Clarifying the optimal PLR cutoff level and independent validation of our findings is warranted.

CONFLICT OF INTEREST

The authors declare no conflict of interest.

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