

Tilburg University

The rank of reduced dispersion matrices

Bekker, P.A.; de Leeuw, J.

Publication date:
1985

Document Version
Publisher's PDF, also known as Version of record

[Link to publication in Tilburg University Research Portal](#)

Citation for published version (APA):
Bekker, P. A., & de Leeuw, J. (1985). *The rank of reduced dispersion matrices*. (Research Memorandum FEW).
Faculteit der Economische Wetenschappen.

General rights

Copyright and moral rights for the publications made accessible in the public portal are retained by the authors and/or other copyright owners and it is a condition of accessing publications that users recognise and abide by the legal requirements associated with these rights.

- Users may download and print one copy of any publication from the public portal for the purpose of private study or research.
- You may not further distribute the material or use it for any profit-making activity or commercial gain
- You may freely distribute the URL identifying the publication in the public portal

Take down policy

If you believe that this document breaches copyright please contact us providing details, and we will remove access to the work immediately and investigate your claim.

R
709
CBM
R

1985-187
7626
1985
187

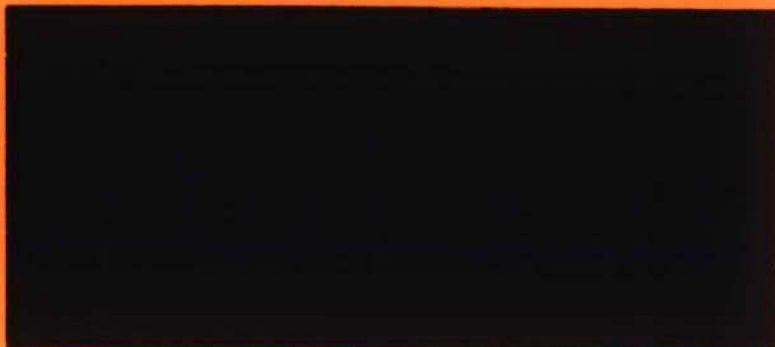
FEW



* C I N O O 3 8 6 *

subfaculteit der econometrie

RESEARCH MEMORANDUM



ECO

TILBURG UNIVERSITY
DEPARTMENT OF ECONOMICS
Postbus 90153 - 5000 LE Tilburg
Netherlands



FEW
187

THE RANK OF REDUCED DISPERSION MATRICES

by

Paul Bekker *

Jan de Leeuw **

June 1985

* Tilburg University; Dept. of Econometrics. Financial support by the Netherlands Organization for the advancement of Pure Research (ZWO) is gratefully acknowledged.

** University of Leiden; Dept. of Data Theory.

Abstract

Psychometricians working in factor analysis and econometricians working in regression with measurement error in all variables are both interested in the rank of dispersion matrices under variation of the diagonal elements. Psychometricians concentrate on cases in which low rank can be attained, preferably rank one, the Spearman case. Econometricians concentrate on cases in which the rank cannot be reduced below the number of variables minus one, the Frisch case. In this paper we give an extensive historical discussion of both fields, we prove the two key results in a more satisfactory and uniform way, we point out various small errors and misunderstandings, and we present a methodological comparison of factor analysis and regression on the basis of our results.

Keywords: factor analysis, errors of measurement, structural regression, functional models, communalities, errors in variables.

1. Introduction

Suppose Σ is a symmetric positive definite matrix of order m . In this paper we study the function $\text{rank}(\Sigma - \Omega)$ as Ω varies over the diagonal matrices satisfying $0 < \Omega < \Sigma$. This inequality notation is convenient shorthand for the requirement that both Ω and $\Sigma - \Omega$ must be positive semidefinite. More in particular we study:

$$(1) \quad \text{mr}(\Sigma) \equiv \min \{ \text{rank}(\Sigma - \Omega) \mid 0 < \Omega < \Sigma; \Omega \text{ diagonal} \}.$$

Investigation of this matrix function is important in at least two data analytic fields. The first field, which is very familiar for most readers of the psychometric literature, is factor analysis. In this context $\text{mr}(\Sigma)$ corresponds to the number of common factors. The older factor analysis literature concentrated on studying conditions for $\text{mr}(\Sigma) = 1$, while later contributions were mainly concerned with finding bounds or estimates of $\text{mr}(\Sigma)$. In our first historical section we shall review the most important algebraic results from the factor analysis literature.

The second field, which has had far less attention in the psychometric literature, is regression with errors of measurement in the variables. This model has been mainly studied in econometrics, with the major emphasis on conditions for $\text{mr}(\Sigma) = m - 1$. We shall also review the most important contributions from econometrics. This will also give us the opportunity to contrast the factor analysis model with the regression model.

2. The Spearman model

As we remarked in the introduction most of the early factor analysis literature concentrated on the characterization of matrices for which the Spearman model with a single common factor was appropriate. In the present context we might say that the early literature concentrated on finding conditions for $\text{mr}(\Sigma) = 1$. We briefly review this work which is riddled with errors and imprecisions.

In his famous paper on general intelligence Spearman (1904, p. 274) used the hierarchy of correlations as a criterion. Test should have the property that, after rearrangement, correlations were decreasing in each row, and that rows were proportional. This turned out to be somewhat too subjective and informal.

Krueger and Spearman (1907, p. 84-85) derived a new criterion from the correction for attenuation formula. It was

$$(2) \quad (\rho_{ik}\rho_{il}\rho_{jk}\rho_{jl})^{\frac{1}{4}} = (\rho_{ij}\rho_{kl})^{\frac{1}{2}},$$

which had to hold true for all quadruples ($i \neq j \neq k \neq l$). Observe that we have formulated the criterion in terms of the correlation matrix R , with elements ρ_{ij} , which are tacitly assumed to be nonnegative. From the Krueger and Spearman formula it is easy to derive

$$(3) \quad \rho_{ik}/\rho_{il} = \rho_{jk}/\rho_{jl}.$$

This formula was published for the first by Burt (1909, p. 159). He did not publish a proof, but he indicated that he derived it from the Krueger - Spearman formula, probably with help from Spearman. The actual (one-line) proof was not published until Spearman (1927, appendix, p. ii). Hart and Spearman (1912, p. 58, footnote) derived (3) from the partial correlation formula of Yule. Garnett (1919a, 1919b) referred to the conditions as Burt's equations, and he stated that there were only $\frac{1}{2}m(m-3)$ independent equations among the $m!/(m-4)!$ possible ones. This was proved in Garnett (1920, p. 245), where the name he had proposed for the conditions was formally withdrawn. Perhaps this was one of the seeds that grew into Burt's later attempts to rewrite the history of factor analysis (Hearnshaw, 1981, chapter 9). The conditions (3) were called the vanishing of the tetrad differences by Spearman and Holzinger (1924, 1925), who also wrote them in the more convenient form

$$(4) \quad \rho_{ik}\rho_{jl} - \rho_{il}\rho_{jk} = 0.$$

It is clear that in the earlier formulations the possibility of negative and zero correlations was sometimes overlooked. In fact one often has the impression that positivity of the correlations was treated

as part of the definition of the Spearman hierarchy. There is a second imprecision, which is perhaps more serious. If we define

$$(5) \quad \text{mr}_*(\Sigma) \equiv \min \{ \text{rank}(\Sigma - \Omega) \mid \Omega \text{ diagonal} \},$$

then the vanishing of all tetrads (or of $\frac{1}{2}m(m-3)$ independent tetrads) is necessary and sufficient for $\text{mr}_*(\Sigma) < 1$. However, in general,

$$(6) \quad \text{mr}_*(\Sigma) < \text{mr}(\Sigma),$$

and there is no guarantee of equality. Remarks to this effect were already made by Garnett, but it was pointed out for the first time by Wilson (1928) and Camp (1932) that the conditions

$$(7) \quad \rho_{jk} > \rho_{ik}\rho_{ji}$$

were necessary as well for $\text{mr}(\Sigma) = 1$.

It is remarkable that the formulation of the conditions for $\text{mr}(\Sigma) = 1$ took about thirty years. In fact, the results can be summarized in a single comprehensive theorem. Once it is formulated, the proof is almost immediate. In order to do so we assume, without loss of generality, that Σ is irreducible, i.e. Σ cannot be brought, by permutations, into block-diagonal form. We also say that Σ is a Spearman matrix if $\text{mr}(\Sigma) = 1$.

Theorem 1. A positive definite, irreducible matrix Σ is a Spearman matrix if and only if, after sign changes of rows and corresponding columns, all its elements are positive and such that

$$(8) \quad \sigma_{ik}\sigma_{j\ell} - \sigma_{i\ell}\sigma_{jk} = 0,$$

and

$$(9) \quad \sigma_{ik}\sigma_{ji} - \sigma_{ii}\sigma_{jk} < 0,$$

for all quadruples $(i \neq j \neq k \neq \ell)$.

Proof is omitted.

There have been various attempts to generalize the approach of this theorem, and the kind of result we have obtained, to various other combinations of m and p (with $p = mr(\Sigma)$: the numbers of common factors). Kelley (1928), Wilson (1929), Wilson and Worcester (1934, 1939) study special cases such as (m,p) equal to $(4,2)$, $(5,2)$, or $(6,3)$. Many different special cases must be distinguished, and very little is achieved in terms of general results.

3. The Ledermann bound

Kelley (1928) also tries to provide much more general results, which are true for all (m,p) . If we write the factor model as $\Sigma = \Lambda\Lambda' + \Omega$, then we find $\frac{1}{2}m(m+1)$ equations in $m + mp - \frac{1}{2}p(p-1)$ unknowns (taking rotational indeterminacy into account). The number of unknowns exceeds the number of equations if

$$(10) \quad p > p(m) \equiv \frac{1}{2}\{2m+1 - (8m+1)^{\frac{1}{2}}\}.$$

Kelley (1928), and later Thurstone (1935), therefore suggest that $mr(\Sigma) < p(m)$ for all Σ . This despite a warning from Wilson. 'There is perhaps no more tricky part of mathematics than that involved in counting equations and variables to determine whether or not the equations can in general be solved. Today this kind of mathematics is, among pure mathematicians, taboo except as a heuristic device.' (Wilson, 1929, p. 156).

Ledermann (1937) has tried to put the bound $mr(\Sigma) < p(m)$ on a somewhat more rigorous footing. Not with much success though. 'Nous offrons un pétale de rose à quiconque énoncera clairement et démontrera sûrement les résultats que Ledermann a voulu nous communiquer.' (Hakim et al., 1976, part II, p. 2). We shall make an attempt. Ledermann writes $mr_*(\Sigma) < p$ as a system of $\frac{1}{2}(m-p)(m+1-p)$ determinantal equations with the elements of Ω as the m unknowns. Of course each determinant can be expanded, which gives a system of $\frac{1}{2}(m-p)(m+1-p)$ polynomial equations in m unknowns. Again the number of unknowns exceeds the number of equations if $p > p(m)$. Ledermann proves, in addition, that at least for

some choices of Σ these equations are independent, i.e. none of them is a consequence of the others. Of course this still does not imply much about their solvability, or about the number of solutions they have.

Perhaps the most general result about the Ledermann bound, which has been proved rigorously, is due recently to Shapiro (1982a). He proves that $\text{mr}_*(\Sigma) > p(m)$ almost surely, i.e. the dispersion matrices Σ for which $\text{mr}_*(\Sigma) < p(m)$ form a set of (Lebesgue) measure zero. Although this result is theoretically of some interest, it does not give any valuable information for specific matrices Σ .

4. Beyond the Ledermann bound

We first mention, as an important step ahead, the work of Albert (1944a, 1944b). He defined $\mu(\Sigma)$, the ideal rank of Σ , to be the largest nonsingular square off-diagonal submatrix. Obviously $\text{mr}_*(\Sigma) > \mu(\Sigma)$. Albert gives necessary and sufficient conditions for equality in his first paper, and he gives a sufficient condition for equality in the second paper. Tumura and Fukutomi (1968) give another sufficient condition.

Guttman (1954, p. 160) observed that 'merely studying the minors outside the main diagonal was not sufficient' for determining $\text{mr}(\Sigma)$, and he showed that $\text{mr}(\Sigma) = m - 1$ for correlation matrices with two different latent roots, the largest of which with a multiplicity of $m - 1$.

Guttman (1956, theorem 1) argued that $\text{mr}(\Sigma) < m - k$ if Σ^{-1} has a $k \times k$ diagonal principal submatrix. In the same paper he presented an interesting inequality which may be reformulated as follows:

Theorem 2. If Σ is irreducible, then

$$(11) \quad \text{mr}(\Sigma) + \text{mr}(\Sigma^{-1}) > m.$$

Proof: (adapted from Guttman (1956)) If $0 < \Omega < \Sigma$, then it follows that $\Omega < D_\Sigma$, where D_Σ is a diagonal matrix with elements as in Σ . As $0 < \Omega < \Sigma$ implies that $\Omega > \Omega \Sigma^{-1} \Omega$ (cf. Bekker et al., 1984), it also follows that $\Omega < D_{\Sigma^{-1}}$. Let $\pi(\cdot)$ denote the number of positive latent

roots, and $\pi^*(\cdot)$ the number of non-negative latent roots, then $0 < \Omega < \Sigma$ implies that

$$(12) \quad \Sigma - \Omega > \Sigma - D \begin{matrix} -1 \\ (\Sigma^{-1}) \end{matrix}$$

so

$$(13) \quad \text{rank}(\Sigma - \Omega) = \pi(\Sigma - \Omega) > \pi\left(\Sigma - D \begin{matrix} -1 \\ (\Sigma^{-1}) \end{matrix}\right) = \pi\left(D \begin{matrix} \frac{1}{2} \\ (\Sigma^{-1}) \end{matrix} \Sigma D \begin{matrix} \frac{1}{2} \\ (\Sigma^{-1}) \end{matrix} - I\right),$$

where the last equality follows from Sylvester's law of inertia. If Σ is irreducible, then $0 < \Omega^* < \Sigma^{-1}$, where Ω^* is diagonal, implies that

$$\Omega^* < D \begin{matrix} -1 \\ (\Sigma^{-1}) \end{matrix}, \text{ so that}$$

$$(14) \quad \Sigma^{-1} - \Omega^* > \Sigma^{-1} - D \begin{matrix} -1 \\ (\Sigma^{-1}) \end{matrix},$$

and thus,

$$(15) \quad \begin{aligned} \text{rank}(\Sigma^{-1} - \Omega^*) &= \pi(\Sigma^{-1} - \Omega^*) > \pi\left(\Sigma^{-1} - D \begin{matrix} -1 \\ (\Sigma^{-1}) \end{matrix}\right) = \\ &= \pi^*\left(D \begin{matrix} -\frac{1}{2} \\ (\Sigma^{-1}) \end{matrix} \Sigma^{-1} D \begin{matrix} -\frac{1}{2} \\ (\Sigma^{-1}) \end{matrix} - I\right). \end{aligned}$$

The theorem follows immediately from (13) and (15).

Q.E.D.

In order to prove that his inequality was the 'best possible' one, Guttman provided examples for which the inequality becomes an equality. Curiously, he didn't use Spearman matrices for this purpose. Obviously, as $\text{mr}(\Sigma) < m - 1$ for any Σ (cf. Guttman, 1954, p. 159), it must hold true that (11) becomes an equality if $\text{mr}(\Sigma) = 1$, or $\text{mr}(\Sigma^{-1}) = 1$. As an interesting consequence of (11) we thus have that $\text{mr}(\Sigma) = m - 1$ if Σ^{-1} is an irreducible Spearman matrix.

The fallacy behind interpreting the Ledermann bound as providing an upper bound to the number of common factors was further discussed in Guttman (1958). It was remarked that a symmetric tridiagonal Σ , with all subdiagonal elements non-zero, had $\text{mr}_*(\Sigma) = m - 1$. Guttman was mistaken in his assertions made in the same paper that for the 'perfect simplex' $\text{mr}_*(\Sigma) = m - 2$, and for the 'quasi-simplex' $\text{mr}_*(\Sigma) = m - 3$. He tried to prove that if Σ^{-1} is tridiagonal, as is true for the 'perfect simplex', then $\text{mr}_*(\Sigma) = m - 2$. This is very strange, since application of the aforementioned theorem 1 in Guttman (1956) shows that if Σ^{-1} is tridiagonal

then even $\text{mr}(\Sigma)$ is in general smaller than $m - 2$.

Tamura and Fukutomi (1968) proved that the tridiagonality of Σ , with non-zero subdiagonal elements, is not only sufficient for $\text{mr}_*(\Sigma) = m - 1$, it is, after permutation, also necessary. Their proof is difficult to understand. Hakim et al. (1976, part II) give another proof, which is more solid but a bit complicated. Fairly simple proofs are available in Fiedler (1969) and Rheinboldt and Sheperd (1974), who discovered the theorem in an entirely different context.

A slightly more interesting theorem, which was proved by Shapiro (1982b), implies the existence of a set of non-zero (Lebesgue) measure satisfying $\text{mr}(\Sigma) = m - 1$. Let

$$(16) \quad \text{mr}(\Sigma)_{\zeta} \equiv \min \{ \text{rank}(\Sigma - \Omega) \mid \Omega \prec \Sigma; \Omega \text{ diagonal} \},$$

so that

$$(17) \quad \mu(\Sigma) \prec \text{mr}_*(\Sigma) \prec \text{mr}(\Sigma)_{\zeta} \prec \text{mr}(\Sigma).$$

Shapiro proved that a necessary and sufficient condition for $\text{mr}(\Sigma) = m - 1$ is that all off-diagonal elements of Σ can be made non-positive by sign changes of rows and corresponding columns.

So far we have found conditions which can be considered as sufficient conditions for $\text{mr}(\Sigma) = m - 1$. Necessary and sufficient conditions are given in the following theorem, which is similar to a result that has been proved by Hakim et al. (1976, part I, corollaire 2.4).

Theorem 3. $\text{mr}(\Sigma) = m - 1$ if and only if for each vector $\gamma' = (\gamma_1, \dots, \gamma_m) \neq 0$ such that $(\Sigma - \Omega)\gamma = 0$, where Ω is diagonal and $0 \prec \Omega \prec \Sigma$, $\gamma_i \neq 0$ for all $i = 1, \dots, m$.

Proof is omitted.

Of course, the conditions in this theorem are not very satisfactory. However, it is an important step towards the formulation of simple necessary and sufficient conditions for $\text{mr}(\Sigma) = m - 1$. These conditions, and also the theorem (cf. Reiersøl, 1941, theorem 12), have been derived in an entirely different tradition, which we will now review.

5. Structural regression

Consider the following 'errors-in-variables model'

$$(18) \quad \xi' \gamma = 0,$$

$$(19) \quad x = \xi + \varepsilon,$$

where the m variables in ξ are not observed, instead the m variables in x are observed. It is assumed that the disturbances, or measurement errors, in ε are mutually independent and also independent of the systematic parts in ξ . The (fixed) m -vector γ is called the structural vector.

Obviously, the model can be considered as a regression model where all variables are subject to measurement error. Indeed, if only one of the errors in ε has a non-zero variance, so that $m - 1$ errors equal zero identically, then the model represents an elementary regression, where one of the variables is regressed on the other variables. As we can do this for each variable, we can also find m different elementary regression vectors γ .

If we assume that $x \sim N(0, \Sigma)$, and $\varepsilon \sim N(0, \Omega)$, where Σ is nonsingular and Ω is diagonal positive semidefinite. Then, of course, $\xi \sim N(0, \Sigma - \Omega)$, so that also $\Sigma - \Omega$ must be positive definite:

$$(20) \quad 0 < \Omega < \Sigma.$$

Furthermore, (18) and (19) may be replaced by the moment equations

$$(21) \quad (\Sigma - \Omega)\gamma = 0.$$

If Ω is known up to a proportionality factor, i.e. $\Omega = \mu \Omega^*$, where Ω^* is some known fixed positive semidefinite matrix, then the equations in (20) and (21) can be used for estimation purposes. That is to say, Σ should be replaced by its sample estimate $\hat{\Sigma}$, μ should be set equal to the smallest root $\hat{\mu}$ of the determinantal equation $\hat{\Sigma} - \hat{\mu} \Omega^* = 0$, and γ may be estimated by $(\hat{\Sigma} - \hat{\mu} \Omega^*)\hat{\gamma} = 0$. This gives consistent estimates, even if the variables are not normally distributed. As an ex-

ample we may consider 'orthogonal regression', where $\Omega^* = I$, which was introduced by Pearson in 1901, or we may consider the m different elementary regressions, where each time Ω has only one non-zero diagonal element. It follows immediately from (21) that the i th elementary regression vector γ must be proportional to the i th column of Σ^{-1} ; the estimate of the i th elementary regression vector is thus proportional to the i th column of $\hat{\Sigma}^{-1}$.

However, in general, the dispersion matrix of the measurement errors is not known up to a proportionality factor, and many diagonal matrices Ω and vectors γ satisfy both (20) and (21). In other words, the model is not identified; which corresponds to the underidentification of a factor model with $m - 1$ factors.

Another problem is that there may exist diagonal matrices Ω satisfying (20) such that $\text{rank}(\Sigma - \Omega) < m - 1$. In that case one may not exclude the possibility that there exist two, or even more, linear relations between the systematic parts in ξ . Consequently, as has been noted by Frisch (1934, p. 191), it would be sheer nonsense, in such cases, to look for significant elementary regression coefficients.

Frisch (1934) was the first to study these problems in some depth in his 'confluence analysis'. In particular he proved that for two observed variables, as in simple regression, the structural regression vector must be a convex linear combination of the two elementary regression vectors. As a result the correct regression line is located between the two elementary regression lines.

Although Frisch conjectured that similar conditions held in the general m -variables case, Koopmans (1937, p. 98-101) was the first to present an m -variable analog of Frisch's result. It says that the structural regression vector is a convex linear combination of the elementary regression vectors, subject to the condition that all elements of Σ^{-1} are strictly positive. It is clear that, as the columns of Σ^{-1} are proportional to the elementary regression vectors, the condition in the theorem can be satisfied, after sign changes, if all elementary regression vectors are located in the same orthant.

Koopmans' proof is complicated. Reiersøl (1941, p. 8) noted the applicability of a theorem by Frobenius, and all later proofs given by Reiersøl (1945), Dhondt (1960), Patefield (1981), Kalman (1982a) and

Klepper and Leamer (1984) make use of the Perron-Frobenius theorem, just as Shapiro (1982b) did when proving his result on $\text{mr}_\zeta(\Sigma) = m - 1$.

The second part of Koopmans' theorem says that if all elements of Σ^{-1} are strictly positive, then each vector in the convex hull of the elementary regression vectors is a structural vector for some diagonal error dispersion matrix Ω satisfying both (20) and (21). Koopmans (1937, p. 103) also claimed to have proved this second part of the theorem. However, as has been pointed out by Kalman (1982a, p. 152), Koopmans' proof was wrong. Later proofs were given by Reiersøl (1945), Kalman (1982a) and Klepper and Leamer (1984); again all authors use the Perron-Frobenius theorem.

Here a formulation of the theorem will be presented which is slightly more general than previous formulations. The theorem will be proved without using the Perron-Frobenius theorem. Furthermore, the two parts of the theorem will be proved almost simultaneously, thereby emphasizing the if and only if argument in the theorem.

It will be convenient to use the following lemma. Let A be a symmetric matrix with strictly positive off-diagonal elements, $A_{ij} > 0$ if $i \neq j$. Let Λ be a diagonal matrix, $\Lambda = \text{diag}(\lambda)$, and let u be a vector of ones, $u' = (1, \dots, 1)$, q is an arbitrary vector.

Lemma 1. $\text{diag}(\Lambda\Lambda u) > \Lambda\Lambda$ if and only if $\lambda_i \lambda_j > 0$ for all i, j .

Proof: $q' \{ \text{diag}(\Lambda\Lambda u) - \Lambda\Lambda \} q = \sum_i \sum_j A_{ij} \lambda_i \lambda_j (q_i^2 - q_i q_j) =$
 $= \sum_{i < j} \sum A_{ij} \lambda_i \lambda_j (q_i - q_j)^2$

(i) if $\lambda_i \lambda_j > 0$, for all i, j , then $A_{ij} \lambda_i \lambda_j (q_i - q_j)^2 > 0$, for all i, j ,

(ii) if for some i, j $\lambda_i \lambda_j < 0$, then choose $q_i = \text{sign}(\lambda_i)$, so that

$$\sum_{i < j} A_{ij} \lambda_i \lambda_j (q_i - q_j)^2 < 0. \quad \text{Q.E.D.}$$

We will also use the result that $0 < \Omega < \Sigma$ is equivalent to $\Omega > \Omega \Sigma^{-1} \Omega$ if Σ is positive definite. The proof is simple as both matrices can be diagonalized simultaneously (cf. Bekker et al., 1984).

Contrary to most other proofs we do not assume that the error dispersion matrix Ω is nonsingular. Without loss of generality we assume

$$(22) \quad \Omega \equiv \begin{bmatrix} \Omega_1 & 0 \\ 0 & 0 \end{bmatrix},$$

where Ω_1 is a $k \times k$ diagonal matrix, $k < m$. Σ and Σ^{-1} are partitioned analogously, i.e.

$$(23) \quad \Sigma \equiv \begin{bmatrix} \Sigma_{11} & \Sigma_{12} \\ \Sigma_{21} & \Sigma_{22} \end{bmatrix}; \quad \Sigma^{-1} \equiv \begin{bmatrix} (\Sigma^{-1})_{11} & (\Sigma^{-1})_{12} \\ (\Sigma^{-1})_{21} & (\Sigma^{-1})_{22} \end{bmatrix} \equiv (\Sigma_1^{-1}, \Sigma_2^{-1}),$$

and also the vector γ has a corresponding partitioning $\gamma' = (\gamma_1', \gamma_2')$.

Theorem 4. Let $(\Sigma^{-1})_{11}$ have strictly positive elements and let

$(\Sigma - \Omega)\gamma = 0$, where Ω is as in (22);

(i) if $0 < \Omega < \Sigma$, then γ lies in the convex hull of Σ_1^{-1} ,

(ii) for each γ in the convex hull of Σ_1^{-1} there exists one and only one Ω_1 such that $0 < \Omega < \Sigma$.

Proof: Define $\lambda \equiv \Omega_1 \gamma_1$, $\Lambda \equiv \text{diag}(\lambda)$, $C \equiv \text{diag}(\gamma_1)$, so that

$$\gamma = \Sigma_1^{-1} \lambda, \quad \Lambda = \Omega_1 C, \quad \text{and } C = \text{diag}((\Sigma^{-1})_{11} \Lambda u).$$

(1)

$$(24) \quad 0 < \Omega < \Sigma$$

$$(25) \quad \Omega_1 > \Omega_1 (\Sigma^{-1})_{11} \Omega_1$$

$$(26) \quad \Omega_1 C > \Omega_1 (\Sigma^{-1})_{11} \Omega_1 C$$

$$(27) \quad \text{diag}(\Lambda (\Sigma^{-1})_{11} \Lambda u) > \Lambda (\Sigma^{-1})_{11} \Lambda$$

$$(28) \quad \lambda_i \lambda_j > 0, \quad i, j = 1, \dots, k,$$

thus $\gamma = \Sigma_1^{-1} \lambda$ lies in the convex hull of Σ_1^{-1} .

(ii) If γ lies in the convex hull of Σ_1^{-1} , then γ_1 must have strictly positive elements, so C is nonsingular. Hence $\Omega_1 = \Lambda C^{-1}$ is unique and (25) and (26) are equivalent, so that (28) implies (24). Q.E.D.

Thus we have proved that if $(\Sigma^{-1})_{11}$ has strictly positive elements and only the first k variables are subject to measurement error, then the set of all structural vectors is the convex hull of the first k elementary regression vectors.

This important theorem can be used to derive necessary and sufficient conditions for $\text{mr}(\Sigma) = m - 1$.

Theorem 5. $\text{mr}(\Sigma) = m - 1$ if and only if Σ^{-1} has strictly positive elements, possibly after sign changes of rows and corresponding columns.

Proof: (i) If Σ^{-1} has strictly positive elements, then, by theorem 4(i), the null-space of $\Sigma - \Omega$ is contained within the convex hull of Σ^{-1} . Consequently, this null-space can be at most one-dimensional.

(ii) If not all elements of Σ^{-1} have compatible signs, then there are two columns of Σ^{-1} , the i th and j th say, such that the (ij) th element of Σ^{-1} is positive, possibly after sign changes of rows and columns, while the i th and j th columns do not lie in the same orthant. That is to say that in the convex hull of these columns there is a vector γ with a zero element. According to theorem 4 (ii) this vector γ is a vector in the null-space of $\Sigma - \Omega$ for some diagonal Ω satisfying (20). Then, according to theorem 3, $\text{mr}(\Sigma) < m - 1$. Q.E.D.

The result in this theorem has been proved before by Reiersøl (1941, theorem 14) and by Kalman (1982a). However, Reiersøl's proof is not very transparent; Kalman simply postulates an Ω such that $\text{rank}(\Sigma - \Omega) < m - 1$ in order to prove the second part of the theorem. Klepper and Leamer (1984) presented theorem 5 in a disguised form. They stated that the coefficients in the normalized structural vector are bounded if and only if all elementary regression vectors lie in the same orthant. Their proof is by reducing the m -variable problem to a 3-variable problem.

Using theorem 4(i) it is easy to derive two different generalizations of theorem 5(i). Again, let Σ_{11} be a $k \times k$ principal submatrix of Σ , and let $(\Sigma^{-1})_{11}$ be the corresponding $k \times k$ submatrix of Σ^{-1} , $k < m$.

Theorem 6. (i) If $(\Sigma^{-1})_{11}$ has strictly positive elements then
 $\text{mr}(\Sigma) > k - 1$.

ii) If $(\Sigma_{11})^{-1}$ has strictly positive elements then
 $\text{mr}(\Sigma) > k - 1$.

Proof is omitted.

6. Discussion

Both factor analysis and structural regression analysis are extremes of a common model which simply says that, apart from unique components or error components, the variables in the analysis are linearly related. In other words, the model says that the covariance matrix $\Sigma - \Omega$ has a deficient rank. In factor analysis attention centres round the low-dimensional range-space of $\Sigma - \Omega = \Lambda\Lambda'$. In structural regression the model is formulated in terms of the one-dimensional null-space: $(\Sigma - \Omega)\gamma = 0$.

Evaluation of $\text{mr}(\Sigma)$, or $m - \text{mr}(\Sigma)$, is important in both fields. In fact, the number $\text{mr}(\Sigma)$ tells us whether the common model should be considered as a factor analysis model or as a structural regression model, or even, as some model in between. Consequently, if a structural regression model is justified, i.e. if the necessary and sufficient conditions for $\text{mr}(\Sigma) = m - 1$ are satisfied, then applying a factor model to the data would be nonsense. Just as Frisch thought it nonsense to look for a single linear relation in case there are two or more linear relations between the variables.

These latter models, where there exist a number, albeit a small number, of linear relations between the variables, have had relatively little attention in the literature. It is only recently that Kalman (1982a, 1982b, 1983, 1984) discussed these models in some detail. In his 1983 paper (p. 119) he claims that 'It is impossible to avoid the conclusion that the lack of progress on and the present confusion surrounding Frisch's ideas are due to mathematical rather than conceptual difficulties.' Indeed, as we have seen, there are necessary and sufficient

conditions for $\text{mr}(\Sigma) = 1$ and $\text{mr}(\Sigma) = m - 1$, however, no such conditions are available for intermediate values of $\text{mr}(\Sigma)$.

Although no complete solutions are known for intermediate values of $\text{mr}(\Sigma)$, there are sufficient conditions. For example, using theorems 5 and 6, it is easy to derive sufficient conditions for $\text{mr}(\Sigma) = m - 2$. On the other hand, there does exist a necessary and sufficient condition for an intermediate value of $\text{mr}(\Sigma)$ in case m is small. Clearly if $m = 3$, then necessary and sufficient conditions are available for all values of $\text{mr}(\Sigma)$. If $m = 4$, then $\text{mr}(\Sigma) = 1$ and $\text{mr}(\Sigma) = 3$ are completely characterized by application of theorems 1 and 5 respectively. Consequently, if $m = 4$, also $\text{mr}(\Sigma) = 2$ is completely characterized, since all (non-diagonal) matrices Σ that do not satisfy the necessary and sufficient conditions for $\text{mr}(\Sigma) = 1$ or $\text{mr}(\Sigma) = 3$, and only those matrices Σ , must have $\text{mr}(\Sigma) = 2$.

Kalman (1984, p. 118) also claims that 'it is possible to give a (rigorous) closed-form solution' for the case $m = 5$ and $\text{mr}(\Sigma) = 3$. As we have difficulty in arriving at this closed-form solution, we would like to offer a 'pétale de rose' to anyone who is able to produce that solution.

REFERENCES

- Albert, A.A., 1944a, The matrices of factor analysis, Proc. Nat. Acad. Sci. 30, 90-95.
- Albert, A.A., 1944b, The minimum rank of a correlation matrix, Proc. Nat. Acad. Sci. 30, 144-148.
- Bekker, P.A., A. Kapteyn, and T.J. Wansbeek, 1984, Measurement error and endogeneity in regression: bounds for ML and VI estimates, in T.K. Dijkstra (ed.), Misspecification Analysis, Springer Verlag, Berlin.
- Burt, C., 1909, Experimental tests of general intelligence, British J. Psychol. 3, 94-177.
- Camp, B.H., 1932, The converse of Spearman's Two-Factor theorem, Biometrika 24, 418-428.
- Dhondt, A., 1960, Sur une généralisation d'un théorème de R. Frisch en analyse de confluence, Cahiers du Centre d'Etude de Recherche Operationelle 2, 37-46.
- Fiedler, M., 1969, A characterization of tridiagonal matrices, Lin. Alg. and Appl. 2, 191-197.
- Frisch, R., 1934, Statistical Confluence Analysis by means of Complete Regression Systems, Publication no 5, University of Oslo Economic Institute, 192 pages.
- Garnett, J.C.M., 1919a, On certain independent factors in mental measurement, Proc. Roy. Soc. London, 96, 91-111.
- Garnett, J.C.M., 1919b, General ability, cleverness, and purpose, British J. Psychol. 9, 345-366.

- Garnett, J.C.M., 1920, The single general factor in dissimilar mental measurements, British J. Psychol. 10, 242-258.
- Guttman, L., 1954, Some necessary conditions for common factor analysis, Psychometrika 19, 149-161.
- Guttman, L., 1956, "Best possible" systematic estimates of communalities, Psychometrika 21, 273-285.
- Guttman, L., 1958, To what extent can communalities reduce rank?, Psychometrika 23, 297-308.
- Hakim, M., E.O. Lochard, J.P. Olivier, and E. Térouanne, 1976, Sur les traces de Spearman, Cahiers de Bureau Universitaire de Recherche Operationelle, Université. Pierre et Marie Curie, Paris, Série Recherche N 25.
- Hart, B. and C.E. Spearman, 1912, General ability, its existence and nature, British J. Psychol. 5, 51-84.
- Hearnshaw, L.S., 1981, Cyril Burt, psychologist, Random House, New York.
- Kalman, R.E., 1982a, System identification from noisy data, in A.R. Bednarek and L. Cesari (eds.), Dynamical Systems II, Academic Press, New York, 331-342.
- Kalman, R.E., 1982b, Identification from real data, in M. Hazewinkel and A.H.G. Rinnooy Kan (eds.), Current developments in the Interface Economics, Econometrics, Mathematics, 161-196, D. Reidel Publishing Co, Dordrecht.
- Kalman, R.E., 1983, Identifiability and modelling in econometrics, in P.R. Krishnaiah (ed.), Developments in Statistics 4, 97-136, Academic Press, New York.
- Kalman, R.E., 1984, We can do something about multicollinearity ! Commun. Statist. Theor. Meth. 13, 115-125.

- Kelley, T.L., 1928, Crossroads in the Mind of Man: a Study of Differentiable Mental Abilities, Stanford University Press.
- Klepper, S. and E.E. Leamer, 1984, Consistent sets of estimates for regressions with errors in all variables, Econometrica, 52, 162-183.
- Koopmans, T.C., 1937, Linear Regression Analysis of Economic Time Series, De Erven F. Bohn NV, Haarlem, 150 pages.
- Krueger, F. and C.E. Spearman, 1907, Die Korrelation zwischen verschiedenen geistigen Leistungsfähigkeiten, Z. Psychol. 54, 50-114.
- Ledermann, W., 1937, On the rank of the reduced correlational matrix in multiple factor analysis, Psychometrika, 2, 85-93.
- Patefield, W.M., 1981, Multivariate linear relationships: maximum likelihood estimation and regression bounds, Journ. of the Roy. Statist. Soc. B 43, 342-352.
- Reiersøl, O., 1941, Confluence analysis by means of lag moments and other methods of confluence analysis, Econometrica 9, 1-24.
- Reiersøl, O., 1945, Confluence analysis by means of instrumental sets of variables, Arkiv för Matematik, Astronomi och Fysik, 32 A, 1-119.
- Rheinboldt, W.C. and R.A. Sheperd, 1974, On a characterization of tri-diagonal matrices by M. Fiedler, Lin. alg. appl. 8, 87-90.
- Shapiro, A., 1982a, Rank reducibility of a symmetric matrix and sampling theory of minimum trace factor analysis, Psychometrika, 47, 187-199.
- Shapiro, A., 1982b, Weighted minimum trace factor analysis, Psychometrika, 47, 243-264.

- Spearman, C.E., 1904, General intelligence objectively measured and defined, Amer. J. Psychol. 15, 201-299.
- Spearman, C.E., 1927, The Abilities of Man, McMillan, London.
- Spearman, C.E., and K.J. Holzinger, 1924, The sampling error in the theory of two factors, British J. Psychol. 15, 17-19.
- Spearman, C.E., and K.J. Holzinger, 1925, Note on the sampling error of tetrad differences, British J. Psychol. 16, 86-89.
- Thurstone, L.L., 1935, The Vectors of Mind, University of Chicago Press.
- Tamura, Y., and K. Fukutomi, 1968, On the identification in factor analysis, Rep. Statist. Appl. Res. JUSE 15, 6-11.
- Wilson, E.B., 1928, On hierarchical correlation systems, Proc. Nat. Acad. Sci. 14, 283-291.
- Wilson, E.B., 1929, Review of Kelley, Crossroads in the mind of man, J. Gen Psychol. 2, 153-169.
- Wilson, E.B., and J. Worcester, 1934, The resolution of four tests, Proc. Nat. Acad. Sc. 20, 189-192.
- Wilson, E.B., and J. Worcester, 1939, The resolution of six tests into three general factors, Proc. Nat. Acad. Sci. 25, 73-79.

IN 1984 REEDS VERSCHENEN

- 138 G.J. Cuypers, J.P.C. Kleijnen en J.W.M. van Rooyen
Testing the Mean of an Asymmetric Population:
Four Procedures Evaluated
- 139 T. Wansbeek en A. Kapteyn
Estimation in a linear model with serially correlated errors when
observations are missing
- 140 A. Kapteyn, S. van de Geer, H. van de Stadt, T. Wansbeek
Interdependent preferences: an econometric analysis
- 141 W.J.H. van Groenendaal
Discrete and continuous univariate modelling
- 142 J.P.C. Kleijnen, P. Cremers, F. van Belle
The power of weighted and ordinary least squares with estimated
unequal variances in experimental design
- 143 J.P.C. Kleijnen
Superefficient estimation of power functions in simulation
experiments
- 144 P.A. Bekker, D.S.G. Pollock
Identification of linear stochastic models with covariance
restrictions.
- 145 Max D. Merbis, Aart J. de Zeeuw
From structural form to state-space form
- 146 T.M. Doup and A.J.J. Talman
A new variable dimension simplicial algorithm to find equilibria on
the product space of unit simplices.
- 147 G. van der Laan, A.J.J. Talman and L. Van der Heyden
Variable dimension algorithms for unproper labellings.
- 148 G.J.C.Th. van Schijndel
Dynamic firm behaviour and financial leverage clienteles
- 149 M. Plattel, J. Peil
The ethico-political and theoretical reconstruction of contemporary
economic doctrines
- 150 F.J.A.M. Hoes, C.W. Vroom
Japanese Business Policy: The Cash Flow Triangle
an exercise in sociological demystification
- 151 T.M. Doup, G. van der Laan and A.J.J. Talman
The $(2^{n+1}-2)$ -ray algorithm: a new simplicial algorithm to compute
economic equilibria

IN 1984 REEDS VERSCHENEN (vervolg)

- 152 A.L. Hempenius, P.G.H. Mulder
Total Mortality Analysis of the Rotterdam Sample of the Kaunas-Rotterdam Intervention Study (KRIS)
- 153 A. Kapteyn, P. Kooreman
A disaggregated analysis of the allocation of time within the household.
- 154 T. Wansbeek, A. Kapteyn
Statistically and Computationally Efficient Estimation of the Gravity Model.
- 155 P.F.P.M. Nederstigt
Over de kosten per ziekenhuisopname en levensduurmodellen
- 156 B.R. Meijboom
An input-output like corporate model including multiple technologies and make-or-buy decisions
- 157 P. Kooreman, A. Kapteyn
Estimation of Rationed and Unrationed Household Labor Supply Functions Using Flexible Functional Forms
- 158 R. Heuts, J. van Lieshout
An implementation of an inventory model with stochastic lead time
- 159 P.A. Bekker
Comment on: Identification in the Linear Errors in Variables Model
- 160 P. Meys
Functies en vormen van de burgerlijke staat
Over parlementarisme, corporatisme en autoritair etatisme
- 161 J.P.C. Kleijnen, H.M.M.T. Denis, R.M.G. Kerckhoffs
Efficient estimation of power functions
- 162 H.L. Theuns
The emergence of research on third world tourism: 1945 to 1970;
An introductory essay cum bibliography
- 163 F. Boekema, L. Verhoef
De "Grijze" sector zwart op wit
Werklozenprojecten en ondersteunende instanties in Nederland in kaart gebracht
- 164 G. van der Laan, A.J.J. Talman, L. Van der Heyden
Shortest paths for simplicial algorithms
- 165 J.H.F. Schilderink
Interregional structure of the European Community
Part II: Interregional input-output tables of the European Community 1959, 1965, 1970 and 1975.

IN (1984) REEDS VERSCHENEN (vervolg)

- 166 P.J.F.G. Meulendijks
An exercise in welfare economics (I)
- 167 L. Elsner, M.H.C. Paardekooper
On measures of nonnormality of matrices.

IN 1985 REEDS VERSCHENEN

- 168 T.M. Doup, A.J.J. Talman
A continuous deformation algorithm on the product space of unit simplices
- 169 P.A. Bekker
A note on the identification of restricted factor loading matrices
- 170 J.H.M. Donders, A.M. van Nunen
Economische politiek in een twee-sectoren-model
- 171 L.H.M. Bosch, W.A.M. de Lange
Shift work in health care
- 172 B.B. van der Genugten
Asymptotic Normality of Least Squares Estimators in Autoregressive Linear Regression Models
- 173 R.J. de Groof
Geïsoleerde versus gecoördineerde economische politiek in een twee-regiomodel
- 174 G. van der Laan, A.J.J. Talman
Adjustment processes for finding economic equilibria
- 175 B.R. Meijboom
Horizontal mixed decomposition
- 176 F. van der Ploeg, A.J. de Zeeuw
Non-cooperative strategies for dynamic policy games and the problem of time inconsistency: a comment
- 177 B.R. Meijboom
A two-level planning procedure with respect to make-or-buy decisions, including cost allocations
- 178 N.J. de Beer
Voorspelprestaties van het Centraal Planbureau in de periode 1953 t/m 1980
- 178a N.J. de Beer
BIJLAGEN bij Voorspelprestaties van het Centraal Planbureau in de periode 1953 t/m 1980
- 179 R.J.M. Alessie, A. Kapteyn, W.H.J. de Freytas
De invloed van demografische factoren en inkomen op consumptieve uitgaven
- 180 P. Kooreman, A. Kapteyn
Estimation of a game theoretic model of household labor supply
- 181 A.J. de Zeeuw, A.C. Meijdam
On Expectations, Information and Dynamic Game Equilibria

- 182 Cristina Pennavaja
Periodization approaches of capitalist development.
A critical survey
- 183 J.P.C. Kleijnen, G.L.J. Kloppenburg and F.L. Meeuwsen
Testing the mean of an asymmetric population: Johnson's modified T
test revisited
- 184 M.O. Nijkamp, A.M. van Nunen
Freia versus Vintaf, een analyse
- 185 A.H.M. Gerards
Homomorphisms of graphs to odd cycles
- 186 P. Bekker, A. Kapteyn, T. Wansbeek
Consistent sets of estimates for regressions with correlated or
uncorrelated measurement errors in arbitrary subsets of all
variables

Bibliotheek K. U. Brabant



17 000 01059752 5