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The relationship of newborn adiposity to fetal growth outcome based on birth weight or the modified neonatal growth assessment score

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Abstract

Objectives—(1) Develop reference ranges of neonatal adiposity using air displacement plethysmography. (2) Use new reference ranges for neonatal adiposity to compare two different methods of evaluating neonatal nutritional status.

Methods—Three hundred and twenty-four normal neonates (35–41 weeks post-menstrual age) had body fat (%BF) and total fat mass (FM, g) measured using air displacement plethysmography shortly after delivery. Results were stratified for 92 of these neonates with corresponding fetal biometry using two methods for classifying nutritional status: (1) population-based weight percentiles; and (2) a modified neonatal growth assessment score (m_3NGAS_{51}).

Results—At the 50th percentile, &BF varied from 7.7% (35 weeks) to 11.8% (41 weeks), while the corresponding 50th percentiles for total FM were 186–436g. Among the subset of 92 neonates, no significant differences in adiposity were found between small for gestational age (SGA), appropriate for gestational age (AGA), and large for gestational age (LGA) groups using population-based weight standards. Classification of the same neonates using m_3NGAS_{51} showed significant differences in mean %BF between corresponding groups.

Conclusions—Population-based weight criteria for neonatal nutritional status can lead to misclassification on the basis of adiposity. A neonatal growth assessment score, that considers the

Declaration of Interest

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growth potential of several anatomic parameters, appears to more effectively classify under-and over-nourished newborns.

Keywords

Air displacement plethysmography; body composition; body fat; fetus; malnutrition; ultrasound

INTRODUCTION

A primary goal of obstetrical care is the timely detection of malnourished fetuses because they have increased the risks for perinatal morbidity, neurodevelopmental disorders, short stature, and metabolic abnormalities in adult life [1]. Pediatricians usually categorize the nutritional state of newborns using birth weight although other techniques such as skin-fold measurements or dual-energy X-ray absorptiometry have also been described for this purpose [2, 3]. Obstetricians have alternatively used sonographic measurements of the head, trunk, and limb to estimate fetal weight [4]. However, this indirect calculation is associated with random errors that typically range between 8.1 and 11.8% [5].

Fetal growth is a result of simultaneous changes in body mass, shape, and composition. Adipose tissue development may be particularly sensitive to variations in fetal growth [6]. Human newborns are uniquely distinct from other mammals because they contain relatively large amounts of fat – up to approximately 15% of their body weight at term [7, 8]. Adipose tissue is an efficient system, to store energy for the process of subsequent neonatal growth [9]. From an evolutionary perspective, mammalian fat provides an energy store that supports rapid fetal brain growth and protects newborns against the possibility of nutritional deprivation [10]. Fetal fat deposition alone represents over half of the caloric accretion during the early third trimester [6]. By late pregnancy, this process increases to about 90% of the caloric accretion at a time when total caloric requirement is 90–100 kcal/kg/day [6]. A more comprehensive understanding of neonatal body composition may facilitate our identification of malnourished fetuses with regards to their specific nutrient and energy requirements.

A poor correlation has been reported between estimated fetal weight (EFW) and neonatal adiposity [11, 12]. Although neonatal fat mass constitutes only 14% of total birth weight, it accounts for 46% of its variance in newborns [13]. Malnourished fetuses can even express abnormal changes in newborn body composition and hormonal concentrations (e.g. IGF-I and IGFBP-3) despite birth weights within the normal range [14, 15]. Such observations are particularly relevant to studies that define intrauterine growth retardation (IUGR) on the basis of fetal abdominal circumference alone [16, 17].

The current investigation uses noninvasive air displacement plethysmography to develop reference ranges for neonatal adiposity in a large number of preterm and term newborns shortly after delivery. These results are used to compare two methods of neonatal nutritional assessment according to adiposity: (1) population-based percentile ranges based on weight only; and (2) a composite scoring system that, in addition to weight, also considers the growth potential of multiple growth parameters.

METHODS

This was a prospective, cross-sectional investigation of newborns delivered to women, who were invited to participate in this study. All women had provided written informed consistent to participate in this study that was approved by the Internal Review Boards at William Beaumont Hospital, Wayne State University, Detroit, MI, USA and the *Eunice*

Kennedy Shriver National Institute of Child Health and Human Development, Bethesda, MD, USA. The protocol enrollment was not consecutive. Inclusion criteria consisted of neonates that were delivered during their third trimester of pregnancy. Pregnancies with poor menstrual dating criteria, multiple gestations, and fetuses with congenital anomalies were excluded from this study.

The normal population sample (n = 15) and smoking (n = 16). Two approaches for nutritional classification were employed: The first one was based on birth weight and postmenstrual age (PMA [18]); the second method used a previously described modified neonatal growth assessment score [19, 20]. This comparison included an additional cohort of 25 neonates from mothers with a history of diabetes (Gestational diabetes = 13; Type 1 diabetes = 9; Type 2 diabetes = 3). Maternal age, menstrual age at time of scan, neonatal gender, ethnicity, and presence of obstetrical complications were documents. Ninety-six neonates from the current cohort have been previously reported in two other investigations: the correlation between fetal biometry and percentage body fat (%BF) in 87 term newborns [14] and eight other fetuses that were measured for a soft tissue parameter known as fractional limb volume [21].

Birth age was based on the first day of the last normal menstrual period and confirmed by either first or early second trimester ultrasound scans. A normal last menstrual period was defined as regular cyclic menses without antecedent oral contraceptive use. Age estimates in the first trimester were based on crown-rump length measurements [22]. Age estimates in the second trimester were determined by using biparietal diameter (BPD), head circumference (HC), abdominal circumference (AC), and femoral diaphysis length (FDL) measurements [23–26]. Sonographic age was used to adjust menstrual age if there was more than 1-week discrepancy between menstrual dating and sonographic assessment.

Neonates were examined within approximately 24–48 h of delivery. For purposes of this investigation, the age of neonatal body composition studies was expressed as "post-menstrual age [27]". Anthropometric assessment included weight and length measurements (StatiometerTM Ellard Instrumentation Ltd., Seattle, WA, USA). Growth assessment was based on measurements of body weight (WT), HC, AC, mid-thigh circumference (ThC), and crown-heel length (CHL) as previously described [28].

Percentage body fat (%BF) and total body fat mass (FM, g) were estimated using direct measurements of volume and body mass based on principles of whole body densitometry (Pea Pod Infant Body Composition System, Life Measurements, Inc., Concord, CA, USA). The standard algorithm compensates for physiological changes in hydration in both term and preterm neonates by adjusting the contribution of total body water to corresponding fat-free density values during the first few days following delivery [29]. The validity, accuracy and reliability of air displacement plethysmography have been previously described for assessment of %BF in adults and neonates [30–32].

Newborns were examined without clothing for air displacement plethysmography studies. Small adjustment for the weight and volume of an umbilical cord clip and name bracelet were compensated by separately subtracting them from the whole body densitometry measurement. Volume measurement artifacts in babies with excessive amounts of hair were minimized from the use of baby oil. After body weight was obtained on an electronic scale, the neonate was placed into a temperature-controlled enclosed chamber that permitted a body volume measurement from well-known pressure/volume relationships. The volume measurement was completed within approximately 2 min. Body density was calculated from these data and applied to the two-compartment model to estimate %BF [8]. Fat mass (g) was calculated by multiplying body mass by the %BF.

Statistical analysis

Standard statistical methods were used to compute means, standard deviations and medians. Mean values were expressed \pm standard deviation (1 SD). Histograms were constructed to inspect data distributions and Kolmogorov-Smirnov tests were used to evaluate normality. Linear regression models were examined with their Pearson correlation coefficients to investigate the relationship between PMA and neonatal adiposity. Where appropriate, group comparisons were performed with ANOVA and Levene's test for equality of variances was used. *p*-values less than 0.05 were considered statistically significant. Among the subset of 92 patients, after using ANOVA to test for differences among groups and Leven's test to verify equality of their variances, Student-Newman-Keuls tests were applied for pairwise comparisons between groups.

Reference ranges were calculated using the methodology outlined by Royston to determine the optimal model and its *z*-scores [33, 34]. Percent BF and absolute FM were the dependent variables; post-menstrual age and newborn weight were the independent variables. The simplest, best model fit was determined and plots of the resulting residuals (difference between observed data and predicted values from the model) were examined to determine whether they conformed to a normal, Belgium-shaped distribution. Kolmogorov-Smirnov tests for normality were applied. Using the method suggested by Altman [35], the standard deviation of the residuals was examined across the range of values for the independent variable. If there were significant heteroscedasticity, then regression of absolute residuals versus the independent variables was used to adjust the standard deviation.

Based on prior estimates of standard deviation for %BF in newborns of approximately 3% [36–38], we estimated that equal sample sizes of approximately 150 for each gender would be required to detect a difference of 1% BF between males and females on the basis of 80% power and p < 0.05, MedCalc (v.12.1.4.0, 2011, MedCalc Software bvba, Mariakerke, Belgium). Both linear and multiple regression models were used to evaluate the outcome variables %BF and FM, with inputs of newborn weight, gestational age, gender, race and pregnancies complicated by gestational diabetes. For multiple regression, stepwise methods were used and variables entered the model if p < 0.05 and removed if p > 0.010. The final model included only those variables with p < 0.05. Data analyses were performed with Excel (Microsoft, 2003), MedCalc (v.12.1.4.0, 2011, MedCalc Software bvba, Mariakerke, Belgium) and S-Plus (v.7.0, 2005, S-Plus Software, Palo Alto, CA, USA).

Body fat estimates were stratified using two different methods of classifying neonatal nutritional status. The first method was based on a population study of 6,690,717 singleton neonates who were delivered between 22 and 44 weeks in the United States [18]. Birth weights and gender were used to define small for gestational age (SGA, birth weight $< 10^{th}$ percentile, AGA, birth weight between 10^{th} and 90^{th} percentile, and LGA, birth weight $> 90^{th}$ percentile). A second classification system was applied to a subset of 92 neonates using both fetal sonographic and neonatal measurements. Figure 1 summarizes key formulas required to evaluate neonatal growth outcome based on the use of each fetus as its own control. Rossavik growth models were used to predict birth characteristics of each newborn on the basis of second trimester fetal biometry (HC, AC, FDL, ThC, FDL). The analysis used an assumption of growth cessation at either birth age [delivery before 38 weeks, PMA] or the actual values of 38 weeks [delivery after 38 weeks [28]]. A Growth Potential Realization Index (GPRI) was then calculated as an outcome variable for neonatal growth status by comparing predicted to actual birth measurements [28].

Deter and colleagues [39] have previously described a composite neonatal growth assessment score (NGAS) that combines weight (WT), head circumference (HC), abdominal circumference (AC), thigh circumference (ThC), and crown-heel length (CHL). The NGAS

score originally assumed that each component had equal importance in determining neonatal growth status although weight and soft tissue development (i.e. thigh circumference) are abnormal more often in fetuses with growth abnormalities. A refinement of this NGAS score, the modified neonatal growth assessment score (m_3NGAS_{51}), was subsequently developed with weighting factors for each growth parameter using principal components analysis [19, 20]. Unlike the conventional approach that relies on population-based standard, the m_3NGAS_{51} evaluates the growth potential of five weighted parameters on the basis of multiple GPRI values. The sum of the products of parameter-specific weighting factors with their fetus-specific GPRI values gives the m_3NGAS_{51} for each neonate. The scores were used to classify neonatal growth outcome into one of three nutritional states: IUGR< normal, and macrosomia [20]. More detailed explanations of the modified neonatal growth assessment score and application of Rossavik growth models are provided as a downloadable "Frequently Asked Questions" Supplementary File (S1).

RESULTS

The primary study cohort consisted of 324 neonates (SGA = 42, AGA = 240, LGA = 42) who were delivered between August 2004 and April 2009. The mean maternal age was 29.3 \pm 6.0 years with a median birth age of 38.9 weeks (interquartile range 37.6 – 39.7 weeks, PMA). Birth weights were normally distributed (mean BQW = 3296 \pm 560 g, range 1935 – 4739 g). Ethnic groups included Caucasian (66.0%), African-American (25.6%), Asian (6.2%), other (2.2%). All neonates were studied within 60 h after birth (range = 1 – 59 h, median age = 22.3 h) with 95% of the pediatric examinations being performed within 48 h after delivery. Average weight loss from delivery time to performance of the neonate body composition study was 116 \pm 79 g (-3.5 \pm 2.4%, paired *t*-test = 26.3, *p* < 0.0001).

Reference ranges for %BF and FM (g) in newborns, delivered between 35 - 41 weeks PMA, are summarized in Table I. The best model for predicting %BF and FM was with simple, linear regression. While the residuals for each model conformed to a normal distribution, the standard deviation of %BF remained constant as PMA increased, but the standard deviation of FM increased significantly with PMA. Therefore, the standard deviation of %BF was constant (3.8%), but the standard deviation of FM increased with PMA (SD = $-399 + 14.3 \times PMA$). At the 50th percentile, mean %BF ranged between 7.7 and 11.8%, whereas total FM ranged from 186 to 436 g (with respective SD of 102 - 187 g). A downloadable Supplementary File includes a breakdown of %BF and FM by gender (S2).

Similarly, using newborn weight as the independent variable for prediction of %BF and FM, the best models were simple, linear regression. Again, the residuals were normally distributed and the standard deviation of %BF was constant across the range of newborn weights examined, while the standard deviation of FM increased proportionately to newborn weight.

Relationship of post-menstrual age to %BF versus total fat mass

Weak linear correlations were found between post-menstrual age (PMA) and either %BF (Figure 2a) or FM (Figure 2b) among neonates born during the third trimester of pregnancy. These correlations were not significantly different by Fisher's *z*-test. Considerable overlap occurred between the SGA, AGA and LGA subgroups. Regression analyses indicate that the AGA and AGS subsets were statistically indistinguishable, although LGA neonates had a slightly greater increase of %BF and FM with increasing PMA.

Relationship of neonatal weight to %BF versus total fat mass

Neonatal body composition studies were performed at a median of 22 h after delivery. Hence, the analyses of FM and %BF were based on body weight at the time of the body composition study, rather than birth weight. Both %BF ($r^2 = 0.370$, p < 0.001) (Figure 3a) and FM ($r^2 = 0.623$, p < 0.0001) (Figure 3b) progressively increased with increasing total weight. Once again, there was considerable overlap between SGA, AGA and LGA groups. The correlation was significantly greater when these results were expressed as FM (Fisher's *z*-test, p < 0.0001), as compared to %BF. Similar to the relationship between PMA and adiposity, regression analyses indicate that the AGA and SGA subsets were statistically indistinguishable. Large for gestational age neonates had a slightly greater increase of %BF and FM with increasing body weight.

Gender and ethnic differences in %BF versus total fat mass

There were essentially equal proportions of male (n = 160, 49.4%) and female neonates (n = 164, 50.6%) in the study sample. Females had greater %BF ($10.8 \pm 4.2\%$) as compared to male neonates ($9.5 \pm 3.6\%$) (*t*-test = 2.98, p = 0.03) during the third trimester (Figure 4a). No significant sex differences were found for FM (351 ± 183 [females] vs. 323 ± 161 g [males]), (*t*-test = 1.47, p = 0.14) (Figure 4b). However, a significant effect of neonatal sex was found for both %BF and FM after controlling for body weight and PMA.

No significant differences in %BF or FM were initially identified between African-American versus Caucasian neonates (other ethnic groups had insufficient sample sizes). Mean FM (g) was 343 ± 182 (Caucasian, n = 214) versus 319 ± 148 (African-American, n = 83), *t*-test = 1.09, p = 0.28); mean %BF was 10.1 ± 4.2 (Caucasian) versus $10.4 \pm 3.6\%$ (African-American), *t*-test = 0.55, p = 0.58. However, significant ethnic differences in %BF and FM were identified after adjusting for neonatal weight, PMA and gender. On average, African-American neonates had $1.9 \pm 0.4\%$ greater %BF (t = 4.82, p < 0.0001) and 55 ± 13 g greater FM (t = 4.26, p < 0.0001) than did Caucasian neonates. These ethnic differences are further considered in a separate analysis for a subset of the same study cohort (Lampl M, et al. Ethnic differences in the accumulation of fat and lean mass in late gestation. Am J Hum Biol, *in press*).

Comparison of newborns from diabetic pregnancies with normal controls

In addition to the 324 neonates of non-diabetic mothers that were used to develop reference ranges, 25 additional neonates (14 females, 11 males) were studied from pregnancies complicated by diabetes. The mean birth weight of this diabetic group (3619 ± 620 g) was greater than neonates from non-diabetic pregnancies (3296 ± 560 g) (t = 2.85, p < 0.01). Both total FM (552 ± 210 vs. 337 ± 173 g, t = 5.88, p < 0.001) and %BF (15.4 ± 4.2 vs. 10.2 ± 4.0%, t = 6.21, p < 0.001) were significantly greater than the normal controls. Mean gestational ages were not significantly different: 38.6 ± 1.6 (normal) versus 38.1 ± 1.3 weeks (diabetes), t = 1.50, p = 0.13). Neonates of diabetic pregnancies had significantly greater %BF and FM after adjusting for newborn weight, gestational age, gender and ethnicity. On average, neonates of diabetic pregnancies had 2.3 ± 0.7% more %BF (t = 3.29, p < 0.001) and 84 ± 24 g greater FM (t = 3.54, p < 0.001) than infants of non-diabetic pregnancies. An important study limitation was the unavailability of precise diagnostic criteria for classifying diabetic pregnancies. A larger and more detailed cohort will be necessary to analyze which conditions of maternal glycemic control are associated with altered adiposity.

Classification of newborn adiposity using population-based percentiles versus modified neonatal growth assessment scores

Neonatal growth outcomes were initially classified using conventional criteria based on cross-sectional percentile ranges that considered birth weight, gender, and PMA [18]: SGA (n = 45, %BF = 10.3 = 11.4 ± 4.4%) (mean ± SD). There was no significant difference in %BF by ANOVA between or among these groups (F ratio = 0.37, p = 0.69). No significant differences in FM were found among these growth outcome groups: (SGA = 332 ± 198 g, AGA = 349 ± 174 g, LGA = 397 ± 221 g, mean ± SD) (ANOVA, F ratio = 1.68, p = 0.19).

A subset of 92 neonates with corresponding prenatal ultrasound measurements permitted a comparison employing the m₃NGAS₅₁including one AGA neonate from a mother with gestational diabetes. These neonates were classified using population-based stratification into small (SGA, n = 11), appropriate (AGA, n = 74) or large for gestational age (LGA, n = 7) nutritional categories. A comparison of %BF identified no significant differences by nutritional category (Figure 5a and 5b). Subsequent classification of the same neonates by m₃NGAS₅₁however, identified significant differences in %BF among the nutritional assessment groups: IUGR (n = 7, %BF = 5.5 ± 2.7%), normal (n = 81, %BF = 9.7 ± 3.7%), and macrosomia (n = 4, %BF = 14.2 ± 1.3%) (ANOVA, *F* ratio = 7.85. *p* 0.001) (Figure 6a [20]). This multivariate score led to a more robust classification of neonates into these nutritional categories, particularly when the results were expressed as total FM (Figure 6b).

Table II summarizes the poor correlation between the population-based assessment and the modified neonatal growth assessment score for classification of nutritional status. Overall agreement between the two classification methods was poor ($\kappa = -0.011$, p > 0.05). Categorical distribution of the data was also not different from chance ($\chi^2 = 1.64$, p = 0.80).

DISCUSSION

Increased fetal adiposity has been proposed as a more sensitive and specific outcome measure at birth among neonates of pregnant women with mild glucose intolerance [40]. This report provides reference ranges of whole body adiposity among neonates delivered of non-diabetic mothers from the largest series of neonatal body fat measurements to date, assessed within hours of delivery. Both %BF and FM increased linearly with PMA and neonatal weight. Although the standard deviation for %BF was constant across the range of PMA and newborn weights, the standard deviation for FM increased with increasing PMA or newborn weight. This dependence was not surprising, since FM itself is defined as the product of %BF and newborn weight. Based on our sample size of 324, the predicted error on the reference range of $5^{\text{th}} - 95^{\text{th}}$ percentiles was 8.5% of the range [33] and was 12.0% when references were calculated by gender. A sample size of approximately 1000 would be required to decrease the error to only 5% of the $5^{\text{th}} - 95^{\text{th}}$ percentile reference range.

Relatively few studies have evaluated neonatal whole body adiposity shortly after birth, especially prior to 38 weeks gestation [14, 15, 17]. Smaller cohorts of newborns have been previously described using air displacement plethysmography [36–38, 41] and dual-energy X-ray absorptiometry [17, 42–44] (Table III). Others have described patterns of regional adiposity with neonatal skin-fold measurements [16, 45–52]. A recent multicenter longitudinal study of breast-fed neonates used air displacement plethysmography to develop gender-specific %BF and fat-free mass curves in 160 neonates up to 6 months after delivery; however, only 35 neonates were studied just after birth [53].

Previous reports have noted that female neonates have a slightly greater proportion of %BF than their male counterparts on the basis of dual-energy X-ray absorptiometry [42–44], skinfold thickness [49, 52], and multi-compartment body composition models [8, 53]. Small but

significant sex differences in %BF disappeared when our results were expressed as FM (g). This reflects the greater body mass of male neonates when compared to their female counterparts. Hence, while female neonates have a higher %BF, their total body mass is less than males, so the product (%BF×body mass) remains similar for both sexes. A much larger sample size would be required to confirm these observations: At least 1500 neonates would need to be studied to demonstrate a difference of only 25 g of FM (≥80% power, p < 0.05 level of significance).

Over four decades ago, the novel introduction of using birth weight for classifying neonatal nutritional status into clinical practice was an important advance in the management of newborns at risk for adverse outcomes [54]. Our findings confirm that %BF and FM are significantly correlated with neonatal body weight [3, 8, 50, 53]. However, a comparison of adiposity according to the traditional population-based classification of nutritional status identified that a substantial number of misclassifications do occur. A pediatric investigation that used dual-energy X-ray absorptiometry study also found that birth weight categorization does not precisely reflect %BF in preterm and term newborns [43]. More frequent misclassifications would be expected to occur in fetuses since the variability of calculated EFW is much greater than actual birth weight.

Accurate assessment of newborn adiposity is important because alterations in early nutritional programming of body composition and subsequent growth patterns may increase the risk of metabolic syndrome during adult life [55, 56]. More sophisticated methods for the early detection of malnourished fetuses could more precisely define the relationships between maternal weight gain or body mass index to neonatal body composition as potential factors for future adult disease risk [57–59]. Differential growth of fetal tissues raise the possibility that the quantification of fat provide a more sensitive and specific marker of abnormal growth than lean body mass [60–62]. Although no single prenatal measurement is currently optimal, it is likely that a combination of parameters will more precisely reflect fetal adiposity and generalized nutritional status. Fractional limb volume is one example of a reproducible fetal soft tissue parameter that could be used for this purpose – as a direct measurement [21, 63, 64], combined with Doppler flow studies, or as part of a more precise fetal weight estimation procedure [65–69].

In summary, population-based weight criteria can lead to misclassifications of neonatal nutritional status on the basis of adiposity. A neonatal growth assessment score, that considers the growth potential of multiple parameters, more effectively classifies under- and over-nourished newborns. The latter approach provides a more comprehensive assessment of neonatal growth and corrects for confounding variables such as differences in age and growth potential. A potential limitation, however, is the need for the second trimester fetal biometry, from a time of assumed normal growth that is required to estimate growth potential. Nonetheless, a more comprehensive assessment of neonatal nutritional status should improve our understanding of fundamental mechanisms that affect intrauterine accumulation of energy stores and their potential impact on postnatal outcome. The clinical significance of our finding will require further examination in larger prospective studies using different population samples.

Supplementary Material

Refer to Web version on PubMed Central for supplementary material.

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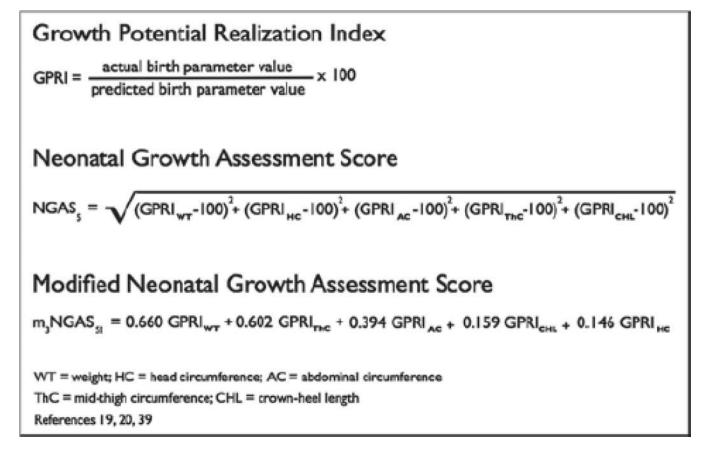


Figure 1.

Equations used for neonatal growth assessment score. WT, weight; HC, head circumference; AC, abdominal circumference; ThC, mid-thigh circumference; CHL, crown-heel length [19, 20, 39].

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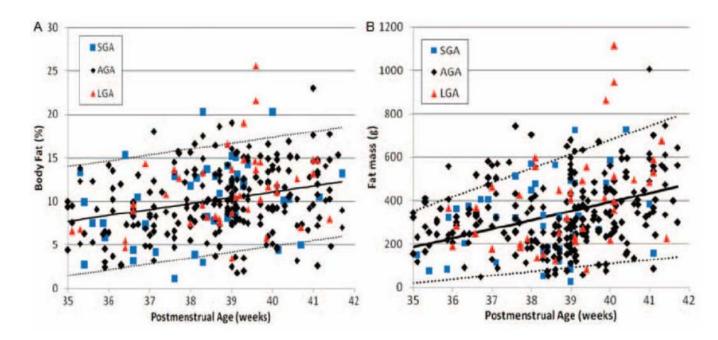


Figure 2.

Body fat versus post-menstrual age expressed as percentage (a) or absolute amount (b) of body fat. Regression equation for percentage body fat = $-15.9 + 0.675 \times \text{weeks}$, $r^2 = 0.075$; for fat mass = $-1270 + 41.6 \times \text{weeks}$, $r^2 = 0.152$, p < 0.001. Regression lines = predicted mean value with 5th and 95th percentiles. While the standard deviation in percentage body fat did not significantly change as post-menstrual age increased, the standard deviation in fat mass increased with increasing post-menstrual age. SGA, small for gestational age; AGA, appropriate for gestational age; LGA, large for gestational age.

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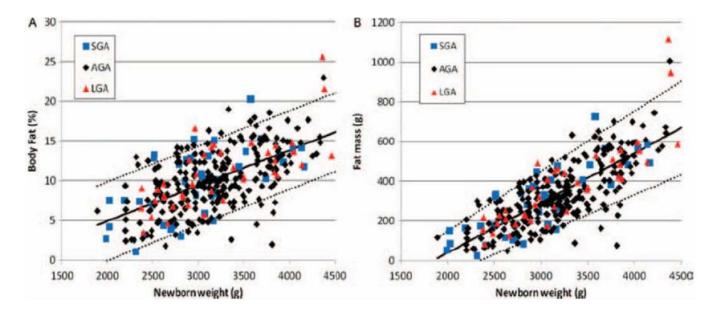


Figure 3.

Body fat versus neonatal weight at the time of air displacement plethysmography expressed as percentage (a) or absolute amount (b) of body fat. Regression equation for percentage body fat = $-4.01 + 0.00447 \times$ weight, $r^2 = 0.370$, p < 0.001; for fat mass = $-461 + 0.251 \times$ weight, $r^2 = 0.623$, p < 0.0001. Regression lines = predicted mean value with 5th and 95th percentiles. While the standard deviation in percentage body fat did not significantly change as newborn weight increased, the standard deviation in fat mass increased with larger newborn weight. SGA, small for gestational age; AGA, appropriate for gestational age.

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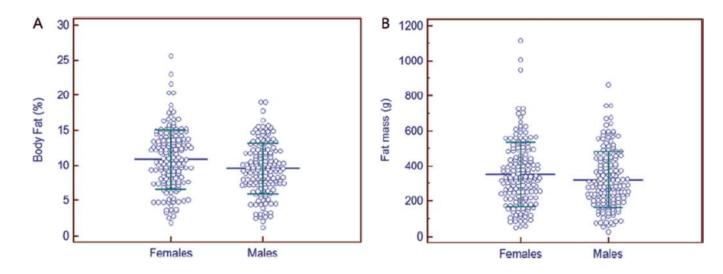


Figure 4.

Effect of gender on percentage body fat (%BF) and fat mass. Females had slightly greater mean %BF (10.8 ± 4.2%) when compared to male neonates (9.5 ± 3.6%) (*t*-tst = 2.98, *p*, 0.003) (a). These difference disappeared when mean fat mass was compared between females (351 ± 183 g) and males (323 ± 161 g) (*t*-test = 1.47, *p* = 0.14) (b). On average, males had 2.3 ± 0.34% less %BF (*t* = 6.78, *p* < 0.0001) and 78 ± 11 g less FM than females (*t* = 6.94, *p* < 0.0001). Both graphs show mean ± 1 SD for each group.

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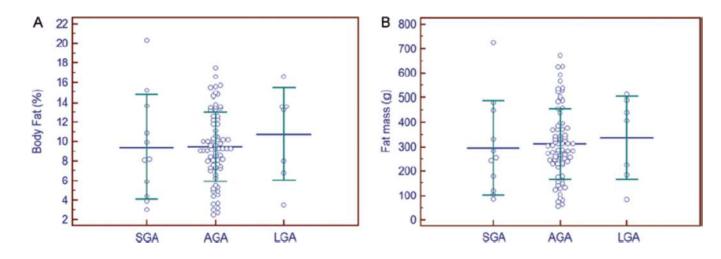


Figure 5.

Nutritional classification of percentage body fat and fat mass using population-based weight standards. No significant differences were found between groups by either percentage body fat (a), (ANOVA, *F* ratio = 0.37, *p* = 0.69), or fat mass (b), (ANOVA, *F* ratio = 0.14, *p* = 0.87); similarly, all pairwise comparisons were not significant. Mean values ± 1 SD. SGA, small for gestational age; AGA, appropriate for gestational age; LGA, large for gestational age.

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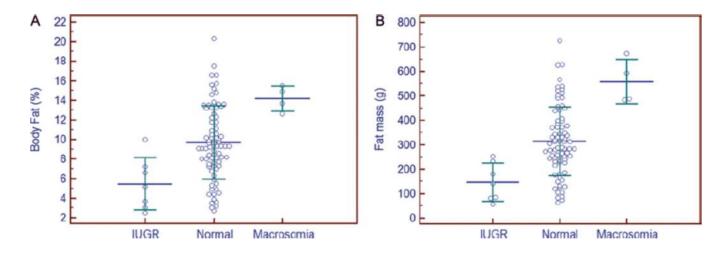


Figure 6.

Nutritional classification of percentage body fat and fat mass using a neonatal growth assessment score (m₃NGAS₅₁). The same subset of 92 neonates from Figure 5 is stratified into different nutritional groups: intrauterine growth retardation (IUGR, n = 7), normal (n = 81) or macrosomia (n = 4). Significant differences were observed for both percentage body fat (a), F ratio = 7.85, p < 0.005, and fat mass (b) (ANOVA, F ratio = 11.9), p < 0.001. Mean value ± 1 SD are shown; all pairwise comparisons were significant at p 0.05. SGA, small for gestational age; AGA, appropriate for gestational age; LGA, large for gestational age.

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| WΚ | u | Weight ± 1 SD | 5 th percentile | 10 th percentile | 25 th percentile | <i>n</i> Weight ± 1 SD S th percentile 10 th percentile 25 th percentile 50 th percentile 75 th percentile 90 th percentile 95 th percentile | 75 th percentile | 90 th percentile | 95 th percentile |
|----|----|----------------|----------------------------|-----------------------------|-----------------------------|---|-----------------------------|-----------------------------|-----------------------------|
| 35 | 25 | 2695 ± 419 | 1.5 (19) | 2.9 (56) | 5.2 (118) | 7.7 (186) | 10.3 (254) | 12.6 (316) | 14.0 (353) |
| 36 | 34 | 2793 ± 340 | 2.1 (37) | 3.5 (79) | 5.8 (150) | 8.4 (228) | 11.0 (306) | 13.3 (376) | 14.7 (418) |
| 37 | 36 | 3037 ± 410 | 2.8 (55) | 4.2 (102) | 6.5 (182) | 9.1 (269) | 11.6 (357) | 13.9 (436) | 15.3 (483) |
| 38 | 99 | 3233 ± 500 | 3.5 (73) | 4.9 (126) | 7.2 (213) | 9.8 (311) | 12.3 (408) | 14.6 (496) | 16.0 (548) |
| 39 | 91 | 3497 ± 504 | 4.2 (91) | 5.6 (149) | 7.9 (245) | 10.4 (352) | 13.0 (459) | 15.3 (556) | 16.7 (613) |
| 40 | 43 | 3536 ± 503 | 4.8 (109) | 6.2 (172) | 8.5 (277) | 11.1 (394) | 13.7 (511) | 16.0 (616) | 17.4 (679) |
| 41 | 29 | 3882 ± 447 | 5.5 (127) | 6.9 (195) | 9.2 (309) | 11.8 (436) | 14.3 (562) | 16.6 (676) | 18.0 (744) |

Fat mass (g) is expressed in parentheses.

Weight, neonate weight (g) at time of body composition study; WK, post-menstrual age in weeks.

Table II

Neonatal nutritional classification: comparison between population standard versus modified neonatal growth assessment score (m_3NGAS_{51})

| <i>n</i> = 92 | SGA | AGA | LGA |
|---------------|-----|-----|-----|
| IUGR | 1 | 6 | 0 |
| Normal | 9 | 65 | 7 |
| Macrosomia | 1 | 3 | 0 |

The correlation between both classification systems was not better than random change on the basis of Kappa agreement (-0.011, p > 0.05) or Chi-Square (1.64, p = 0.80). Population standard[18].

SGA, small for gestational age; AGA, appropriate for gestational age; LGA, large for gestational age. Modified neonatal growth assessment score IUGR, intrauterine growth retardation [20].

Table III

Newborn infant body composition-total body fat stores

| | Lee (2012) Koo (2004) | Schmeizie (2007) | verkauskiene (2007) | Fields (2009) | Erickkson (2010) | Erickkson (2010) Carberry (2010) | Roggero (2010) |
|------------------------------|-------------------------|-------------------|------------------------|---------------------------------|---------------------|----------------------------------|------------------|
| Method ADP | DXA | DXA | DXA | ADP | ADP | ADP | ADP |
| Subjects 324 | 120 | 124 | 159 | 117 | 108 | 45 | 59 |
| Study age ^a 22 h | 3.0 days | 4.1 days | 72 h | 20 days | 8 days | 0–4 days | 3 days |
| MA (range) 38.6 ±1.6 (35–41) | 5-41) 36.9 ±2.5 (30-42) | 39.7 ±1.4 (37–42) | 39.1 ±1.6 (33–42) | 39.3 ± 1.1 (≥37) | 40±1.1(<i>S</i> 7) | $40 \pm 1.1 \; (37 - 42)$ | 39.2 ±1.3 (37-42 |
| Percentage body fat | | | | | | | |
| Females 10.8 ±4.2 | 2 14.3 ±4.0 | 14.6 ±5.4 | 19.2 ±6.8 | 15.1 ±5.9 | 13.4 ± 3.7 | 10.1 ± 4.4 | 8.7 ±3.1 |
| Males 9.5 ±3.6 | 5 (Both genders) | (Both genders) | (Both genders) | 12.7 ±4.5 | 12.5 ±.0 | 9.4 ±3.4 | 8.9 ±2.3 |
| Fat mass (g) | | | | | | | |
| Females 351± 183 | .3 477 ±210 | 510 ±290 | 653 ± 304 | 588 ±288 | 484 ± 173 | 331±162 | 260 ± 120 |
| Males 323± 161 | (Both genders) | (Both genders | (Both genders) | 516 ±227 | 484 ± 203 | 341± 149 | 290 ± 90 |