

## Validity and reliability of the Block98 food-frequency questionnaire in a sample of Canadian women

Beatrice Boucher<sup>1,\*</sup>, Michelle Cotterchio<sup>1,2</sup>, Nancy Kreiger<sup>1,2,3</sup>, Victoria Nadalin<sup>1</sup>, Torin Block<sup>4</sup> and Gladys Block<sup>5</sup>

<sup>1</sup>Division of Preventive Oncology, Cancer Care Ontario, 620 University Avenue, Toronto, Ontario, Canada, M5G 2L7; <sup>2</sup>Department of Public Health Sciences, University of Toronto, Toronto, Ontario, Canada; <sup>3</sup>Department of Nutritional Sciences, University of Toronto, Toronto, Ontario, Canada; <sup>4</sup>Block Dietary Data Systems, Berkeley, CA, USA; <sup>5</sup>Department of Public Health, University of California, Berkeley, CA, USA

Submitted 14 December 2004; Accepted 18 May 2005

### Abstract

**Objective:** To assess the validity and reliability of the most recent adaptation of Block's full-diet food-frequency questionnaire (FFQ) among a sample of Canadian women.

**Design:** Participants completed a self-administered FFQ (FFQ1), two unannounced 24-hour recalls (weekday and weekend) and a second FFQ (FFQ2) between October 2003 and February 2004. FFQs and recalls were analysed for 32 nutrients using Block Dietary Data Systems and the University of Minnesota's Nutrient Data System. Mean and median intakes were computed, along with crude and deattenuated Pearson correlation coefficients between FFQ1 and the average of two recalls (validity) and between FFQ1 and FFQ2 (reliability).

**Setting:** Ontario, Canada.

**Subjects:** A random population-based sample ( $n = 166$ ) of women aged 25 to 74 years.

**Results:** One hundred and fifteen (69%) women completed FFQ1, 96 completed FFQ1 and both recalls, and 93 completed both FFQs, about 56 days apart. Mean intakes were similar for most nutrients. FFQ reliability was high, with Pearson correlation coefficients having a median of 0.75, ranging from 0.57 to 0.90 (macronutrients) and from 0.65 to 0.88 (micronutrients from supplements and food). FFQ validity was moderate to high, with deattenuated Pearson correlation coefficients having a median of 0.59, ranging from 0.11 to 0.73 (macronutrients) and from 0.50 to 0.76 (micronutrients from supplements and food). Our micronutrient correlations were similar to or higher than those of other studies that included supplements. Two correlations  $< 0.40$  were associated with fats.

**Conclusions:** The validity and reliability of this full-diet version of the Block FFQ were moderate to high, supporting its use in future studies among Canadian women.

**Keywords**  
Food-frequency questionnaire  
Dietary recall  
Validity  
Reliability  
Women  
Canada

Food-frequency questionnaires (FFQs) provide the most practical and economical method for collecting data on 'usual' dietary intake in population-based epidemiological studies. In aetiological research, it is preferable to use FFQs that query the whole diet rather than a limited number of foods for specific hypotheses, since comprehensiveness will improve the ability to adjust for confounding nutrients and energy intake, when appropriate, and to explore future dietary hypotheses<sup>1-3</sup>.

Many comprehensive FFQs have been developed for epidemiological research and, among these, the Block FFQ is one of the most widely used<sup>2,4-6</sup>. Originally developed at the US National Cancer Institute in 1984<sup>1,5,7</sup>, it has undergone a number of revisions to reflect changes in consumption and improvements in FFQ design<sup>4,5,7-9</sup>. Although a new FFQ is in development (Torin Block,

personal communication), the most recent full-diet version at the time of this study was developed in 1998 to incorporate dietary and questionnaire changes suggested by American national consumption data (from the Third National Health and Nutrition Examination Survey (NHANES III)) and cognitive research<sup>9,10</sup>.

Different versions of the Block FFQ have been validated by Block or others to allow comparison with other FFQs<sup>4,11-13</sup> or to assess FFQ performance for specific research hypotheses or groups<sup>1,14-20</sup>. This approach follows guidelines to validate an FFQ after its development, or whenever it has been substantially changed or used with a new population<sup>2-4,21</sup>. Validation involves assessing how well an FFQ measures what it was designed to measure and requires comparison with a 'more accurate' reference measure of intake, usually multiple

\*Corresponding author: Email [beatrice.boucher@cancercare.on.ca](mailto:beatrice.boucher@cancercare.on.ca)

dietary recalls, records or biomarkers. Validation correlations vary with the nutrient, but typically range from 0.40 to 0.70<sup>3,4,22,23</sup>.

The current study is the first to report on the validation of the most recent full-diet Block FFQ in common use, Block98, and compares intakes from the FFQ with short-term dietary recall data in a North American population. Additionally, this is the first time the Block FFQ has been validated in Canada, despite previous use in Canadian research<sup>24,25</sup>. We assessed the agreement of 32 nutrient intakes estimated from the FFQ with corresponding estimates from a second FFQ and the average of two 24-hour dietary recalls.

## Methods

### *Sample recruitment and study design*

During the autumn of 2003, random digit dialling methods were used to recruit a random population-based sample of Ontario women aged 25 to 74 years, stratified to represent the age distribution of women in Ontario. After excluding non-working or non-residential numbers and ineligible women (those who did not know English or were outside the age range), 235 women were identified, of whom 166 (71%) agreed to participate. Data collection began in October 2003 and continued to February 2004.

Subjects were mailed the first FFQ (FFQ1), food portion guides, a short background questionnaire regarding age, marital status, ethnicity, education, height, weight and smoking history, and an availability form to assist study staff in contacting respondents for the interviews. Approximately 14 days after FFQ1 was completed, the first of two unannounced telephone-administered recalls was done, with a second one an average of 18 days later. About 26 days after the second recall, a second FFQ (FFQ2) was completed. The first and second FFQs were completed an average of 56 days apart. The study was approved by the Health Sciences Research Ethics Board at the University of Toronto.

### *Dietary assessment*

#### *FFQ*

The 1998 version of the full-diet Block FFQ (Block98) is an eight-page, scannable, quantitative instrument that includes 109 food and beverage items (including alcohol); three multiple and nine single vitamin and mineral supplement items; additional questions to assess fats, fat-modified foods, types of milk and cereal; and summary cereal, fruit and vegetable questions. The FFQ food list was checked against other full-diet FFQs validated in Canada<sup>26,27</sup> for any important differences. Food items considered unusual in Canada (e.g. grits, vitamin brand names) were deleted or altered, and a few items (e.g. winter squash, vitamin D supplement) were added along with a list of soy foods (to address

phyto-oestrogens in future research), to yield 126 food and beverage items in the final analysis. The FFQ asked how often each food or beverage was usually consumed, and offered nine continuous responses ranging from 'never' to 'every day' for most foods. Four portion size choices were given using standard units (e.g. tablespoons or slices), or by referring to a separate sheet of photographs representing three-dimensional bowl and plate portions ranging from 1/4 to 2 cups. FFQs were self-administered by participants, and have been described as taking 30–40 min to complete (Torin Block, personal communication). Questionnaires were checked for completeness upon return, and telephone calls were made to collect missing information. FFQs were scanned and analysed by Block Dietary Data Systems using a nutrient content database based on the US Department of Agriculture (USDA) Nutrient Database for Standard Reference, national food consumption data (NHANES III and the Continuing Survey of Food Intakes by Individuals) and values from the published literature<sup>1,9</sup>. Summary questions on the FFQ for cereals, fruit and vegetables were used to adjust nutrient intakes associated with these foods<sup>15</sup>.

#### *Dietary recalls*

Unannounced recalls, asking respondents to describe all foods, drinks and portion sizes consumed during the previous 24 h, were done over the telephone by one dietitian trained to use a Windows-based interview and nutrient analysis system, the Nutrient Data System for Research (NDS-R)<sup>28</sup>, version 4.06 (2003, Nutrition Coordinating Center, University of Minnesota, MN, USA, Food and Nutrient Database 34). Training included attending a 2-day NDS-R workshop, and conducting 23 practice interviews evaluated by the study nutritionist for consistency and completeness.

A 16-page portion size booklet showing two-dimensional, actual-size food and drink models<sup>29</sup> was mailed to respondents and used to guide the interviews<sup>30</sup>. Standardised data collection was facilitated through NDS-R's scripted, multiple-pass approach<sup>31</sup>, to which questions about alcohol and vitamin and mineral supplements were added to improve intake measurement<sup>7,32,33</sup> and comparability with the FFQ, where these items were also queried. When a respondent identified a supplement or food that was not included in NDS-R, a generic or best-fit alternative was used after matching ingredients identified by the respondent or, when unknown, by searching the Health Canada Drug Product Database<sup>34</sup> or manufacturer websites. This best-fit approach was taken to improve the accuracy of intake reports because it was not known how thoroughly NDS-R included brands available in Canada. All interviews were checked for completeness and use of standardised methods by the study dietitian,

and a random sample of 10% were duplicate-checked by the study nutritionist.

The two unannounced recalls included a weekday and a weekend day for each respondent. Weekend days included Friday to Sunday to capture food and alcohol consumption patterns different from those on weekdays (Monday to Thursday)<sup>3,35–37</sup>. It was not possible to address seasonality within the time frame of this study.

### Statistical analysis

All analyses were done using SAS, version 8.02 (SAS Institute Inc., Cary, NC, USA, 1999–2001). Descriptive statistics were computed to describe response rates, demographic characteristics and average daily nutrient intakes. Statistical analysis was restricted to 32 nutrients common to the FFQ and recall databases, a list derived in consultation with Block Dietary Data Systems and the Nutrition Coordinating Center for NDS-R. Nutrient intakes were reported in the same units by both databases, except for alcohol, which was reported as grams of alcohol by NDS-R but grams of alcoholic beverages by Block Dietary Data Systems. Since we were unable to separate food from supplements in the recalls, all micronutrient intakes include food and supplements combined. Energy adjustment (density method) was done for major macronutrients, but not for micronutrients since food values (associated with energy) could not be separated from supplement values (not associated with energy).

Crude data from the FFQs and recalls were log-transformed to improve normality, and outliers (intakes greater than three standard deviations ( $>3SD$ ) from the mean in FFQs or recalls) were removed before calculating intake means, medians and Pearson correlation coefficients. On average, less than one outlier was excluded per nutrient, ranging from zero for vitamin C (all methods) to three for thiamine (recalls). One-way analysis of variance was used to test for significant differences between mean intakes. Reliability was assessed using Pearson correlation coefficients between FFQ1 and FFQ2, while validity was assessed using Pearson correlation coefficients between FFQ1 and the average of the two 24-hour recalls. All validity coefficients were corrected for attenuation due to random error in within-person variability, to allow a reasonable estimate of true correlation, given that two days of recall were collected<sup>3,38</sup>.

### Results

Of the 166 women who agreed to participate in this study, 115 (69%) completed FFQ1; of these, 105 (91%) completed the first dietary recall, 96 (83%) completed the second dietary recall and 93 (81%) completed FFQ2. Among those women who completed both dietary recalls and one or two FFQs ( $n = 96$ ) (Table 1), the majority were married (74%), Caucasian (82%), non-smokers (80%), with college or university education (61%).

Mean, SD and median daily nutrient intakes for the two FFQs and the average of two diet recalls are shown in Table 2. Comparing FFQ1 and FFQ2, mean intakes were similar across all nutrients except for energy and *trans* fatty acids, where FFQ1 gave significantly higher estimates. Mean intakes from FFQ1 and the recalls were similar for total energy (1709 and 1844 kcal) although the FFQ gave significantly lower estimates for carbohydrate, protein, saturated fat and cholesterol, and higher estimates for total fat (percentage of energy), *trans* fatty acids and polyunsaturated fat. Mean intakes of 13 of the 17 micronutrients (food and supplements combined) were not different between FFQ1 and the recalls, although the FFQ gave significantly higher estimates for vitamin A,  $\beta$ -carotene and vitamin C, and a lower estimate for selenium.

Pearson correlation coefficients between the two FFQs (reliability) and between FFQ1 and the recalls (validity) are presented in Table 3. Most correlations fell within the range of 0.40 to 0.70, and most for macronutrients increased when described in terms of energy density. Correlation coefficients to estimate reliability between FFQ1 and FFQ2 were found to be relatively high, with a median of 0.75, and ranged from 0.57 (percentage energy from protein) to 0.90 (percentage energy from alcohol) for macronutrients, and from 0.65 (iron) to 0.88 (vitamin C) for micronutrients.

All correlations to estimate validity increased with deattenuation by 0.03 to 0.29. Deattenuated correlation

**Table 1** Characteristics of women who completed two dietary recalls and at least FFQ1 in the Block98 FFQ validation study, 2003–2004 ( $n = 96$ )

Characteristic	Frequency, $n$ (%)
Age	
25–39	25 (26.0)
40–49	23 (24.0)
50–59	22 (22.9)
60+	26 (27.1)
Marital status	
Never married	11 (11.5)
Married/common law	71 (74.0)
Separated/divorced	10 (10.4)
Widowed	4 (4.2)
Ethnicity	
Caucasian	79 (82.3)
Black	4 (4.2)
South Asian	2 (2.1)
Southeast Asian	2 (2.1)
Aboriginal	1 (1.0)
Other	8 (8.3)
Highest education	
Grade 1–8	5 (5.2)
Grade 9–13	23 (24.0)
Vocational school	9 (9.4)
College or university	49 (51.0)
Graduate degree	10 (10.4)
Smoking status	
Current	19 (19.8)
Former	30 (31.2)
Never	47 (49.0)

FFQ – food-frequency questionnaire.

**Table 2** Mean, SD and median daily nutrient intakes estimated by FFQ1, FFQ2 and the average of two 24-hour recalls among Canadian women, 2003–2004\*

Nutrient	FFQ1 (n = 115)		FFQ2 (n = 93)		Average of two 24-hour recalls (n = 96)	
	Mean (SD)	Median	Mean (SD)	Median	Mean (SD)	Median
Energy (kcal)	1708.7 (585.4)	1655.4	1539.2 (549.6)¶	1444.6	1843.6 (611.2)	1722.4
Alcohol (g)†	80.1 (123.3)	22.3	86.4 (136.8)	23.6	6.2 (11.3)	0.1
Alcohol (% of energy)†	3.1 (5.2)	0.9	3.8 (6.5)	1.1	2.3 (4.1)	0.0
Carbohydrate (g)	209.1 (88.0)§	195.7	186.5 (79.5)	178.4	234.2 (86.3)	224.7
Carbohydrate (% of energy)	49.5 (8.0)	50.0	48.7 (8.8)	48.0	51.1 (8.8)	50.4
Protein (g)	67.0 (25.0)§	62.8	61.9 (25.7)	59.7	73.6 (26.9)	69.2
Protein (% of energy)	15.8 (3.1)	15.8	15.9 (3.2)	15.6	16.1 (3.5)	16.3
Total fat (g)	66.8 (27.7)	62.4	60.7 (25.1)	57.5	67.3 (28.7)	63.4
Total fat (% of energy)	35.6 (7.6)§	35.4	35.8 (6.9)	36.0	32.3 (7.2)	33.1
Saturated fat (g)	20.0 (8.5)§	18.4	18.4 (7.6)	17.8	23.5 (11.6)	21.3
Polyunsaturated fat (g)	16.3 (8.5)§	14.4	14.6 (7.6)	12.2	13.1 (7.7)	12.0
Monounsaturated fat (g)	25.1 (10.9)	22.5	22.9 (10.0)	21.3	25.1 (11.1)	23.3
Trans fatty acids (g)	4.9 (2.5)§	4.7	4.2 (2.4)¶	3.7	4.0 (2.5)	3.4
Cholesterol (mg)	189.0 (75.8)§	171.8	171.4 (76.5)	159.0	254.2 (145.7)	220.2
Dietary fibre (g)	18.2 (9.7)	16.1	17.5 (9.8)	15.2	19.3 (9.0)	17.2
Vitamin A (IU)‡	13903.8 (9825.6)§	11 111.8	13 263.4 (9439.6)	10 744.4	11 062.6 (8363.7)	8394.5
β-Carotene (μg)‡	6377.8 (7196.6)§	4056.7	5174.6 (5021.8)	3537.1	4248.1 (3959.9)	2756.0
Thiamine (B <sub>1</sub> ) (mg)‡	3.1 (3.4)	2.0	2.6 (3.0)	1.6	2.6 (2.5)	1.9
Riboflavin (B <sub>2</sub> ) (mg)‡	3.4 (3.4)	2.4	3.0 (3.0)	2.0	2.8 (2.6)	2.0
Niacin (mg)‡	36.8 (34.0)	26.1	31.6 (30.5)	22.3	30.0 (18.9)	24.0
Vitamin B <sub>6</sub> (mg)‡	3.0 (2.0)	2.5	2.7 (1.9)	2.1	3.0 (2.0)	2.3
Folate (μg)‡	560.9 (317.4)	508.4	519.8 (296.1)	457.1	582.2 (319.9)	545.8
Vitamin B <sub>12</sub> (μg)‡	6.8 (5.1)	5.1	6.2 (4.8)	4.5	9.7 (10.6)	5.9
Vitamin C (mg)‡	358.1 (396.2)§	209.3	352.9 (482.9)	169.0	240.2 (344.4)	101.7
Vitamin D (μg)‡	8.5 (5.9)	7.5	8.3 (6.2)	7.3	8.8 (7.1)	7.1
Vitamin E (mg)‡	62.7 (96.5)	20.3	59.1 (96.3)	13.0	61.5 (123.0)	14.2
Calcium (mg)‡	1159.6 (635.6)	952.6	1131.7 (711.6)	985.8	1026.1 (563.9)	898.2
Copper (mg)‡	1.9 (1.1)	1.5	1.8 (1.1)	1.4	1.8 (1.1)	1.5
Iron (mg)‡	22.5 (18.6)	17.2	21.0 (18.1)	16.4	17.7 (8.5)	15.3
Magnesium (mg)‡	326.3 (132.7)	313.6	312.9 (142.6)	292.7	342.5 (137.4)	320.9
Selenium (μg)‡	97.3 (55.6)§	85.4	87.2 (45.9)	79.6	116.6 (52.7)	104.9
Zinc (mg)‡	17.5 (14.7)	12.7	16.2 (12.3)	12.3	15.3 (9.3)	11.2

SD – standard deviation; FFQ – food-frequency questionnaire.

\* All data are based on crude intakes with log-transformed outliers removed.

† Based on grams of alcohol in recalls, and grams of alcoholic beverages in FFQ.

‡ Total intake based on supplements and food.

§ Means from FFQ1 and recalls are significantly different ( $P < 0.05$ ).¶ Means from FFQ1 and FFQ2 are significantly different ( $P < 0.05$ ).

coefficients between FFQ1 and recalls were generally moderate to high with an overall median of 0.59. For macronutrients, the median was 0.46, ranging from 0.11 (cholesterol) to 0.73 (percentage energy from carbohydrate); for micronutrients, the median was 0.65, ranging from 0.50 (iron) to 0.76 (folate). Only two of the 32 deattenuated correlation coefficients were less than 0.40, and both were associated with fats – cholesterol and monounsaturated fat.

Although reporting alcohol in different units hindered our ability to compare mean intakes from the FFQ and recalls, moderate to high reliability (0.86) and validity, especially when energy-adjusted (0.69), were observed for alcohol intake.

## Discussion

Our estimates of reliability often exceeded the usual range for macronutrients (0.50 to 0.70)<sup>2,3</sup> and were similar to

those reported for micronutrients from food and supplements (0.53 to 0.83)<sup>39,40</sup>. Our estimates of validity were moderate for macronutrients compared with others (0.40 to 0.70)<sup>3,4,22,23</sup>, and moderate to high for micronutrients compared with those including food and supplements (0.21 to 0.83)<sup>15,39–43</sup>, and support the use of our FFQ in studies with Canadian women. Validity correlations less than 0.40 suggest FFQ shortcomings, based on our assumption that two recalls represent true intake.

### Comparison of average intakes with other studies

Although ranking of individuals is the primary objective of most epidemiological studies<sup>3,22</sup>, nutrient intakes at the group level help assess how comparably each method describes the group mean<sup>44</sup> and the comprehensiveness of the FFQ with respect to more detailed recalls<sup>3,45</sup>.

Our average macronutrient and energy intakes from recalls were comparable to those reported among

**Table 3** Pearson correlation coefficients between nutrients estimated by FFQ1 and FFQ2 (reliability) and by FFQ1 and the average of two 24-hour recalls (validity) among Canadian women, 2003–2004\*

Nutrient	FFQ1 and FFQ2 ( <i>n</i> = 93)		FFQ1 and average of two 24-hour recalls ( <i>n</i> = 96)	
	Non-deattenuated Pearson correlation coefficient (95% CI)		Deattenuated Pearson correlation coefficient (95% CI)	
Energy (kcal)	0.74 (0.64, 0.82)	0.34 (0.15, 0.51)	0.44 (0.15, 0.73)	
Alcohol (g)†	0.86 (0.80, 0.91)	0.29 (0.10, 0.47)	0.52 (0.07, 0.97)	
Alcohol (% of energy)†	0.90 (0.85, 0.93)	0.40 (0.21, 0.55)	0.69 (0.14, 1.00)	
Carbohydrate (g)	0.80 (0.72, 0.87)	0.41 (0.22, 0.56)	0.51 (0.23, 0.79)	
Carbohydrate (% of energy)	0.78 (0.68, 0.85)	0.49 (0.32, 0.63)	0.73 (0.24, 1.00)	
Protein (g)	0.72 (0.61, 0.81)	0.30 (0.10, 0.47)	0.41 (0.11, 0.71)	
Protein (% of energy)	0.57 (0.41, 0.69)	0.29 (0.09, 0.46)	0.46 (0.24, 1.00)	
Fat (g)	0.72 (0.60, 0.81)	0.29 (0.10, 0.47)	0.41 (0.10, 0.72)	
Fat (% of energy)	0.70 (0.57, 0.79)	0.43 (0.25, 0.58)	0.61 (0.24, 0.98)	
Saturated fat (g)	0.72 (0.61, 0.81)	0.28 (0.09, 0.46)	0.41 (0.09, 0.73)	
Polyunsaturated fat (g)	0.72 (0.60, 0.81)	0.32 (0.13, 0.49)	0.42 (0.13, 0.71)	
Monounsaturated fat (g)	0.69 (0.57, 0.79)	0.24 (0.04, 0.42)	0.35 (0.03, 0.67)	
<i>Trans</i> fatty acids (g)	0.74 (0.63, 0.82)	0.37 (0.18, 0.53)	0.53 (0.18, 0.88)	
Cholesterol (mg)	0.69 (0.56, 0.78)	0.07 (−0.14, 0.27)	0.11 (−0.21, 0.43)	
Dietary fibre (g)	0.84 (0.76, 0.89)	0.52 (0.36, 0.65)	0.62 (0.33, 0.91)	
Vitamin A (IU)‡	0.82 (0.74, 0.88)	0.40 (0.21, 0.55)	0.55 (0.22, 0.88)	
β-Carotene (μg)‡	0.76 (0.65, 0.83)	0.27 (0.07, 0.44)	0.49 (0.05, 0.93)	
Thiamine (B <sub>1</sub> ) (mg)‡	0.75 (0.65, 0.83)	0.68 (0.55, 0.77)	0.75 (0.47, 1.00)	
Riboflavin (B <sub>2</sub> ) (mg)‡	0.75 (0.64, 0.83)	0.66 (0.53, 0.76)	0.74 (0.45, 1.00)	
Niacin (mg)‡	0.75 (0.64, 0.83)	0.64 (0.50, 0.74)	0.74 (0.43, 1.00)	
Vitamin B <sub>6</sub> (mg)‡	0.79 (0.70, 0.86)	0.63 (0.49, 0.74)	0.70 (0.43, 0.97)	
Folate (μg)‡	0.78 (0.68, 0.85)	0.63 (0.50, 0.74)	0.76 (0.41, 1.00)	
Vitamin B <sub>12</sub> (μg)‡	0.72 (0.60, 0.81)	0.55 (0.39, 0.68)	0.65 (0.36, 0.94)	
Vitamin C (mg)‡	0.88 (0.82, 0.92)	0.61 (0.46, 0.72)	0.69 (0.41, 0.97)	
Vitamin D (μg)‡	0.76 (0.66, 0.83)	0.48 (0.31, 0.63)	0.54 (0.29, 0.79)	
Vitamin E (mg)‡	0.81 (0.72, 0.87)	0.68 (0.56, 0.78)	0.71 (0.48, 0.94)	
Calcium (mg)‡	0.80 (0.71, 0.86)	0.56 (0.40, 0.68)	0.71 (0.35, 1.00)	
Copper (mg)‡	0.83 (0.76, 0.89)	0.53 (0.36, 0.66)	0.60 (0.34, 0.86)	
Iron (mg)‡	0.65 (0.51, 0.75)	0.42 (0.23, 0.57)	0.50 (0.23, 0.77)	
Magnesium (mg)‡	0.82 (0.74, 0.88)	0.53 (0.37, 0.66)	0.63 (0.34, 0.92)	
Selenium (μg)‡	0.69 (0.57, 0.79)	0.41 (0.23, 0.56)	0.56 (0.23, 0.89)	
Zinc (mg)‡	0.75 (0.64, 0.83)	0.51 (0.34, 0.64)	0.62 (0.33, 0.91)	

CI – confidence interval; FFQ – food-frequency questionnaire.

\* All data are based on crude intakes with log-transformed outliers removed.

† Based on grams of alcohol in recalls, and grams of alcoholic beverages in FFQ.

‡ Total intake based on supplements and food.

women in recent validation studies in Canada<sup>26</sup> and the USA<sup>4</sup>, where energy intake hovered around 1700–1800 kcal. Average intakes from the FFQ were also similar for most macronutrients and energy compared with those estimated using the Block95 FFQ<sup>4</sup>, but were sometimes lower than those using an FFQ developed for the Canadian Study of Lifestyle and Health<sup>26</sup>, possibly due to different respondent and study factors. In contrast, our average micronutrient intakes were usually much higher in both the FFQ and recalls since neither of the other studies<sup>4,26</sup> included supplements, although these may substantially increase intake over that from food alone<sup>7,46,47</sup>. As an example, our average vitamin E intakes were at least twice those reported by others<sup>4,26</sup>. We expect high supplement use in our sample since all subjects were female, and the majority were Caucasian, university-educated and non-smokers – all factors positively associated with supplement use<sup>46,48–51</sup>. A review of our data showed the majority were supplement users: 25% reported no supplements on both the FFQ and recalls. We were unable to separate food from supplements

in the recall data to review their contributions separately; however, when we compared our results with those reporting combined food and supplement intakes among Caucasian women, our average intakes were usually similar<sup>15,40,47</sup>.

### **Comparison of average intakes from FFQ1 and recalls: a focus on micronutrients**

Mean intakes from FFQ1 and recalls were similar for most nutrients, although the FFQ gave different estimates for some macronutrients and higher estimates for micronutrients. Our lower FFQ estimates for protein, carbohydrate, cholesterol and saturated fats have been reported before, as have higher estimates of polyunsaturated fat and percentage energy from fat<sup>4,6,11,17–19,43</sup>, and may result from limitations of the FFQ food list, response options or respondents to report diet accurately<sup>21,22,43</sup>. Recent biomarker studies suggest that FFQs underestimate energy and protein more severely than recalls<sup>52–54</sup>, and although this may not have resulted in a significant

difference in total energy intake, it may help explain our lower FFQ macronutrient estimates.

FFQs that include a large number of items to represent specific food groups encourage higher estimates of intake<sup>6,8,11,21,55,56</sup>. Since vitamin A,  $\beta$ -carotene and vitamin C are associated with fruit and vegetable consumption<sup>57,58</sup>, higher estimates for these micronutrients may be partially due to the large number of fruit and vegetable items on our FFQ: the Block98 FFQ included 28 fruit or vegetable foods among its original 109 items.

Higher FFQ estimates of vitamins C and A and  $\beta$ -carotene have been reported previously with the Block FFQ compared with recalls or records, both among women<sup>1,4,18,20</sup> and men<sup>1,14</sup> and whether using earlier versions of the full-diet<sup>1,4,14,18</sup> or reduced FFQ<sup>20</sup>. This is not, however, unique to the Block FFQ, and has been seen with other FFQs<sup>21,26,39,43,59</sup>.

Fruit and vegetable adjustment questions were introduced to help correct this effect<sup>15</sup>, but we and others<sup>20</sup> observed higher estimates despite adjustment, suggesting either incomplete correction or the role of other factors, such as the inclusion of supplements or limitations of the reference method. Higher Block FFQ estimates had been observed whether intakes were based on food only<sup>1,4,14,18</sup>, food with adjustment<sup>20</sup>, food and supplements<sup>18</sup>, or food and supplements with adjustment (our study).

Although the inclusion of supplements improves intake measurement and classification<sup>7,42</sup>, their contribution may be so great relative to food<sup>7,32,33,46</sup> that any correction based on food alone may be masked, particularly when other sources of dietary assessment error are not adequately reduced. It is notable that another study, which included food and supplements with adjustment, did not show higher FFQ estimates compared with records<sup>15</sup>. This may be partly explained by the inclusion of more days in the reference method and hence a better assessment of usual diet and supplement use than in our study, affecting both individual and group estimates<sup>60</sup>. Block *et al.* collected three 4-day records per participant<sup>15</sup> compared with our two 24-hour recalls. Although one day may adequately capture average group energy intake in a sample of 100 women, many more days would be needed for the more variable nutrients from food, such as vitamin A<sup>60</sup>. More than four days, and perhaps as many as 14, may be needed to adequately identify supplement use<sup>42</sup>.

### **Correlation of intakes from FFQ1 and recalls (validity)**

How well the FFQ captures usual diet affects average intakes and correlations between the FFQ and recalls. Under our validity assumptions, FFQ limitations mentioned earlier account for low correlation coefficients. However, since no dietary assessment

method is error-free<sup>21,22,61</sup>, our reference method based on two recalls must also be examined.

For nutrients with small within-person to between-person ratios, only a few days of recall are needed to capture usual intake, whereas those with large ratios require many more days, especially when evaluating individual rather than group intake<sup>22,37,62–65</sup>. Among macronutrients, fat variables have the highest ratios, especially cholesterol; of the micronutrients, vitamin A and carotene have the highest ratios<sup>37,60,64,65</sup>. It is not surprising that correlation coefficients lower than 0.40 were associated with two of the most variable macronutrients – cholesterol (0.11) and monounsaturated fat (0.35). Additionally, although two days of recall may be adequate to estimate validity, correlations between FFQ and recalls have been shown to increase with the number of recall days<sup>15,22,40</sup>.

Comparing our deattenuated correlation coefficients with those found in other studies among women<sup>4,26</sup>, correlations for energy were comparable. Those for many macronutrients, however, were higher than ours (most notably for cholesterol: 0.11 vs. 0.57 and 0.55), although less so when using an earlier version of our FFQ<sup>4</sup>. This trend was reversed, however, for micronutrients, where we observed higher correlation coefficients, sometimes dramatically so for those commonly consumed as supplements<sup>33,46,47,49,51</sup>, such as vitamin E (0.71 vs. 0.37 and 0.28). Since the inclusion of supplements improves micronutrient correlations, we also compared our results with the few studies among Caucasian women that included food and supplements combined. Many of our correlations continued to be higher, particularly for B vitamins, as did our average micronutrient correlation (0.65 vs. 0.58)<sup>15,40,41</sup>.

### **Discrepancy between validity correlations for macronutrients and micronutrients**

The discrepancy between lower correlations associated with macronutrients but higher correlations associated with micronutrients may be explained by shortcomings and strengths of our FFQ, as well as our study. Many of the confidence intervals around correlations were wide, especially for macronutrients, due to our small sample size, and this somewhat limited the interpretation of our findings. Additionally, usual diet and the full extent of dietary variation may not have been captured by two days of recall. Correlations may have improved with additional days of recall<sup>15,22,40</sup>, as in Jain *et al.* (three days)<sup>26</sup> and Subar *et al.* (four days)<sup>4</sup>, and over different seasons<sup>4</sup>. This shortcoming in our recall data would affect macronutrients, especially the most variable ones such as fats. A supplementary analysis highlights this limitation and its effects (data not shown). When grams of food consumed daily were examined, 18% of our recall data fell within the top 2.5% of the consumption

distribution in a large nationally representative sample of Americans (NHANES 1999–2000)<sup>66</sup>, suggesting unusually high intake reports over our two days of recall. Some of these high reports may be due to misunderstandings of the two-dimensional portion size measures used for recalls<sup>67</sup>. For one respondent, estimates using a picture of a glass instead of a tablespoon added up to 40 tablespoons of cream in coffee; for another, the portion size picture chosen for meat in one dish added to 0.57 kg of ground beef. These high reports would have been minimally affected by removing less than 1% of our data as outliers. When these high reports (18%) were removed from the analysis, deattenuated correlations for most fat variables increased from less than or close to 0.40 to 0.50 or greater.

In contrast, validity correlations associated with all micronutrients, including the most variable, were higher in our study than in those based on more days of recall but which excluded supplement data<sup>4,26</sup>. Our correlations were also higher than those reported in studies including supplement data<sup>15,40–43</sup>. Four possible explanations arise within the context of our sample of women and their supplement use. First, given the important contribution of supplements relative to food, the addition of supplement data not only improves micronutrient intake measurements at the group and individual levels but also correlations between different methods<sup>42</sup>, even when food intake data may be limited. Second, although supplement use is common among women in our study, there may be considerable intra-individual heterogeneity in use that improved correlations<sup>59</sup>. Third, by adding questions to the recall interviews about supplement intake, we encouraged reports similar to those prompted by the FFQ list of supplements. (Similarly, the addition of alcohol recall questions may have contributed to the moderate to high alcohol correlations.) Fourth, the Block98 FFQ included modifications<sup>9</sup> which may have encouraged higher micronutrient correlations than those seen with the Block95 FFQ<sup>4</sup>. Modifications included the incorporation of additional vitamin supplements, seasonal and canned fruit, an additional frequency option (a few times a year), different portion options, and a portion guide whose three-dimensional representations may have offered cognitive advantages to estimating intake<sup>30,67</sup>.

### **Other effects on validity**

Other factors may have affected validity although the direction of their effects is uncertain. Although both NDS-R and the Block Dietary Data Systems databases are based on the USDA Nutrient Database for Standard Reference<sup>1,9,28</sup>, differences in inclusions and accuracy<sup>21,22,68,69</sup> may affect recall and FFQ analyses, and comparisons between them. Since neither database was modified for Canadian foods, the current validation study was

facilitated by basing analyses on US data only. However, future Canadian studies using the Block98 FFQ and its affiliated database may need to be interpreted cautiously, particularly when examining foods or nutrients affected by fortification regulations that differ between Canada and the USA, such as for folate, vitamin D, calcium, zinc, iron or vitamin A<sup>70</sup>. Similarly, it is uncertain how the different FFQ and recall food portion models affected the accuracy and comparability of intake reports<sup>67</sup>, and how this might have affected validity.

While our study sample was similar to the general population in terms of including visible minorities<sup>71</sup> and smokers<sup>72</sup>, it differed substantially on education. Although our recruitment of university-educated respondents (61%) likely assisted FFQ and recall completion and validity<sup>22</sup>, this selection bias may hinder the generalisability of our findings to the broader population of Canadians, where only 33% are so highly educated<sup>73</sup>. Participant characteristics may have also affected our energy estimates, since factors such as body mass index (BMI) are associated with energy underreporting<sup>74</sup>. In a group of Canadian women where 35% were low energy reporters (LERs) by FFQ, mean energy intake among non-LER women was higher than ours<sup>75</sup>. Since our sample had a higher mean BMI (26.4 vs. 23.6 kg m<sup>-2</sup>, data not shown) this may have increased the extent of underreporting in our group. Although energy intakes from our FFQ1 and recalls were similar, both are likely underestimates given that underreporting is extensive in all methods, particularly FFQs<sup>52–54</sup>. Although macronutrients are reported differently by LERs and non-LERs, it is not known to what extent each macronutrient contributes to energy underestimation, how well different dietary methods capture this, or how this affects validity<sup>52–54,74</sup>.

An additional issue, and one whose effects cannot be assessed here, is the suggestion that traditional validation studies are flawed by systematic intake-related and person-specific biases (e.g. overreporting of healthy foods) which, by being large and shared by FFQs and report-based reference measures such as recalls, undermine critical validation assumptions of unbiasedness and independence<sup>76</sup>. Using various FFQs from the European Prospective Investigation into Cancer and Nutrition (EPIC), including a modified Willett FFQ used in the UK arm, models correcting for these biases have suggested that validation studies traditionally overestimate FFQ accuracy and validity, and modify the ability to detect diet–disease relationships<sup>76,77</sup>, although it is not clear how this would specifically alter our results. Future validation studies may benefit from the inclusion of biomarker reference methods such as urinary nitrogen, potassium or doubly labelled water to better evaluate measurement error associated with protein, potassium or energy, respectively<sup>54,76</sup>, and its impact on the design and interpretation of epidemiological studies where FFQs are applied.

## Acknowledgements

This research was supported by the Canadian Breast Cancer Research Alliance with special funding support of the Canadian Breast Cancer Foundation no. Ontario Chapter (CBCRA grant no. 013572).

The authors wish to thank Sarah Winterton for assistance with dietary assessment and study co-ordination.

## References

- Block G, Hartman AM, Dresser CM, Carroll MD, Gannon J, Gardner L. A data-based approach to diet questionnaire design and testing. *American Journal of Epidemiology* 1986; **124**: 453–69.
- Cade J, Thompson R, Burley V, Warm D. Development, validation and utilisation of food-frequency questionnaires – a review. *Public Health Nutrition* 2002; **5**: 567–87.
- Willett W. *Nutritional Epidemiology*, 2nd ed. New York: Oxford University Press, 1998.
- Subar AF, Thompson FE, Kipnis V, Midthune D, Hurwitz P, McNutt S, *et al.* Comparative validation of the Block, Willett and National Cancer Institute food frequency questionnaires: the Eating at America's Table Study. *American Journal of Epidemiology* 2001; **154**: 1089–99.
- Stark A. An historical review of the Harvard and the National Cancer Institute food frequency questionnaires: their similarities, differences, and their limitations in assessment of food intake. *Ecology of Food and Nutrition* 2002; **41**: 35–74.
- Wirfält AKE, Jeffery RQ, Elmer PJ. Comparison of food frequency questionnaires: the reduced Block and Willett questionnaires differ in ranking on nutrient intakes. *American Journal of Epidemiology* 1998; **148**: 1148–56.
- Block G, Coyle LM, Hartman AM, Scoppa SM. Revision of dietary analysis software for the Health Habits and History Questionnaire. *American Journal of Epidemiology* 1994; **139**: 1190–6.
- Block G. Invited commentary: comparison of the Block and Willett food frequency questionnaires. *American Journal of Epidemiology* 1998; **148**: 1160–1.
- Block G. Invited commentary: another perspective on food frequency questionnaires. *American Journal of Epidemiology* 2001; **154**: 1103–4.
- Block G, Wakimoto P, Block T. A revision of the Block dietary questionnaire and database, based on NHANES III data [online]. Available at [http://www.nutritionquest.com/B98\\_Dev.pdf](http://www.nutritionquest.com/B98_Dev.pdf). Accessed 29 October 2004.
- Caan BJ, Slattery ML, Potter J, Quesenberry CP, Coates AO, Schaffer DM. Comparison of the Block and Willett self-administered semiquantitative food frequency questionnaires with an interviewer-administered dietary history. *American Journal of Epidemiology* 1998; **149**: 1137–47.
- Tylavsky FA, Sharp GB. Misclassification of nutrient and energy intake from use of closed-ended questions in epidemiologic research. *American Journal of Epidemiology* 1995; **142**: 342–51.
- Thompson FE, Subar AF, Brown CC, Smith AF, Sharbaugh CO, Jobe JB, *et al.* Cognitive research enhances accuracy of food frequency questionnaire reports: results of an experimental validation study. *Journal of the American Dietetic Association* 2002; **102**: 212–25.
- Sobell J, Block G, Koslowe P, Tobin J, Andres R. Validation of a retrospective questionnaire assessing diet 10–15 years ago. *American Journal of Epidemiology* 1989; **130**: 173–87.
- Block G, Woods M, Potosky A, Clifford C. Validation of a self-administered diet history questionnaire using multiple diet records. *Journal of Clinical Epidemiology* 1990; **43**: 1327–35.
- Coates RJ, Eley JW, Block G, Gunter EW, Sowell AL, Grossman C, *et al.* An evaluation of a food frequency questionnaire for assessing dietary intake of specific carotenoids and vitamin E among low-income black women. *American Journal of Epidemiology* 1991; **134**: 658–71.
- Tucker KL, Bianchi LA, Maras J, Bermudez OI. Adaptation of a food frequency questionnaire to assess diets of Puerto Rican and non-Hispanic adults. *American Journal of Epidemiology* 1998; **148**: 507–18.
- Mayer-Davis EJ, Vitolins MZ, Carmichael SL, Hemphill S, Tsaroucha G, Rushing J, *et al.* Validity and reproducibility of a food frequency interview in a multi-cultural epidemiological study. *Annals of Epidemiology* 1999; **9**: 314–24.
- Kim J, Chan MM, Shore RE. Development and validation of a food frequency questionnaire for Korean Americans. *International Journal of Food Sciences and Nutrition* 2002; **53**: 129–42.
- Kumanyika SK, Mauger D, Mitchell DC, Phillips B, Smiciklas-Wright H, Palmer JR. Relative validity of food frequency questionnaire nutrient estimates in the Black Women's Health Study. *Annals of Epidemiology* 2003; **13**: 111–8.
- Margetts BM, Nelson M. *Design Concepts in Nutritional Epidemiology*, 2nd ed. New York: Oxford University Press, 1977.
- Block G, Hartman AM. Issues in reproducibility and validity of dietary studies. *American Journal of Clinical Nutrition* 1989; **50**: 1133–8.
- Byers T. Food frequency dietary assessment: how bad is good enough? *American Journal of Epidemiology* 2001; **154**: 1087–8.
- Whiting SJ, Boyle JL, Thompson A, Mirwald RL, Faulkner RA. Dietary protein, phosphorus and potassium are beneficial to bone mineral density in adult men consuming adequate dietary calcium. *Journal of the American College of Nutrition* 2002; **21**: 402–9.
- Goodwin PJ, Ennis M, Pritchard KI, Koo J, Trudeau ME, Hood N. Diet and breast cancer: evidence that extremes in diet are associated with poor survival. *Journal of Clinical Oncology* 2003; **21**: 2500–7.
- Jain MG, Rohan TE, Soskolne CL, Kreiger N. Calibration of the dietary questionnaire for the Canadian Study of Diet, Lifestyle and Health cohort. *Public Health Nutrition* 2003; **6**: 79–86.
- Bright-See E, Catlin G, Godin G. Assessment of the relative validity of the Ontario Health Survey food frequency questionnaire. *Journal of the Canadian Dietetic Association* 1994; **55**: 33–8.
- Shakel SF, Sievert YA, Buzzard IM. Sources of data for developing and maintaining a nutrient database. *Journal of the American Dietetic Association* 1988; **88**: 1268–71.
- Van Horn L, Stumbo P, Moag-Stahlberg A, Obarzanek E, Hartmuller V, Farris R, *et al.* The Dietary Intervention Study in Children (DISC): dietary assessment method for 8- to 10-year olds. *Journal of the American Dietetic Association* 1993; **93**: 1396–403.
- Chambers EIV, Godwin SL, Vecchio F. Cognitive strategies for reporting portion sizes using dietary recall procedures. *Journal of the American Dietetic Association* 2000; **100**: 891–7.
- Yanek LR, Moy TF, Raqueno JV, Becker DM. Comparison of the effectiveness of a telephone 24-hour dietary recall method vs an in-person method among urban African-American women. *Journal of the American Dietetic Association* 2000; **100**: 1172–7.

- 32 Patterson RE, Kristal AR, Levy L, McLerran D, White E. Validity of methods used to assess vitamin and mineral supplement use. *American Journal of Epidemiology* 1998; **148**: 643–9.
- 33 Satia-Abouta J, Patterson RE, King IB, Stratton KL, Shattuck AL, Kristal AR, *et al.* Reliability and validity of self-report of vitamin and mineral supplement use in the Vitamins and Lifestyle Study. *American Journal of Epidemiology* 2003; **157**: 944–54.
- 34 Health Canada. Drug product database [online]. Available at <http://www.hc-sc.gc.ca/hpb/drugs-dpd/>. Accessed 29 October 2004.
- 35 Haines PS, Hama MY, Guilkey DK, Popkin BM. Weekend eating in the United States is linked with greater energy, fat, and alcohol intake. *Obesity Research* 2003; **11**: 945–9.
- 36 Brathen G, Brodtkorb E, Sand T, Helde G, Bovim G. Weekday distribution of alcohol consumption in Norway: influence on the occurrence of epileptic seizures and stroke. *European Journal of Neurology* 2000; **7**: 413–21.
- 37 Beaton GH, Milner J, Corey P, McGuire V, Cousins M, Stewart E, *et al.* Sources of variation in 24-hour dietary recall data: implications for nutrition study design and interpretation. *American Journal of Clinical Nutrition* 1979; **32**: 2546–59.
- 38 Rosner B, Willett WC. Interval estimates for correlation coefficients corrected for within-person variation: implications for study design and hypothesis testing. *American Journal of Epidemiology* 1988; **127**: 377–86.
- 39 Kelemen LE, Anand SS, Vuksan V, Yi Q, Teo KK, Devanesen S, *et al.* Development and evaluation of cultural food frequency questionnaires for South Asians, Chinese, and Europeans in North America. *Journal of the American Dietetic Association* 2003; **103**: 1178–84.
- 40 Mares-Perlman JA, Klein BEK, Klein R, Ritter LL, Fisher MR, Freudenheim JL. A diet history questionnaire ranks nutrient intakes in middle-aged and older men and women similarly to multiple food records. *Journal of Nutrition* 1993; **123**: 489–501.
- 41 Patterson RE, Kristal AR, Tinker LF, Carter RA, Bolton MP, Agur-Collins T. Measurement characteristics of the Women's Health Initiative food frequency questionnaire. *Annals of Epidemiology* 1999; **9**: 178–87.
- 42 Messerer M, Johansson SE, Wolk A. The validity of questionnaire-based micronutrient intake estimates is increased by including dietary supplement use in Swedish men. *Journal of Nutrition* 2004; **134**: 1800–5.
- 43 Rimm EB, Giovannucci EL, Stampfer MJ, Colditz GA, Litin LB, Willett WC. Reproducibility and validity of an expanded self-administered semiquantitative food frequency questionnaire among male health professionals. *American Journal of Epidemiology* 1992; **135**: 1114–26.
- 44 Barr SI, Murphy SP, Poos MI. Interpreting and using the Dietary References Intakes in dietary assessment of individuals and groups. *Journal of the American Dietetic Association* 2002; **102**: 780–8.
- 45 Rimm EB, Giovannucci EL, Stampfer MJ, Colditz GA, Litin LB, Willett WC. Authors' response to 'Invited commentary: some limitations of semiquantitative food frequency questionnaires'. *American Journal of Epidemiology* 1992; **135**: 1133–6.
- 46 White E, Patterson RE, Kristal AR, Thornquist M, King I, Shattuck AL, *et al.* VITamins And Lifestyle cohort study: study design and characteristics of supplement users. *American Journal of Epidemiology* 2004; **159**: 83–93.
- 47 Troppmann L, Gray-Donald K, Johns T. Supplement use: is there any nutritional benefit? *Journal of the American Dietetic Association* 2002; **102**: 818–25.
- 48 Palaniappan U, Starkey IJ, O'Loughlin JO, Gray-Donald K. Fruit and vegetable consumption is lower and saturated fat intake is higher among Canadians reporting smoking. *Journal of Nutrition* 2001; **131**: 1952–8.
- 49 Hoggatt KJ, Bernstein L, Reynolds P, Anton-Culver H, Deapen D, Peel D, *et al.* Correlates of vitamin supplement use in the United States: data from the California Teachers Study cohort. *Cancer Causes & Control* 2002; **13**: 735–40.
- 50 Radimer K, Bindewald B, Hughes J, Ervin B, Swanson C, Picciano MF. Dietary supplement use by US adults: data from the National Health and Nutrition Examination Survey, 1999–2000. *American Journal of Epidemiology* 2004; **160**: 339–49.
- 51 Foote JA, Murphy SP, Wilkens LR, Hankin JH, Henderson BE, Kolonel LN. Factors associated with dietary supplement use among healthy adults of five ethnicities. *American Journal of Epidemiology* 2003; **157**: 888–97.
- 52 Subar AF, Kipnis V, Troiano RP, Midthune D, Schoeller DA, Bingham S, *et al.* Using biomarkers to evaluate the extent of dietary misreporting in a large sample of adults: the OPEN study. *American Journal of Epidemiology* 2003; **158**: 1–13.
- 53 Schatzkin A, Kipnis V, Carroll RJ, Midthune D, Subar AF, Bingham S, *et al.* A comparison of a food frequency questionnaire with a 24-hour recall for use in an epidemiological cohort study: results for the biomarker-based Observing Protein and Energy Nutrition (OPEN) study. *International Journal of Epidemiology* 2003; **32**: 1054–62.
- 54 Freedman LS, Midthune D, Carroll RJ, Krebs-Smith S, Subar A, Troiano RP, *et al.* Adjustments to improve the estimation of usual dietary intake distributions in the population. *Journal of Nutrition* 2004; **134**: 1836–43.
- 55 Krebs-Smith SM, Heimendinger J, Subar AF, Patterson BH, Pivonka E. Using food frequency questionnaires to estimate fruit and vegetable intake: association between the number of questions and total intakes. *Journal of Nutrition Education* 1995; **27**: 80–5.
- 56 Amanatidis S, Mackerras D, Simpson JM. Comparison of two frequency questionnaires for quantifying fruit and vegetable intake. *Public Health Nutrition* 2001; **4**: 233–9.
- 57 Block G, Sinha R, Gridley G. Collection of dietary-supplement data and implications of analysis. *American Journal of Clinical Nutrition* 1994; **59**: 232S–9S.
- 58 Block G, Norkus E, Hudes M, Mandel S, Helzlsouer K. Which plasma antioxidants are most related to fruit and vegetable consumption? *American Journal of Epidemiology* 2001; **154**: 1113–8.
- 59 Johansson I, Hallmans G, Wikman A, Biessy C, Riboli E, Kaaks R. Validation and calibration of food-frequency questionnaire measurements in the Northern Sweden Health and Disease cohort. *Public Health Nutrition* 2002; **5**: 487–96.
- 60 Basiotis PP, Welsh SO, Cronin FJ, Kelsay JL, Mertz W. Number of days of food intake records required to estimate individual and group nutrient intakes with defined confidence. *Journal of Nutrition* 1987; **117**: 1638–41.
- 61 Beaton GH. Approaches to analysis of dietary data: relationship between planned analyses and choice of methodology. *American Journal of Clinical Nutrition* 1994; **59**: 253S–61S.
- 62 Kushi LH. Gaps in epidemiologic research methods: design considerations for studies that use food-frequency questionnaires. *American Journal of Clinical Nutrition* 1994; **59**: 180S–4S.
- 63 Beaton GH, Milner J, McGuire V, Feather TE, Little JA. Source of variance in 24-hour dietary recall data: implications for nutrition study design and interpretation. Carbohydrate sources, vitamins, and minerals. *American Journal of Clinical Nutrition* 1983; **37**: 986–95.
- 64 Nelson M, Black AE, Morris JA, Cole TJ. Between- and within-subject variation in nutrient intake from infancy to old age: estimating the number of days required to rank dietary intakes with desired precision. *American Journal of Clinical Nutrition* 1989; **50**: 155–67.

- 65 Hartman AM, Brown CC, Palmgren J, Pietinen P, Verkasalo M, Myer D, *et al.* Variability in nutrient and food intakes among older middle-aged men: implications for design of epidemiologic and validation studies using food recording. *American Journal of Epidemiology* 1990; **132**: 999–1012.
- 66 National Health and Nutrition Examination Survey [online]. Available at <http://www.cdc.gov/nchs/nhanes.htm>. Accessed 29 October 2004.
- 67 Godwin SL, Chambers E, Cleveland L. Accuracy of reporting dietary intake using various portion-size aids in-person and via telephone. *Journal of the American Dietetic Association* 2004; **104**: 585–94.
- 68 Lee RD, Nieman DC, Rainwater M. Comparison of eight microcomputer dietary analysis programs with the USDA Nutrient Data Base for Standard Reference. *Journal of the American Dietetic Association* 1995; **95**: 858–67.
- 69 McCullough ML, Karanja NM, Lin PH, Obarzanek E, Phillips KM, Laws RL, *et al.* Comparison of 4 nutrient databases with chemical composition data from the Dietary Approaches to Stop Hypertension trial. *Journal of the American Dietetic Association* 1999; **99**: S45–53.
- 70 Csizmadia I, Subar A, Kahle L, Ullman R, Dawe U, Zimmerman T, *et al.* Adaptation of the National Cancer Institute's Diet History Questionnaire and nutrient database for Canadian populations. *Journal of Nutrition* 2004; **134**: 3532S.
- 71 Statistics Canada. 2001 Census: Analysis Series. Canada's ethnocultural portrait: the changing mosaic [online], 2003. Available at <http://www.gov.on.ca/FIN/english/demographics/cenhi6e.htm>. Accessed 29 October 2004.
- 72 Health Canada. Adult smoking in Canada. Annual Canadian Tobacco Use Monitoring Survey (CTUMS), February–December 2000 [online]. Available at [http://www.hc-sc.gc.ca/hecs-sesc/tobacco/pdf/CTUMS-2000e%20\(Adult\).pdf](http://www.hc-sc.gc.ca/hecs-sesc/tobacco/pdf/CTUMS-2000e%20(Adult).pdf). Accessed 29 October 2004.
- 73 Statistics Canada. Census of population, population 15 years and over by highest degree, certificate or diploma (1986–2001 censuses) [online]. Available at <http://www.statcan.ca/english/Pgdb/educ42.htm>. Accessed 29 October 2004.
- 74 Livingstone MBE, Black AE. Markers of the validity of reported energy intake. *Journal of Nutrition* 2003; **133**: 895S–920S.
- 75 Bedard D, Shatenstein B, Nadon S. Underreporting of energy intake from a self-administered food-frequency questionnaire completed by adults in Montreal. *Public Health Nutrition* 2004; **7**: 675–81.
- 76 Kipnis V, Midthune D, Freedman L, Bingham S, Day NE, Riboli E, *et al.* Bias in dietary-report instruments and its implications for nutritional epidemiology. *Public Health Nutrition* 2002; **5**: 915–23.
- 77 Kipnis V, Midthune D, Freedman LS, Bingham S, Schatzkin A, Subar A, *et al.* Empirical evidence of correlated biases in dietary assessment instruments and its implications. *American Journal of Epidemiology* 2001; **153**: 394–403.