1	Vertical distribution of buoyant <i>Microcystis</i> blooms in a
2	Lagrangian particle tracking model for short-term forecasts
3	in Lake Erie
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17	Key Points
18	Microcystis vertical distribution is a dynamic balance between turbulence and buoyancy

19 Appropriate time step and numerical scheme avoid artifacts in random walk models

20 Vertical mixing with buoyancy improved simulation of bloom spatial distribution

21 Abstract

22 Cyanobacterial harmful algal blooms (CHABs) are a problem in western Lake Erie, and in 23 eutrophic fresh waters worldwide. Western Lake Erie is a large (3000 km²), shallow (8 m mean 24 depth), freshwater system. CHABs occur from July to October, when stratification is intermittent 25 in response to wind and surface heating or cooling (polymictic). Existing forecast models give 26 the present location and extent of CHABs from satellite imagery, then predict two-dimensional 27 (surface) CHAB movement in response to meteorology. In this study, we simulated vertical 28 distribution of buoyant Microcystis colonies, and 3D advection, using a Lagrangian particle 29 model forced by currents and turbulent diffusivity from the Finite Volume Community Ocean 30 Model (FVCOM). We estimated the frequency distribution of *Microcystis* colony buoyant velocity from measured size distributions and buoyant velocities. We evaluated several random-31 32 walk numerical schemes to efficiently minimize particle accumulation artifacts. We selected the 33 Milstein scheme, with linear interpolation of the diffusivity profile in place of cubic splines, and 34 varied the time step at each particle and step based on the curvature of the local diffusivity profile 35 to ensure that the Visser time step criterion was satisfied. Inclusion of vertical mixing with 36 buoyancy significantly improved model skill statistics compared to an advection-only model, and 37 showed greater skill than a persistence forecast through simulation day 6, in a series of 26 38 hindcast simulations from 2011. The simulations and in-situ observations show the importance of

39	subtle thermal structure, typical of a polymictic lake, along with buoyancy in determining vertical
40	and horizontal distribution of Microcystis.
41	Index Terms
42	4239 Limnology, 4842 Modeling, 4855 Plankton
43	Keywords
44	Numerical modeling, Ecological forecasting, Lake Erie, Microcystis, Harmful algal bloom,
45	cyanobacteria
46	
47	1 Introduction
48	Harmful algal blooms (HABs) are a global problem that is linked to anthropogenic eutrophication
49	of inland and coastal waters, and may be exacerbated by climate change [O'Neil, J. et al., 2012].
50	Lake Erie has experienced recurring blooms of toxin-producing cyanobacteria since the mid
51	1990s [Brittain, S. M. et al., 2000; Michalak, A. M. et al., 2013; Wynne, T. T. and R. P. Stumpf,
52	2015]. Lake Erie is the most productive, warm, and shallow of the Laurentian Great Lakes of
53	North America. In the open waters of Lake Erie, cyanobacterial harmful algal blooms (CHABs)
54	generally occur from July to October, and are dominated by the species Microcystis aeruginosa,
55	which produces the group of hepatotoxin compounds known as microcystins [Rinta-Kanto, J. M.
56	et al., 2009]. CHABs occur primarily in the shallow western basin, which receives the main
57	hydraulic load from the Detroit River in the north and the main nutrient load from the Maumee

58 River [Kane, D. D. et al., 2014] in the southwest (Fig. 1a). Occasionally, CHABs are transported

through the islands into the deeper central basin, while the eastern basin is largely free of CHABs
(Fig. 1a)[*Wynne, T. T. and R. P. Stumpf*, 2015]. A bloom of record-breaking intensity and extent
occurred in 2011. Analysis by Michalak et al. [2013] indicated that the conditions of meteorology
and agricultural land use that caused the 2011 record bloom are likely to recur, and the 2015
bloom subsequently exceeded the severity of the 2011 record bloom [*Stumpf, R. P. and T. T. Wynne*, 2015].

65 Forecasts of CHAB abundance and spatial distribution are useful to water treatment plant 66 operators, anglers, recreational boaters, and beach users. Lake Erie is a source of drinking water to 11 million people [French, R. et al., 2011]. In 2014, the City of Toledo issued a do-not-drink 67 68 order that affected a half million residents for two days as a result of contamination of treated 69 water by microcystins [Henry, T., 2014]. In addition to spatial forecasts, forecasts of Microcystis 70 vertical distribution are of interest to water treatment plant operators because intake structures are 71 usually located sub-surface, so the risk of toxins in their raw water may be greater during mixing 72 events than when *Microcystis* colonies are concentrated on the surface. In addition to providing 73 drinking water, Lake Erie supports economically valuable sport and commercial fisheries as well as recreation and tourism. 74

Short-term and seasonal CHAB forecasts have been developed for Lake Erie. Seasonal forecasts predict the annual bloom severity using statistical models based on the cumulative March to July phosphorus load from the Maumee River, which are able to explain > 90% of the interannual variance in bloom severity [*Obenour, D. R. et al.*, 2014]. Bloom severity is defined as the lake-wide cyanobacterial biomass averaged over the 30 days containing the maximum biomass

80 [Obenour, D. R. et al., 2014; Stumpf, R. P. et al., 2012], and has been estimated using both 81 satellite [Stumpf, R. P. et al., 2012] and plankton tow data [Bridgeman, T. B. et al., 2013]. The 82 seasonal forecast is used by water treatment plant managers for seasonal planning, to determine 83 recommended phosphorus load reductions to meet commitments under the Great Lakes Water 84 Quality Agreement [GLWQA Annex 4, 2015], and is distributed to over 1600 subscribers. Short-85 term forecasts are distinct from the seasonal forecast in the greater importance of physical 86 transport processes over biological mechanisms in explaining short-term bloom variability. Experimental short-term forecasts have been developed for Lake Erie that indicate the present 87 88 location and extent of CHABs from satellite imagery, then predict future movement of the CHAB 89 using forecast meteorology, a hydrodynamic model, and a Lagrangian particle tracking model to 90 simulate horizontal advection of neutrally-buoyant particles at the water surface [Wynne, T. T. et 91 al., 2013; Wynne, T. T. et al., 2011].

92 Skill assessment of short-term CHAB forecasts is needed so that forecast users may have an 93 appropriate level of confidence in forecast data, and for development of improved models. 94 However, CHAB forecast skill assessment can be challenging, and methods are not well 95 established. Wynne et al. [2011] showed that a forecast model more accurately predicted 96 horizontal movement of the bloom centroid (center of mass) than a persistence forecast over a 97 ten-day simulation for an August 2003 event. A persistence forecast is a benchmark used for 98 model skill assessment in which no change is assumed from the initial observed location. The 99 same quantitative skill assessment method could not be applied in the case of an August 2008 100 event because cloud cover and vertical mixing obscured parts of the bloom in subsequent images

101	and prevented an accurate estimate of the centroid location [Wynne, T. T. et al., 2011]. In a
102	detailed analysis of the August 2008 bloom in Lake Erie, Wynne et al. [2010] found that wind
103	speed was a significant predictor of apparent bloom intensity in satellite imagery, and presented
104	evidence to support the hypothesis that variation in the mixing depth of the buoyant
105	cyanobacterial colonies was the underlying mechanism causing changes in surface concentration.
106	In a qualitative analysis of the first three years of the experimental Lake Erie CHAB forecast
107	(2008-2010), Wynne et al. [2013] concluded that the forecast provided useful information, but
108	could be improved by a means to fill in cloud-covered areas in the satellite images, and by
109	simulation of vertical mixing of the buoyant cyanobacteria.
110	To simulate concentration profiles of buoyant particles produced by vertical mixing in stratified
111	turbulence, the partial differential equations governing advection and diffusion may be solved
112	from an Eulerian or a Lagrangian perspective; each approach has strengths and weaknesses. Real
113	phytoplankton communities have properties that vary among individuals within a population,
114	such as size, specific gravity, light exposure history, and nutrient quotas. A strength of
115	Lagrangian particle models is that nearly continuous distributions of properties can be
116	represented by allowing properties to vary by particle [e.g., Ross, O. N. and J. Sharples, 2008],
117	while in Eulerian models property distributions are usually represented by a limited number of
118	discrete property classes [e.g., Medrano, E. A. et al., 2013].
119	A weakness of the Lagrangian approach is that artificial accumulation of particles can occur in
120	low diffusivity areas in random-walk turbulence simulations for the case of spatially (vertically)

121 non-uniform diffusivity if an inappropriate numerical scheme or time step is used [Visser, A.,

122 1997]. These artifacts can be easy to misinterpret as features of interest. Several studies have 123 applied Lagrangian particle models to the 1D column case with steady, idealized diffusivity 124 profiles [Gräwe, U. et al., 2012; Ross, O. N. and J. Sharples, 2008; Visser, A., 1997]. In 3D 125 ocean and lake models, use of random-walk vertical mixing schemes can be challenging because 126 the required time step varies spatially and temporally over the model domain. A fixed time step 127 may be selected that is adequate in most places and times [Huret, M. et al., 2007], or in some 128 cases a well-mixed profile was imposed in shallow regions where a small time step would be 129 required to avoid formation of artifacts [Gilbert, C. et al., 2010]. Our application required 130 realistic rather than idealized diffusivity profiles for a polymictic lake in which conditions are 131 alternately turbulent and stratified, and to simulate concentration profiles rather than specify a 132 well-mixed condition, even in shallow areas, therefore some development of the random-walk 133 approach was required.

134 We describe a short-term forecast system for CHAB abundance and distribution in Lake Erie that 135 takes a similar approach to that of Wynne et al. [2011], but uses updated hydrodynamic and 136 Lagrangian particle tracking models, and includes a means of filling in cloud-covered areas of 137 satellite images using model data from a previous run. In addition, we describe a method to 138 simulate vertical distribution of buoyant cyanobacteria in stratified turbulence. We evaluated the 139 performance of several random-walk turbulence numerical schemes in terms of computational 140 efficiency and their ability to minimize artifacts in simulations with vertically-varying diffusivity 141 typical of a large polymictic lake. We compared simulated vertical distributions of cyanobacteria 142 to observed profiles in Lake Erie. Finally, we show that model skill statistics were improved by

including vertical mixing with buoyancy in hindcast simulations of the spatial distribution of the2011 bloom in Lake Erie.

145

146 **2** Methods

147 2.1 Hydrodynamic model

148 The Finite Volume Community Ocean Model (FVCOM, v. 3.2) is an unstructured grid, finite-

149 volume, free surface, three-dimensional primitive equation ocean model that solves the

150 momentum, continuity, temperature, salinity, and density equations [Chen, C. et al., 2003].

151 Turbulence closure is implemented through the MY-2.5 scheme for vertical mixing [Galperin, B.

152 et al., 1988], and the Smagorinsky scheme for horizontal mixing [Smagorinsky, J., 1963].

153 FVCOM has been adapted and implemented for the Great Lakes in several recent studies

154 [Anderson, E. J. et al., 2015; Anderson, E. J. and D. J. Schwab, 2013; Anderson, E. J. et al.,

155 2010; Bai, X. et al., 2013], yielding accurate predictions of temperature, water levels, and

156 currents. In particular, FVCOM has been applied to Lake Erie for extreme storm prediction

157 [Anderson, E. J. et al., 2015] and serves as the oceanographic model underlying NOAA's next-

158 generation Lake Erie Operational Forecast System (LEOFS), a real-time short-term

159 hydrodynamic forecast model (<u>http://tidesandcurrents.noaa.gov/ofs/leofs/leofs.html</u>).

160 The FVCOM-based LEOFS model was applied for this study with bathymetry interpolated from

161 the NOAA National Geophysical Data Center.

162 (www.ngdc.noaa.gov/mgg/greatlakes/greatlakes.html). The unstructured grid consisted of 6,106

163	nodes and 11,509 elements (Fig. 1). Spatial resolution was 2 km in the central basin, 1.5 km in
164	the western basin and 0.5 km in Maumee Bay and the islands (Fig. 1b) to improve representation
165	of currents in these key areas for CHAB transport, with 20 uniform vertical sigma layers.
166	Dynamic water levels (6-minute) were prescribed at open boundaries at the Detroit River and the
167	Niagara River, taken from adjustments to the NOAA/NOS gauges at Gibraltar (9044020) and
168	Buffalo (9063020), to drive the primary inflow and outflow, respectively (Fig. 1a). Atmospheric
169	forcing conditions were generated using station-based interpolation methods as in the NOAA
170	Great Lakes Coastal Forecasting System (Beletsky et al., 2003; Schwab and Beletsky, 1998).
171	Hourly meteorological forcing variables of wind speed, wind direction, air temperature, dewpoint
172	temperature, and cloud cover were interpolated over Lake Erie from several land-based
173	meteorological stations and offshore NOAA/NDBC buoys (45004, 45132, and 45142), when
174	available. Wind speeds were adjusted to 10-m height and empirical relationships were used to
175	adjust land-based meteorological variables for over-lake modification (Beletsky et al., 2003;
176	Schwab and Beletsky, 1998).
177	Hydrodynamic model simulations were based on the real-time LEOFS nowcast, which was
178	initialized on January 1, 2004 with uniform temperature of 2 °C. For the 2011 scenario presented
179	here, the hydrodynamic model simulation was initialized on January 1, 2011 with initial
180	conditions provided by the simulation from the previous year, and produced hourly output of
181	three-dimensional currents, water temperature, turbulent diffusivity, and 2D water level
182	fluctuations.

183 **2.2** Lagrangian particle tracking model

- 184 Lagrangian particle tracking was accomplished using a Fortran program developed previously to
- 185 study transport of larval cod in the Gulf of Maine [Churchill, J. H. et al., 2011; Huret, M. et al.,
- 186 2007], which is distributed as part of the FVCOM code package
- 187 (http://fvcom.smast.umassd.edu/), and has previously been applied in the Great Lakes [Anderson,
- 188 E. J. and M. S. Phanikumar, 2011]. Advection of particles was determined by

$$\frac{d}{dt}\boldsymbol{X}(t) = \boldsymbol{V}(\boldsymbol{X}(t), t) \tag{1}$$

189 where X(t) is the three-dimensional particle position at time t, and V(X(t),t) is the three-

190 dimensional, time varying velocity field. Linear interpolation in space and time was used to

191 obtain V(X(t),t) from hourly-archived FVCOM output. The contribution of advection to the

192 particle position was updated by integrating Eq. (1) using an explicit fourth-order Runge-Kutta

193 scheme with a time step, $\Delta t = 600$ s. Vertical mixing due to turbulent eddy diffusivity was

194 optionally simulated using one of several random-walk schemes, described below. We used

195 reflected boundary conditions at vertical and horizontal boundaries in all simulations.

196 **2.3 Well-mixed condition test**

We compared the performance of several random walk vertical mixing numerical schemes using a well-mixed condition test case [*Visser*, *A.*, 1997] that can be used to evaluate whether a given numerical scheme and time step will produce artifacts in simulated concentration profiles. In our case, the well-mixed condition simulation was performed by initiating 1000 particles, uniformly distributed through the water column, then simulating 1D vertical mixing with neutral buoyancy 10 using time series of Lake Erie diffusivity profiles output from FVCOM for the month of August,
203 2011 (Table 1). Random noise in the simulated concentration profiles decreases with increasing
number of particles, and 1000 particles were found to be adequate in 1D simulations [*Ross, O. N.*205 *and J. Sharples*, 2004]. These simulations are expected to give uniform concentration at all time
regardless of the diffusivity profile, consistent with the Eulerian solution to the 1D diffusion
207 equation with initial uniform concentration [*Visser, A.*, 1997].

To evaluate performance in the well-mixed condition test in a way that is directly relevant to our application, we defined a signal-to-noise ratio, *SN*, based on simulated surface concentration,

$$SN(t) = C(t)/|1.0 - C_{wm}(t)|$$
(2)

where $C_{wm}(t)$ is the time-dependent concentration in the 1-m thick surface layer, normalized to 210 211 the column-mean concentration, in a well-mixed condition simulation with neutral buoyancy. A 212 constant value of 1.0 is expected for $C_{wm}(t)$, therefore $|1.0 - C_{wm}(t)|$ represents the magnitude 213 of "noise" due to artifacts and random fluctuations due to calculating concentration by counting 214 discrete particles in a control volume. The "signal" is provided by the analogous surface 215 concentration, C(t), from an identical simulation with buoyancy. A large value of SN indicates 216 that artifacts are small compared to the effect of interest, which is the fluctuation in surface 217 concentration due to buoyancy. SN can be made arbitrarily large by using a small time step and a 218 large number of particles, but at the expense of computational time. We selected a value of 5 as a 219 goal minimum value for SN, based on the "Rose criterion" for the detection limit of the human 220 eye for image features [Rose, A., 1948]. Accordingly, we used the frequency of occurrence of SN

- 221 < 5 over the hourly records in a simulation as the performance criterion by which to evaluate the
- 222 numerical schemes and time step criteria.

223 **2.4 Random walk vertical mixing schemes**

We evaluated the Visser [1997] scheme, as-implemented by Huret et al. [2007]

$$z(t+\delta t) = z(t) + w_b \delta t + K'(z(t)) \delta t + R \sqrt{\frac{2K(\tilde{z})\delta t}{\sigma^2}}$$
(3)

where z(t) is the vertical position of the particle at time t, δt is the vertical random walk time step, w_b is a floating/sinking/swimming vertical velocity component, K is the vertical turbulent diffusivity, K' = dK/dz, R is a random variable sampled from a uniform distribution with zero mean and standard deviation σ , and $\tilde{z} = z(t) + 0.5K'(z(t))\delta t$ is a vertical position displaced from the particle position as a function of the diffusivity gradient. Following Ross and Sharples [2004], a cubic spline interpolation was used to obtain a continuous, differentiable diffusivity profile.

In addition to the Visser scheme, described above, we evaluated random walk vertical mixing schemes with higher order accuracy including the Milstein, Strong 1.5 (S1.5), and Platen twostep (PC2) schemes that were implemented in Fortran for the General Ocean Turbulence Model by Gräwe [2011]. After evaluating the performance of ten random walk schemes [*Gräwe, U.*, 2011], Grawe et al. [2012] recommended the use of either the Milstein scheme or higher-order schemes such as S1.5 or PC2. The order of accuracy (rate of convergence) of numerical approximations to stochastic differential equations is separated into weak and strong cases, where

the weak case relates to the accuracy of the ensemble particle distribution, while the strong case
relates to the accuracy of particle trajectories [*Gräwe*, *U*., 2011]. The weak order of accuracy of
the schemes that we evaluated was 1 for Visser and Milstein, and 2 for S1.5 and PC2. The strong
order of accuracy was 1 for Milstein, 1.5 for S1.5, and was not defined for Visser or PC2 [*Gräwe*, *U*., 2011]. The Milstein scheme is

$$z(t+\delta t) = z(t) + w_b \delta t + 0.5K'(z(t))[\Delta W^2 + \delta t] + \Delta W \sqrt{2K(z(t))}$$
⁽⁴⁾

where ΔW is a random variable drawn from a Gaussian distribution with zero mean and standard 244 deviation $\sqrt{\delta t}$. Because the Milstein scheme is first order, linear interpolation of the diffusivity 245 profile to obtain K(z(t)) and K'(z(t)) are sufficient, and the added computational expense 246 247 compared to the Visser scheme (Eq. 3) is minimal. The second order schemes S1.5 and PC2 248 retain additional terms from the stochastic Taylor expansion, including higher order derivatives 249 that require the application of cubic splines to the diffusivity profile, and S1.5 requires an 250 additional random variable, therefore the second order schemes have greater computational 251 expense. We refer to Gräwe [2011] for a full explanation of the S1.5 and PC2 schemes.

252 **2.5 Random walk time step**

Visser [1997] introduced a time step criterion, $\delta t \ll \min(1/|K''|)$, where K'' is the second derivative of the diffusivity profile, so that the diffusivity profile is reasonably approximated by a first-order Taylor series expansion over the range of particle displacement. Ross and Sharples [2004] found that

$$\delta t = \frac{1}{100|K''|} \tag{5}$$

257 is acceptable in most applications. To ensure the use of an appropriate time step throughout a 3D 258 model domain, and to avoid the need to specify an appropriate δt in advance, we modified the 259 code of Huret et al. [2007] to allow an appropriate value of δt to be calculated and applied for 260 each particle at each δt . To evaluate Eq. (5) independently of spline interpolation, we calculated 261 K'' directly at the FVCOM sigma levels using a centered finite difference approximation, and 262 extended the profile beyond the surface and bottom by reflection. We further limited 0.01 $s \leq$ $\delta t \geq \Delta t$. We tracked the minimum δt during simulations to confirm that the lower limit of 0.01 263 264 seconds was rarely applied.

265 **2.6** Measurement of *Microcystis* colony size distribution

266 We assigned buoyant velocity in the model based on measured size distributions and regressions 267 of buoyant velocity versus colony diameter. We measured Microcystis colony diameter of Lugol 268 preserved samples collected from western Lake Erie in the summers of 2012, 2013, and 2014. In 269 2012 and 2013 colony diameters were measured by microscopy (Table 1). In 2014 we used the 270 FlowCam (Fluid Imaging Technologies). The FlowCam captures images of individual colonies 271 and estimates their equivalent spherical diameter by image analysis. Wang et al. [2015] showed 272 that counts and colony diameters of Microcystis given by FlowCAM and microscopic image 273 analysis diameters were nearly identical. Colonies in Lake Erie are typically $> 50 \ \mu m$ 274 [Vanderploeg, H. A. et al., 2001] and buoyant. Samples were preserved with 1% formalin upon 275 collection, immediately refrigerated, and analyzed within 24 hours. FlowCam analyses were 14

performed with a 2× objective and 1 × 3 mm field of view flow cell with samples diluted as per
manufacturer recommendations to avoid capturing more than one image per trigger event.
Fluorescent triggering mode was used to avoid imaging detrital material that might be confused
with *Microcystis*. Samples were diluted in 0.2 µm filtered algal culture media [e.g., *Vanderploeg, H. A. et al.*, 2001] and injected into the FlowCAM with a 60-mL syringe, which was constantly
turned over so as to prevent the buoyant colonies from aggregating in the syringe. The image
analysis algorithm was calibrated to identify the colony outline including the mucilage.

283 **2.7** Measurement of *Microcystis* colony buoyant velocity

284 Microcystis colony buoyant velocity was measured using microscopic videography [Bundy, M. H. et al., 1998; Strickler, J. R., 1985], a method in principle similar to that of Nakamura et al. 285 286 [1993]. Surface water samples were collected on 15 and 21 July 2015 (Table 1) at station WE15 287 (Fig. 1b, -83.0, 41.6) during the early afternoon. Water was placed in 1-L glass bottles in an 288 incubator outdoors with a neutral density filter to cut light intensity to 50% of surface irradiance. 289 Measurements of colony velocities were made throughout the morning and afternoon of the next 290 day. Digital video clips were captured of individual colonies rising through a 2-cm \times 2-cm cross 291 section × 30-cm tall glass cuvette filled with ambient lake water inside of a water jacket in a 292 temperature controlled environmental room maintained at the lake temperature. Image capture 293 and analysis software (Templo Motus, Vicon Motus, Contemplas, GmbH, Germany) was used to 294 measure the velocity of the rising colonies. Diameters of the colonies were determined from 295 image analysis of video images using Image-Pro software (Media Cybernetics, Rockville, MD).

Water samples were diluted as needed with 0.2- μm filtered lake water to avoid turbulence
induced by multiple rising colonies.

298 **2.8** Measured vertical profiles of temperature and cyanobacterial concentration

299 Vertical profiles of temperature and cyanobacterial concentration (reported in µg chlorophyll a L ¹) were measured using the FluoroProbe (bbe Moldaenke, GmbH), which uses spectral 300 301 fluorometry to partition total chlorophyll into multiple phytoplankton classes on the basis of their 302 characteristic pigments (green algae, cyanobacteria, diatoms/dinoflagellates, cryptophytes), with 303 correction for possible interference by colored dissolved organic matter [Catherine, A. et al., 304 2012; Kring, S. A. et al., 2014]. Standard factory calibration settings for representative algal 305 classes were used. Profiles were measured at 11 stations (Fig. 1) weekly from June through 306 September, 2015 (Table 1). Profiles were selected for model skill assessment in which the 307 cyanobacterial chlorophyll was greater than chlorophyll from other algal classes, which occurred 308 13 July to 28 September.

309 2.9 Satellite remote sensing data

To initialize model simulations, and for model skill assessment, we used a series of images of cyanobacterial blooms in Lake Erie from July to October 2011 (Table 1) that were derived from the Medium Resolution Imaging Spectrometer (MERIS) [*Wynne, T. T. and R. P. Stumpf,* 2015]. MERIS standard level 2 data sets (in units of sr⁻¹) were used with a spectral shape algorithm based around 680 nm [*Wynne, T. et al.,* 2008] to obtain the cyanobacterial index (CI). CI varies linearly with biomass, with a value of 10^{-3} sr⁻¹ corresponding to approximately 10^{5} cells mL⁻¹

316 [*Stumpf, R. P. et al.*, 2012], which is the World Health Organization's threshold of significantly 317 increased risk for human health effects [*Chorus, I. and J. Bartram*, 1999]. For our analysis, we 318 converted CI to chlorophyll concentration using an empirical relationship derived from field 319 radiometry and grab sample extracted chlorophyll from eutrophic lakes in Florida (R^2 0.95, std. 320 error 7.7 µg L⁻¹, range 16 to 115 µg L⁻¹). The relationship was also verified for satellite-derived 321 CI, and gave a relative root-mean square error of 27% [*Tomlinson, M. C. et al.*, 2016].

$$Chl = 12570 Cl + 10$$
 (6)

322 A value of $CI = 10^{-3} \text{ sr}^{-1}$ is approximately equivalent to 23 µg L⁻¹ chlorophyll, which we used as a 323 threshold to define the presence of a CHAB.

324 2.10 Hindcast simulations

325 Daily satellite images for the period July to October 2011 were evaluated, and 26 images were 326 selected that had > 50% cloud-free views of western Lake Erie. A ten-day model simulation was 327 initialized from each image by assigning surface chlorophyll concentration values to FVCOM 328 nodes by nearest neighbor interpolation. Concentration was converted to Lagrangian particles by specifying a chlorophyll mass per particle $(10^{10} \mu g \text{ Chl particle}^{-1})$ and placing the specified 329 330 number of particles within a control volume. The FVCOM node-centered tracer control elements 331 were used as control volumes (Fig. 1). Vertical layers were specified as constant-thickness (1 m) 332 z-layers.

333 Preliminary hindcast skill assessment indicated a need to censor satellite data within a buffer
334 region of shorelines due to frequent false positives in these areas, which was likely caused by

335	contamination of the water signal from relatively bright surrounding land or by bottom
336	reflectance. The buffer width was set to 1 km south of Stony Point (41.94 °Lat.) and east of
337	Catawba Island (-82.85 °Lon.) and to 1.5 km elsewhere. The buffer was not applied in Maumee
338	Bay because CHABs were often present [Wynne, T. T. and R. P. Stumpf, 2015], reducing the
339	likelihood of false positives. Buffered or missing data areas were assigned values by nearest
340	neighbor if pixels containing valid data were available within 2 km. If no valid pixels were
341	available within 2 km of a node one of two approaches was used: 1) it was assumed that no
342	CHAB was present ($Chl = 0$), or 2) model output was carried forward to fill in the no-data area if
343	a previous model run was available covering the time period.
344	Two types of simulations were run, 2D and 3D. In 2D simulations, the surface chlorophyll
345	concentration was applied to the surface 1 m, with $Chl = 0$ below, and random walk vertical
346	mixing was turned off. In 3D simulations, surface chlorophyll concentration was applied over the
347	surface mixed layer (SML) depth (see below), and random walk vertical mixing was simulated.

348 Both 2D and 3D simulations included 3D advection.

349 **2.11** Estimation of the surface mixed layer depth

It was necessary to estimate the surface mixed layer (SML) depth for buoyant *Microcystis* colonies for the purpose of initializing the vertical distribution of particles (chlorophyll concentration) in 3D simulations from satellite-derived surface chlorophyll concentration. The vertical distribution of buoyant particles in the water column depends on buoyancy in addition to turbulent diffusivity (e.g., Fig. 2d,e); therefore, we used the Lagrangian particle model to estimate the initial vertical distribution of *Microcystis* colonies rather than using diffusivity or temperature 18

356 profiles from FVCOM directly. Initial vertical profiles were simulated at a subset of FVCOM 357 nodes (stations) because a large number of particles is needed to obtain a well-resolved profile, which would be computationally intensive to simulate at all 6,106 nodes. The SML depth was 358 359 estimated by running 1D column vertical random walk simulations at 11 station locations (Fig. 360 1b) that were selected to provide representative coverage of the western basin where CHABs are 361 most common, with additional stations added at representative deeper locations. 1D simulations 362 were initialized with 1000 particles uniformly distributed over the column 36 hours prior to the 363 initialization time of the 3D model (satellite image time) to allow the particle distribution 364 sufficient time to adapt to the varying diffusivity. Random walk vertical mixing was forced by 365 hourly diffusivity output from FVCOM. The SML depth for Microcystis colonies was estimated 366 as the depth at which the 1D concentration profile decreased to half the surface concentration, 367 and the satellite-derived surface concentration was applied from the surface to this depth; this approach provides an unbiased estimate of the total column biomass for the case of a uniform 368 369 concentration profile or a profile that can be approximated by a linear decrease. SML depth was 370 interpolated spatially to the FVCOM nodes by the nearest neighbor method.

371 **2.12 Model skill statistics**

372 Comparison of model results to in-situ profile data was conducted using column-integrated 373 quantities to minimize the effect of noise in the profile data on the statistics. Turbulent diffusivity 374 is strongly influenced by the static stability of the water column, which can be quantified using 375 the potential energy anomaly, ϕ ,

$$\phi = \frac{1}{h} \int_0^h (\hat{\rho} - \rho) gz dz; \quad \hat{\rho} = \frac{1}{h} \int_0^h \rho dz \tag{7}$$

376 where ρ is the local density, *h* is the water column depth, and *g* is acceleration due to gravity 377 [*Simpson, J. and D. Bowers*, 1981; *Wiles, P. et al.*, 2006]. Vertical distribution of concentration 378 was characterized by calculating the center of mass of the normalized concentration profile, σ_m

$$\sigma_m = \frac{1}{C} \sum_{k=1}^{k_b} c_k \sigma_k; \ C = \sum_{k=1}^{k_b} c_k$$
(8)

where $\sigma = z/h$ is the normalized vertical coordinate, *c* is the concentration at layer *k* normalized to the column-mean concentration, $k_b = 20$ is the number of uniformly-spaced σ layers in the model grid. The observed concentration profile was averaged over the σ layers of the model grid for the purpose of comparison to the model profiles.

Skill assessment in hindcast simulations was conducted by comparing model results to remote sensing images that were within the model simulation period. Each hindcast simulation was initialized from a satellite image, and two to four subsequent images were typically available within the simulation period for skill assessment. Skill assessment was conducted using a binary categorical variable (CHAB, no CHAB), and pixel-by-pixel comparisons of model to remote sensing observations were conducted. FVCOM tracer control elements (Fig. 1) were used as the spatial segmentation (pixels).

- 390 Our approach to skill assessment statistics followed Hogan and Mason [2012]. Two statistics
- 391 were calculated from the elements of the contingency table, which are the number of *a*, correctly

392 predicted events (hits); *b*, false events (false alarms); *c*, false negatives (misses); and *d*, correct 393 non-events. The frequency bias (B) gives the ratio of the number of forecasts of occurrence to the 394 number of observed occurrences,

$$B = \frac{a+b}{a+c} \tag{9}$$

and the Pierce Skill Score (PSS) gives the hit rate minus the false alarm rate.

$$PSS = \frac{ad - bc}{(b+d)(a+c)}$$
(10)

An unbiased forecast has a frequency bias B = 1.0. PSS values range from -1.0 to 1.0, with positive values indicating that the hit rate was greater than the false positive rate, and therefore the model had greater skill than a random forecast or constant CHAB or no-CHAB prediction [*Hogan, R. J. and I. B. Mason*, 2012].

To provide a reference forecast for skill comparison, we defined a "persistence" forecast as the assumption of no change from the satellite image that was used to initialize the model, which represents the best available information to a forecast user in the absence of a useful model. We took the further steps of filling in missing data in the persistence forecast with the most recent satellite data for each spatial segment, and applying the same shoreline buffering procedure that was used to initialize the model.

To test whether the model had significantly greater skill than the persistence forecast, we used the bootstrap method described by Hogan and Mason [2012] to estimate the confidence interval around the difference in skill score of the model compared to the persistence forecast. Starting

409	with a series of <i>n</i> triplets of observations, model predictions, and persistence predictions, we
410	created 1000 different bootstrap samples, each of length n , by taking random samples with
411	replacement from the series. We then calculated the difference in <i>PSS</i> , ΔPSS , for each bootstrap
412	sample. Finally, we estimated the 95% confidence interval as the 0.025 to 0.975 quantiles of the
413	ensemble of 1000 values of ΔPSS . While analytical formulas are available to estimate the
414	uncertainty in PSS, the bootstrap method accounts for effects of spatial and temporal
415	autocorrelation in environmental data, which effectively reduce the number of independent
416	observations to be < n [Hogan, R. J. and I. B. Mason, 2012].

417

418 **3 Results and Discussion**

419 Western Lake Erie is polymictic, meaning that it does not continuously stratify during the 420 summer owing to shallow bathymetry and exposure to wind. Temperature profiles simulated by 421 FVCOM show periods of temporary stratification that are strongest during calm afternoons when 422 the surface is warmed by the sun and warm summer air (Figure 2a, 19-20 Aug.). At night, 423 cooling of the surface often causes deepening of the surface mixed layer by convection. This diel 424 cycle can be overpowered by shear-induced mixing during windy periods (Fig. 2a, 21-22 Aug.). 425 The temperature difference over the water column during periods of stratification is small (Fig. 426 2a), but the static stability is sufficient to cause turbulent diffusivity to vary by orders of 427 magnitude over a depth range of a few meters (Fig. 2b). Random walk turbulence schemes are 428 susceptible to formation of artificial accumulations of particles in the presence of strong gradients 429 in diffusivity [Visser, A., 1997].

430 **3.1 Vertical random walk schemes**

431 We tested the random walk schemes using a 1D well-mixed condition simulation (see Methods) 432 that was run using hourly time series of diffusivity profiles output from FVCOM for the month of 433 August 2011 at stations representative of the range of conditions that occur during the summer 434 CHAB season in western Lake Erie (Fig. 1). The range of conditions represented by the 435 diffusivity profiles can be characterized by defining a Peclet number that represents the ratio of mixing time scale to floating/sinking time scale of the water column, $Pe = w_b h/\overline{K}$, where h is 436 the water column depth and \overline{K} is the column mean eddy diffusivity [Ross, O. N. and J. Sharples, 437 438 2004]. Values of Pe >> 1 indicate that w_b has a strong influence on particle concentration 439 profiles. Combining the time series of diffusivity profiles that was used in the well-mixed condition simulations with $w_{\rm b} = 70 \ \mu {\rm m \ s}^{-1}$ (see below), values ranged on the order of 0.01 < Pe <440 100. A small time step is usually required for $Pe \ll 1$ because small h and large \overline{K} will produce 441 442 strong gradients in diffusivity for realistic diffusivity profiles (i.e., $K \sim 0$ at the bottom or at the 443 thermocline), and therefore, small values of the Visser time step criterion (Eq. 3). At the 3-m 444 deep station (WE6) the variable time step occasionally was limited by the specified minimum and 445 maximum values of 0.01 and 600 s, with typical hourly means of 0.2-3 s. At the 13-m deep 446 station (NDBC45005) in the central basin the water column was continuously stratified, and 447 longer time steps could be used; hourly minimum values of the variable time step were typically 0.2-2 s, mean values were 3-30 s and maximum hourly values were constrained by the upper 448 449 limit of 600 s.

450 Example time series of concentration profiles output from a well-mixed condition simulation are 451 shown in Figure 2. A typical particle accumulation artifact is visible in Fig. 2c, where the 452 normalized concentration deviated from the expected constant value of unity. The artifact formed 453 when high diffusivity in the surface mixed layer on 24 August (Fig. 2b) caused particles to jump 454 across the sharp diffusivity gradient into the area of low diffusivity in the lower half of the water 455 column without the opportunity for the gradient term (Eq. 3-4) to push the particles back toward 456 the high diffusivity SML. Improved performance can be seen in Fig. 2d, where the variable time 457 step was reduced during the high diffusivity event on 24 August, thereby limiting the maximum 458 particle displacements and reducing the magnitude of the artifact.

459 The Gräwe Milstein scheme [Gräwe, U., 2011] with the variable time step provided the best 460 combination of computational efficiency and accuracy of the random-walk numerical schemes 461 tested. The shortest run time was achieved by the Gräwe Milstein scheme (Fig. 3a), which was 462 unique in the use of linear interpolation of diffusivity to the particle position, while the other 463 schemes used cubic splines at a greater computational cost. The second order schemes, PC2 and 464 S1.5, required greater computational effort to calculate additional terms and had the longest run 465 times. The run time of the variable time step simulation was similar to that of the fixed 1 s time 466 step for the Visser and Milstein schemes because the average of the variable time step was 467 similar to the value of the fixed 1 s time step (Fig. 3a). The Milstein schemes showed improved 468 accuracy compared to the Visser scheme with the fixed 1 s time step (Fig. 3b), consistent with the 469 finding of Gräwe [2011]. All schemes were more accurate with the variable time step than with 470 the fixed 1 s time step (Fig. 3b) because the Visser time step criterion (Eq. 5) was always

471 satisfied in the case of the variable time step. The second order schemes did not offer sufficiently472 improved accuracy to compensate for their greater computational effort (Fig. 3a,b).

In contrast to our result, Gräwe (2011) found that the second order schemes did offer improved 473 474 accuracy that justified the additional computational effort, but for the case of idealized diffusivity 475 profiles specified at high vertical resolution and for a realistic test case of a tidally-mixed bay 476 with model diffusivity output at 200 levels. In our case of a shallow polymictic lake, diffusivity 477 profiles were highly irregular with sharp gradients (Fig. 2b), and diffusivity was output at only 20 478 levels. We found that spline fits often had spurious features between the levels at which 479 diffusivity was specified by FVCOM that were not representative of physically realistic 480 diffusivity profiles. Higher-order random walk schemes depend on higher-order derivative terms 481 from the spline fits to the diffusivity profiles [Gräwe, U., 2011], which may not be accurate in the 482 case of a non-representative spline fit. The higher order schemes might produce better results if 483 we were to output diffusivity at a large number of levels, but that would come at the expense of 484 greater computational effort in the hydrodynamic model. For our application, the Gräwe Milstein 485 scheme produced satisfactory results and did not require a spline fit, so it was selected for further 486 work.

487 **3.2** *Microcystis* colony size distribution and buoyant velocity

The parameter w_b represents the *Microcystis* colony terminal velocity resulting from the balance
of forces between buoyancy and fluid drag. Our approach was to specify a *Microcystis* colony
size distribution, then apply an empirical relationship between w_b and colony diameter to obtain a
frequency distribution of w_b for use in the model. According to Stoke's law for the terminal
25

492 velocity of a floating/sinking spherical particle in a fluid, one might expect the relationship 493 between w_b and colony diameter to give a straight line on a log-log plot with a slope of 2. 494 However, Nakamura et al. [1993] showed that *Microcystis* colony specific gravity approaches 495 that of the surrounding fluid as colony diameter increases owing to the fractal geometry of the 496 colonies and the increasing volume of void spaces filled with the surrounding water; the result is 497 that the slope of the log-log plot is < 2.

498 The *Microcystis* colony diameter frequency distribution measured by FlowCam was unimodal 499 with a median of 117 µm and a maximum of 740 µm (Fig. 4a, Station WE12, 4 Aug 2014). The 500 size frequency distribution measured by microscopy on samples collected at stations WE 2,4,6, 501 and 8 in July – October 2013 and June – July 2014 gave a similar size distribution to that of the 4 502 August 2014 sample (Fig. 4a). It is likely that the colony size distribution varies to some extent 503 spatially and temporally [e.g., *Lin*, *L. et al.*, 2014], and our estimate could be refined through 504 additional measurements. Even so, the consistency between our two estimated size distributions 505 gives some indication of representativeness.

506 Our measured values of Lake Erie *Microcystis* colony buoyant velocity, w_b , were similar to those 507 of Nakamura et al. [1993] for colonies larger than 200 µm in their sample collected from a lake in 508 Japan on 3 August 1990, and generally less than their 18 September sample (Fig. 4b). We were 509 not able to resolve colonies smaller than 200 µm by our method; however, large colonies account 510 for the majority of biomass and toxin concentration. For example, colonies > 112 µm accounted 511 for 93% of *Microcystis* cells (biomass) in Lake Erie samples [*Chaffin, J. D. et al.*, 2011], and 512 colonies > 100 µm showed the highest proportion of microcystin-producing genotypes, highest

513 microcystin cell quotas, and highest microcystin production rate, compared to smaller colony size 514 classes in Lake Wannsee, Germany [Kurmaver, R. et al., 2003]. In addition to colony size, 515 Microcystis buoyancy is a function of Microcystis strain and light exposure history, as it affects 516 gas vacuole and carbohydrate content of the cells [Ibelings, B. W. et al., 1991; Xiao, Y. et al., 517 2012]. Further research is necessary to define buoyant velocities over a wide size range of Lake 518 Erie Microcystis under a variety of environmental conditions. Our results, while limited in size 519 range, do show similarity between the $> 200 \,\mu\text{m}$ values of Nakamura et al. [1993] and samples 520 from two different dates in Lake Erie, and support using the lower estimate of buoyancy (Fig. 4b, 521 N93 3 Aug.) from Nakamura et al. [1993].

522 For the model simulations, we assigned buoyant velocities to Lagrangian particles by random 523 sampling with replacement from the frequency distributions shown in Figure 4c, which were 524 obtained by applying the regression lines (Fig. 4b) from the data of Nakamura et al. [1993] to the 525 diameter frequency distribution from the 8 August 2014 sample from Lake Erie (Fig. 4a). We 526 tested the sensitivity of 1D model simulations to the two buoyant velocity frequency distributions 527 shown in Figure 2c, and refer to these hereafter as "N93 3 Aug" and "N93 18 Sep". Example 528 time series of concentration profiles simulated with the low estimate of buoyancy (N93 3 Aug) 529 are shown in Figure 2e.

530 **3.3** Vertical profiles of cyanobacterial concentration and temperature

We tested the ability of the random-walk model with buoyancy to simulate realistic *Microcystis*concentration profiles by comparing measured profiles of cyanobacterial chlorophyll

concentration from Lake Erie (predominantly *Microcystis*) to corresponding 1D simulations. On
20 July, the concentration profile showed strong accumulation within the surface two meters
(Fig. 5a), which corresponded to a surface mixed layer defined by a thermocline at 2-m depth
(Fig. 5b). A second profile was measured on 9 September, which showed concentration
enrichment within a 3-m thick surface mixed layer (Fig. 5c), which was similarly defined by a
thermocline at 3-m depth (Fig. 5d).

539 The temperature difference across the thermocline in both cases was only about 1 °C (Fig. 5b,d), 540 but the FVCOM simulations indicated that such subtle stratification features can have a strong 541 influence on diffusivity (e.g., Fig. 2a,b). The accuracy of temperature simulations in 542 hydrodynamic models is often only within a few degrees, which brings into question whether the 543 subtle stratification features that are influencing the Microcystis vertical distribution can be 544 reasonably simulated by a hydrodynamic model. Even though the simulated temperature profiles 545 have a warm bias of 1-2 °C, at these locations and times, they show thermoclines at multiple 546 levels that are similar to the observed profiles (Fig. 5b,d). The deeper thermocline may have 547 formed due to convective deepening of the SML overnight, followed by surface warming during 548 the day that produced the shallower thermocline; the profiles were captured in the afternoon. It is 549 the static stability of the profile rather than the absolute temperature that is important in 550 simulation of the diffusivity, and the static stability of the simulated and observed profiles is in 551 reasonable agreement (Fig. 5b,d). Over the full set of 69 profiles, the frequency distribution of 552 static stability simulated by FVCOM was in reasonable agreement with the observed frequency

distribution, although the model was biased slightly less stable than the observations (Fig. 6b,Table 2).

555 Simulated normalized concentration profiles of buoyant particles showed enrichment within the 556 surface mixed layer, similar to the observed profiles (Fig. 5a,c). We calculated the center of mass, 557 σ_m , of the concentration profile as a column-integrated indicator of the vertical distribution of 558 concentration (horizontal lines in Fig. 5a,c). Concentration was weighted toward the surface (σ_m > -0.5) in > 80 % of the observed profiles (Fig. 6a), which is consistent with the assumption to 559 560 treat *Microcvstis* colonies as buoyant particles in the model. 561 We selected the lower estimate of buoyancy (N93 3 Aug) for use in the hindcast simulations. The 562 simulated frequency distribution of σ_m was in reasonable agreement with the observed 563 distribution for both the low and high estimates of $w_{\rm b}$, although the low estimate was closer to the 564 observations (Fig. 6a, Table 2). Similarly, the direct measurements of w_b also indicated better 565 agreement with the lower estimate of buoyancy (Fig. 4b).

3.4 Hindcast simulations of CHAB intensity and distribution

567 Having shown that 1D random walk simulations reasonably approximated the changing vertical 568 distributions of buoyant *Microcystis* colonies in response to varying turbulence, we went on to 569 test whether inclusion of this mechanism in the forecast model improved model skill. Hindcast 570 simulations were initiated from each of the 26 quality satellite images of CHAB distribution for 571 the 2011 CHAB season.

572	In one example, a hindcast simulation was initialized on 6 August, which was a calm day (wind \leq
573	5 m s ⁻¹) with an intense CHAB event throughout the central western basin (Fig. 7a). On the
574	following day, wind increased (5-10 m s ⁻¹), and a second satellite image indicated reduced
575	surface CHAB intensity and distribution (Fig. 7d). The 3D simulation captured the reduced
576	surface CHAB intensity on 7 August, while the 2D model did not, which can be seen
577	qualitatively by comparing Figures 7e and 7f, and was indicated quantitatively by reduced
578	frequency bias (B) of the 3D simulation compared to the 2D simulation (3D B = 1.10 ; 2D B =
579	1.34). On day 9 (15 Aug.), the simulated CHAB distribution was distinctly different between the
580	2D and 3D models (Fig. 7h,i). In comparison to the 2D model, the 3D model CHAB distribution
581	was more similar to the observed distribution (3D PSS = 0.56 ; 2D PSS = 0.41), having less
582	CHAB coverage in the central basin east of Sandusky and more continuous coverage along the
583	coast from Monroe to Toledo. Both 2D and 3D models simulated the advection of CHAB to Port
584	Clinton (Fig 7b,c and h,i), which was minimally affected by CHAB on 6 August and fully
585	covered on 15 August (Fig. 7a,g).
586	In a second example, a hindcast simulation was initialized on 29 August, which was a date with
587	only partial coverage by the satellite image, leaving no data over much of the western basin (Fig.
588	8a). Output from a previous model run was used to initialize the CHAB distribution in the

589 western basin (Fig. 8b,c). On simulation day four (2 Sept.) a second partial satellite image

590 indicated extensive CHAB coverage in the western basin (Fig. 8d), consistent with both models.

591 Both 2D and 3D models underestimated the CHAB coverage, although the 3D model better

592 matched the observed coverage (2D B = 0.81, PSS = 0.76; 3D B = 0.90, PSS = 0.84; Fig. 8d,e,f).

The partial image on 2 September did not show the extensive CHAB outbreak into the central basin east of Learnington, Ontario, although it was simulated by both 2D and 3D models (Fig. 8e,f), and was revealed the following day in the 3 September satellite image (Fig. 8g). The 3D model better simulated the CHAB distribution on simulation day 5 (3 Sept.) than the 2D model (2D B = 0.79, PSS = 0.68; 3D B = 0.99, PSS = 0.80; Fig. 8g,h,i).

598 The examples in Figures 7 and 8 show, that both 2D and 3D models capture some observed 599 events that may be attributed to advection, but the 3D model performed better in several cases. 600 The 3D model is initialized with a better estimate of total biomass than the 2D model because an 601 estimate of the surface mixed layer depth for buoyant *Microcystis* colonies is used to assign the 602 depth over which the satellite-derived surface concentration is applied. In addition, the 3D model 603 is able to simulate changing surface concentration in response to changing mixed layer depth. 604 Finally, the 3D model produced different final CHAB spatial distribution than the 2D model, 605 which likely results from the more accurate vertical distribution within a complex 3D flow field.

606

3.5 Summary of hindcast skill statistics

Skill statistics were summarized by simulation day to evaluate how long the model can be run
from initialization before skill begins to decline. The Pierce skill score (PSS) gave positive values
for the 2D model, 3D model, and the persistence forecast on simulation days 1-10 (Fig. 9).
Positive values of PSS indicate that the hit rate was greater than the false positive rate, and
therefore the model had greater skill than a random forecast or constant CHAB or no-CHAB
prediction [*Hogan, R. J. and I. B. Mason*, 2012]. The frequency bias was less than 1.0 on 8 of 10
forecast days, indicating that both models had an overall bias toward under-prediction, although

614 not consistently so (Fig. 9a,b). The 95% confidence intervals on the difference in PSS indicated 615 that the 3D model displayed significantly greater skill in the hindcast simulations than the 616 persistence forecast through simulation day 6, and was not significantly worse than the 617 persistence forecast through day 10 (Fig. 10). The 3D model had significantly greater skill than 618 the 2D model over the full simulation period. The 2D model had significantly less skill than the 619 persistence forecast on all simulation days (Fig. 10).

620 It may be surprising that the persistence forecast displayed a reasonable level of skill. This can be 621 explained in that the spatial distribution of CHABs in Lake Erie has a number of persistent 622 features. For example CHABs often persist in the southern and western portions of the western 623 basin, while they are rarely present in the Detroit River plume and in the central basin east of the 624 islands, as indicated by 13 years of Lake Erie CHAB spatial patterns compiled by Wynne and 625 Stumpf [2015]. The model does not necessarily preserve these persistent features. For example 626 CHABs may be erroneously flushed from Maumee Bay in long simulations, although this 627 happened to a lesser extent in the 3D model than in the 2D model (Fig. 7g,h,i). The skill of the 628 persistence forecast indicates that the most recent satellite image is a reasonable indication of the 629 CHAB distribution for several days after.

Skill statistics based on pixel-by-pixel comparisons, and use of a persistence forecast as a
benchmark, provide a useful point of comparison among models, but do not capture all aspects of
model performance. For example, the large simulated plume that extended into the central basin
on 3 September (Fig. 8g,h,i) does not exactly match the observed plume in terms of shape and
position. This pattern mismatch detracted from pixel-by-pixel skill statistics, but both models

635 provided information regarding the existence of this transport event before it could be seen in 636 satellite imagery; information that would be useful to forecast users even if the shape of the plume is not entirely accurate. A persistence forecast can score reasonably well in skill statistics 637 638 that compare spatial patterns, but cannot provide any information on likely transport trajectories. 639 Therefore, even though the 2D model had less skill than the persistence forecast (Fig. 10), this 640 does not indicate that the 2D model has no value because it may provide useful information on 641 likely transport trajectories. It is a challenge to formulate skill statistics that test for accuracy in simulation of transport events, largely because it is difficult to identify and quantify transport 642 643 events by comparing among subsequent satellite images. Wynne et al. [2011] attempted to 644 calculate skill statistics based on movement of the bloom centroid; however, this approach is 645 likely to work only for special cases because accurate calculation of the bloom centroid is 646 sensitive to missing data (cloud cover) and the bloom often consists of multiple patches that may 647 move in different directions rather than one distinct patch. Formulation of appropriate skill 648 statistics for CHAB forecasts is an area for further work.

649 **3.6 Ecological significance**

Aside from the specific application of CHAB forecasting, the observations and simulations shown here provide interesting insights on the physical processes that influence phytoplankton ecology in a polymictic lake. Previous studies of Lake Erie circulation and thermal structure considered the western basin to be largely unstratified [e.g., *Beletsky, D. et al.*, 2013], but our study highlighted the importance of fine-scale thermal structure in the western basin in a biological context. Surface mixed layer depth varies hour by hour due to subtle features in the

656 temperature profile caused by the diel cycle of surface heating and cooling, further modified by 657 varying wind stress. Colony buoyancy is sufficient to keep *Microcystis* concentrated within the 658 constantly changing surface mixed layer depth. The thermal structure is subtle in comparison to 659 the continuous seasonal stratification that occurs in deeper lakes, and in the central and eastern 660 basins of Lake Erie, but important nonetheless to the Microcystis vertical distribution. The 661 position of Microcystis colonies in the water column is critical to their light exposure, nutrient 662 acquisition, and ultimately to their ability to dominate the phytoplankton community, and 663 produce toxic blooms.

664

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677 **5 References**

- Anderson, E. J., A. J. Bechle, C. H. Wu, D. J. Schwab, G. E. Mann, and K. A. Lombardy (2015),
- 679 Reconstruction of a meteotsunami in Lake Erie on 27 May 2012: Roles of atmospheric conditions
- 680 on hydrodynamic response in enclosed basins, J. Geophys. Res.: Oceans, 120, 8020-8038,
- 681 doi:10.1002/2015JC010883.
- Anderson, E. J., and M. S. Phanikumar (2011), Surface storage dynamics in large rivers:
- 683 Comparing three-dimensional particle transport, one-dimensional fractional derivative, and
- multirate transient storage models, *Water Resour. Res.*, 47(9), W09511.
- Anderson, E. J., and D. J. Schwab (2013), Predicting the oscillating bi-directional exchange flow in the Straits of Mackinac, *J. Great Lakes Res.*, *39*(4), 663-671.
- Anderson, E. J., D. J. Schwab, and G. A. Lang (2010), Real-Time Hydraulic and Hydrodynamic
- Model of the St. Clair River, Lake St. Clair, Detroit River System, J. Hydraul. Eng., 136, 507.
- Bai, X., J. Wang, D. J. Schwab, Y. Yang, L. Luo, G. A. Leshkevich, and S. Liu (2013), Modeling
- 690 1993–2008 climatology of seasonal general circulation and thermal structure in the Great Lakes
- 691 using FVCOM, Ocean Model., 65, 40-63.
- 692 Beletsky, D., N. Hawley, and Y. R. Rao (2013), Modeling summer circulation and thermal 693 structure of Lake Erie, *J. Geophys. Res.: Oceans*, *118*(11), 6238-6252.
- 694 Bridgeman, T. B., J. D. Chaffin, and J. E. Filbrun (2013), A novel method for tracking western 695 Lake Erie Microcystis blooms, 2002–2011, *J Great Lakes Res*, *39*(1), 83-89.
- 696 Brittain, S. M., J. Wang, L. Babcock-Jackson, W. W. Carmichael, K. L. Rinehart, and D. A.
- 697 Culver (2000), Isolation and characterization of microcystins, cyclic heptapeptide hepatotoxins
- from a Lake Erie strain of *Microcystis aeruginosa*, J. Great Lakes Res., 26(3), 241-249.
- Bundy, M. H., T. F. Gross, H. A. Vanderploeg, and J. R. Strickler (1998), Perception of inert
- particles by calanoid copepods: behavioral observations and a numerical model, *J. Plankton Res.*,
 20(11), 2129-2152.
- 702 Catherine, A., N. Escoffier, A. Belhocine, A. Nasri, S. Hamlaoui, C. Yéprémian, C. Bernard, and
- M. Troussellier (2012), On the use of the FluoroProbe®, a phytoplankton quantification method
- based on fluorescence excitation spectra for large-scale surveys of lakes and reservoirs, *Wat. Res.*, 46(6), 1771-1784.
- 706 Chaffin, J. D., T. B. Bridgeman, S. A. Heckathorn, and S. Mishra (2011), Assessment of
- 707 Microcystis growth rate potential and nutrient status across a trophic gradient in western Lake
- 708 Erie, J. Great Lakes Res., 37(1), 92-100.
- 709 Chen, C., H. Liu, and R. C. Beardsley (2003), An unstructured grid, finite-volume, three-
- 710 dimensional, primitive equations ocean model: application to coastal ocean and estuaries, J.
- 711 Atmos. Oceanic Technol., 20(1), 159-186.

- 712 Chorus, I., and J. Bartram (1999), Toxic cyanobacteria in water: a guide to their public health
- consequences, monitoring and management *Rep.*, 400 pp, World Health Organization, London.
- 714 Churchill, J. H., J. Runge, and C. Chen (2011), Processes controlling retention of spring-spawned
- 715 Atlantic cod (Gadus morhua) in the western Gulf of Maine and their relationship to an index of
- recruitment success, Fish. Oceanogr., 20(1), 32-46.
- 717 French, R., L. Cargnelli, and M. C. Doyle (2011), Lake Erie Binational Nutrient Management
- 718 Strategy: Protecting Lake Erie by Managing Phosphorus *Rep.*, 30 pp, Lake Erie Lakewide
- 719 Management Plan Nutrient Management Task Group,
- 720 <u>http://www.epa.gov/lakeerie/binational_nutrient_management.pdf</u>.
- Galperin, B., L. Kantha, S. Hassid, and A. Rosati (1988), A quasi-equilibrium turbulent energy
 model for geophysical flows, *J. Atmos. Sci.*, 45(1), 55-62.
- 723 Gilbert, C., W. Gentleman, C. Johnson, C. DiBacco, J. Pringle, and C. Chen (2010), Modelling
- dispersal of sea scallop (*Placopecten magellanicus*) larvae on Georges Bank: the influence of
- depth-distribution, planktonic duration and spawning seasonality, *Prog. Oceanogr.*, 87(1), 37-48.
- 726 GLWQA Annex 4 (2015), Recommended Phosphorus Loading Targets for Lake Erie, Annex 4
- 727 Objectives and Targets Task Team Final Report to the Nutrients Annex Subcommittee,
- 728 <u>http://www.epa.gov/glwqa</u>, 70.
- Gräwe, U. (2011), Implementation of high-order particle-tracking schemes in a water column
 model, *Ocean Model.*, *36*(1), 80-89.
- 731 Gräwe, U., E. Deleersnijder, S. H. A. M. Shah, and A. W. Heemink (2012), Why the Euler
- scheme in particle tracking is not enough: the shallow-sea pycnocline test case, *Ocean Dynam.*,
 62(4), 501-514.
- Henry, T. (2014), Toledo seeks return to normalcy after do not drink water advisory lifted, *The Toledo Blade*, *5 August*.
- Hogan, R. J., and I. B. Mason (2012), Deterministic forecasts of binary events, *Forecast Verification: A Practitioner's Guide in Atmospheric Science, Second Edition*, 31-59.
- Huret, M., J. Runge, C. Chen, G. Cowles, Q. Xu, and J. Pringle (2007), Dispersal modeling of
- fish early life stages: sensitivity with application to Atlantic cod in the western Gulf of Maine,
- 740 Marine Ecol.: Prog. Ser., 347, 261-274.
- 741 Ibelings, B. W., L. R. Mur, and A. E. Walsby (1991), Diurnal changes in buoyancy and vertical
- 742 distribution in populations of Microcystisin two shallow lakes, *Journal of Plankton Research*,
- 743 *13*(2), 419-436.
- Kane, D. D., J. D. Conroy, R. P. Richards, D. B. Baker, and D. A. Culver (2014), Re-
- eutrophication of Lake Erie: Correlations between tributary nutrient loads and phytoplankton
- 746 biomass, J. Great Lakes Res., 40(3), 496-501.
- 747 Kring, S. A., S. E. Figary, G. L. Boyer, S. B. Watson, and M. R. Twiss (2014), Rapid in situ
- measures of phytoplankton communities using the bbe FluoroProbe: evaluation of spectral
 - 36

- calibration, instrument intercompatibility, and performance range, Can. J. Fish. Aquat. Sci.,
- 750 *71*(7), 1087-1095.
- 751 Kurmayer, R., G. Christiansen, and I. Chorus (2003), The abundance of microcystin-producing
- 752 genotypes correlates positively with colony size in Microcystis sp. and determines its microcystin
- net production in Lake Wannsee, *Applied and Environmental Microbiology*, 69(2), 787-795.
- Lin, L., G. Appiah-Sefah, and M. Li (2014), Using a laser particle analyzer to demonstrate
- relationships between wind strength and *Microcystis* colony size distribution in Lake Taihu,
- 756 China, J. Freshwater Ecol., 30(3), 425-433, doi:10.1080/02705060.2014.976666.
- 757 Medrano, E. A., R. Uittenbogaard, L. D. Pires, B. van de Wiel, and H. Clercx (2013), Coupling
- 758 hydrodynamics and buoyancy regulation in *Microcystis aeruginosa* for its vertical distribution in
- 759 lakes, *Ecol. Model.*, 248, 41-56, doi:10.1016/j.ecolmodel.2012.08.029.
- 760 Michalak, A. M., E. J. Anderson, D. Beletsky, S. Boland, N. S. Bosch, T. B. Bridgeman, J. D.
- 761 Chaffin, K. Cho, R. Confesor, and I. Daloğlu (2013), Record-setting algal bloom in Lake Erie
- caused by agricultural and meteorological trends consistent with expected future conditions,
- 763 Proc. Natl. Acad. Sci. U. S. A., 110(16), 6448-6452.
- Nakamura, T., Y. Adachi, and M. Suzuki (1993), Flotation and sedimentation of a single
 Microcystis floc collected from surface bloom, *Water Res.*, 27(6), 979-983.
- O'Neil, J., T. W. Davis, M. A. Burford, and C. Gobler (2012), The rise of harmful cyanobacteria
 blooms: the potential roles of eutrophication and climate change, *Harmful Algae*, *14*, 313-334,
 doi:10.1016/j.hal.2011.10.027.
- 769 Obenour, D. R., A. D. Gronewold, C. A. Stow, and D. Scavia (2014), Using a Bayesian
- hierarchical model to improve Lake Erie cyanobacteria bloom forecasts, *Water Resour. Res.*, 50,
 7847-7860, doi:10.1002/2014WR015616.
- 772 Rinta-Kanto, J. M., E. A. Konopko, J. M. DeBruyn, R. A. Bourbonniere, G. L. Boyer, and S. W.
- 773 Wilhelm (2009), Lake Erie Microcystis: relationship between microcystin production, dynamics
- of genotypes and environmental parameters in a large lake, *Harmful Algae*, 8(5), 665-673.
- Rose, A. (1948), The sensitivity performance of the human eye on an absolute scale, *J. Opt. Soc. Am.*, 38(2), 196-208, doi:10.1364/JOSA.38.000196.
- 777 Ross, O. N., and J. Sharples (2004), Recipe for 1-D Lagrangian particle tracking models in
- space-varying diffusivity, *Limnol. Oceanogr.: Methods*, 2(9), 289-302.
- Ross, O. N., and J. Sharples (2008), Swimming for survival: a role of phytoplankton motility in a
 stratified turbulent environment, *J. Marine Syst.*, 70(3), 248-262.
- Simpson, J., and D. Bowers (1981), Models of stratification and frontal movement in shelf seas,
- 782 Deep-Sea Res., 28(7), 727-738.
- 783 Smagorinsky, J. (1963), General circulation experiments with the primitive equations: I. the basic
- 784 experiment, *Mon. Weather Rev.*, *91*(3), 99-164.

- 785 Strickler, J. R. (1985), Feeding currents in calanoid copepods: Two new hypotheses, in
- 786 *Physiological Adaptations of Marine Animals*, edited by M. S. Laverack, pp. 459-485, Society
- 787 for Experimental Biology, UK.
- 788 Stumpf, R. P., and T. T. Wynne (2015), Experimental Lake Erie Harmful Algal Bloom Bulletin
- 789 (Bulletin 27), www2.nccos.noaa.gov/coast/lakeerie/bulletin.
- 500 Stumpf, R. P., T. T. Wynne, D. B. Baker, and G. L. Fahnenstiel (2012), Interannual variability of 501 cyanobacterial blooms in Lake Erie, *PLoS One*, 7(8), e42444, doi:10.1371/journal.pone.0042444.
- 792 Tomlinson, M. C., R. P. Stumpf, T. T. Wynne, D. Dupuy, R. Burks, J. Hendrickson, and R. S.
- Fulton III (2016), Relating chlorophyll from cyanobacteria-dominated inland waters to a MERIS bloom index, *Remote Sensing Letters*, 7(2), 141-149.
- 795 Vanderploeg, H. A., J. R. Liebig, W. W. Carmichael, M. A. Agy, T. H. Johengen, G. L.
- Fahnenstiel, and T. F. Nalepa (2001), Zebra mussel (Dreissena polymorpha) selective filtration
- 797 promoted toxic *Microcystis* blooms in Saginaw Bay (Lake Huron) and Lake Erie, *Can. J. Fish.*
- 798 Aquat. Sci., 58(6), 1208-1221.
- Visser, A. (1997), Using random walk models to simulate the vertical distribution of particles in a
 turbulent water column, *Mar. Ecol. Prog. Ser.*, 158, 275-281.
- 801 Wang, C., X. Wu, C. Tian, Q. Li, Y. Tian, B. Feng, and B. Xiao (2015), A quantitative protocol
- 802 for rapid analysis of cell density and size distribution of pelagic and benthic Microcystis colonies
- 803 by FlowCAM, J. Appl. Phycol., 27(2), 711-720.
- Wiles, P., L. van Duren, C. Häse, J. Larsen, and J. Simpson (2006), Stratification and mixing in the Limfjorden in relation to mussel culture, *J. Marine Syst.*, *60*(1), 129-143.
- Wynne, T., R. Stumpf, M. Tomlinson, R. Warner, P. Tester, J. Dyble, and G. Fahnenstiel (2008),
 Relating spectral shape to cyanobacterial blooms in the Laurentian Great Lakes, *Int. J. Remote Sens.*, 29(12), 3665-3672.
- 809 Wynne, T. T., and R. P. Stumpf (2015), Spatial and Temporal Patterns in the Seasonal
- Bitribution of Toxic Cyanobacteria in Western Lake Erie from 2002–2014, *Toxins*, 7(5), 16491663.
- 812 Wynne, T. T., R. P. Stumpf, M. C. Tomlinson, and J. Dyble (2010), Characterizing a
- 813 cyanobacterial bloom in western Lake Erie using satellite imagery and meteorological data,
- 814 Limnol. Oceangr., 55(5), 2025-2036.
- 815 Wynne, T. T., R. P. Stumpf, M. C. Tomlinson, G. L. Fahnenstiel, J. Dyble, D. J. Schwab, and S.
- J. Joshi (2013), Evolution of a cyanobacterial bloom forecast system in western Lake Erie:
- 817 Development and initial evaluation, J. Great Lakes Res., 39, 90-99.
- 818 Wynne, T. T., R. P. Stumpf, M. C. Tomlinson, D. J. Schwab, G. Y. Watabayashi, and J. D.
- 819 Christensen (2011), Estimating cyanobacterial bloom transport by coupling remotely sensed
- imagery and a hydrodynamic model, *Ecol. Appl.*, 21(7), 2709-2721.

- Xiao, Y., N. Gan, J. Liu, L. Zheng, and L. Song (2012), Heterogeneity of buoyancy in response to light between two buoyant types of cyanobacterium *Microcystis*, *Hydrobiologia*, 679(1), 297-
- 823 311.

826 Table 1. Dates of measurements and model simulations.

Measurement or simulation	Date	
Well-mixed condition simulations	Aug., 2011	
Satellite images, 2D, 3D simulations	26 dates, July to Oct., 2011	
Colony size distribution (FlowCam)	4 Aug., 2014	
Colony size distribution (microscopy)	Weekly sampling July to Sep., 2012 and 2013	
Buoyant velocity measurements	15 and 21 July, 2015	
Vertical profiles (FluoroProbe)	Weekly sampling July to Sep., 2015	

- 830 Table 2. Statistics evaluating the skill of the Lagrangian particle model in simulating the vertical
- 831 distribution of cyanobacterial chlorophyll concentration (center of mass of the normalized
- 832 concentration profile) and of FVCOM in simulating temperature profiles (potential energy
- 833 anomaly). The statistics are the mean bias, root mean square deviation (RMSD), and Pearson
- 834 correlation coefficient (*r*).

	Bias	RMSD	r
Center of mass, N93 18 Sep, σ	0.04	0.09	0.56
Center of mass, N93 3 Aug, σ	-0.01	0.08	0.53
Potential energy anomaly, J m ⁻³	-0.20	0.72	0.83

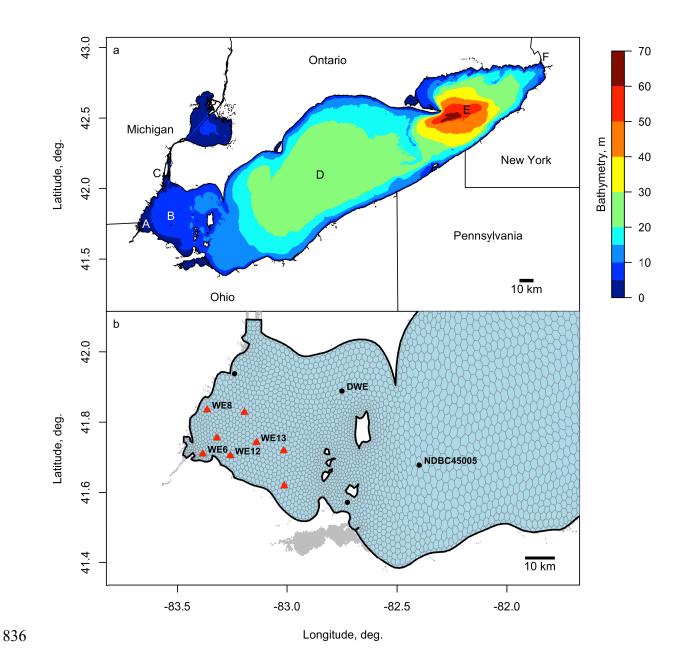


Figure 1. a) Geographic location and bathymetry of Lake Erie, showing bordering U.S. states and
the Canadian province of Ontario. Features of interest are identified, including A) Maumee Bay,
B) Western Basin, C) Gibraltar on the Detroit River, D) Central Basin, E) Eastern Basin, and F)

- 840 Buffalo on the Niagara River. b) An enlarged view of the western portion of the FVCOM model
- 841 domain. FVCOM domain boundaries are indicated with a heavy black line, and node-centered
- 842 tracer control elements with gray lines. Surface mixed layer depth was estimated in hindcast
- simulations at the stations identified with symbols. Well-mixed condition simulations were
- 844 conducted at the named stations. Profiles of temperature and cyanobacterial chlorophyll
- 845 concentration were measured at the stations indicated by red triangles.
- 846

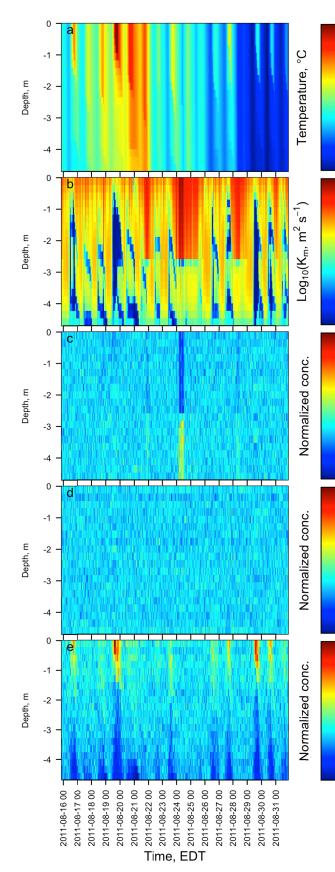


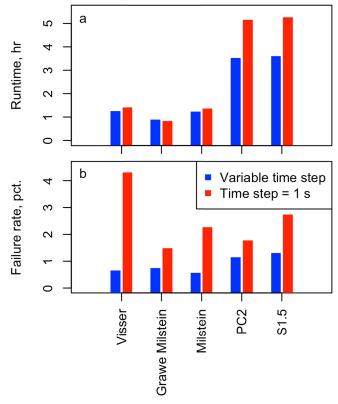
Figure 2. Time series of temperature 28.0 27.5 (a) and diffusivity (b) profiles output 27.0 26.5 from FVCOM for 16-31 August 2011 26.0 25.5 at station WE8. Time series of 25.0 normalized concentration profiles -2 from 1-D vertical random walk simulations without buoyancy (wellmixed condition test) using the Visser 3.0 scheme with fixed 1 s time step (c), 2.5 and Visser scheme with variable time 2.0 1.5 step (d) set according to the Visser 1.0 criterion (Eq. 4). A 1-D simulation 0.5 0.0 3.0 with buoyancy (N93 3 Aug) using the 2.5 Visser scheme and variable time step 2.0 (e). Artificial accumulation of 1.5 1.0 particles in a low diffusivity area with 0.5 associated depletion in the upper half 0.0 3.0 2.5 of the water column can be seen on 2.0 24 August (c), with improved 1.5 1.0 performance from the variable time 0.5 step (d). 0.0

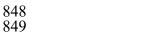
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Numerical scheme

850 Figure 3. Performance comparison of the five vertical random walk numerical schemes and the 851 variable time step scheme. One-dimensional vertical mixing simulations with 1000 particles were 852 conducted for the month of August, 2011, at the six stations indicated in Figure 1. a) Total run 853 time for the six simulations. b) percent occurrence of failure to meet the quality criterion out of 854 4464 hourly records. The quality criterion was a signal to noise ratio > 5 where the signal was 855 simulated normalized surface concentration with buoyancy (N93 3 Aug) and the noise was 856 absolute deviation of the normalized surface concentration from 1.0 in a well-mixed condition 857 test (without buoyancy). Concentration was normalized to the mean column concentration.

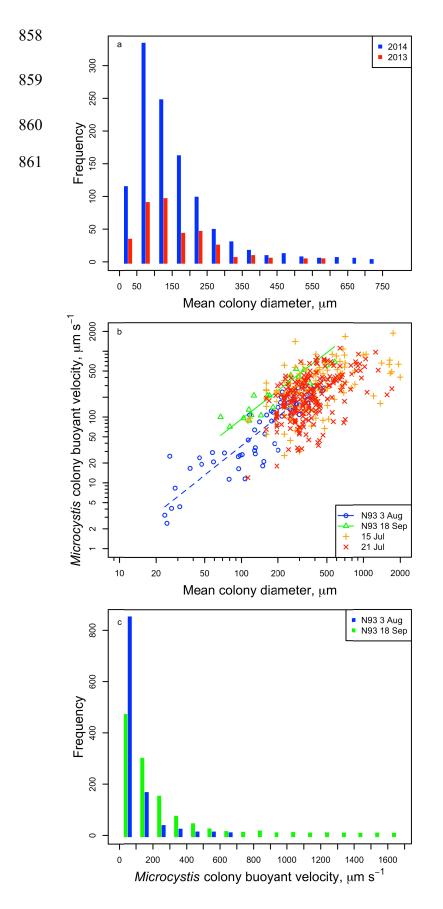
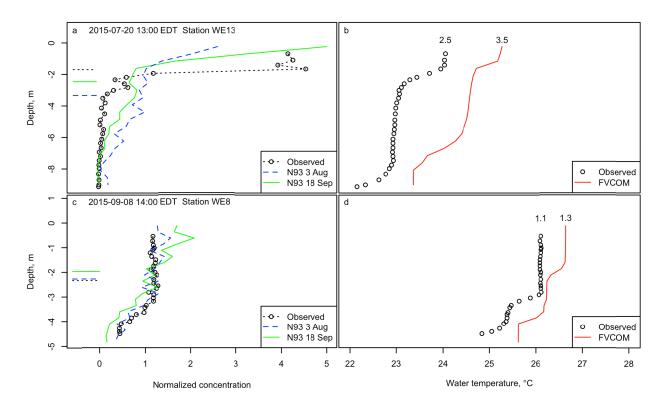


Figure 4. a) Histogram of *Microcystis* colony size distribution for samples collected in Lake Erie in 2013 and 2014. b) *Microcystis* colony buoyant velocity for samples collected in Lake Erie on 15 and 21 July, 2015, and data digitized from Nakamura et al. (1993, their Fig. 3). c) Histogram of *Microcystis* colony buoyant velocity resulting from application of the regression lines in Fig. 2b to the 2015 size distribution in Fig. 2a.





863 Figure 5. Two examples of observed vertical profiles of normalized concentration of

864 cyanobacterial chlorophyll and temperature with simulated values plotted for comparison.

- 865 Horizontal lines indicate the center of mass, σ_m , of the normalized concentration profiles, and the
- 866 potential energy anomaly, ϕ , J m⁻³, is given above the temperature profiles.

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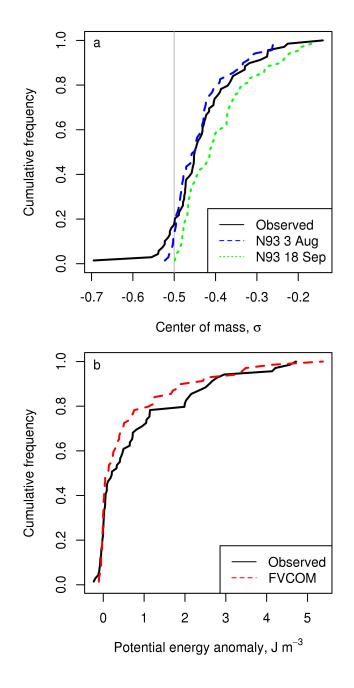
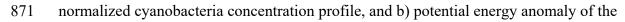


Figure 6. Cumulative frequency distributions of simulated and observed a) center of mass of the



- 872 temperature profile for 69 profiles collected in July September of 2015 at stations indicated in
- 873 Figure 1.

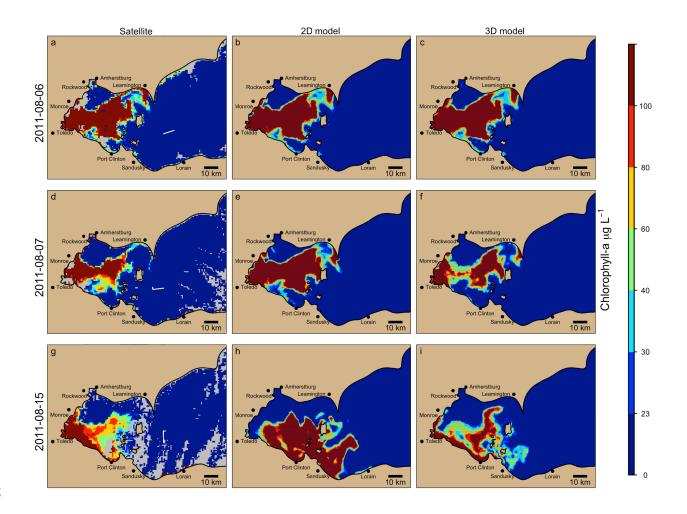
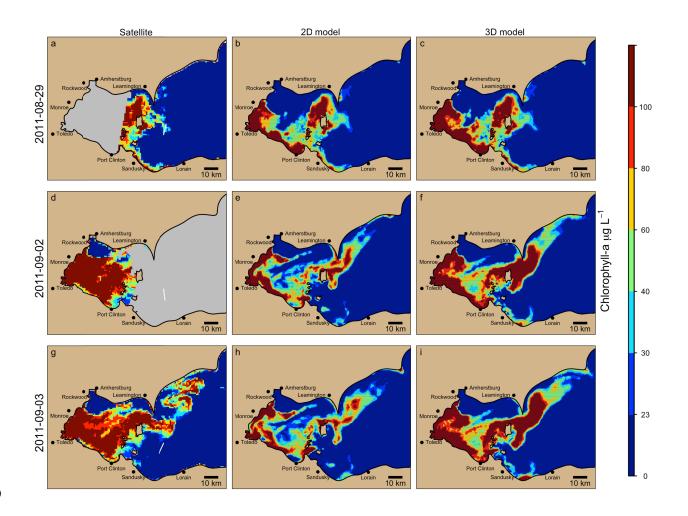




Figure 7. Comparison of 2D and 3D hindcast simulations initialized from satellite-derived 876 877 cyanobacterial chlorophyll concentration (a) on 6 August 2011. Subsequent satellite images (d,g) 878 were used for model skill assessment. Gray color indicates missing data. Wind speed, averaged over the preceding 12 hours, was $< 5 \text{ m s}^{-1}$ on 6 August and increased to 5-10 m s⁻¹ on 7 and 15 879 880 August (wind barbs a,d,g). In the 2D simulation, particles were initiated in the surface 1m and 881 vertical mixing was not simulated (advection only). In the 3D simulation, particles were 882 initialized over the surface mixed layer, as determined by preliminary simulations of 1D mixing 883 with buoyancy, and the model was run with 3D advection in addition to vertical mixing with

- 884 buoyancy. While both 2D and 3D models simulated CHAB advection toward Port Clinton
- observed on 15 August, the 3D model better simulated reduced intensity and coverage observed
- 886 on 7 and 15 August due to higher winds, and continued CHAB coverage near Toledo and
- 887 Monroe on 15 August.
- 888



889

Figure 8. Comparison of 2D and 3D hindcast simulations initialized from satellite-derived cyanobacterial chlorophyll concentration (a) on 29 August 2011. Symbols and model setup are explained in Figure 8. While both 2D and 3D models simulated CHAB advection east of Leamington into the central basin observed on 3 September, the 3D model better simulated CHAB intensity and extent in the western basin observed on 3 September. Wind speed was < 5 m s⁻¹ during the simulation period, but was north at 10 m s⁻¹ until 12 hours prior to the initial image.

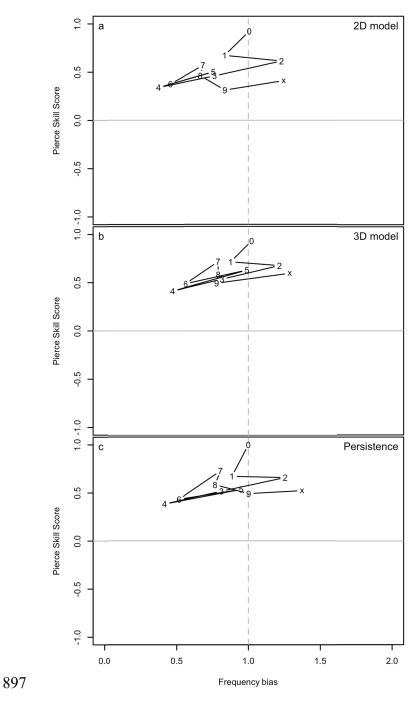


Figure 9. Skill-bias plots for the 2D model (advection only), 3D model (advection, vertical mixing, and buoyancy), and persistence forecasts for the 2011 hindcast simulations. The plot symbol indicates the simulation day (0 = initial, x = day10). Pierce Skill Score (PSS) is the hit rate minus the false detection rate, and frequency bias (B) is the ratio of forecast hits to observed hits. Positive PSS indicates greater skill than a random forecast. Frequency bias of 1.0 indicates the same number of CHAB pixels were predicted as observed.

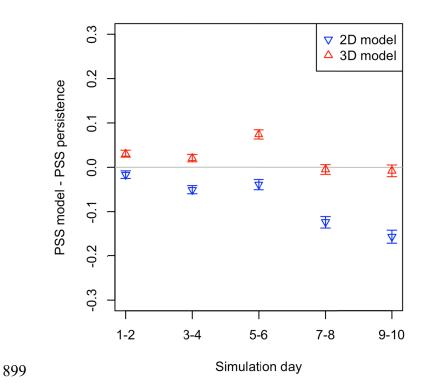
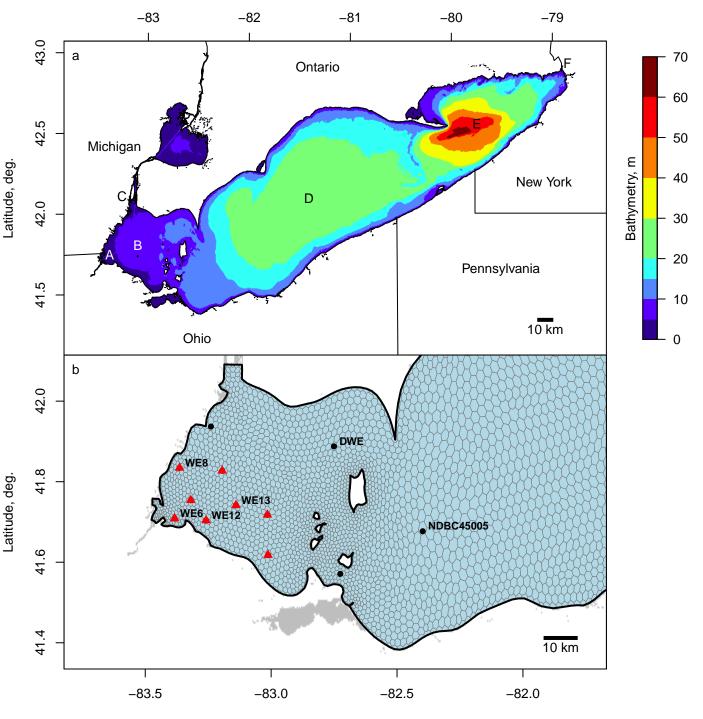


Figure 10. Pierce skill score (PSS) of the model minus PSS of the persistence forecast. Positive values indicate greater skill for the model than for the persistence forecast. Error bars indicate the 902 95% bootstrap confidence interval on the difference in PSS for the 26 hindcast simulations from the 2011 CHAB season, grouped into two-day intervals. The 3D model (including vertical mixing with buoyancy) had greater skill than the 2D model (advection only) and greater skill than the persistence forecast through day 6 and comparable skill out to day 10.

Figure 1. Figure



Longitude, deg.

Figure 2. Figure

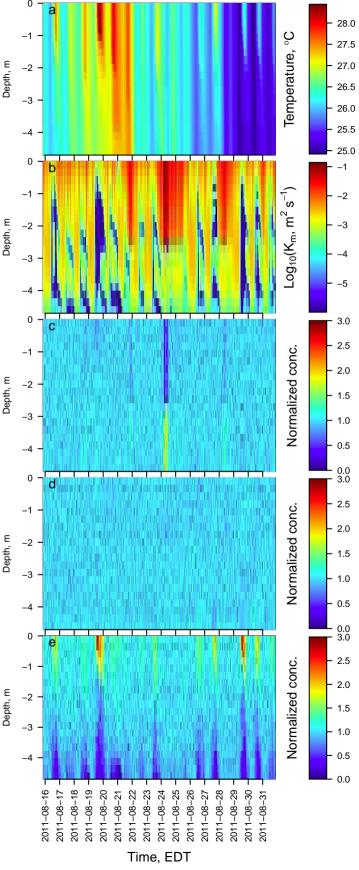


Figure 3. Figure

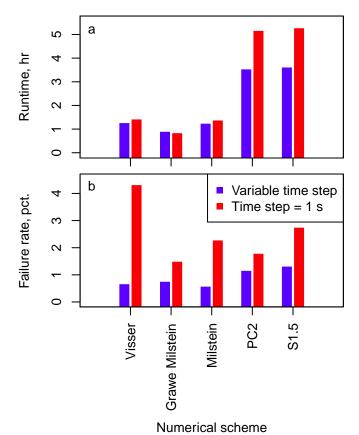
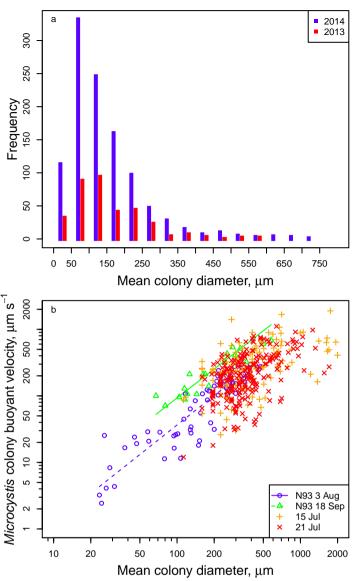


Figure 4. Figure



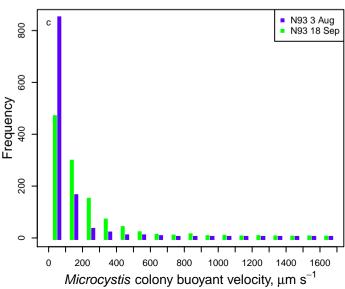
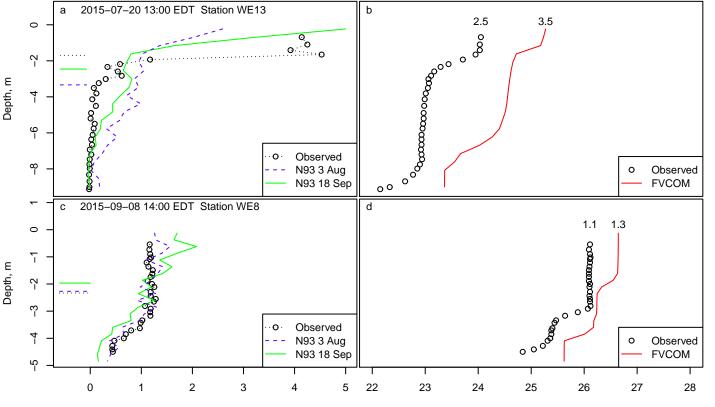


Figure 5. Figure



Normalized concentration

Water temperature, °C

Figure 6. Figure

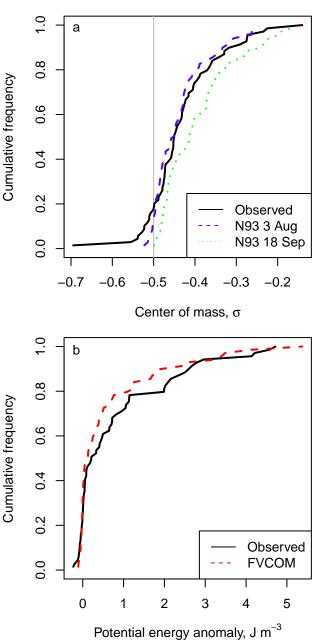


Figure 7. Figure

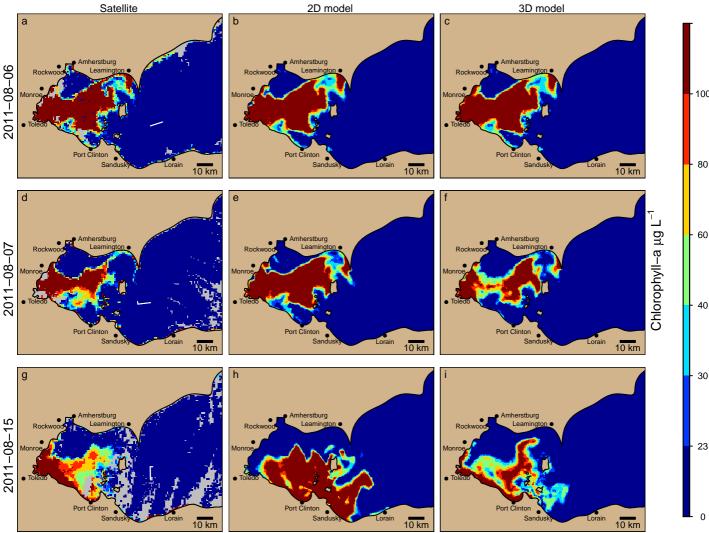


Figure 8. Figure

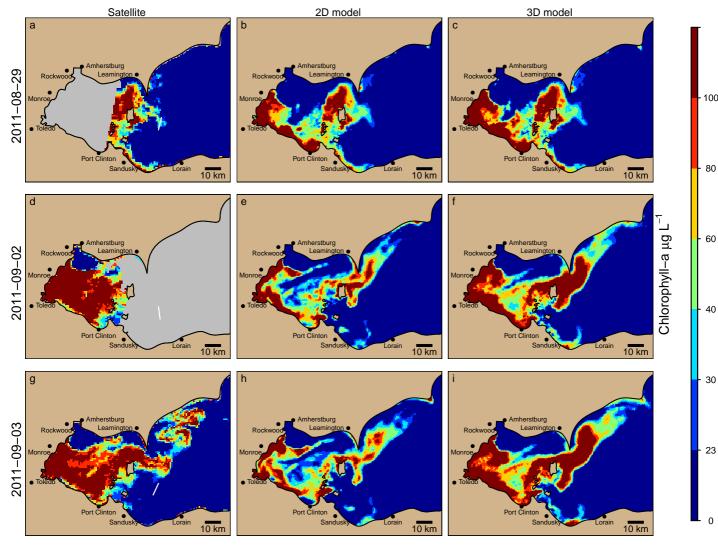
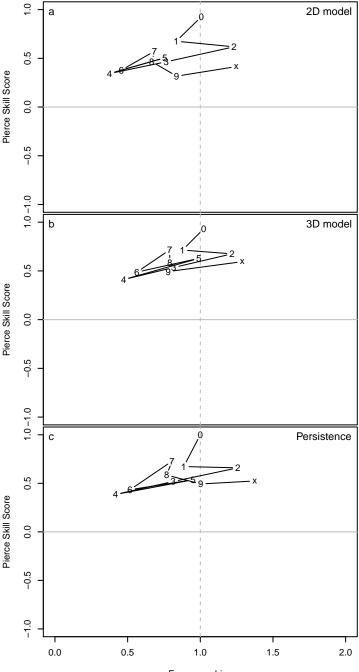
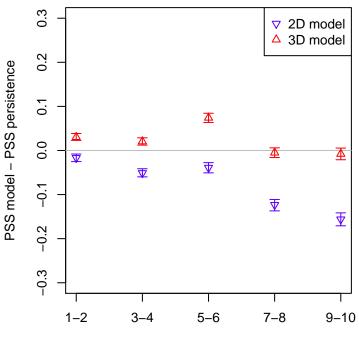


Figure 9. Figure



Frequency bias

Figure 10. Figure



Simulation day