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Within- and between-individual (co)variance partitioning reveals limited pleiotropic effects of testosterone on immune function, sexual signaling, and parental investment — Source link [2]

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13 Abstract

14 How and why individuals differ from each other is a central question in behavioral and evolutionary ecology, because selection particularly acts on this among-individual variation. It is therefore 15 important to accurately partition phenotypic variances into their within- and between-individual 16 components. Partitioning covariances into both components can also inform about underlying 17 mechanistic pathways that potentially interlink trait expression. In the current study we applied such a 18 (co)variance partitioning approach to test key predictions of two central hypotheses in behavioral 19 ecology, namely the immunocompetence handicap hypothesis and the challenge hypothesis. To this 20 21 end, we assessed potential pleiotropic effects of testosterone on male sexual signaling, immune 22 function and parental care. We here repeatedly measured a set of relevant traits in 47 breeding pairs of 23 captive canaries (Serinus canaria). We found that a within-individual increase in female testosterone 24 level suppressed immune function. Furthermore, testosterone levels were positively related to male 25 song repertoire size as an important component of sexual signaling at the between-male level. These 26 were, however, the only relevant significant correlations. Overall, our data do therefore not 27 convincingly support the hypotheses tested, and suggest rather limited hormonal pleiotropic effects of testosterone on immune function, parental care and male sexual signaling, at least in our study system. 28

29

30 Keywords

- 31 challenge hypothesis, hormonal pleiotropy, immunocompetence handicap hypothesis, parental care,
- 32 sexual signaling, suites of traits, trade-off

33 Significance statement

34 Phenotypic variances and covariances can nowadays be partitioned easily into within- and betweenindividual components. These components inform about trait repeatability and the extent to which 35 multiple traits form phenotypic suites of traits, as well as about their joint underlying mechanistic 36 pathways. Testosterone for example, may be key to mediate the expression of suites of traits in many 37 vertebrate species. We here used captive canaries and repeatedly measured male and female traits, 38 relevant in the context of endocrinology, sexual signaling, immune function and parental investment. 39 For both sexes, we report particularly weak correlations between testosterone and all other measured 40 traits at both within- and between-individual levels. Our case study thus questions the pleiotropic 41 effects of testosterone, but exemplifies the applicability and relevance of (co)variance partitioning in 42 43 behavioral ecological research.

44 Introduction

45

Steroid hormones such as testosterone (T) regulate the development of secondary sexual traits like 46 weaponry, bright coloration or complex songs in many male vertebrates (reviewed by Adkins-Regan 47 48 2005; Kempenaers et al. 2008), which are typically positively related to male reproductive success. 49 Their expression is thought to honestly signal the quality of the bearer, as only high quality individuals are supposed to sustain the costs associated with their production and maintenance (Grafen 1990; 50 51 Zahavi and Zahavi 1997). Indeed, high plasma T levels, and thus enhanced sexual signaling, has been 52 associated with a number of costs (summarised by Kempenaers et al. 2008). The two most often mentioned costs are a reduction in pathogen and parasite resistance (cf. immunocompetence handicap 53 hypothesis: Folstad and Karter 1992; Roberts et al. 2004) and a reduction of parental care (cf. a key 54 prediction of the challenge hypothesis: Wingfield et al. 1990; Hirschenhauser and Oliveira 2006; 55 Wingfield 2016). These concurrent positive and negative effects of T on life-history traits, along with 56 the involved physiological trade-offs have been extensively studied, though nearly exclusively in 57 58 males (but see Ketterson et al. 2005; Zysling et al. 2006). This makes T an important within-individual 59 mechanistic component causing hormone-mediated suites of traits (a mechanism known as hormonal 60 pleiotropy, e.g. Hau 2007; Ketterson et al. 2009) that may be adaptive because it enables an organism 61 to adjust its phenotype to a novel environment via phenotypic shifts (McGlothlin and Ketterson 2008). But the evolution of a single trait may be hampered when entwined with the expression and thus 62 selection on other traits, via a common physiological mechanism (e.g. hormonal pleiotropy) or via a 63 64 genetic mechanism (e.g. gene pleiotropy or linkage disequilibrium), which also depends on the magnitude and sign of the correlation (Wolf and Weissing 2012). 65

66 Phenotypic correlations among traits occur both at a within- and between-individual level. Within-67 individual correlations indicate the similarity in plastic responses of two focal traits across repeated 68 measurements. In other words, these correlations specify the extent of simultaneous changes in 69 expression of phenotypic traits in response to a common internal or environmental factor. Traits may 70 thus change in concert, which could be driven by a common mechanistic pathway. Between-individual 71 correlations then indicate the extent of permanent environmental (e.g. maternal effects) and genetic 72 correlations (e.g. by gene pleiotropy or linkage disequilibrium) between focal traits, and may cause multiple traits to respond to selection as a unit (cf. suites of traits; Sinervo and Svensson 2002; 73 Dochtermann and Dingemanse 2013). Covariance partitioning of phenotypic traits may therefore be a 74 75 first indicative tool to reveal the potential for constrained evolutionary trajectories. Such constraining 76 effects can be substantial, as has been shown for a variety of behavioral (Wolf and Weissing 2012) and 77 endocrine (Ketterson et al. 2009) traits. Phenotypic covariances can nowadays easily be partitioned into within- and between-individual components, which allows for testing basic predictions from 78 79 central hypotheses such as the immunocompetence handicap and the challenge hypotheses.

80 Decomposing phenotypic variance into within- and between-individual components (Dingemanse and 81 Dochtermann 2013), is also relevant for our understanding of the (response to) selection of a given 82 trait. Between-individual variation describes how much individuals differ from each other in their 83 average phenotype, and mostly relates to both 'permanent' environmental factors (e.g. maternal effects) and heritable genetic effects (Badyaev and Uller 2009; Dingemanse and Dochtermann 2013). 84 Selection particularly acts on this among-individual variation However, traits often vary substantially 85 within individuals too, which is particularly true for state-dependent physiological and behavioral 86 87 traits (Rands and Johnstone 2006; Williams 2008; Westneat et al. 2014). Such variation within individuals reflects the phenotypic plasticity of an individual's genotype responding to variable 88 environmental and internal conditions (and to some extent stochastic processes and measurement 89 error; Dingemanse and Dochtermann 2013; Baugh et al. 2014; Cleasby et al. 2015). The extent to 90 91 which within- or between-individual variation contributes to explain the total phenotypic variation is 92 commonly characterized by the repeatability of a trait (Lessells and Boag 1987; Bell et al. 2009).

93

94 In the current study we aimed to provide an empirical example of how (co)variance partitioning can be 95 applied to test central ecological and evolutionary hypotheses. We focused on the steroid hormone

96 testosterone, which plays a pivotal role in behavioral ecology. More precisely, previous studies highlighted the dual role of T within individuals, in which the extent of sexual signaling is traded off 97 98 against immune function (immunocompetence handicap hypothesis, Folstad and Karter 1992; Roberts et al. 2004) or parental care (as part of the challenge hypothesis, Wingfield et al. 1990; Hirschenhauser 99 and Oliveira 2006). The majority of these studies focused on the within-individual correlated 100 101 immunological and behavioral responses to experimentally enhanced T levels (e.g. Duffy et al. 2000; 102 Mougeot et al. 2004). In fact, the same expectations hold for between-individual trade-offs (van Oers 103 et al. 2011). In this case, individuals with relatively high T levels should be characterized by a suite of 104 correlated behavioral and physiological traits. We here assessed the hormonal pleiotropic effects of T 105 (Hau 2007; McGlothlin and Ketterson 2008) on three important traits, i.e. immune function, sexual 106 signaling and parental investment, which were repeatedly measured using captive canaries (Serinus 107 canaria) as a model species. We then applied mixed modeling tools (Dingemanse and Dochtermann 108 2013) to partition the (co)variances of these traits at two hierarchical levels, i.e. within-individuals and 109 between-individuals. In summary, we aimed (1) at assessing the repeatability of each investigated trait 110 under constant environmental conditions (2) at elucidating the mechanistic role of peak plasma T in mediating within-individual trait correlations and (3) at describing between-individual trait 111 correlations as a preliminary indication for the presence of suites of correlated, possibly co-adapted 112 113 phenotypic traits.

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115 Methods

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117 Study species and animal keeping

During spring 2014, we used a total of 47 adult male and 47 adult female Fife Fancy canaries (*Serinus canaria*) that originated from our own outbred pedigreed population. Males and females were housed separately in two large indoor aviaries, were exposed to a long light regime to induce reproductive activity (14:10, L:D), and had access to seeds and water *ad libitum*. Egg food was provided twice a

122 week. After five weeks, males were moved to individual cages (50x64x40 cm, GEHU cages, The 123 Netherlands) for the next three weeks to stimulate territorial behaviors. Subsequently, each male was 124 paired by allocating a female to its cage. Nest cups and nesting materials were provided. Progress on 125 nest building, egg laying and incubation was monitored daily (Iserbyt et al. 2015a; Iserbyt et al. 2015b). A first clutch with at least one egg was produced in 45 nests and a second clutch (see further) 126 127 with at least one egg was produced in 42 nests. Hatching was synchronized within broods to facilitate 128 cross-fostering (Estramil et al. 2013). Successful hatching of at least one egg occurred in 37 nests. 129 Enriched egg food was provided on a daily basis after hatching (day 0). Nestlings were cross-fostered 130 at an early stage (day 0) to standardize parental work load across nests with all foster nests containing 131 four unrelated nestlings. Nestlings were cross-fostered with respect to nestling age (max. 12h age 132 difference between nestlings) to minimize size differences among nestlings and, if possible, with 133 respect to egg order (nestlings hatched from eggs with different egg order) to equalize maternal effects 134 across nests. Within each nest, nestlings were individually marked with a unique color on their back, 135 using non-toxic markers (Artline®70N), which was reapplied when necessary. They were ringed with 136 a unique code on day 7. A total of 36 pairs successfully raised offspring, with on average 3.54 ± 0.12 (SE) fledglings per nest. 137

138

139 Physiological measurements

140 Plasma testosterone levels

Blood (\pm 100 µl) was sampled from the alar wing vein for measuring plasma testosterone (T) levels near the expected seasonal peak in T concentrations (Johnsen 1998; Parisot et al. 2005; Ketterson et al. 2005; Goymann et al. 2007). For males, blood was sampled before the pairs were formed, which corresponds to the period of territorial establishment. For females this was done on the day that the first egg was laid (see Fig. 1). Blood was always sampled between 10am and 1pm to reduce potential diurnal variation in T levels (Kempenaers et al. 2008). Plasma concentrations of T were measured using an enzyme-linked immunoassay ELISA kit (Enzo Life Sciences, ADI-900-065), full details of

the analytical procedure can be found in Iserbyt et al. (2015a). Analytical sensitivity of the ELISA-148 assay is 5.67 pg/ml and shows 100% cross-reactivity with testosterone, 14.63% with 19-149 150 hydroxytestosterone, 7.2% with androstenedione and <1% with other steroid compounds. The assay 151 kit was validated by ensuring parallelism of serial dilutions of plasma samples with the standard 4parameter logistic curve. Plasma samples were diluted 1:15 in the supplied assay buffer and 152 concentrations of T were assessed with the standard curve for each assay separately. Each plasma 153 154 sample was analyzed in duplicate (within-assay measurement error: 0.9%). Furthermore, a set of randomly chosen samples (N = 39) was measured twice in different assays, indicating an across-assay 155 measurement error of 3.2%. Similarly, within- and across-assay coefficients of variability based on 156 standard dilutions were 2.51% (N = 50) and 16.4% (N=10), respectively. 157

158

159 *Immune functioning*

Although the effect of blood sampling on an individual's health state is probably very limited (e.g. 160 most males resume singing within one hour and most females continued incubation), a three-day 161 162 recovery period was allowed before the assessment of immune status. An immune response was induced by injecting phytohaemagglutinin (PHA, 30µg dissolved in 30µl PBS [0.01M]) 163 subcutaneously in the left wing web (patagium). Injections took place between 10am and 1pm. 164 165 Patagium thickness was measured with a pressure-sensitive micrometer (Mitutoyo, accuracy 0.01mm), 166 first prior to injection and a second time exactly 24h post-injection. On both days, measurements were 167 taken six times independently (Smits et al. 1999). A mixed modelling approach, taking into account sex differences and repeated sampling of individuals indicated that 30.0% of all variance is explained 168 169 by measurement error. The mean of the six independent measurements was therefore used to reduce 170 this error. The wing web swelling in response to an injection of PHA is a commonly used proxy for the contemporary individual immune status as it is indicative for the accumulation of circulating T-171 172 lymphocytes to an injected mitogen (Smits et al. 1999).

174 Behavioral measurements

175 Song behavior

176 Bird song, a sexually selected trait, was recorded for three hours (mean \pm 1SE: 187 \pm 23 min) between 177 9 am and 3 pm, one or two days before blood sampling. To this end an omnidirectional AV-Jefe 178 TCM141 tie clip microphone was used, which was connected to a M-audio microtrack II recorder 179 (recording settings: WAV, 44,100 Hz, 16 Bits, Mono). The microphone was placed centrally on the 180 trellised front of the cage and as such was pointed towards the cage's interior. The entire cage front 181 was covered with clear perspex to reduce background noise. Spectrographic analyses were performed 182 using Avisoft-SASLab Pro 5.2.07 (Avisoft, Berlin, Germany; spectrogram parameters: FFT length 183 512, frame size 100% and overlap of 50 %). The first ten and last five recorded minutes of the recording were neglected to avoid potential human disturbance effects. Number of song bouts and 184 number of unique syllables per bout were both counted for the entire recording. Song bouts were at 185 least 1.5 seconds and were delineated by a pause of at least 0.4s (Leitner et al. 2001). Forty-four 186 187 individuals were successfully recorded. Birdsong is a multifaceted behavior that provides conspecifics with information about the signaler's genetic quality or condition (reviewed by Møller and 188 189 Pomiankowski 1993; Gil and Gahr 2002). Among the different song components we here quantified 190 three key song properties. First, repertoire size is scored as the average number of unique syllables per 191 bout, of which its magnitude or complexity is thought to rise with (genetic) quality (e.g. Hasselquist et 192 al. 1996, but see Byers and Kroodsma 2009). Second, song effort (i.e. the relative time spent singing during the recording; Madison et al. 2015), is considered as a performance-related trait and hence 193 194 reflects a male's current condition. Finally, song consistency quantifies the ability to consistently 195 repeat specific song elements and is known to contain a variety of dynamic information on the male 196 quality, including age, health status and motor skills (Rivera-Gutierrez et al. 2010; Sakata and 197 Vehrencamp 2012; Vehrencamp et al. 2013). To assess consistency, a specific type of syllable was 198 chosen that is used by most individuals in the population and is defined as a "common syllable" (Vergauwen et al. 2014). Ten of these common syllables were selected for each male, all belonging to 199 200 different song bouts. The syllables were uploaded in Avisoft-CORRELATOR version 3.1 and the

spectrograms were subsequently compared using an analysis of spectrographic cross-correlation. The resulting matrix of pairwise comparisons was averaged to obtain a single consistency estimate per individual.

204

205 <u>Parental provisioning</u>

206 Parental provisioning behavior was assessed from video recordings that were taken between 207 10am and 1pm when the nestlings were 10 and 11 days old (see e.g. Estramil et al. 2013). In order to 208 standardize hunger levels across nests, we handfed Orlux (Versele-Laga) to all nestlings until 209 satiation. Nestlings were then returned to their nest and food was removed for one hour prior to video 210 recording, which lasted 2h. Duration (s) of all observed feeding events were scored for males and females separately. For males, we further distinguished feeding events towards chicks (direct feeding) 211 and towards their partner (i.e. mate-feeding: Galván and Sanz 2011; also known as allo-feeding: 212 Estramil et al. 2014). Provisioning was estimated as the total duration of feeding behavior corrected 213 214 for the length of the recorded video and is presented as minutes per hour. Provisioning estimates like 215 this are commonly used proxies to describe parental care in birds (Mutzel et al. 2013; Bowers et al. 216 2014; Iserbyt et al. 2015b).

To minimize observer bias, blinded methods were used when all behavioral data were recorded andanalyzed.

219

220 Morphometrics and maternal investment

Standard morphometric measurements were taken prior to blood sampling. Tarsus length was measured with a digital caliper (Digimatic, Mitutoyo; accuracy 0.01 mm) and body mass with a digital balance (Kern EMB200-2; accuracy 0.01g). Tarsus length remains nearly constant in adult birds, the within-individual variation therefore nearly exclusively represents measurement error (females: 7.2%; males: 6.4%). Additionally, we quantified the clutch size and egg mass as a measure of femalereproductive investment.

227

228 Repeated measurements

All traits under study were measured twice to assess their repeatability, with the timing of sampling 229 being adjusted to the specific characteristics of the trait (see Fig. 1). For males, song recordings, blood 230 231 sampling, immune challenge and morphological measurements were spaced out by 10 days during the period of territorial establishment. Female traits (testosterone level, immune challenge, 232 morphometrics) were measured and sampled on the first day of egg laying of respectively the first and 233 second clutch (Mean \pm SE interval: 35 \pm 1 days; N = 42). For females that failed to produce a clutch 234 235 (first clutch: N = 2; second clutch: N = 5), these measurements were taken seven days after the peak of measurements in our experiment. Parental provisioning was scored when parents were feeding 236 offspring of the first clutch on day 10 and day 11, as hatching of the second clutch was inhibited. 237 238 Female reproductive traits were measured for both clutches while paired with the same male. Except 239 for egg mass, data collected from females that failed to produce a first and/or second clutch were considered as biologically meaningful and are hence included in the statistical analyses (see further). 240 Sample sizes for within- and between-individual correlations (Tables 2 and 3) may show slight 241 242 deviations from the sample sizes mentioned above, due to technical difficulties during the first or the 243 repeated physiological or behavioural measurement.

244

245 Statistical analyses

All analyses were performed in SAS 9.3 (SAS Institute Inc., Cary, NC, U.S.A.), following the equations and the exact scripts as published by Dingemanse and Dochtermann (2013). Normality assumptions were checked for all variables via Shapiro Wilk tests. To meet these assumptions, testosterone levels and mate feeding were log-transformed. All variables were centered by conversion to standardized z-scores (similar to Baugh et al. 2014; Thys et al. 2017). Specifically, univariate mixed
models including random intercepts were used to partition within- and between-individual variances.
Repeatability is then calculated as the proportion of total variance explained by between-individual
variation (Lessells and Boag 1987). Similar to the convention on effect sizes of correlations (Møller
and Jennions 2002; Garamszegi et al. 2012), we consider repeatability values <0.3 as low, between 0.3
and 0.5 as intermediate and >0.5 as high.

256 Next, bivariate mixed models with random intercepts for the factor 'individual' were used to partition 257 within- and between-individual covariances among all pairwise trait combinations (equations 7a,b 258 detailed in Dingemanse and Dochtermann 2013). Specifically, these models generate two matrixes. 259 The first matrix provides the between-individual variances for both individual traits, the between-260 individual covariance and the between-trait covariance. The second matrix provides within-individual 261 and within-trait variances, as well as the within-individual and between-trait covariances. Within- and 262 between-individual correlations are then calculated separately as the respective covariance divided by 263 the square root of both multiplied within-individual (or between-individual) trait variances (see equations 7c,d in Dingemanse and Dochtermann 2013). We chose to perform multiple bivariate mixed 264 models, rather than one multivariate mixed model, because individuals with one or more missing 265 266 values were entirely excluded by the statistical program. Thus the bivariate mixed models made 267 optimal use of the given data, by including the highest possible sample size for each trait combination (see Table 2 and 3). These models were performed for males and females separately. Significance 268 269 levels were determined via likelihood-ratio tests. For this, the absolute difference in 2*LogLikelihood from an original model and its respective alternative model was compared against a χ^2 -distribution. 270 271 For the univariate analyses, this alternative model excluded the random individual effect. For the 272 alternative models in the bivariate analyses, either the between- or within-covariance component was fixed to zero (Dingemanse and Dochtermann 2013). 273

274

275 **Results**

277 Within- and between-individual variation

All but one of the investigated traits were significantly repeatable in both males and females (see Table 1; Fig. 2). The immune response was least repeatable, but approached significance in males (r=0.26 ± 0.15; P= 0.083) and in females ($r= 0.27 \pm 0.15$; P=0.054; Fig. 2b). As expected, morphological traits were highly repeatable (male: $r \ge 0.90$; female: $r \ge 0.84$). All behavioral traits varied more between than within individuals (provisioning behavior: $r \ge 0.56$, Fig. 2c; song activity: $r \ge 0.62$). Likewise, all female reproductive traits were highly repeatable (clutch size: $r= 0.54 \pm 0.17$; average egg mass: $r=0.72 \pm 0.19$).

285

286 Within- and between-individual correlations

The only significant within-individual correlation in females was found between peak plasma T levels and immune response ($r_0 = -0.19$; $\chi^2_1 = 4.2$; P = 0.040; see Table 2; Fig. 3e). Specifically, this negative correlation indicates that an increase of the immune response from the first to the second measurement corresponds with a decrease in plasma T levels. The magnitude and sign of this correlation was similar, but not significant, in males ($r_0 = -0.19$; $\chi^2_1 = 1.7$; P = 0.19; see Table 3; Fig. 3e). In fact, none of the tested within-individual correlations were significant in males (see Table 3).

Peak plasma T levels correlated negatively with body mass in a between-female comparison (r_{ind} = -0.42; $\chi^2_1 = 5.13$; P = 0.024). In general, female body mass was positively associated with tarsus length (r_{ind} = 0.33; $\chi^2_1 = 4.9$; P = 0.034), clutch size (r_{ind} = 0.51; $\chi^2_1 = 7.49$; P = 0.006), and egg mass (r_{ind} = 0.36; $\chi^2_1 = 6.10$; P = 0.013). Heavy females also tended to provide less food to their nestlings (r_{ind} = -0.36; $\chi^2_1 = 3.51$; P = 0.061). Furthermore, a negative between-individual correlation was observed between immune response and tarsus length (r_{ind} = -0.49; $\chi^2_1 = 3.90$; P = 0.048). No other significant between-individual correlations were observed for females (see Table 2; Fig. 3b,c). 300 Male plasma testosterone levels did not correlate significantly with any of the other traits (betweenindividual correlations), except for a positive relationship with repertoire size (r_{ind} = 0.59; χ^2_1 = 4.0; P = 301 0.046; Fig. 3a). Song activity ($r_{ind}=0.37$; $\chi^2_1 = 5.8$; P = 0.015) and body mass ($r_{ind}=0.32$; $\chi^2_1 = 4.6$; P = 302 0.032) were positively correlated with tarsus length. Male body mass was also positively associated 303 with the immune response ($r_{ind} = 0.41$; $\chi^2_1 = 3.9$; P = 0.048). The only significant negative between-304 individual correlation was found between song consistency and mate feeding (r_{ind} = -0.52; χ^2_1 = 4.6; P 305 306 = 0.032). No other significant between-individual correlations were observed for males (see Table 3; 307 Fig. 3b,c).

308

309 Discussion

Covariance partitioning is a particularly useful statistical tool to decompose phenotypic variances and covariances into their within- and between-individual components. These components contain different information about the extent to which multiple traits are linked via a common mechanism or form phenotypic suites of traits. This approach has only recently been applied within an ecological framework, and it may also be highly suited to test central hypotheses in behavioral ecology, as we exemplify in this study.

316

317 Within-individual correlations

318 Within-individual correlations inform about potential mechanistic pathways, because they indicate how quantitative changes in time for one trait correspond with changes in one or several other traits 319 320 (Dingemanse and Dochtermann 2013). In the current study, we found weak evidence for correlated temporal changes across traits within individual males and females. In fact, none of these correlations 321 322 were significant after correcting for a potential type-I error (false positives, see also last part of this discussion). However, these weak or absent within-individual correlations may also be caused by a 323 324 type-II error (false negatives), or due to the particularly stable laboratory conditions between repeated 325 measurements. That said, female peak testosterone levels correlated negatively with their immune 326 response to the PHA challenge. Despite the relatively high measurement error in this latter trait, this 327 observation is in line with predictions from the immunocompetence handicap hypothesis (Folstad and 328 Karter 1992). A similar, though insignificant within-individual trade-off was observed in males, given 329 the power issue as described above. It is worth emphasizing that the mechanistic role of T has been extensively studied, though nearly exclusively in males (but see Ketterson et al. 2005; Zysling et al. 330 2006). However, females are also known to circulate significant levels of androgens, to possess 331 332 androgen receptors, as well as androgen metabolizing enzymes (Staub and De Beer 1997). Despite increased research interest, the adaptive value of female T and its causal effects on the expression of 333 334 phenotypic and behavioral traits, as well as on immune function remain equivocal (Owen-Ashley et al. 2004; Ketterson et al. 2005; Jawor et al. 2006; Goymann et al. 2007; O'Neal et al. 2008; Iserbyt et al. 335 336 2015a). A more elaborate discussion on the immunocompetence handicap hypothesis and challenge 337 hypothesis is given below.

338

339 Between-individual correlations

Elevated T levels are thought to be beneficial for males as they enhance sexual signaling (inter-sexual 340 341 selection), competitiveness (intra-sexual selection) and thus reproductive success via increased mating 342 opportunities (Oliveira 2004; Kempenaers et al. 2008). We indeed found a strong positive (between-343 individual) correlation between male T levels and repertoire size. Repertoire size has traditionally been considered as a sexually selected song trait in many bird species (Gil and Gahr 2002), including 344 canaries (Gil et al. 2004). Indeed, male repertoire size has been positively related to female 345 346 preferences (Searcy 1992; but see Byers and Kroodsma 2009), extra-pair paternity (Hasselquist et al. 347 1996) and shown to influence egg yolk composition (Gil et al. 2004). However, it has been suggested by Byers and Kroodsma (2009) that repertoire size may not have originated through sexual 348 selection per se, but may rather represent a by-product of social selection. It has been hypothesized 349 350 that multiple song characteristics reflect different messages (Gil and Gahr 2002) and hence, selection for increasingly complex social communication may have favored large repertoire sizes. This multiple 351

message hypothesis may at least in part explain why we did not observe any significant correlationsamong our three different song parameters.

354

355 Current theory predicts that sexually selected traits should be costly in order to function as honest 356 indicators of quality in mate choice (Grafen 1990; Zahavi and Zahavi 1997; Gil and Gahr 2002). 357 These costs may lead to trade-offs, the two most prominent in this context are captured by the 358 immunocompetence handicap hypothesis (Folstad and Karter 1992) and the challenge hypothesis 359 (Wingfield et al. 1990). These mechanistic trade-offs are typically studied at a within-individual level 360 by measuring the immunological and behavioral responses to experimentally enhanced T levels (e.g. 361 Duffy et al. 2000; Mougeot et al. 2004). At a between-individual level however, males with relatively high T values could be characterized by a suite of correlated behavioral and physiological traits, a 362 mechanism known as hormonal pleiotropy (Reviewed by Hau 2007; McGlothlin and Ketterson 2008; 363 364 Ketterson et al. 2009). A recent empirical example comes from great tits (*Parus major*) where artificial 365 selection lines for divergent personalities ('fast- bold' and 'slow-shy' explorers) showed a correlated 366 response in plasma levels of testosterone, immune response (van Oers et al. 2011) and corticosterone 367 (Baugh et al. 2012), which in turn may alter several metabolic functions. Our data also indicate that 368 the expression of repertoire size is T-dependent (see also Van Hout et al. 2012), but we could not 369 confirm this for any other song parameters (here: song activity and consistency) and neither did we 370 find evidence for the immunocompetence handicap hypothesis or a trade-off with parental care in males as predicted by the challenge hypothesis. After accounting for type-I errors and considering the 371 372 probability of type-II errors, we may have to conclude that the weak and limited number of between-373 individual correlations with T questions the role of hormonal pleiotropy in shaping an individual's phenotype, at least in our study system. 374

Tables 2 and 3 also depict several other interesting between-individual correlations. Although discussing these correlations in detail goes beyond the aims of this manuscript, and bearing in mind potential type-I and type-II errors (see above and details below), it is worth noting that our

morphological traits (body mass and tarsus length) were significantly related to multiple physiological and behavioral traits. This suggests that specific suites of morphological, physiological and behavioral traits exists which are partly linked by a common mechanism, like gene pleiotropy or linkage disequilibrium. It is thus possible that past artificial selection in this domesticated canary type may have caused correlated responses for several other linked phenotypic traits.

383

384 Within- versus between-individual variation

385 All other studied traits except the immune response were significantly repeatable, including traits that 386 are considered to be labile and state-dependent (Westneat et al. 2014). These are traits related to 387 female reproduction (clutch size r = 0.54; egg mass r = 0.71), behavior (range r = 0.35-0.87) and 388 endocrinology (peak T level male r = 0.36; female r = 0.31). Such overall high and significant 389 between-individual variation relative to within-individual variation is important from an evolutionary 390 perspective because repeatability estimates are often considered as an upper bound estimate of 391 heritability (Lessells and Boag 1987; Dohm 2002; Greives et al. 2016), which forms a prerequisite for evolutionary responses to selection. We observed small sex differences in repeatability estimates, 392 which contrasts with the general pattern across taxa that males are more repeatable than females for 393 example in behavioral traits (Bell et al. 2009). It has been hypothesized that T-related male sexually 394 395 selected traits should to be more predictable, providing females with reliable cues during mate choice (summarized by Bell et al. 2009). Females are also considered as the more responsive sex, e.g. 396 397 towards the needs of the brood (Johnstone and Hinde 2006; Nakagawa and Schielzeth 2010; Iserbyt et al. 2015b), inherently implying lower repeatability of female parental care compared with males. 398 399 However, we did not find evidence for this. The absence of sex-differences in repeatability in our 400 study may in part be explained by the stable laboratory environment and the relatively short interval 401 between repeated measurements (i.e. days to weeks).

402 Our study may also have practical implications for future sampling designs. Given the relatively high 403 repeatability in most of our traits, a single snapshot in time might be sufficient to quantify (between-404 individual) physiological and behavioral trait variation, at least within a defined period of the 405 reproductive cycle (e.g. during territory defense or during egg laying) and under constant 406 environmental conditions (Trösch et al. 2017). However, artificial selection during domestication may 407 also have reduced sensitivity to environmental conditions (Price 1999), which could also explain the 408 relatively high repeatability.

409

410 *Covariance partitioning in behavioral ecological studies – a preliminary conclusion*

411 Covariance partitioning is a particularly useful statistical tool, but it requires large sample sizes to draw firm conclusions. The statistical power depends on the sample size (here: maximum 47 412 individuals per sex), the number of measurements per individual (here: 2), the repeatability of a trait 413 (here: on average in females: r = 0.49; males: r = 0.50), and the observed magnitude of the between-414 415 individual correlation (here: on average for females: $r_{ind} = 0.28$; males: $r_{ind} = 0.20$). With these 416 parameters, the predicted power is about 0.3 and it would require 200 individuals per sex to reach a 417 power threshold of 0.8 (according to Dingemanse and Dochtermann 2013). Thus our study may face a 418 high probability of type-II error, and several true correlations may have remained undetected. 419 However, the statistical power is proportional to the magnitude of r_{ind} (Dingemanse and Dochtermann 420 2013), which makes it very difficult to achieve the required sample size when between-individual 421 correlations are low, as is often the case (see Tables 2 and 3). Thus whether typical behavioral 422 ecological studies reach the power threshold of 0.8 in cases of low r_{ind} remains to be shown. We 423 would also like to add a final cautionary note concerning the potential inflation of false positive correlations (i.e. type-I errors) as a result of multiple testing with the bivariate mixed models. 424 Applying Bonferroni corrections to these models would lower the significance level to $\alpha = 0.002$ and α 425 426 = 0.005 for between- and within-female correlations and to α = 0.001 and α = 0.003, respectively for 427 males. None of the correlations in table 2 and 3 were found to be significant following these 428 corrections.

429 Covariance partitioning, once the above mentioned aspects are taken into account, is nevertheless very430 promising as it allows to elucidate mechanistic pathways within individuals and informs about the

431 between-individual covariance structure among traits of interest (for a recent example, see Thys et al. 2017). Future studies elucidating the mechanistic pathways within individuals could benefit from 432 433 combining this statistical approach with experimental manipulations, for example by comparing 434 behavioral and physiological traits before and after elevating T levels (e.g. Ketterson et al. 2005; Lynn 2008; but see Goymann and Wingfield 2014). It may also be a logical next step to apply a behavioral 435 reaction norm concept (Dingemanse et al. 2010). In this approach, traits are compared across different 436 437 environmental conditions, for example, before and after exposure to a territorial intruder (Wingfield 1985). Finally, relating T profiles with various physiological and behavioral traits across successive 438 439 stages of reproduction, thus capturing temporal dynamics of within-individual correlations, may form 440 another promising line of research (also discussed in Iserbyt et al. 2015a).

441

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450 Compliance with ethical standards

451 Ethical standards

452 All experiments were carried out in accordance with the guidelines of the Ethical Committee of the453 University of Antwerp, Belgium (ID: 2014-72).

454 **Conflict of interest**

455 The authors declare no conflict of interest.

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634

636 Tables

Table 1: Overview of the measured traits. Mean (\pm SD) values are given for untransformed and638unstandardized estimates. N represents the number of individuals for which both independent639measurements could be taken. The variable sample size used for each trait are caused by technical640problems during one of both measurements. The presented within (V_0) - and between (V_{ind}) individual641variation is based on univariate mixed models, for which the variables were standardized to z-scores.642Hence, between-individual variation here also represent repeatability for each trait. P-values of643significantly repeatable traits are highlighted in bold.

Parameter	Mean ± SD	Ν	$V_0 \pm SE$	$V_{ind} \pm SE$	Chi ²	Р
(a) Female						
Testosterone (ng/ml)	1.01 ± 0.57	47	0.69 ± 0.14	0.31 ± 0.15	4.9	0.027
Wing web swell (mm)	0.55 ± 0.18	47	0.73 ± 0.15	0.27 ± 0.15	3.7	0.054
Tarsus length (mm)	17.83 ± 0.58	47	0.07 ± 0.02	0.93 ± 0.20	90.9	<0.0001
Body mass (g)	22.90 ± 2.43	47	0.18 ± 0.04	0.82 ± 0.20	48.3	<0.0001
Clutch size (#)	3.61 ± 1.20	47	0.46 ± 0.09	0.54 ± 0.17	16.3	<0.0001
Egg mass (g)	1.80 ± 0.14	42	0.28 ± 0.06	0.72 ± 0.19	30.5	<0.0001
Provisioning (min/h)	2.45 ± 0.98	36	0.38 ± 0.09	0.62 ± 0.20	16.6	<0.0001
(b) Male						
Testosterone (ng/ml)	4.80 ± 2.53	42	0.63 ± 0.13	0.37 ± 0.16	6.3	0.012
Wing web swell (mm)	0.41 ± 0.18	44	0.74 ± 0.15	0.26 ± 0.15	3.0	0.083
Tarsus length (mm)	18.16 ± 0.62	44	0.06 ± 0.01	0.94 ± 0.19	93.0	<0.0001
Body mass (g)	18.38 ± 1.45	45	0.07 ± 0.01	0.93 ± 0.19	84.8	<0.0001
Provisioning (min/h)	1.59 ± 0.67	36	0.43 ± 0.10	0.57 ± 0.19	14.1	<0.0001
Mate provisioning (min/h)	0.64 ± 0.72	36	0.13 ± 0.03	0.87 ± 0.22	47.8	<0.0001
Repertoire size (syll./bout)	6.82 ± 1.79	39	0.55 ± 0.12	0.45 ± 0.17	9.0	0.003
Song activity (%)	3.62 ± 3.11	44	0.37 ± 0.07	0.63 ± 0.17	22.3	<0.0001
Consistency	0.46 ± 0.04	30	0.64 ± 0.16	0.36 ± 0.19	4.1	0.043

Table 2: Output of the bivariate mixed models for females showing between-individual correlations (± 645 646 SE) above, and within- individual correlations (± SE) below the diagonal. Given that our data are 647 converted to z-scores, the regression coefficient (β) equals the correlation coefficient (r) in each 648 model. Hence, each value indicates the unit of change, sign, and magnitude for each correlation. The 649 value in brackets indicate the sample size used within each bivariate model. Slight variations in sample 650 size used for each model are caused by technical problems during one of both measurements. Within-651 individual correlations between tarsus length and all other traits are not presented, assuming that this trait is nearly fixed in adult birds. Similarly, within-individual correlations between female 652 provisioning (Pr) and all other traits are not indicated, given the asynchrony in our sampling scheme 653 (Figure 1), hence these correlations are considered biologically irrelevant. Significant correlations are 654 655 indicated in bold and asterisks indicate the following: (*) $0.1 \le p \le 0.05$; * $0.05 \le p \le 0.01$; ** $0.01 \le p$ < 0.001. Applying Bonferroni corrections (α level: P = 0.002 - 0.005) to reduce type-I errors resulted 656 in insignificant correlations for all pairwise comparisons. 657

	Т	WS	Та	BM	CS	EM	Pr
Testosterone (T)	-	0.02±0.15 (47)	-0.14±0.15 (47)	-0.42*±0.15 (47)	0.26±0.15 (47)	-0.13±0.15 (45)	-0.06±0.16 (36)
Web swell (WS)	-0.19*±0.15 (46)	-	-0.49*±0.14 (47)	0.30±0.15 (47)	0.58(*)±0.14 (47)	0.32±0.14 (45)	0.32±0.17 (36)
Tarsus (Ta)	-	-	-	0.33*±0.15 (47)	0.18±0.14 (47)	0.21±0.14 (45)	-0.14±0.17 (36)
Body mass (BM)	-0.09±0.29 (46)	0.09±0.29 (47)	-	-	0.51**±0.13 (47)	0.36*±0.13 (45)	-0.36(*)±0.16 (36)
Clutch size (CS)	0.03±0.18 (46)	-0.18±0.23 (47)	-	-0.02±0.23 (47)	-	0.27±0.12 (45)	0.15±0.15 (36)
Egg mass (EM)	-0.11±0.21 (41)	0.04±0.19 (42)	-	0.15±0.12 (42)	-0.05±0.19 (42)	-	0.08±0.17 (35)

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Table 3: Output of the bivariate mixed models for males. Within-individual correlations between male provisioning (Pr) or mate-provisioning (MP) and all other traits are not indicated, given the asynchrony in our sampling scheme (Figure 1). For further details, see caption table 2. Bonferroni corrections (α level: P = 0.001 – 0.003) resulted in insignificant correlations for all pairwise comparisons.

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	Т	WS	Та	BM	R	SA	С	Pr	MP
Testosterone (T)	-	0.25±0.16 (43)	-0.25±0.16 (43)	-0.04±0.15 (43)	0.59*±0.16 (41)	0.34±0.15 (43)	0.25±0.16 (38)	-0.14±0.15 (36)	0.04±0.15 (36)
Web swell (WS)	-0.19±0.17 (43)	-	0.29±0.15 (44)	0.41*±0.15 (44)	0.05±0.15 (42)	-0.08±0.15 (44)	-0.31±0.16 (39)	0.38±0.17 (36)	-0.03±0.17 (36)
Tarsus (Ta)	-	-	-	0.32*±0.14 (47)	-0.18±0.16 (42)	0.37*±0.15 (44)	-0.23±0.17 (39)	-0.05±0.17 (36)	0.18±0.17 (36)
Body mass (BM)	0.00±0.05 (43)	0.23±0.50 (44)	-	-	0.10±0.16 (42)	-0.04±0.16 (44)	-0.28±0.16 (39)	-0.33±0.15 (36)	0.06±0.16 (36)
Repertoire (R)	0.02±0.15 (38)	0.01±0.20 (39)	-	0.11±0.06 (39)	-	0.2±0.16 (42)	0.15±0.17 (39)	-0.1±0.21 (35)	0.08±0.22 (35)
Song activity (SA)	0.00±0.12 (43)	0.00±0.21 (44)	-	-0.06±0.36 (44)	0.08±0.20 (39)	-	0.12±0.15 (39)	-0.08±0.18 (36)	0.2±0.17 (36)
Consistency (C)	-0.05±0.18 (29)	0.07±0.20 (30)	-	0.07±0.53 (30)	-0.04±0.19 (30)	0.17(*)±0.23 (30)	-	0.12±0.19 (33)	-0.52*±0.18 (33)
Provisioning (Pr)	-	-	-	-	-	-	-	-	0.08±0.17 (36)
Mate provisioning (MP)	-	-	-	-	-	-	-	-0.05±0.31 (35)	-

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Fig. 1: Indicative timeline of the experimental procedure. Phenotypic measurements are indicated by
letters. The apostrophe (^c) indicates the second (repeated) measurement. A: song recordings (males).
B: blood sampling (testosterone levels), body mass and tarsus length measurements (males). C:
immune challenge (males). D: blood sampling (testosterone levels), body mass and tarsus length
measurements (females). E: immune challenge (females). F: parental provisioning (males and
females).

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Fig. 2: Graphical representation of the repeatability analyses for four key traits in our study. Values are centered around 0 for both measurements (m1 and m2) separately. Positive values indicate relatively high values within our study population, and *vice versa*. Regression fits are drawn when significant in the univariate mixed model. Filled dots and solid regression lines represent males, open circles and dashed lines represent females.

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683 Fig. 3: Graphical representation of the tested hypotheses at the between- (a-c) and within- (d, e) 684 individual level. Testosterone levels are related to repertoire size (a, d) which tests the sexual signaling hypothesis, to wing web swelling in response to an injection of PHA (b, e) which tests the 685 immunocompetence hypothesis, and to offspring provisioning (c) which tests the trade-off hypothesis 686 with parental care. The within-individual correlation between testosterone level and offspring 687 provisioning is not presented given the asynchrony in our sampling scheme (Figure 1). Filled dots 688 represent males and open dots represent females. Trait values are centered around 0 for each sex 689 690 separately. Positive and negative values in graphs a-c indicate respectively relatively high and low average values (mean of measurement 1 and 2) within our study population. Positive and negative 691 692 values in graphs d and e indicate respectively increased or decreased values for each trait from

- 693 measurement 1 to 2. Regression fits and 95% confidence bands are only drawn if found to be
- 694 significant in the bivariate mixed model.